

1 **The built environment as determinant of childhood obesity: a**  
2 **systematic literature review**

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24 **Abbreviations:**

25 PRISMA Preferred Reporting Items for systematic Reviews and Meta-Analysis

26 PROSPERO International Prospective Register of Systematic Reviews

27 BMI Body mass index

28 BMI z-score BMI standardised for age and sex

29 PIAMA Prevention and incidence of asthma and mite allergy birth cohort

30 IOTF International Obesity Task Force

31 NO<sub>x</sub> Nitrogen oxides

32 STAMINA Standard Model Instrumentation for Noise Assessments

33 PM<sub>2.5</sub> Particulate matter with diameter less than 2.5 μm

34 NO<sub>2</sub> Nitrogen dioxide

35 PM<sub>10</sub> Particulate matter with diameter less than 10 μm

36 SO<sub>2</sub> Sulphur dioxide

37 ESCAPE European Study of Cohorts for Air Pollution Effects

38 CDC Centers for Disease Control and Prevention

39 WHO World Health Organization

40 UK United Kingdom

41 NDVI Normalized Difference Vegetation Index

## 42 **Abstract**

43 We evaluated the epidemiological evidence on the built environment and its link to childhood obesity,  
44 focusing on environmental factors such as traffic noise and air pollution, as well as physical factors  
45 potentially driving obesity-related behaviours, such as neighbourhood walkability and availability  
46 and accessibility of parks and playgrounds. Eligible studies were i) conducted on human children  
47 below the age of 18 years, ii) focused on body size measurements in childhood, iii) examined at least  
48 one built environment characteristic, iv) reported effect sizes and associated confidence intervals, and  
49 v) were published in English language. A  $z$ -Test, as alternative to the meta-analysis, was used to  
50 quantify associations due to heterogeneity in exposure and outcome definition. We found strong  
51 evidence for an association of traffic-related air pollution (nitrogen dioxide and nitrogen oxides  
52 exposure;  $p < 0.001$ ) and built environment characteristics supportive of walking (street intersection  
53 density;  $p < 0.01$  and access to parks;  $p < 0.001$ ) with childhood obesity. We identified a lack of studies  
54 which account for interactions between different built environment exposures or verify the role and  
55 mechanism of important effect modifiers such as age.

## 56 **Introduction**

57 The prevalence of childhood obesity has more than tripled over the last four decades. Latest figures  
58 suggest that up to 30 percent of children in Europe are with overweight or obesity.<sup>1</sup> The growing rate  
59 of children with overweight and obesity is the most important preventable public health crisis of the  
60 21<sup>st</sup> century, with serious health, social and economic implications. Obesity in childhood often  
61 persists into adulthood with severe consequences for health. An expanding set of chronic diseases has  
62 been linked to childhood obesity including increased risk of developing cardiovascular disease, type  
63 2 diabetes and certain cancers, as well as diminished mental health.<sup>2-5</sup>

64 Obesity is preventable and reversible. Restricting energy intake and increasing energy expenditure  
65 have previously been the focus of prevention and treatment strategies. Most efforts and initiatives  
66 have, however, so far been unsuccessful at a population level and a broadened approach is warranted.<sup>6</sup>  
67 The causes of obesity are multifactorial ranging from individual, household, to policy settings. In this  
68 context, place-based obesogenic factors are increasingly being recognised as important determinants  
69 of obesity, including the social context, the environment individuals live in, and behaviours linked to  
70 modern, urban living.<sup>7</sup> In order to target place-based mitigation approaches, interventions and policy  
71 implementations, a clear understanding of the spatial context in which obesity determinants act is  
72 needed.<sup>8</sup>

73 The place we live in has increasingly been recognised as a strong determinant of health, including  
74 obesity.<sup>9</sup> In this context, the term ‘built environment’ has been coined to describe the physical and  
75 built infrastructure in which people live, learn, work, play, socialise and travel.<sup>10</sup> Within urban  
76 settings, the natural infrastructure is an integral part of the wider concept of the built environment.  
77 The built environment has strong influences on residents’ behaviours, with physical activity and  
78 sedentary lifestyles being the most widely studied.<sup>11</sup> Additionally, environmental pollution linked to  
79 the built environment such as air pollution and traffic also have strong impacts on urban health.<sup>12</sup>

80 This systematic review synthesises the empirical evidence on the built environment as determinant  
81 of childhood obesity. We focused on environmental factors including traffic noise and air pollution,  
82 as well as physical factors potentially driving obesity-related behaviours, including neighbourhood  
83 walkability and availability and accessibility of parks and playgrounds. Supported by a rigorous  
84 quality assessment and a focus on objectively measured built environment characteristics, we provide  
85 a quantitative synthesis of the updated evidence base with an emphasis on conceptual and  
86 methodological aspects, and public health implications.

87

## 88 **Methods**

### 89 **Search strategy**

90 We followed the Preferred Reporting Items for Systematic reviews and Meta-Analysis (PRISMA)  
91 guidelines and registered the protocol with the International Prospective Register of Systematic  
92 Reviews (PROSPERO) database (registration number CRD42020170337). We used a comprehensive  
93 and reproducible search strategy to identify peer-reviewed journal articles in the English language,  
94 published from inception until February 2020, focusing on three databases: EMBASE, MEDLINE  
95 and Web of Science. A preliminary search identified relevant keywords and MeSH terms at the  
96 intersection of two concept clusters: “childhood obesity” and “built environment” (Table S1).

97

### 98 **Eligibility criteria**

99 Studies were eligible for inclusion if they met the following criteria: (1) Population: Children and/or  
100 adolescents under the age of 18 years; (2) Exposure: Objectively measured environmental and  
101 physical features of the built environment potentially linked to the onset of obesity; (3) Outcomes:  
102 Objectively measured and self-reported body mass index (BMI), or BMI standardised for age and sex  
103 (BMI  $z$ -score); (4) Study design: Observational studies (cross-sectional and longitudinal)  
104 quantitatively assessing associations of outcome and exposure. We excluded studies which assessed

105 the built environment as confounder only, those which used self-reported perceived features of the  
106 built environment, and studies using controlled experiments in manipulated settings (Table S2). We  
107 also excluded studies with an explicit focus on the food environment as this was outside the scope of  
108 the review. After the removal of duplicates, articles were screened independently by two reviewers  
109 (D.M. and D.F.) against the eligibility criteria, using the online tool Covidence.<sup>13</sup>

110

### 111 **Data extraction**

112 Data extraction was performed independently by two reviewers (D.M., E.H.), discrepancies were  
113 mediated by D.F. Information was extracted on study characteristics (first author, year, study design,  
114 study area, sample size), participant characteristics (age, sex), exposure (built environment  
115 characteristic, data collection method), outcome measures (outcome, data collection methods,  
116 measure of association), individual- and area-level confounders and main findings (direction and  
117 magnitude of association, statistical significance).

118

### 119 **Quality assessment**

120 The quality of the eligible studies was assessed independently by two reviewers (D.M., E.H.),  
121 discrepancies were mediated by D.F. We used a modified Newcastle-Ottawa scale for quality  
122 assessment,<sup>14</sup> which we adapted for the assessment of observational studies. The elements used for  
123 the assessment include (1) representativeness of the exposed population, (2) selection of the non-  
124 exposed population, (3) objective ascertainment of the exposure, (4) sample size, (5) appropriateness  
125 of considered confounding factors, (6) assessment of the outcome, and (7) statistical test used for  
126 analysis (Table S3). Stars were assigned for each criterion with a maximum of twelve stars. A score  
127 of 0-4 was defined as poor quality, 5-8 as fair quality and 9-12 as good quality. Publication bias was  
128 assessed using a funnel plot.

129

### 130 **Data synthesis**

131 Due to the heterogeneity in exposure metrics and methodologies used across eligible studies, a meta-  
132 analysis was not possible. Instead, we used an alternative methodology to assess and synthesise the  
133 strength of associations, the weighted  $z$ -Test.<sup>15</sup> This approach has previously been used for systematic  
134 reviews on the built environment and health<sup>16, 17</sup> and is based on the number of studies with findings  
135 in the expected direction and their level of significance. For each study, we assigned a  $z$ -value based  
136 on the level of statistical significance ( $\alpha$ ) and direction of association (expected direction of  
137 association based on research hypothesis vs. unexpected direction of association). If associations were  
138 in the expected direction, then  $z = 1.96$  for  $\alpha = 0.05$ , and  $z = 1.64$  for  $\alpha = 0.10$ ; if associations were in  
139 the unexpected direction, then  $z = -1.96$  for  $\alpha = 0.05$ , and  $z = -1.64$  for  $\alpha = 0.10$ ;  $z = 0.00$  was assigned  
140 to null (statistically not significant) associations with  $p > 0.10$ . We summed the  $z$ -value for each  
141 reported finding and weighted these by the quality assessment score for each study, divided by the  
142 square root of the sum of squared quality assessment scores. To determine the strength of association  
143 for each built-environment-outcome combination, a two-tailed  $p$ -value was computed for each  
144 weighted  $z$ -value with interpretation of weak evidence if  $p < 0.05$ , strong evidence if  $p < 0.01$  and very  
145 strong evidence if  $p < 0.001$ .<sup>16</sup> To avoid overrepresentation of individual studies reporting built  
146 environment-outcome associations by different subgroups (e.g., boys/girls, geographic area, age  
147 group), we applied fractional weights to each finding so that the sum of the weights across all reported  
148 associations was 1.<sup>17</sup> For example, if a study reported a positive association of fine particulate matter  
149 with childhood obesity, but that association was significant ( $\alpha = 0.05$ ) only in boys ( $z = 1.96$ ) and  
150 not in girls ( $z = 0.00$ ), the  $z$ -value assigned to the study was  $1.96 * 0.5 + 0 * 0.5 = 0.98$ . Following  
151 the standard set for meta-analysis, associations for each built environment feature-outcome  
152 combination were only synthesised if five or more studies reported such associations.

## 153 **Results**

154 Results are presented separately for each built environment characteristics: (1) traffic noise, (2) air  
155 pollution, (3) neighbourhood walkability, and (4) accessibility and availability of parks and  
156 playgrounds. PRISMA flow diagrams are shown in Figures S1 to S4, respectively.

157 Our search initially identified 1192 studies with some studies included in more than one built  
158 environment domain. After the removal of duplicates and applying screening criteria, we included  
159 four studies on traffic noise and childhood obesity, 14 studies on air pollution, 19 studies on  
160 neighbourhood walkability and 28 studies on accessibility and availability of parks and playgrounds.  
161 Data extracted for all studies meeting eligibility criteria are presented in Tables S4a to S4d. We did  
162 not find evidence for publication bias (Figure S5).

163

### 164 **Childhood obesity and traffic noise**

#### 165 *Study characteristics*

166 The four studies investigating effects of traffic noise on childhood obesity were recent (2016-2019)  
167 longitudinal studies from Northern Europe (Table 1).<sup>18-21</sup> Two studies used national birth cohorts,  
168 <sup>19,20</sup> the others longitudinal studies with national coverage. Sample sizes ranged from 3,963 to 40,974  
169 participants. All studies assessed exposure to noise through standard modelling methods, linked to  
170 the home addresses of the subjects. Three studies used an implementation of the Nordic prediction  
171 method for road traffic noise, one study a national noise standard<sup>18</sup>. Methodologies between studies  
172 were generally comparable. The Swedish study<sup>20</sup> obtained height and weight from school and health  
173 records and, in part measurements, while the three other studies used height and weight from  
174 questionnaires. The Norwegian study<sup>21</sup> accounted for age and sex in the model via interaction terms  
175 to explore the effect of noise on BMI trajectory, while all others studies either used a age/sex  
176 standardisation of BMI (BMI  $z$ -score) and/or categorised BMI based on sex and age-specific cut-offs



177 for overweight and obese from the International Obesity Task Force (IOTF). All studies accounted  
178 for age, sex and maternal education in analysis, in addition to other study-specific confounders  
179 including maternal BMI prior pregnancy,<sup>19-21</sup> parental smoking,<sup>18-20</sup> neighbourhood socioeconomic  
180 status<sup>18</sup> and physical activity.<sup>20</sup> One study controlled further for urbanisation and nitrogen oxides  
181 (NO<sub>x</sub>).<sup>19</sup> Studies used either linear mixed models,<sup>18,21</sup> multiple regression<sup>19</sup> or quantile regression<sup>20</sup>  
182 with increasing levels of adjustment. All studies were of high quality with scores of 9 to 10 out of the  
183 maximum 12 stars (see Table S5).

184

### 185 *Summary of findings*

186 Due to the small number of studies, meta-analysis was not applied, and findings descriptive. Impacts  
187 of traffic noise on childhood obesity were observed in three studies, but overall results were mixed  
188 and varied by life stage (see Table S4a). Positive associations of road-traffic noise exposure during  
189 pregnancy and the risk of being with overweight in school-age children (7/8 years) were observed in  
190 Denmark and Norway,<sup>19,21</sup> but not Sweden.<sup>20</sup> For the same age group, no impact of childhood noise  
191 exposure on weight was found.<sup>18-21</sup> Wallas et al. (2019), however, studied the effect of traffic noise  
192 exposure during adolescence and found a strong association with adolescence BMI between the ages  
193 of 8 and 16 years, which was slightly stronger for girls.<sup>20</sup>

194

### 195 **Childhood obesity and air pollution**

#### 196 *Study characteristics*

197 The majority ( $n = 11$ ) of the 14 reviewed studies were longitudinal studies, the others cross-sectional.  
198 Half of the eligible studies were conducted in the U.S. ( $n = 7$ ), followed by European ( $n = 5$ ) and  
199 Asian studies ( $n = 2$ ). The largest sample size was 30,056 children in a cross-sectional study.<sup>22</sup>  
200 Longitudinal studies were smaller, also due to a loss to follow-up.<sup>23</sup> Most studies ( $n = 8$ ) were

201 conducted in urban settings, resulting in ~80% of participants residing in urban areas. Most studies  
202 ( $n = 9$ ) focused on childhood, only three studies on adolescents.<sup>18,24,25</sup> Studies analysed a wide range  
203 of air pollutants in relation to childhood obesity. The most studied pollutant was particulate matter  
204 with diameter less than  $2.5 \mu\text{m}$  ( $\text{PM}_{2.5}$ ) ( $n = 8$ ), followed by  $\text{NO}_x$  ( $n = 7$ ), nitrogen dioxide ( $\text{NO}_2$ ) ( $n$   
205  $= 6$ ),  $\text{PM}_{10}$  ( $n = 5$ ), sulphur dioxide ( $\text{SO}_2$ ) ( $n = 2$ ), ozone ( $n = 1$ ) and black carbon ( $n = 1$ ). All studies  
206 assessed air pollution exposure at the home address, one study also at school.<sup>26</sup> Five studies modelled  
207 air pollution exposure using dispersion models,<sup>25,27-30</sup> five others used Land Use Regression.<sup>18,26,31-33</sup>  
208 Two studies interpolated measurement data from multiple monitoring stations using inverse distance  
209 weighting<sup>24,34</sup> and two studies linked measurements from the nearest monitoring station.<sup>22, 35</sup> BMI  
210 was used as main outcome in six studies<sup>24,25,27,30,31,34</sup>, two longitudinal studies used BMI  
211 trajectories<sup>30,32</sup> and seven studies used weight status classification. Different growth charts and  
212 guidelines were used to standardize BMI to adjust for age and sex (BMI  $z$ -score). The most common  
213 was the Centers for Disease Control and Prevention (CDC) growth chart, used in five US  
214 studies<sup>24,25,30,31,35</sup> and one study from China<sup>22</sup>. Three studies used the World Health Organization  
215 (WHO) growth reference data and one the IOTF indications. Two studies utilized national standards  
216 from the UK<sup>29</sup> and Sweden.<sup>28</sup> The majority of studies adjusted for age and sex, one study used Tanner  
217 stage,<sup>24</sup> and one only studied four-year old children.<sup>28</sup> Three studies did not adjust for age but used  
218 age and sex standardised BMI measures.<sup>26,29,35</sup> Covariates varied widely across studies and included  
219 parental socioeconomic status, maternal BMI, birth weight, parental smoking and passive smoking  
220 exposure. All studies had a quality rating of good, ranging from 9 to 11 stars (see Table S5).

221

## 222 ***Summary of findings***

223 To synthesise findings using the  $z$ -Test, we combined  $\text{NO}_2$  and  $\text{NO}_x$  results,  $\text{PM}_{2.5}$  and  $\text{PM}_{10}$  were  
224 considered separately (Table 2). No  $z$ -statistics was derived for  $\text{SO}_2$ , ozone and black carbon due to  
225 the small number of studies. Of the eleven studies which looked at  $\text{NO}_2/\text{NO}_x$ , five reported

226 significant associations with BMI-derived outcomes,<sup>18,22,24,25,27</sup> four studies did not find  
227 significant results.<sup>26,28,29,33</sup> Two of the studies had mixed results, one found an effect only in boys,<sup>34</sup>  
228 in one study the effect dependent on the exposure period.<sup>30</sup> Overall, the association of NO<sub>2</sub>/NO<sub>x</sub>  
229 exposure on childhood obesity was strong with a two-tailed *p*-value from the weighted *z*-value being  
230 *p* = 0.003. Overall, there was no statistically significant effect of PM<sub>2.5</sub> on childhood obesity with *p*  
231 = 0.10. Five out of the eight studies investigating PM<sub>2.5</sub> did not find any significant effect, two showed  
232 a positive association,<sup>24,35</sup> and one found an effect only in boys.<sup>31</sup> Only one of the five studies looking  
233 at PM<sub>10</sub> reported an effect,<sup>22</sup> reflected by the *p*-value of 0.15. SO<sub>2</sub> and ozone were associated with  
234 increased prevalence of obesity in one of the studies<sup>34</sup> but in another study a higher SO<sub>2</sub> in utero and  
235 in childhood was associated with lower BMI at ~13 and ~15 years.<sup>22</sup> Four studies did not find any  
236 significant evidence of a link between air pollution and childhood obesity,<sup>26,28,32,33</sup> two of which were  
237 conducted in areas of modest level air pollution.<sup>28,32</sup>

238

## 239 **Childhood obesity and neighbourhood walkability**

### 240 *Study characteristics*

241 Most of the 19 included studies used a cross-sectional study design (*n* = 14), four were longitudinal  
242 and one study included both a longitudinal and a cross-sectional approach.<sup>37</sup> Most studies were  
243 conducted in the USA (*n* = 12), three were conducted in Canada<sup>38-40</sup> and the others in Germany,<sup>41</sup>  
244 UK,<sup>23</sup> Spain<sup>42</sup> and Israel.<sup>43</sup> Four studies were based on large population samples (*n* > 35,000)  
245 reflecting the cross-sectional study designs, one was a longitudinal study conducted in the USA with  
246 a small loss during follow-up.<sup>44</sup> Six studies relied on medium sample sizes (9,440 < *n* < 14,084) and  
247 the other nine on small sample sizes (*n* < 1,000). Four studies were focused on children (< 7 years  
248 old), six included only adolescents and the nine studies included both categories. Among the included  
249 studies several methodologies were used to quantify neighbourhood walkability. The most common  
250 method (*n* = 10) was the walkability index based on the approach developed by Frank et al. (2006).<sup>45</sup>

251 The original method by Frank et al. (2006) incorporated land use mix, street connectivity, net  
252 residential density, and retail floor area ratios, giving street connectivity twice the weight of the other  
253 three variables. Often studies used modified versions of the walkability index, i.e. giving street  
254 connectivity the same weight as the other variables, using destinations as proxy for land use, not  
255 accounting for the retail floor area ratio or including additional elements such as access to facilities  
256 and parks. The three main components of the walkability index (land use mix, street connectivity, net  
257 residential density) were often individually analysed. Two studies used the Walk Score,<sup>38,46</sup> a web-  
258 based tool (www.walkscore.com) which relies mainly on the distance to various amenities, but also  
259 includes population density and road metrics such as block length and intersection density. One study  
260 adopted a different approach by deriving a walkability index composed of land use mix, sidewalks,  
261 sidewalk buffers, sidewalk/street lighting, other sidewalk elements, traffic lights, pedestrian signal at  
262 traffic lights, marked crosswalks, pedestrian crossing and other signage and public transport.<sup>47</sup> Except  
263 one study which analysed percentage of body fat as outcome,<sup>48</sup> all studies used BMI-derived  
264 outcomes (BMI *z*-score, BMI trajectories, overweight and obesity prevalence), two of which analysed  
265 waist circumference<sup>40</sup> and skinfold thickness<sup>49</sup> as additional measures. Sex was always considered as  
266 covariate and age was missing only in one study.<sup>36</sup> Other covariates relating to individual, household  
267 and neighbourhood confounders were included in the models. The most used were race/ethnicity,  
268 parental education, and neighbourhood socioeconomic status. In general, studies were of good  
269 quality, with scores ranging from 8 to 11. The main factors which penalized some of the studies were  
270 small sample size, low representativeness of the general population and the use of self-reported  
271 data.<sup>44,47,50</sup>

272

### 273 *Summary of findings*

274 There was limited evidence that the walkability index is linked to childhood obesity ( $p = 0.28$ ), with  
275 only one out of ten studies finding significant associations<sup>41</sup> (Table 2). Two further studies showed

276 mixed results based on sex (effect on bodyweight status in girls, but not boys)<sup>39</sup> and geographic area  
277 (healthy BMI associated with higher levels of walkability in one of three studied cities).<sup>43</sup> The Walk  
278 score was associated with decreased BMI  $z$ -score in rural but not urban youths in one study,<sup>46</sup> but did  
279 not show any significant association in another study.<sup>38</sup> The walkability index based on street element  
280 characteristics, however, did identify a significant association with childhood obesity. With regards  
281 to individual walkability indicators, street intersection density was the most widely used indicator ( $n$   
282 = 7). Three studies found significant associations with childhood obesity,<sup>37,44,51</sup> one study found a  
283 weak positive association,<sup>52</sup> mixed results were found in two studies, with effects observed in girls  
284 but not boys<sup>39</sup> and one out of three studied cities.<sup>43</sup> The  $z$ -Test revealed strong evidence to support a  
285 link between street intersection density and obesity measures ( $p = 0.005$ ). Out of six studies analysing  
286 associations with population density, only one study found an effect of lower residential density being  
287 linked to higher BMI  $z$ -score<sup>37</sup> and one study found an effect only in girls. Overall, the evidence did  
288 not suggest a link between population density and childhood obesity ( $p = 0.23$ ). Land use mix was  
289 only analysed in four studies, with one study finding a significant association.

290

## 291 **Childhood obesity and accessibility and availability of parks and playgrounds**

### 292 *Study characteristics*

293 The dominant study design of the 28 included studies was cross-sectional ( $n = 20$ ), the others  
294 longitudinal ( $n = 8$ ).<sup>18,29,54-58</sup> One of the longitudinal studies conducted a quasi-experiment which  
295 considered a pre-park and post-park time frame and dividing the children into those who live near the  
296 park (the exposure group) and those who live further from the park (the control group) to examine  
297 how exposure to a newly built park translates to changes in BMI  $z$ -score over time.<sup>59</sup> Almost half of  
298 the studies were conducted in the USA ( $n = 13$ ), ten studies were conducted in Europe, four of which  
299 in the UK<sup>29,57,60,61</sup>, two studies from Germany<sup>62,63</sup> and Spain<sup>64,65</sup> and one from Netherlands<sup>18</sup> and  
300 Lithuania<sup>66</sup>. The sample sizes ranged from 93 to 41,283. Seven studies used small cohorts with less

301 than 1,000 subjects,<sup>48,49,53,54,65,67,68</sup> most studies used medium size cohorts ( $n = 15$ ) not exceeding  
302 7,000 participants, four studies included larger samples over 10,000 participants<sup>69-72</sup> and two studies  
303 included very large samples of around 40,000 subjects.<sup>52,73</sup> Five studies considered a wide age range  
304 up to 18 years. Seven studies included children under the age of 9 years,<sup>49,60,61,63,66,70,74</sup> and four  
305 studies included exclusively adolescents of at least 10 years.<sup>62,65,73,75</sup> Twelve studies included both  
306 children and adolescents with ages ranging from 4 to 18 years.

307 Most studies analysed park accessibility and availability based on children's place of residence, two  
308 studies focus on the school environment.<sup>63,72</sup> The definition of the sphere of influence was in 14  
309 studies based on circular or network buffers ranging from 100 to 3,000 meters in radius from the place  
310 of residence, one study which considered a ten miles (16,000 meters) radius.<sup>54</sup> Eight studies based  
311 their analysis on official administrative or statistical boundaries and three studies analysed distance  
312 from the nearest park, without defining a sphere of influence.<sup>52,59,65</sup> The remaining studies used  
313 neighbourhood area without further specifications on the delimitations. The most used exposure  
314 metric was the relative amount of park surface in the sphere of influence ( $n = 11$ ). Other studies  
315 quantified exposure through the dichotomous variable presence/absence of parks, the number or  
316 density of parks, and the distance from the nearest park. Four studies used the satellite-derived  
317 Normalized Difference Vegetation Index (NDVI) to quantify the greenness of the surrounding  
318 environment. The definition of park/greenspace was inconsistent across studies. Most studies  
319 identified areas intended as urban free-usable greenspace. Some studies identified specific features  
320 (e.g. children playgrounds), others used a broad approach (e.g. NDVI), considering the total amount  
321 of vegetation without distinct function.

322 The outcomes analysed were BMI  $z$ -score, BMI trajectories, BMI percentiles and weight status.  
323 Anthropometric measures were rarely used: waist circumference ( $n = 3$ ),<sup>56,76,77</sup> waist-to-height ratio  
324 ( $n = 1$ ),<sup>56</sup> sum of skinfold ( $n = 1$ )<sup>49</sup> and percentage body fat ( $n = 2$ ).<sup>48,77</sup>

325 The quality of the studies was either fair ( $n = 6$ ) or good ( $n = 22$ ). The main reasons for fair quality  
326 were small sample sizes, self-reported outcomes (height and weight) or study population scarcely  
327 representative of the population (see Table S5).

328

### 329 **Summary of findings**

330 Due to the great variability in exposure metrics, we synthesised findings across the following  
331 exposure categories: distance to the nearest park ( $n = 9$ ), park area ( $n = 10$ ), number of parks ( $n = 8$ )  
332 and presence/absence of parks ( $n = 5$ ). Only three studies analysed NDVI which was insufficient for  
333 meta-analysis according to our criteria (Table 2). The  $z$ -Test and related  $p$ -value suggest that there  
334 was insufficient evidence to support an association of distance to park and childhood obesity ( $p =$   
335  $0.170$ ). Out of the nine studies only one found a significant association.<sup>67</sup> Two studies concluded with  
336 mixed findings: one study found a significant association in boys of all ages and girls of high school  
337 age, but not in younger girls,<sup>52</sup> one study found an significant association in children living in urban  
338 areas but not those in rural areas.<sup>18</sup> The  $p$ -value suggested weak evidence of an association with  
339 percentage of park area ( $p = 0.014$ ). Three studies found significant associations, six studies found  
340 no statistically significant effects and two studies had mixed results, with effects only found in boys  
341 and older children. The  $p$ -value showed little evidence of an effect of number or density of parks on  
342 childhood obesity ( $p = 0.148$ ). One study found a significant association, five studies did not find  
343 significant associations, and one studies reported mixed results with effects only in girls.<sup>69</sup> The  
344 intervention study did, however, find an effect in the intervention group which could not be replicated  
345 in the control group.<sup>54</sup> We identified strong evidence on the presence of a park within the sphere of  
346 influence and childhood obesity ( $p < 0.001$ ). Out of the five studies, four studies found statistically  
347 significant effects. Results from the three studies which explored the effect of greenness via the NDVI  
348 suggest a potential association in the more proximal environment of less than 250 metres.<sup>18,64,66</sup> Three

349 studies specifically focussed on playgrounds and none of them found statistically significant  
350 associations.

351

## 352 **Discussion**

### 353 **Impact of built environment characteristics on childhood obesity**

354 We systematically reviewed the epidemiological evidence on the influence of four built environment  
355 characteristics on obesity outcomes in children: traffic noise, air pollution, neighbourhood walkability  
356 and accessibility and availability of parks and playgrounds. To our knowledge, this is the first  
357 systematic review on this topic that applied a systematic synthesis of findings to evaluate the strength  
358 of the available evidence.

359 Studies were generally of high quality, using objectively measured outcome and exposure measures  
360 and adjusting for relevant confounders. Some studies, however, had small sample sizes which were  
361 not necessarily representative of the overall population. Overall, 42% of studies used longitudinal  
362 data, however, the small number of longitudinal studies investigating effects of neighbourhood  
363 walkability and parks accessibility should be emphasized.

364 We found very strong evidence of association of BMI-derived obesity outcomes with NO<sub>2</sub>/NO<sub>x</sub>  
365 ( $p < 0.001$ ) and presence/absence of parks in the neighbourhood ( $p < 0.001$ ), strong evidence with  
366 intersection density ( $p < 0.01$ ) and some evidence with the amount of park area in the neighbourhood  
367 ( $p < 0.05$ ). There was little evidence of an effect on childhood obesity in relation to PM<sub>2.5</sub>, PM<sub>10</sub>,  
368 walkability index, residential density, distance to the nearest park, number of parks and access to  
369 playgrounds.

370 Air pollution has been shown to decrease birth weight<sup>78</sup> and might independently affect weight in  
371 childhood through epigenetic and behavioural adaptation. Some hypotheses on the mechanism



372 involved in the exposure both during pregnancy and childhood were highlighted in previous  
373 publications: prenatal growth restrictions can lead to growth spurts in early childhood with  
374 implications on increased weight into later childhood and adolescence;<sup>79</sup> heavy traffic roads, an  
375 important sources of air pollution, might deter active transport and reduce physical activity.<sup>80</sup> Our  
376 findings point towards this direction with traffic-related air pollutants NO<sub>2</sub> and NO<sub>x</sub> having a strong  
377 impact on increased weight in childhood, but not particulate matter (PM<sub>2.5</sub> and PM<sub>10</sub>) which is driven  
378 to a lesser degree by local traffic.<sup>18</sup> Another explanation could be the biochemical mechanism which  
379 emphasizes the role of NO<sub>2</sub> as active oxidant involved in many physiological pathways in the human  
380 body which might impact consequently the onset of obesity.<sup>81</sup>

381 Despite evidence suggesting a link between walkability and physical activity<sup>82</sup>, we found little  
382 evidence of neighbourhood walkability decreasing BMI-derived outcomes. Intersection density is the  
383 only indicator of walkability which showed strong evidence of a negative association with childhood  
384 obesity. The central role of this measure in the walkability index has already been highlighted in the  
385 original equation by Frank et al. (2006), which gave street connectivity twice the weight of the other  
386 variables. Given the same source, road traffic, future studies should explore the effect of collinearity  
387 between the walkability components and other traffic-related factors such as traffic noise and air  
388 pollution. Studies on walkability were mainly conducted in the United States with only small number  
389 of studies from Europe North American cities have a different urban structure compared to European  
390 cities and results might not be directly comparable and transferable. This should be explored further  
391 in future studies.

392 We also found strong evidence for the presence (or accessibility) of parks with decreased prevalence  
393 of childhood obesity, while studies focusing on playgrounds did not find significant associations.  
394 This is supported by findings from Bird et al. (2016) who concluded that parks that emphasize  
395 unstructured activities (i.e. with few team sport installations) were associated with lower percentage  
396 of truncal fat among children at risk of being with obesity.

397

### 398 **Methodological considerations**

399 Some of the included studies investigated more than one built environment characteristics. Several  
400 studies explored walkability and parks.<sup>29,48,49,52,70,84</sup> Among the studies which considered walkability  
401 and greenspaces, walkability was not statistically significant, except intersection density in boys in  
402 one of the studies<sup>48</sup>, and greenspace was at least partially associated with weight outcomes in all  
403 studies. No multi-exposure interactions were evaluated in these studies, except for a Pearson  
404 correlation coefficient between intersection density and park space, which did not show collinearity<sup>48</sup>.  
405 Overall, we found a lack of studies which explore the interaction between multiple exposures on  
406 childhood obesity. Bloemsa (2019)<sup>18</sup> investigated the combined effect of noise, air pollution and  
407 park accessibility. They found that the association of NO<sub>2</sub> with overweight remained after adjustment  
408 for noise and greenspace, but the associations between greenspace and overweight weakened  
409 substantially after adjustment for NO<sub>2</sub>, indicating that NO<sub>2</sub> is driving the relationship. To better  
410 understand the complex relationship of multiple built environment characteristics on childhood  
411 obesity more evidence is required.

412 Our review highlighted a strong presence of effect modifiers. Sex was the most studied effect modifier  
413 but there was no consistency across studies. Two studies reported an increased effect in boys for the  
414 association between air pollution exposure and BMI<sup>31,34</sup>, but one of the studies found also an opposite  
415 effect considering waist-to-hip ratio as anthropometric measure, which was statistically significant  
416 only in girls<sup>31</sup>. Walkability and intersection density were found to be associated with body weight  
417 status in girls but not in boys in one of the studies<sup>39</sup>, but in another study, a high level of street  
418 connectivity was related to lower percentage of body fat only in boys<sup>48</sup>. The association between park  
419 accessibility and obesity was gender-dependent in five studies, of which three showed more  
420 significant effects on boys<sup>52,55,56</sup> and two on girls<sup>48,69</sup>. Overall, sex affected the results in nine studies,

421 concluding with an increased effect in boys in five studies, in girls in three studies and with opposite  
422 effects depending on the considered anthropometric measure in one of the studies. Age was another  
423 common effect modifier, showing differential results in five studies. In one study the exposure to road  
424 traffic noise was associated with increased BMI from school age to adolescence, but not at earlier  
425 ages, the relation increased in the older age groups.<sup>20</sup> Age also modified the association between  
426 greenspace exposure and BMI in four studies (two of them were based on the same sample), always  
427 with increased effects in older children<sup>29, 52, 55, 56</sup>. Another effect modifier was urbanisation, with one  
428 study finding a negative association between walk score and BMI z-score for youths in rural settings  
429 and a positive association among urban youths<sup>46</sup>, whereas in another study children living in a urban  
430 area had a negative association of the distance to the nearest park with weight status and no  
431 association for those living in rural areas<sup>18</sup>. No studies analysed effect modification by socioeconomic  
432 status, an important omission which could potentially highlight important pathways to health  
433 inequalities.

434

435 This systematic review assessed the strength of the evidence and identified the role of different  
436 elements of the built environment on childhood obesity, consolidated associations and indicating  
437 areas in need of further evidence. Our review has some limitations. Due to the observational nature  
438 of included studies, no direct causal relationships can be inferred from the results. The absence of  
439 sample size restriction in the selection of studies allowed the inclusion of very small cohorts with  
440 results potentially not being transferable beyond the specific setting. The fact that some of the studies  
441 used self-reported outcomes (weight and height) could also influence the quality of the results due to  
442 the introduction of error and bias in the outcome measures. Finally, it was not possible to conduct a  
443 meta-analysis due to the large heterogeneity in study results which could have influenced the validity  
444 of our findings. Previous reviews on the effect of the physical and built environment on childhood

445 obesity, however, expressed the results only through descriptive synthesis or narrative review. The  
446 use of the z-Test is a strength which allows us to assess and quantify the strength of the associations.

447

## 448 **Conclusion**

449 In summary, we found strong evidence for an association of traffic-related air pollution (nitrogen  
450 dioxide and nitrogen oxides exposure;  $p < 0.001$ ) and built environment characteristics supportive of  
451 walking (street intersection density;  $p < 0.01$  and access to parks;  $p < 0.001$ ) with childhood obesity.  
452 Studies on traffic noise had mixed results and were too few to be included in the z-Test analysis.  
453 Future studies should consider the interactions between different environmental exposures and verify  
454 the role of age and sex as an effect modifier.

455

456

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