1 TITLE

2 Climatic drivers of (changes in) bat migration phenology at Bracken Cave (USA)

3 RUNNING TITLE

4 Climatic drivers of bat migration phenology

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13 AUTHOR CONTRIBUTIONS

- 14 B.H. conceptualized the study. P.S. and C.W. contributed the phenology dataset. B.H. designed and
- 15 conducted the analyses. B.H. wrote the paper, with contributions of P.S., C.W., F.L. and S.B.

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21 ABSTRACT

Climate change is drastically changing the timing of biological events across the globe. Changes 22 23 in the phenology of seasonal migrations between the breeding and wintering grounds have been observed across biological taxa, including birds, mammals, and insects. For birds, strong links 24 25 have been shown between changes in migration phenology and changes in weather conditions at the wintering, stopover, and breeding areas. For other animal taxa, the current understanding 26 27 of, and evidence for, climate (change) influences on migration still remains rather limited, 28 mainly due to the lack of long-term phenology datasets. Bracken Cave in Texas (USA) holds one 29 of the largest bat colonies of the world. Using weather radar data, a unique 23-year (1995-2017) 30 long time series was recently produced of the spring and autumn migration phenology of 31 Brazilian free-tailed bats (Tadarida brasiliensis) at Bracken Cave. Here, we analyse these 32 migration phenology time series in combination with gridded temperature, precipitation, and 33 wind data across Mexico and southern USA, to identify the climatic drivers of (changes in) bat 34 migration phenology. Perhaps surprisingly, our extensive spatiotemporal search did not find 35 temperature to influence either spring or autumn migration. Instead, spring migration 36 phenology seems to be predominantly driven by wind conditions at likely wintering or spring 37 stopover areas during the migration period. Autumn migration phenology on the other hand, 38 seems to be dominated by precipitation to the east and north-east of Bracken Cave. Long-term 39 changes towards more frequent migration-favourable wind conditions have, furthermore, 40 allowed spring migration to occur 16 days earlier. Our results illustrate how some of the 41 remaining knowledge gaps on the influence of climate (change) on bat migration and abundance 42 can be addressed using weather radar analyses.

43 INTRODUCTION

44 Every year, many billions of animals migrate across the globe in search of conditions that increase 45 survival and reproductive success. In doing so, they contribute to the ecosystem functioning and 46 provide invaluable ecosystem services, such as pest control and pollination, at the many locations 47 along their migratory route (Bauer et al., 2019; Bauer & Hoye, 2014; Bowlin et al., 2010). Many 48 populations of migratory animals have been steeply declining (Vickery et al., 2014; Wilcove & 49 Wikelski, 2008), compromising their role in the ecosystems along their way. Knowledge on the timing 50 of migration is a crucial factor in determining the relative impacts of migratory animals on those 51 ecosystems, but also for implementing efficient conservation actions (Bauer & Hoye, 2014; Dechmann 52 et al., 2017; Wilcove & Wikelski, 2008). Over the past decades, many species have changed their 53 migratory timing in response to climate change (Thackeray et al., 2016). The direction and magnitude 54 of these changes, however, vary across geographic locations and taxa (Charmantier & Gienapp, 2014; 55 Chmura et al., 2019; Hurlbert & Liang, 2012).

- 56 To understand why species and populations differ in their phenological response to climate change, we
- 57 first need to identify the environmental drivers of migration timing (Haest et al., 2018b, 2019; Shaw,
- 58 2016). Such an endeavour would not only advance our fundamental understanding of animal
- 59 migration, but, more importantly, allow us to predict the consequences of ongoing and future climate
- 60 change (Bowlin et al., 2010; Pettit & O'Keefe, 2017), identify the species most at risk (Bauer et al.,
- 61 2011; Hurlbert & Liang, 2012; Wilcove & Wikelski, 2008), and support their conservation. While
- 62 drivers of migration have been extensively studied in some taxa, e.g., birds, other taxa, e.g., bats and
- 63 insects, are remarkably understudied (Bauer et al., 2011; Liechti & McGuire, 2017; Moussy et al.,
- 64 2013; Popa-Lisseanu & Voigt, 2009; Wilcove & Wikelski, 2008).
- 65 The majority of bat species are considered sedentary (Hutterer et al., 2005). For many temperate bat 66 species, this includes a winter hibernation period close to the summer habitat, even at higher latitudes (Fleming, 2019). With increasing latitude, however, an increasing number of species travel 67 68 considerable distances (i.e. some hundreds up to perhaps three thousand kilometers) between winter 69 and summer habitats. Particularly for bat species living in highly seasonal (temperate) environments, 70 migration is an essential part of their ecology (Fleming & Eby, 2003; Hutterer et al., 2005). Because of 71 their cryptic nocturnal activity patterns and often secretive roosting, bat behaviour (including 72 migration) has remained notoriously difficult to study (Fleming, 2019; Krauel & McCracken, 2013; Liechti & McGuire, 2017; Smith & McWilliams, 2016; Weller et al., 2016). In Europe and North 73 74 America, some knowledge on migratory and other movement patterns has been gathered over the past 75 century using ringing (or banding) (Ellison, 2008; Hutterer et al., 2005; Petersons, 2004). Further 76 insights into rough spatiotemporal movement patterns have also been gained from museum specimens 77 (Cryan, 2003), stable isotope analyses (Britzke et al., 2009; Lehnert et al., 2018), acoustics (Rydell et 78 al., 2014; Smith & McWilliams, 2016), and genetic analyses (Russell et al., 2005; Russell & 79 McCracken, 2006). Gaining detailed insights into the migratory behaviour of individuals or 80 populations has, however, proven difficult due to the lack of appropriate monitoring tools or 81 technology to either: (a) mark and track individual bats (Holland & Wikelski, 2009; Krauel & 82 McCracken, 2013; Roby et al., 2019); or (b) study the spatiotemporal dynamics of entire populations. 83 Many bat species are rather small (and light-weight), limiting possibilities for long-term tracking 84 devices (Moussy et al., 2013; but see Weller et al., 2016). Monitoring (seasonal) bat abundance at a 85 single location such as large roosting colonies (e.g. maternity caves) is also far from straightforward (McCracken, 2003). In fact, studies taking advantage of technological (and algorithmic) developments, 86 87 e.g. thermal imaging, have indicated that quantitative estimates in historical abundance census records 88 are probably highly questionable (Hristov et al., 2010). Over the past decade, a handful of radio-89 tagging studies have produced the first insights into relationships between weather and bat migration 90 phenology that go beyond the anecdotal (Dechmann et al., 2017; Jonasson & Guglielmo, 2019; 91 McGuire et al., 2012; Roby et al., 2019). Reliable long-term datasets on bat migration phenology are, 92 however, extremely rare (Stepanian & Wainwright, 2018; but see Pettit & O'Keefe, 2017 for a notable

- exception). As such, potential effects of climate change on bat migration phenology have hithertoremained largely speculative.
- 95 Using weather radar data, a unique 23-year (1995-2017) continuous time series was recently produced of nightly population estimates of Brazilian free-tailed bats at Bracken Cave in Texas (USA) 96 97 (Stepanian & Wainwright, 2018). Brazilian free-tailed bats are one of the most abundant (Davis et al., 98 1962) and probably best-studied bat species (Russell & McCracken, 2006). Long-term, often severe, 99 declines in their abundance have repeatedly been reported (McCracken et al., 1994), although the 100 magnitude of declines may have been overestimated due to inaccurate historical census data (Hristov 101 et al., 2010). While much research has been dedicated to Brazilian free-tailed bats, many questions 102 remain, especially pertaining to their variation in abundance and timing of life-history activities 103 (Stepanian & Wainwright, 2018). Every summer, millions of female Brazilian free-tailed bats gather 104 in large maternity colonies across the USA, with Bracken Cave being one of the biggest. Very little is 105 known, however, on the exact locations and the ecology of Brazilian free-tailed bats during winter or 106 migration periods, with the proportion of the North American population accounted for in suspected 107 winter roosts estimated at less than 1% (López-González & Best, 2006) or perhaps up to 5% 108 (Wiederholt et al., 2013).
- Here, we use yearly spring and autumn migration phenology at Bracken Cave over the period 1995-2017, derived from the nightly abundance estimated by weather radar (Stepanian & Wainwright, 2018), in combination with gridded temperature, precipitation and wind data to first identify which weather at which location and over which time window is most likely influencing inter-annual phenology of Brazilian free-tailed bats. Subsequently, we show to what extent these drivers may have caused the observed long-term temporal trends in migration phenology.

115 MATERIALS AND METHODS

116 Spring and autumn migration phenology at Bracken Cave

117 Weather surveillance radars regularly detect large numbers of Brazilian free-tailed bats as they emerge 118 from their roosts at dusk and take flight into the airspace (Frick et al., 2012; Stepanian & Wainwright, 119 2018). A previous study used radar observations of the dusk exodus flights of the bat colony 120 inhabiting Bracken Cave (Texas, USA) to produce nightly colony population estimates over a 121 continuous period spanning 20 March 1995 through 30 November 2017 (Stepanian & Wainwright, 122 2018). These high temporal resolution population estimates revealed cyclic seasonal changes in the 123 Bracken Cave colony size that are indicative of migration phenology, and were used to extract quantitative annual phenophases corresponding with spring and autumn migration timing (Stepanian 124 125 & Wainwright, 2018). The resulting dataset (Stepanian et al., 2020) provides annual dates on which 126 50% of the mean summer population is first detected at the cave (i.e., a quantitative metric for the onset of spring migratory arrivals to the cave), as well as the final dates on which 50% of the mean 127

- 128 summer population is detected (i.e., a quantitative metric for autumn dispersal out of the cave). These
- spring and autumn migration timings were obtained for 23 consecutive years (1995-2017; Figure 1).
- 130 Spring migration showed a clear advancement over the study period (-0.64 days/year, standard error =
- 131 0.29, p-value t-test = 0.04), while autumn migration timing remained rather stable (-0.24 days/year,



132 standard error = 0.49, p-value t-test = 0.63).

Figure 1 Spring (a) and autumn (b) migration timing of the bats at Bracken Cave over the period 1995-2017 as determined by Stepanian & Wainwright (2018). Lines represent estimated linear temporal trends (and grey areas the 95% confidence intervals). Spring migration (a) temporal slope coefficient = -0.64, standard error = 0.29, pvalue (t-test) = 0.04; Autumn migration (b) temporal slope coefficient = -0.24, standard error = 0.49, p-value (ttest) = 0.63.

138 <u>Weather data</u>

139 We used the R package RNCEP (Kemp, van Loon, et al., 2012) to gather National Center for 140 Environmental Prediction (NCEP) Reanalysis I data (Kalnay et al., 1996; Kanamitsu et al., 2002) on temperature, precipitation, and wind in an area from about 119° to 83° W, and 11° to 38° N (see maps 141 142 in Appendix S1), which includes the entire countries of Honduras, El Salvador, Guatemala, Belize, and Mexico, and twenty south-western states of the USA partly or entirely. Ocean grid cells were 143 144 masked from the analysis. While bats have been observed to migrate offshore (Hüppop & Hill, 2016), 145 this is unlikely for bats passing through Bracken Cave (Wiederholt et al., 2013). The spatial resolution of a grid cell was 1.905° (latitude) x 1.875° (longitude) for temperature/precipitation and 2.5° x 2.5° 146 for wind. At the northern edge of the analysed area, these resolutions correspond to 212 (latitude) x 147 148 164 (longitude) km, and 278 (latitude) x 217 (longitude) km, respectively. At the southern range of the 149 studied area, the areas covered by the grid cells are somewhat larger, i.e. 212 (latitude) x 204 150 (longitude) km for temperature/precipitation and 278 (latitude) x 272 (longitude) km for wind. We 151 derived eight variables from the NCEP data: mean daily air temperature at 2 meters above ground 152 level, daily accumulated precipitation at surface, daily mean wind direction at the 925, 850, and 700-153 hPa pressure levels, and daily mean wind assistance at the same three pressure levels (Appendix S1). 154 The 925, 850, and 700-hPa atmospheric pressure levels roughly correspond to altitudes of 750, 1500, 155 and 3000 m above-sea-level (asl), depending on geographic location and environmental conditions. In 156 the time window analysis, the wind directions were used to calculate the number of days that wind at that location was in the direction towards, or coming from, Bracken Cave (at each of the atmospheric 157 pressure levels), hence resulting in eleven variables in total being analysed. We did this by counting 158 159 every day with a mean wind direction between -45° and 45° of the angle between Bracken Cave and 160 the centre of the focal grid cell. Depending on the location of the grid cell relative to Bracken Cave, 161 these winds were then interpreted as tail- or headwinds. In terms wind direction effects, we thus 162 analysed two different potential hypotheses separately, i.e. days with headwinds delay migration and 163 tailwinds advance migration. Additionally, we also analysed wind assistance to check for a joint effect 164 of wind direction and wind speed. We calculated the wind assistance at each of the three atmospheric 165 pressure levels with the RNCEP package (Kemp, van Loon, et al., 2012) using the "M.Groundspeed" 166 equation (Kemp, Shamoun-Baranes, et al., 2012). While various approaches exist to calculate wind 167 assistance, we deemed the "M.Groundspeed" approach most appropriate because Brazilian free-tailed 168 bats have been observed (albeit during foraging flights) to maintain a relatively constant groundspeed 169 irrespective of wind conditions (McCracken et al., 2016).

Determining the most likely combinations of "weather variable – location – time window" that influence migration phenology

To determine the combinations of "weather variable - location - time window" that most likely 172 influence bat spring and autumn migration phenology, we used a method that has recently been shown 173 174 effective on similar time series of bird migration phenology (Haest et al., 2018b, 2019, 2020b). The 175 method consists of two consecutive analyses: (1) a broad search for all combinations of "weather 176 variable – location – time window" that show a relationship with the migration phenology that is 177 unlikely to be due to chance only (but might still be a false positive due to spatiotemporal correlation 178 within and between the weather variables); and (2) a set of refined analyses to narrow down the 179 candidates from the first step to the most likely influences.

180 More specifically, in the first step, a time-window analysis is performed for each weather variable on 181 each grid cell to search for a (continuous) time window of any length that correlates better with the migration phenology data than is expected by chance. The time window search is performed using the 182 183 climwin R package (Bailey & van de Pol, 2016; van de Pol et al., 2016). To determine the best-184 performing time window, the method compares AICc model values for each time window to a base 185 reference model. We used a base reference model consisting of a linear temporal trend (i.e. a model 186 with year as the independent variable) to avoid spurious correlations due to shared temporal trends 187 (Haest et al., 2018a; Noriega & Ventosa-Santaulària, 2007). To estimate the probability of obtaining a

188 similarly performing 'best' time window due to chance alone, the time window search is repeated on 189 randomizations of the weather data. Ideally, at least a hundred randomizations are run to approximate 190 the \triangle AICc distribution of best-performing time windows obtainable by chance alone, which, however, quickly becomes highly resource-intensive. Therefore, we used the alternative probability statistic P_{c} 191 192 of the *climwin* package, which uses as little as five randomizations to estimate the probability of obtaining a similarly performing best time window due to chance alone (Bailey & van de Pol, 2016; 193 194 van de Pol et al., 2016). The P_c statistic ranges from 0 to 1, with values closer to 0 indicating a higher 195 probability that the relationship is not due to chance. For a sample size of 23 (as the number of years 196 in our study) and a cut-off value of $P_c < 0.5$, Bailey & van de Pol (2016) estimated the rate of false 197 positives and negatives to be about 0.14 and 0.10, respectively. These rates, however, apply to simulated relationships ranging from 0.2 to 0.8 in strength (R²). The stronger the simulated 198 relationship with the response variable, the lower these rates become. To lower the probabilities of 199 false positives in our analysis, we set the P_c threshold to 0.3 (instead of 0.5) (see Bailey & van de Pol, 200 201 2016). For ease of reference, all settings and decision rules of the time windows analyses are 202 summarized in Table S1 and S2. From the output of the time window analyses, we created $\Delta AICc$, R², 203 and regression slope maps for each weather variable, representing the values of the best model for 204 each grid cell, excluding the grid cells for which $P_c > 0.3$ (Figure S1 to S22). In many cases, 205 neighbouring cells in these maps had similar time windows and $\Delta AICc$ for the best-performing model 206 due to spatiotemporal autocorrelation in the weather variables. As a result, there are spatial gradients present in these maps. We considered the cells with regional AAICc maxima as the most 207 208 representative of the potential relation between the weather variable for that (larger, correlated) area 209 and the migration phenology at Bracken Cave. This first step of the analysis resulted in 43 and 56 210 potential "weather variable - location - time window" influences for spring and autumn migration, 211 respectively (Figure S1 to S22).

212 In the second part of the analysis, we extracted those candidate signals from the potential influences 213 that are most likely to have the strongest influence on the migration phenology. To do so, we first use 214 a sequence of different variable filtering approaches to remove the (relatively) least likely candidates, followed by an ensemble of (relative) "variable importance" methods to determine the most likely 215 216 final "weather variable – location – time window" influences. Throughout this analysis, we no longer 217 include the temporal trend, i.e. the "year" variable, as spurious correlations due to shared trends were already excluded in the first part of the analysis. Instead, we now assess to what extent the relationship 218 219 between the identified potential effects of weather and the observed migration phenology still holds 220 when not accounting for temporal trends. In a first filtering, we compared AICc values for models 221 with the identified effects of weather as the only independent variable to an intercept-only model. We 222 removed two potential influences for spring migration (and none for autumn migration), because the 223 AICc difference with the intercept-only model was less than two units. Subsequently, we checked for 224 collinearity between the remaining candidate weather effects. For spring and autumn migration, we

- removed 21 and 28 candidates, respectively, because they had a Pearson correlation > 0.7 (Dormann et
- al., 2013) with another candidate that had a larger Δ AICc with an intercept-only model. Next, we used
- the *boruta* method to further remove 6 and 14 candidate influences for spring and autumn migration,
- respectively, that had a variable importance that is likely to be obtained by chance (Kursa & Rudnicki,
- 229 2010). Finally, we used an ensemble of "variable importance" methods (Burnham & Anderson, 2002;
- 230 Grömping, 2006, 2015; Kursa & Rudnicki, 2010) to identify the most likely "weather variable -
- 231 location time window" influences on bat spring and autumn migration phenology (Figure 2; Table
- 232 S3 and S4).

233 <u>Climatic contributions to trends in phenology</u>

A weather variable that affects inter-annual variability in migration timing can only result in a temporal trend in migration phenology if it also shows a temporal trend. The overall contribution of a weather variable to the observed temporal trend in migration phenology can be calculated using the chain rule (Haest et al., 2019, 2020b; McLean et al., 2018):

climate contributions to
trend in migration phenology
$$= \sum_{i=1}^{n} \left(\frac{\partial MPD}{\partial Climate_{i}} \times \frac{dClimate_{i}}{dTime} \right),$$

with n being the total number of influencing weather variables for a given species; $\partial MPD/\partial Climate_i$ the regression coefficients of a multiple linear regression between (spring or autumn) migration passage date and all of the identified final weather variables; and $dClimate_i/dTime$ the regression coefficient of a simple linear regression between the respective weather variable and time, i.e. years. Standard errors were calculated following error propagation rules for multiplication (Taylor, 1997). Note that this approach ignores by definition any other (e.g. non-climatic) factors that might possibly affect changes in migration phenology over time.

245 **RESULTS**

246 <u>Identified</u>, weather variable – location – time window" influences

We found that during spring migration tailwinds in central North Mexico and West Texas to the west of Bracken Cave (Figure 2) explain about 84% of the observed variance in spring migration phenology (Table 1). The tailwinds were negatively related to spring migration phenology (Table 2). Thus, more days with tailwinds in central North Mexico and West Texas result in earlier spring passage or arrival at Bracken Cave. For each (additional) day of tailwind during spring migration, bats passed on average about 2 to 2.2 days earlier at Bracken Cave. The central North Mexico tailwind occurred at the 850hPa pressure level (i.e. roughly 1500 m asl), while the western Texas tailwind was at the 925-hPa level

254 (i.e. roughly 750 m asl).

- For autumn migration, summer and autumn precipitation to the east-northeast of Bracken Cave (Figure 2) explained about 83% of the observed variance. The two precipitation signals had contrasting effects (Figure 4 and Table 2), i.e. precipitation in Northeast Texas / Southeast Oklahoma / Arkansas / West Louisiana results in later average autumn migration, while precipitation in eastern Texas and Louisiana results in earlier autumn migration. Additionally, we found strong statistical support for an influence of headwind at the 700-hPa level (i.e. approximately 3000 m asl) in north-western Mexico (Sonora / Baja California / Chihuahua) on autumn migration phenology.
- 262**Table 1** Explained variance (adjusted R^2) and predictive performance (predictive R^2) for spring and autumn263migration phenology using a linear model with the final identified weather signals (see Figure 1), but not the264temporal trend. Adjusted R^2 is defined as in Miles (2005). Predictive R^2 was calculated as leave-one-year-out.

Season	Weather variables	Adjusted R ²	Predictive R ²
Spring	Both wind variables	0.84	0.81
Autumn	Both precipitation + the wind variable	0.88	0.86
Autumn	Both precipitation variables	0.83	0.79

265



Figure 2 Location and timing of the identified most important weather variables that are likely to influence (a) spring and (b) autumn migration timing at Bracken Cave. The timelines in each plot represent the period of the

single best time window (left) and of the medians for the time window opening and closing of the 95%

- 269 confidence interval of all time windows (right). P: precipitation; HW: headwind; TW: tailwind. The location of
- 270 Bracken Cave is marked with a star. The white background triangles in the time window subfigures represent the
- 271 migration period at Bracken Cave, i.e. the period between the earliest and latest estimated average migration
- time at Bracken Cave over the entire study period (see Stepanian and Wainwright, 2018).

Although our analyses included a large area of potential effects of weather - from Nicaragua in the south, Nevada in the northwest, and Kentucky to the northeast (Appendix S1) - four of the five identified influences lie within distances up to 1100 km around Bracken Cave, i.e. well within the

- 276 currently known (maximum) migration distance of Brazilian free-tailed bats (Cockrum, 1969; Glass,
- 277 1982). The times during which the weather variables influence migration span the period from about
- two months prior to the earliest and up to the latest estimated mean migration time, in both spring and
- autumn. Due to the inherent temporal autocorrelation of weather variables, the exact timing of the time
- 280 windows remains somewhat uncertain (Figure 2).
- 281 Predictive R² (calculated using leave-one-year-out) were not much lower (0.03 for spring and 0.04 for
- autumn) than adjusted R^2 (Miles, 2005), indicating that the final identified effects of weather are
- robust, i.e. the models do not suffer from overfitting. We did not find any support for influences of
- 284 either temperature or wind assistance on spring or autumn migration.
- 285 **Table 2** Effect sizes and standard errors (SE) for the identified most influential weather variables, estimated
- using the full linear model consisting of the migration phenology as the dependent response variable and all of the identified weather variables (but not the temporal trend) as the independent variables.

Season	Weather variable	Location	Effect Size	SE
Spring	tailwind 925-hPa	North Coahuila de Zaragoza / East New Mexico / West Texas	-2.20	0.45
	tailwind 850-hPa	Central North Mexico (Chihuahua)	-2.06	0.36
Autumn	precipitation	Northeast Texas / Southeast Oklahoma / Arkansas / West Louisiana	0.06	0.01
	precipitation	East Texas / Louisiana	-0.41	0.07
	headwind 700-hPa	Sonora / Baja California / Chihuahua	-2.46	0.75

288 <u>Contributions of climatic influences to temporal trends in migration phenology</u>

The wind conditions in north Mexico and western Texas did not only have a strong effect on interannual variability in spring migration phenology, they also showed temporal trends over the 1995-2017 study period (Figure 3, Table S5). More days with tailwinds at both locations have led to a systematically earlier mean spring migration at Bracken Cave, with advancements of approximately 5 and 12 days over the period 1995-2017 (Figure 3 and Table S5).

A decrease in the (positively associated) summer and autumn precipitation in Northeast Texas, Southeast Oklahoma, Arkansas, and West Louisiana seems to have pushed towards a slight advancement in autumn migration timing at Bracken Cave, while lower spring precipitation in East Texas and Louisiana seems to have resulted in a slight delay (Figure 4, Table S5). Please note, however, that autumn migration timing at Bracken Cave did not show a significant temporal trend over the study period (Stepanian & Wainwright, 2018; Figure 1).



Figure 3 Overview of (a) the effect of the identified weather signals on spring migration phenology of Brazilian free-tailed bats at Bracken Cave; (b) the temporal trends in these weather variables; and (c) the resulting contribution to the overall temporal trends in spring migration phenology. In (c), the dashed lines represent the calculated trend contributions using the chain rule (see Table S5), and the full lines the overall observed trends in spring migration phenology (see Figure 1). Dark grey ribbons in (c) are the 95% confidence interval for the estimated trend contribution, and light grey for the overall temporal trend.



Figure 4 Overview of (a) the effect of the identified weather signals on autumn migration phenology of Brazilian free-tailed bats at Bracken Cave; (b) the temporal trends in these weather variables; and (c) the resulting contribution to the overall temporal trends in autumn migration phenology. In (c), the dashed lines represent the calculated trend contributions using the chain rule (see Table S5), and the full lines the overall observed trends in autumn migration phenology (see Figure 1). Dark grey ribbons in (c) are the 95% confidence interval for the estimated trend contribution, and light grey for the overall temporal trend.

312 **DISCUSSION**

313 Migration and population structure of Brazilian free-tailed bats at Bracken Cave

Our results suggest that in spring Brazilian free-tailed bats that use Bracken Cave as a spring stopover or maternity location move eastward. These bats seem to originate from central North Mexico (Figure

316 2) and western Texas, and these areas could be wintering grounds, spring stopover, mating locations,

317 or a mixture of all of these (see also Wiederholt et al., 2013). Brazilian free-tailed bats mate in spring. 318 Where bats exactly congregate to mate has, however, remained largely elusive (Keeley & Keeley, 319 2004; Svoboda et al., 1985). It has been suggested that they gather in transitional roosts in Mexico 320 shortly before or during northward migration (McCracken et al., 1994; Russell & McCracken, 2006; 321 Wilkins, 1989), although copulation has also been observed during March and April in central Texas (Keeley & Keeley, 2004). The combined location and timing of our spring wind influences, i.e. 322 323 December to April, suggests that these might be the locations (and times) where mating (mainly) 324 occurs. Individuals from many different maternity roosts (ranging from west to east USA) perhaps

mingle at these locations during spring migration, which may also explain the observed lack of genetic
structuring in the North American populations (McCracken et al., 1994; Russell & McCracken, 2006).

327 Surprisingly, we did not find weather at Bracken Cave itself to influence autumn migration phenology. 328 Instead, migration phenology at Bracken Cave was dominated by precipitation to the east and 329 northeast of Bracken Cave (Figure 2). These areas might represent maternity colonies of bats that 330 stopover in Bracken Cave, post-parturition feeding grounds of adult females and juveniles (of perhaps 331 also bats from the Bracken Cave maternity colony), or a mixture of these (see also McCracken et al., 332 2018). Alternatively, the weather at these locations could also affect migratory decisions at Bracken 333 Cave indirectly, e.g. through insect migrations originating from these areas (Krauel et al., 2015, 2018). 334 Adult female Brazilian free-tailed bats (likely) leave their maternity colonies (such as Bracken Cave) 335 after they weaned their young (Davis et al., 1962; Hristov et al., 2010; Krauel & McCracken, 2013), 336 and often even move northwards (Glass, 1982; Russell & McCracken, 2006). Autumn movements in 337 the direction of and to, as well as spring stopover at Bracken Cave by bats from (maternity) colonies in 338 Oklahoma have also been observed before from banding recoveries (Glass, 1959, 1982).

339 The contrasting directions of the precipitation effects on autumn migration phenology may seem 340 contradictory (Figure 2 and Table 2). However, if, for example, precipitation affects migration 341 phenology through insect abundance, differences in insect abundance at the two locations may affect 342 bat survival and result in different abundances (Frick et al., 2010) or reproductive rates (Adams, 2010) 343 at each location. If the populations at each location tend to migrate through Bracken Cave at different 344 times, these differences in abundances between the two populations may, hence, change the overall 345 observed autumn migration phenology at Bracken Cave in different directions. Alternatively, 346 precipitation in spring and autumn may have contrasting effects, i.e. higher precipitation in autumn 347 leads to later autumn migration, whereas higher spring precipitation leads to earlier autumn migration.

348 The influence of headwind in northwest Mexico (Figure 2) seems to challenge the current knowledge

349 on migration patterns of Brazilian free-tailed bats (but see ring recovery maps in Villa & Cockrum,

350 1962; and Wiederholt et al., 2013). Although we have tailored our approach towards avoiding false

351 positives, we can not exclude them entirely (Bailey & van de Pol, 2016). Nevertheless, such long-

352 distance West-East movements have, albeit perhaps more rarely, been observed in banding studies,

353 including an individual that was banded in Nevada and recovered in Kansas (Baker, 1978; Cockrum,

1969; Fleming & Eby, 2003; Svoboda et al., 1985). Additionally, the ecology of Brazilian free-tailed bats during winter is hardly known, including exact locations and time periods. While their summer ecology is generally better known, much of the knowledge originates from females and first-year juveniles while the ecology of males is much less known.

Bats that use Bracken Cave in spring and autumn do not necessarily originate from the same populations, i.e. some populations may use it as a spring transient roost, and others in autumn. While winter site fidelity has been suggested in other (hibernating) migratory bat species (Lehnert et al., 2018), migratory routes and stopover routes need not be identical in spring and autumn. In general, our results suggest more of an East-West, instead of North-South, migratory movement of Brazilian freetailed bats passing at Bracken Cave (Figure 2). Interestingly, a similar observation was recently made, also contradicting expectation, for Indiana bats (*Myotis sodalis*) (Roby et al., 2019).

365 Drivers of bat migration phenology

366 The importance of climate and atmospheric conditions for bat migration ecology has long been recognized (Liechti & McGuire, 2017). Until recently, however, the proposed influences of weather on 367 368 bat migration were almost exclusively based on anecdotal observations of weather conditions during 369 (migratory) flight activity (e.g. Baker, 1978; Davis et al., 1962; Petersons, 2004). Temperature, 370 precipitation, pressure, wind conditions, and lunar illumination have all been suggested as likely 371 migratory drivers (Dechmann et al., 2017; Pettit & O'Keefe, 2017; Roby et al., 2019; Smith & 372 McWilliams, 2016). While perhaps all of these weather conditions may, to a certain degree, have a 373 direct impact on (migratory) flight activity, the exact cues, drivers, and underlying mechanisms of 374 seasonal bat migration have not been identified due to a lack of long-term (i.e. multi-annual) data and 375 detailed knowledge on spatiotemporal migratory patterns (Pettit & O'Keefe, 2017).

376 We found wind conditions, i.e. the frequency of days with tailwinds during spring migration, to 377 explain most of the inter-annual variability in spring migration phenology of bats at Bracken Cave 378 (Figure 2 and Table 1). Wind in favourable directions also influenced spring migration departure 379 probability of European common noctule bats (Nyctalus noctula) (Dechmann et al., 2017). Effects of 380 winds are probably direct, i.e. by influencing departure decisions or migratory flight progress. 381 Brazilian free-tailed bats appear to adjust their airspeed to maintain similar ground speeds regardless 382 of wind support, and ground speeds also do not seem influenced by the direction of prevailing winds 383 (McCracken et al., 2016). Additionally, we did not find any support for an influence of wind 384 assistance on migratory timing. This suggests that advanced or delayed spring migrations with 385 tailwinds and headwinds, respectively, do not result from increased or decreased distances covered 386 but, instead, from decisions to continue migration or stay at a (transient) roost. Combined with the 387 findings of Dechmann et al. (2017) and Pettit & O'Keefe (2017), this indicates that bats rely primarily 388 on day (or night) length as the cue for initiating spring migration, but that the decision to effectively 389 depart or continue the spring migratory journey is strongly determined by wind conditions. Even

though (insectivorous) bats generally live energetically demanding lives, both migration and reproduction (i.e. late pregnancy and lactation) periods are even more energetically demanding (Sommers et al., 2019). Choosing favourable wind conditions for migration is a simple effective mechanism to save energy during movement and arrive in good body condition at the maternity grounds.

While knowledge on flight altitudes of bats during migration specifically is sparse, the highest 395 396 densities of Brazilian free-tailed bats during foraging flights have been shown to occur at around 400 397 to 600 m above-ground-level (agl) (McCracken et al., 2008). For spring migration, we found a 398 tailwind influence at the 850- and 925-hPa pressure level, i.e. approximately 1500 and 750 m asl. The 399 higher ground elevations (Figure S23) at the approximate location (Figure 2) of the 850-hPa tailwind 400 effect compared to those at the 925-hPa tailwind location, indicate that the difference in pressure 401 levels or elevations asl between these effects probably does not reflect differences in migration 402 altitudes, but instead mainly differences in ground elevations. Hence, while Brazilian free-tailed bats 403 have been observed at altitudes exceeding 3000 m agl (Williams et al., 1973), our results point 404 towards flight altitudes during (spring) migration similar to those during foraging flights (McCracken 405 et al., 2008). The autumn wind influence at the 700-hPa or approximately 3000 m asl altitude, 406 however, perhaps indicates higher flight altitudes during autumn migration.

407 We found precipitation to be the main driver of autumn migration phenology of bats at Bracken Cave 408 (Figure 2 and Table 1). While precipitation has often been suggested to affect bat migration, it has yet 409 to be determined how this exactly works and how important the effect of precipitation is relative to 410 other variables. One possibility is a direct negative influence of precipitation on (migratory) flight 411 activity (McGuire et al., 2012; Pettit & O'Keefe, 2017; Voigt et al., 2011). Another might be the 412 indirect (positive) influence of precipitation on insect abundance (Hristov et al., 2010), and perhaps 413 particularly the abundance of migratory insects (Krauel et al., 2018; Lee & McCracken, 2005). 414 Furthermore, precipitation may also influence migratory timing through other indirect pathways, e.g. 415 carry-over effects of the timing of reproduction (Grindal et al., 1992; Linton & Macdonald, 2018) or 416 effects on reproductive rates (Linton & Macdonald, 2018). We found a positive association between 417 autumn precipitation and migration phenology, i.e. passage or departure at Bracken Cave is later in 418 years with more autumn precipitation (Figure 2, Table S5). This positive relationship conforms to 419 expectations for a (negative) direct effect of precipitation on (migratory) flight activity, as well as an 420 indirect effect through either increased insect abundance or delayed insect migration that may result in 421 delayed bat migration. Unfortunately, very little is known about the drivers and mechanisms 422 influencing migratory behaviour (e.g. timing) of insects in general (Satterfield et al., 2020), and even 423 less so for autumn migration specifically (Krauel et al., 2015). Increased insect abundance is, however, 424 often linked to higher precipitation levels (Krauel et al., 2018), and migration theory predicts the 425 optimal time for bat migration to be determined by (changes in) insect abundance (Hedenström, 2009; 426 Krauel & McCracken, 2013). The effect of insect abundance on Brazilian free-tailed bat migration has 427 previously been suggested to work through the effect wind has on migratory insects (Krauel et al., 428 2015). Our results suggest that it is mainly precipitation that affets overall insect abundance, including 429 those of migratory populations. Finally, the timing of our precipitation effects in autumn suggests that 430 it does not work through carry-over effects of delayed parturition or increased reproductive success 431 (Figure 2). Wet springs, however, have also been associated with reduced insect abundance (Krauel et 432 al., 2015; Lee & McCracken, 2005; Pair & Westbrook, 1995), which could also result in earlier 433 autumn migration.

434 We found no influence of temperature. Brazilian free-tailed bats seem to be highly physiologically and 435 morphologically adapted to living under relatively extreme (both warm and cold) temperature 436 conditions but occur mostly in relatively warm geographic areas (Reichard et al., 2010). Moreover, a 437 small overwintering population has been shown to inhabit Bracken Cave across the duration of the 23-438 year period, demonstrating the physiological ability for bats to even survive the winter at this site 439 (Stepanian & Wainwright, 2018). This may explain why migration in Brazilian free-tailed bats is not 440 driven by temperature but by precipitation and wind - contrary to species in regions with more extreme 441 inter-seasonal temperature differences, e.g. at higher latitudes. Some studies of migratory bats (at 442 higher latitudes) in northern Europe or North America, have indeed suggested that temperature plays 443 an important role in migration timing (Jonasson & Guglielmo, 2019; Muthersbaugh et al., 2019; Pettit 444 & O'Keefe, 2017; Roby et al., 2019; Rydell et al., 2014; Smith & McWilliams, 2016), although 445 temperature was often investigated alone, neglecting possible precipitation and wind effects. However, 446 it is possible that wind and precipitation become primary drivers of migratory timing in regions where 447 a minimum temperature threshold for survivability is met throughout the migration period. Note that 448 we analysed mean daily air temperatures in this study, and influences of other temperature-based 449 metrics, e.g. daily minimum temperature, cannot be entirely excluded.

450 In birds, it has repeatedly been shown that the relative influence of different weather variables on 451 migration phenology is strongly species-, but also context-dependent (Gordo, 2007; Haest et al., 452 2018b, 2019; Shaw, 2016), and a similarly diverse response to weather has been suggested for bats 453 (Muthersbaugh et al., 2019). Therefore, future studies on other bat species and populations may find 454 our results of wind being highly important for spring migration and precipitation for autumn migration 455 timing to be specific for Brazilian free-tailed bats. However, we think that the migration of bats 456 adheres to general principles that determine which weather variables will be relevant for migration 457 timing: (1) optimization of (metabolic) migratory costs by choosing the atmospheric conditions 458 (including wind, temperature, and precipitation) that favour efficient energy use; and (2) adjustment to 459 (current or expected) changes in food availability, e.g. insect prey (Bauer et al., 2011; Fleming & Eby, 460 2003; Krauel & McCracken, 2013; Pettit & O'Keefe, 2017). Depending on the geographic, climatic, 461 and individual context, the relative importance of these principles may vary, and thus also the relative 462 importance of the weather (and other) variables that influence specific species/populations.

463 Climate change, bat migration, and conservation of bats

464 For climate change to cause changes in the timing of a biological event, weather does not only need to influence the biological event, i.e. cause inter-annual variability, but the relevant weather variable also 465 466 needs to show a distinct temporal trend. Our results indicate that the spring migration phenology of 467 Brazilian free-tailed bats at Bracken Cave advanced because wind conditions favourable for migration 468 occurred more frequently (Figure 3). Interestingly, favourable wind conditions are expected to 469 continue becoming more prevalent over the next century in this part of the USA (La Sorte et al., 470 2019), raising questions on whether and how this will continue to affect bat migration timing at 471 Bracken Cave.

472 Brazilian free-tailed bats have been proposed as a key indicator species for climate-change impacts on 473 global migratory bat species (Newson et al., 2009). Yet, even in this relative intensively-studied bat 474 species, large uncertainties remain on (changes in) colony-specific population sizes, and hence even 475 more so on overall species abundance (McCracken, 2003; Stepanian & Wainwright, 2018). Even less 476 is known on how weather and climate change may cause inter-annual differences or long-term trends 477 in bat population sizes. Effects of weather on processes directly affecting population sizes, such as 478 reproductive success (Linton & Macdonald, 2018) and timing (Grindal et al., 1992), have been suggested in some bat species. Since long-term data across species and large scales are still largely 479 480 lacking, the specific pathways through which weather affects abundances remain unidentified 481 (Sherwin et al., 2013). Over the period 1995-2017, the mean summer population of Brazilian free-482 tailed bats at Bracken Cave did not show any significant temporal trend (Stepanian & Wainwright, 483 2018), suggesting that the trends in phenology have not (yet) had consequences on population levels of 484 Brazilian free-tailed bats at Bracken Cave. Our results illustrate how long-term time series derived 485 from weather radar data can provide unique insights into climate change impacts on Brazilian free-486 tailed and other, especially cave-dwelling, bat colonies. In this paper, we focused on identifying the 487 effects of climate (change) on migration phenology, but a similar approach could be used to assess the 488 potential relationship between changes in inter-annual population sizes and climate (change).

489 DATA AVAILABILITY STATEMENT

The phenology data necessary to replicate the results of this study have been deposited in the Zenodo repository, <u>https://doi.org/10.5281/zenodo.3945907</u> (Stepanian et al., 2020). The code from (Haest et al., 2020a) can be used to replicate the analysis. NCEP Reanalysis data provided by the NOAA/OAR/ESRL PSL, Boulder, Colorado, USA, and downloaded from their Web site at <u>https://psl.noaa.gov/</u>.

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Supporting Information for:

Climatic drivers of (changes in) bat migration phenology at Bracken Cave (USA)

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Appendix S1 Supplementary Figure S1 to S22 Supplementary Tables S1 to S5 SI References

Weather variable	NCEP variable	Spatial Resolution (in degrees)	Number of analysed grid cells	Data pre-processing and comments
temperature	,air.2m [*] , i.e. air temperature at 2 meters (above ground level)	1.905° x 1.875° (latitude x longitude) (T62 Gaussian grid)	179	We calculated daily mean temperatures from the four 6-hour temperature values.
Precipitation	"prate.sfc", i.e. precipitation rate at surface level	1.905° x 1.875° (latitude x longitude) (T62 Gaussian grid)	179	Precipitation rate data were converted to mm/day.
Wind direction	(East-West) "uwnd" and (North-South) "vwnd" at the 925, 850, and 700-hPa pressure levels	2.5° x 2.5°	111	The 925, 850, and 700-hPa pressure levels roughly corresponds to 750, 1500 and 3000 m altitude above-sea-level. For each pressure level, we calculated daily mean wind directions using the average of the four 6-hour wind values for each of the two wind components.
Wind assistance	(East-West) ,µwnd" and (North-South) ,ywnd" at the 925, 850, and 700-hPa pressure levels	2.5° x 2.5°	111	For each pressure level (i.e. 925, 850, and 700-hPa), we calculated daily mean wind assistance using the average of the four 6-hour wind values for each of the two wind components. Wind assistance was calculated with the RNCEP package (Kemp <i>et al.</i> , 2012a) using the "M.Groundspeed" equation (Kemp <i>et al.</i> , 2012b). To do so, we assumed a groundspeed of 5.7 m/s in still air conditions (based on results from McCracken <i>et al.</i> , 2016), and a preferred direction of movement from each grid cell centre towards Bracken Cave.

Appendix S1 Properties and pre-processing of the weather data that were acquired from the NCEP Reanalysis I database.



Figure S1 Per species Δ AICc, adjusted R², and regression coefficient maps of the identified best time windows for temperature influence on spring migration phenology. The Δ AICc values are the difference of the AICc value of the model with the selected best time window for each grid cell with the AICc value of the baseline model (arrival/passage = $\alpha + \beta$ *year). Δ AICc values are only shown for grid cells that had a probability Pc value < 0.3, i.e., grid cells for which the relation between the identified time window and arrival/passage dates had a Δ AICc that is less likely to obtain due to chance. The adjusted R² values are for the models that have as independent variables both the best identified time window for temperature and the "year" term to account for trends. The regression coefficient maps show the regression coefficient (days/°C) of the best identified time window for temperature when "year" terms to account for trends are also included in the model. Yellow dots with annotated numbers indicate the candidate grid cells selected as potentially influencing spring migration phenology at Bracken Cave. The ID values are used consistently throughout all supplementary figures and tables for ease of reference.



Figure S2 Per species Δ AICc, adjusted R², and regression coefficient maps of the identified best time windows for precipitation influence on spring migration phenology. The Δ AICc values are the difference of the AICc value of the model with the selected best time window for each grid cell with the AICc value of the baseline model (arrival/passage = $\alpha + \beta$ *year). Δ AICc values are only shown for grid cells that had a probability Pc value < 0.3, i.e., grid cells for which the relation between the identified time window and arrival/passage dates had a Δ AICc that is less likely to obtain due to chance. The adjusted R² values are for the models that have as independent variables both the best identified time window for precipitation and the "year" term to account for trends. The regression coefficient maps show the regression coefficient (days/mm) of the best identified time window for precipitation when "year" terms to account for trends are also included in the model. Yellow dots with annotated numbers indicate the candidate grid cells selected as potentially influencing spring migration phenology at Bracken Cave. The ID values are used consistently throughout all supplementary figures and tables for ease of reference.



Figure S3 Per species $\triangle AICc$, adjusted R², and regression coefficient maps of the identified best time windows for the influence on spring migration phenology of the number of days with winds at the 925-hPa pressure level coming from the direction of Bracken Cave. The AAICc values are the difference of the AICc value of the model with the selected best time window for each grid cell with the AICc value of the baseline model (arrival/passage = $\alpha + \beta$ *year). Δ AICc values are only shown for grid cells that had a probability Pc value < 0.3, i.e., grid cells for which the relation between the identified time window and arrival/passage dates had a Δ AICc that is less likely to obtain due to chance. The adjusted R^2 values are for the models that have as independent variables both the best identified time window for number of days with wind coming from Bracken Cave and the "year" term to account for trends. The regression coefficient maps show the regression coefficient (days/day) of the best identified time window for the number of days with wind coming from Bracken Cave when "year" terms to account for trends are also included in the model. Yellow dots with annotated numbers indicate the candidate grid cells selected as potentially influencing spring migration phenology at Bracken Cave. The ID values are used consistently throughout all supplementary figures and tables for ease of reference.



Figure S4 Per species $\triangle AICc$, adjusted R², and regression coefficient maps of the identified best time windows for the influence on spring migration phenology of the number of days with winds at the 925-hPa pressure level going in the direction of Bracken Cave. The \triangle AICc values are the difference of the AICc value of the model with the selected best time window for each grid cell with the AICc value of the baseline model (arrival/passage = $\alpha + \beta^*$ year). \triangle AICc values are only shown for grid cells that had a probability Pc value < 0.3, i.e., grid cells for which the relation between the identified time window and arrival/passage dates had a Δ AICc that is less likely to obtain due to chance. The adjusted R^2 values are for the models that have as independent variables both the best identified time window for number of days with wind going in the direction of Bracken Cave and the "year" term to account for trends. The regression coefficient maps show the regression coefficient (days/day) of the best identified time window for the number of days with wind going in the direction of Bracken Cave when "year" terms to account for trends are also included in the model. Yellow dots with annotated numbers indicate the candidate grid cells selected as potentially influencing spring migration phenology at Bracken Cave. The ID values are used consistently throughout all supplementary figures and tables for ease of reference.



Figure S5 Per species $\triangle AICc$, adjusted R², and regression coefficient maps of the identified best time windows for the influence on spring migration phenology of the number of days with winds at the 850-hPa pressure level coming from the direction of Bracken Cave. The AAICc values are the difference of the AICc value of the model with the selected best time window for each grid cell with the AICc value of the baseline model (arrival/passage = $\alpha + \beta$ *year). Δ AICc values are only shown for grid cells that had a probability Pc value < 0.3, i.e., grid cells for which the relation between the identified time window and arrival/passage dates had a Δ AICc that is less likely to obtain due to chance. The adjusted R^2 values are for the models that have as independent variables both the best identified time window for number of days with wind coming from Bracken Cave and the "year" term to account for trends. The regression coefficient maps show the regression coefficient (days/day) of the best identified time window for the number of days with wind coming from Bracken Cave when "year" terms to account for trends are also included in the model. Yellow dots with annotated numbers indicate the candidate grid cells selected as potentially influencing spring migration phenology at Bracken Cave. The ID values are used consistently throughout all supplementary figures and tables for ease of reference.


Figure S6 Per species \triangle AICc, adjusted R², and regression coefficient maps of the identified best time windows for the influence on spring migration phenology of the number of days with winds at the 850-hPa pressure level going in the direction of Bracken Cave. The \triangle AICc values are the difference of the AICc value of the model with the selected best time window for each grid cell with the AICc value of the baseline model (arrival/passage = $\alpha + \beta^*$ year). \triangle AICc values are only shown for grid cells that had a probability Pc value < 0.3, i.e., grid cells for which the relation between the identified time window and arrival/passage dates had a Δ AICc that is less likely to obtain due to chance. The adjusted R² values are for the models that have as independent variables both the best identified time window for number of days with wind going in the direction of Bracken Cave and the "year" term to account for trends. The regression coefficient maps show the regression coefficient (days/day) of the best identified time window for the number of days with wind going in the direction of Bracken Cave when "year" terms to account for trends are also included in the model. Yellow dots with annotated numbers indicate the candidate grid cells selected as potentially influencing spring migration phenology at Bracken Cave. The ID values are used consistently throughout all supplementary figures and tables for ease of reference.



Figure S7 Per species $\triangle AICc$, adjusted R², and regression coefficient maps of the identified best time windows for the influence on spring migration phenology of the number of days with winds at the 700-hPa pressure level coming from the direction of Bracken Cave. The AAICc values are the difference of the AICc value of the model with the selected best time window for each grid cell with the AICc value of the baseline model (arrival/passage = $\alpha + \beta$ *year). Δ AICc values are only shown for grid cells that had a probability Pc value < 0.3, i.e., grid cells for which the relation between the identified time window and arrival/passage dates had a Δ AICc that is less likely to obtain due to chance. The adjusted R^2 values are for the models that have as independent variables both the best identified time window for number of days with wind coming from Bracken Cave and the "year" term to account for trends. The regression coefficient maps show the regression coefficient (days/day) of the best identified time window for the number of days with wind coming from Bracken Cave when "year" terms to account for trends are also included in the model. Yellow dots with annotated numbers indicate the candidate grid cells selected as potentially influencing spring migration phenology at Bracken Cave. The ID values are used consistently throughout all supplementary figures and tables for ease of reference.



Figure S8 Per species $\triangle AICc$, adjusted R², and regression coefficient maps of the identified best time windows for the influence on spring migration phenology of the number of days with winds at the 700-hPa pressure level going in the direction of Bracken Cave. The \triangle AICc values are the difference of the AICc value of the model with the selected best time window for each grid cell with the AICc value of the baseline model (arrival/passage = $\alpha + \beta^*$ year). \triangle AICc values are only shown for grid cells that had a probability Pc value < 0.3, i.e., grid cells for which the relation between the identified time window and arrival/passage dates had a Δ AICc that is less likely to obtain due to chance. The adjusted R^2 values are for the models that have as independent variables both the best identified time window for number of days with wind going in the direction of Bracken Cave and the "year" term to account for trends. The regression coefficient maps show the regression coefficient (days/day) of the best identified time window for the number of days with wind going in the direction of Bracken Cave when "year" terms to account for trends are also included in the model. Yellow dots with annotated numbers indicate the candidate grid cells selected as potentially influencing spring migration phenology at Bracken Cave. The ID values are used consistently throughout all supplementary figures and tables for ease of reference.



Figure S9 Per species $\triangle AICc$, adjusted R², and regression coefficient maps of the identified best time windows for the influence of wind assistance at the 925-hPa pressure level on spring migration phenology. The Δ AICc values are the difference of the AICc value of the model with the selected best time window for each grid cell with the AICc value of the baseline model (arrival/passage = $\alpha + \beta^*$ year). $\Delta AICc$ values are only shown for grid cells that had a probability Pc value < 0.3, i.e., grid cells for which the relation between the identified time window and arrival/passage dates had a \triangle AICc that is less likely to obtain due to chance. The adjusted R² values are for the models that have as independent variables both the best identified time window for wind assistance and the "year" term to account for trends. The regression coefficient maps show the regression coefficient (days/m*s⁻¹) of the best identified time window for the wind assistance when "year" terms to account for trends are also included in the model. Yellow dots with annotated numbers indicate the candidate grid cells selected as potentially influencing spring migration phenology at Bracken Cave. The ID values are used consistently throughout all supplementary figures and tables for ease of reference.



Figure S10 Per species Δ AICc, adjusted R², and regression coefficient maps of the identified best time windows for the influence of wind assistance at the 850-hPa pressure level on spring migration phenology. The Δ AICc values are the difference of the AICc value of the model with the selected best time window for each grid cell with the AICc value of the baseline model (arrival/passage = $\alpha + \beta^*$ year). $\Delta AICc$ values are only shown for grid cells that had a probability Pc value < 0.3, i.e., grid cells for which the relation between the identified time window and arrival/passage dates had a \triangle AICc that is less likely to obtain due to chance. The adjusted R² values are for the models that have as independent variables both the best identified time window for wind assistance and the "year" term to account for trends. The regression coefficient maps show the regression coefficient (days/m*s⁻¹) of the best identified time window for the wind assistance when "year" terms to account for trends are also included in the model. Yellow dots with annotated numbers indicate the candidate grid cells selected as potentially influencing spring migration phenology at Bracken Cave. The ID values are used consistently throughout all supplementary figures and tables for ease of reference.



Figure S11 Per species Δ AICc, adjusted R², and regression coefficient maps of the identified best time windows for the influence of wind assistance at the 700-hPa pressure level on spring migration phenology. The Δ AICc values are the difference of the AICc value of the model with the selected best time window for each grid cell with the AICc value of the baseline model (arrival/passage = $\alpha + \beta^*$ year). $\Delta AICc$ values are only shown for grid cells that had a probability Pc value < 0.3, i.e., grid cells for which the relation between the identified time window and arrival/passage dates had a $\triangle AICc$ that is less likely to obtain due to chance. The adjusted R^2 values are for the models that have as independent variables both the best identified time window for wind assistance and the "year" term to account for trends. The regression coefficient maps show the regression coefficient (days/m*s⁻¹) of the best identified time window for the wind assistance when "year" terms to account for trends are also included in the model. Yellow dots with annotated numbers indicate the candidate grid cells selected as potentially influencing spring migration phenology at Bracken Cave. The ID values are used consistently throughout all supplementary figures and tables for ease of reference.



Figure S12 Per species Δ AICc, adjusted R², and regression coefficient maps of the identified best time windows for temperature influence on autumn migration phenology. The Δ AICc values are the difference of the AICc value of the model with the selected best time window for each grid cell with the AICc value of the baseline model (arrival/passage = $\alpha + \beta$ *year). Δ AICc values are only shown for grid cells that had a probability Pc value < 0.3, i.e., grid cells for which the relation between the identified time window and arrival/passage dates had a Δ AICc that is less likely to obtain due to chance. The adjusted R² values are for the models that have as independent variables both the best identified time window for temperature and the "year" term to account for trends. The regression coefficient maps show the regression coefficient (days/°C) of the best identified time window for temperature when "year" terms to account for trends are also included in the model. Yellow dots with annotated numbers indicate the candidate grid cells selected as potentially influencing autumn migration phenology at Bracken Cave. The ID values are used consistently throughout all supplementary figures and tables for ease of reference.



Figure S13 Per species Δ AICc, adjusted R², and regression coefficient maps of the identified best time windows for precipitation influence on autumn migration phenology. The Δ AICc values are the difference of the AICc value of the model with the selected best time window for each grid cell with the AICc value of the baseline model (arrival/passage = $\alpha + \beta$ *year). Δ AICc values are only shown for grid cells that had a probability Pc value < 0.3, i.e., grid cells for which the relation between the identified time window and arrival/passage dates had a Δ AICc that is less likely to obtain due to chance. The adjusted R² values are for the models that have as independent variables both the best identified time window for precipitation and the "year" term to account for trends. The regression coefficient maps show the regression coefficient (days/mm) of the best identified time window for precipitation when "year" terms to account for trends are also included in the model. Yellow dots with annotated numbers indicate the candidate grid cells selected as potentially influencing autumn migration phenology at Bracken Cave. The ID values are used consistently throughout all supplementary figures and tables for ease of reference.



Figure S14 Per species \triangle AICc, adjusted R², and regression coefficient maps of the identified best time windows for the influence on autumn migration phenology of the number of days with winds at the 925-hPa pressure level coming from the direction of Bracken Cave. The AAICc values are the difference of the AICc value of the model with the selected best time window for each grid cell with the AICc value of the baseline model (arrival/passage = $\alpha + \beta$ *year). Δ AICc values are only shown for grid cells that had a probability Pc value < 0.3, i.e., grid cells for which the relation between the identified time window and arrival/passage dates had a Δ AICc that is less likely to obtain due to chance. The adjusted R^2 values are for the models that have as independent variables both the best identified time window for number of days with wind coming from Bracken Cave and the "year" term to account for trends. The regression coefficient maps show the regression coefficient (days/day) of the best identified time window for the number of days with wind coming from Bracken Cave when "year" terms to account for trends are also included in the model. Yellow dots with annotated numbers indicate the candidate grid cells selected as potentially influencing autumn migration phenology at Bracken Cave. The ID values are used consistently throughout all supplementary figures and tables for ease of reference.



Figure S15 Per species \triangle AICc, adjusted R², and regression coefficient maps of the identified best time windows for the influence on autumn migration phenology of the number of days with winds at the 925-hPa pressure level going in the direction of Bracken Cave. The \triangle AICc values are the difference of the AICc value of the model with the selected best time window for each grid cell with the AICc value of the baseline model (arrival/passage = $\alpha + \beta^*$ year). \triangle AICc values are only shown for grid cells that had a probability Pc value < 0.3, i.e., grid cells for which the relation between the identified time window and arrival/passage dates had a Δ AICc that is less likely to obtain due to chance. The adjusted R^2 values are for the models that have as independent variables both the best identified time window for number of days with wind going in the direction of Bracken Cave and the "year" term to account for trends. The regression coefficient maps show the regression coefficient (days/day) of the best identified time window for the number of days with wind going in the direction of Bracken Cave when "year" terms to account for trends are also included in the model. Yellow dots with annotated numbers indicate the candidate grid cells selected as potentially influencing autumn migration phenology at Bracken Cave. The ID values are used consistently throughout all supplementary figures and tables for ease of reference.



Figure S16 Per species Δ AICc, adjusted R², and regression coefficient maps of the identified best time windows for the influence on autumn migration phenology of the number of days with winds at the 850-hPa pressure level coming from the direction of Bracken Cave. The AAICc values are the difference of the AICc value of the model with the selected best time window for each grid cell with the AICc value of the baseline model (arrival/passage = $\alpha + \beta$ *year). Δ AICc values are only shown for grid cells that had a probability Pc value < 0.3, i.e., grid cells for which the relation between the identified time window and arrival/passage dates had a Δ AICc that is less likely to obtain due to chance. The adjusted R^2 values are for the models that have as independent variables both the best identified time window for number of days with wind coming from Bracken Cave and the "year" term to account for trends. The regression coefficient maps show the regression coefficient (days/day) of the best identified time window for the number of days with wind coming from Bracken Cave when "year" terms to account for trends are also included in the model. Yellow dots with annotated numbers indicate the candidate grid cells selected as potentially influencing autumn migration phenology at Bracken Cave. The ID values are used consistently throughout all supplementary figures and tables for ease of reference.



Figure S17 Per species Δ AICc, adjusted R², and regression coefficient maps of the identified best time windows for the influence on autumn migration phenology of the number of days with winds at the 850-hPa pressure level going in the direction of Bracken Cave. The \triangle AICc values are the difference of the AICc value of the model with the selected best time window for each grid cell with the AICc value of the baseline model (arrival/passage = $\alpha + \beta^*$ year). \triangle AICc values are only shown for grid cells that had a probability Pc value < 0.3, i.e., grid cells for which the relation between the identified time window and arrival/passage dates had a Δ AICc that is less likely to obtain due to chance. The adjusted R^2 values are for the models that have as independent variables both the best identified time window for number of days with wind going in the direction of Bracken Cave and the "year" term to account for trends. The regression coefficient maps show the regression coefficient (days/day) of the best identified time window for the number of days with wind going in the direction of Bracken Cave when "year" terms to account for trends are also included in the model. Yellow dots with annotated numbers indicate the candidate grid cells selected as potentially influencing autumn migration phenology at Bracken Cave. The ID values are used consistently throughout all supplementary figures and tables for ease of reference.



Figure S18 Per species Δ AICc, adjusted R², and regression coefficient maps of the identified best time windows for the influence on autumn migration phenology of the number of days with winds at the 700-hPa pressure level coming from the direction of Bracken Cave. The AAICc values are the difference of the AICc value of the model with the selected best time window for each grid cell with the AICc value of the baseline model (arrival/passage = $\alpha + \beta$ *year). Δ AICc values are only shown for grid cells that had a probability Pc value < 0.3, i.e., grid cells for which the relation between the identified time window and arrival/passage dates had a Δ AICc that is less likely to obtain due to chance. The adjusted R^2 values are for the models that have as independent variables both the best identified time window for number of days with wind coming from Bracken Cave and the "year" term to account for trends. The regression coefficient maps show the regression coefficient (days/day) of the best identified time window for the number of days with wind coming from Bracken Cave when "year" terms to account for trends are also included in the model. Yellow dots with annotated numbers indicate the candidate grid cells selected as potentially influencing autumn migration phenology at Bracken Cave. The ID values are used consistently throughout all supplementary figures and tables for ease of reference.



Figure S19 Per species \triangle AICc, adjusted R², and regression coefficient maps of the identified best time windows for the influence on autumn migration phenology of the number of days with winds at the 700-hPa pressure level going in the direction of Bracken Cave. The \triangle AICc values are the difference of the AICc value of the model with the selected best time window for each grid cell with the AICc value of the baseline model (arrival/passage = $\alpha + \beta^*$ year). \triangle AICc values are only shown for grid cells that had a probability Pc value < 0.3, i.e., grid cells for which the relation between the identified time window and arrival/passage dates had a Δ AICc that is less likely to obtain due to chance. The adjusted R^2 values are for the models that have as independent variables both the best identified time window for number of days with wind going in the direction of Bracken Cave and the "year" term to account for trends. The regression coefficient maps show the regression coefficient (days/day) of the best identified time window for the number of days with wind going in the direction of Bracken Cave when "year" terms to account for trends are also included in the model. Yellow dots with annotated numbers indicate the candidate grid cells selected as potentially influencing autumn migration phenology at Bracken Cave. The ID values are used consistently throughout all supplementary figures and tables for ease of reference.



Figure S20 Per species Δ AICc, adjusted R², and regression coefficient maps of the identified best time windows for the influence of wind assistance at the 925-hPa pressure level on autumn migration phenology. The Δ AICc values are the difference of the AICc value of the model with the selected best time window for each grid cell with the AICc value of the baseline model (arrival/passage = $\alpha + \beta^*$ year). $\Delta AICc$ values are only shown for grid cells that had a probability Pc value < 0.3, i.e., grid cells for which the relation between the identified time window and arrival/passage dates had a $\triangle AICc$ that is less likely to obtain due to chance. The adjusted R^2 values are for the models that have as independent variables both the best identified time window for wind assistance and the "year" term to account for trends. The regression coefficient maps show the regression coefficient (days/m*s⁻¹) of the best identified time window for the wind assistance when "year" terms to account for trends are also included in the model. Yellow dots with annotated numbers indicate the candidate grid cells selected as potentially influencing autumn migration phenology at Bracken Cave. The ID values are used consistently throughout all supplementary figures and tables for ease of reference.



Figure S21 Per species Δ AICc, adjusted R², and regression coefficient maps of the identified best time windows for the influence of wind assistance at the 850-hPa pressure level on autumn migration phenology. The Δ AICc values are the difference of the AICc value of the model with the selected best time window for each grid cell with the AICc value of the baseline model (arrival/passage = $\alpha + \beta^*$ year). $\Delta AICc$ values are only shown for grid cells that had a probability Pc value < 0.3, i.e., grid cells for which the relation between the identified time window and arrival/passage dates had a $\triangle AICc$ that is less likely to obtain due to chance. The adjusted R^2 values are for the models that have as independent variables both the best identified time window for wind assistance and the "year" term to account for trends. The regression coefficient maps show the regression coefficient (days/m*s⁻¹) of the best identified time window for the wind assistance when "year" terms to account for trends are also included in the model. Yellow dots with annotated numbers indicate the candidate grid cells selected as potentially influencing autumn migration phenology at Bracken Cave. The ID values are used consistently throughout all supplementary figures and tables for ease of reference.



Figure S22 Per species \triangle AICc, adjusted R², and regression coefficient maps of the identified best time windows for the influence of wind assistance at the 700-hPa pressure level on autumn migration phenology. The Δ AICc values are the difference of the AICc value of the model with the selected best time window for each grid cell with the AICc value of the baseline model (arrival/passage = $\alpha + \beta^*$ year). $\Delta AICc$ values are only shown for grid cells that had a probability Pc value < 0.3, i.e., grid cells for which the relation between the identified time window and arrival/passage dates had a $\triangle AICc$ that is less likely to obtain due to chance. The adjusted R^2 values are for the models that have as independent variables both the best identified time window for wind assistance and the "year" term to account for trends. The regression coefficient maps show the regression coefficient (days/m*s⁻¹) of the best identified time window for the wind assistance when "year" terms to account for trends are also included in the model. Yellow dots with annotated numbers indicate the candidate grid cells selected as potentially influencing autumn migration phenology at Bracken Cave. The ID values are used consistently throughout all supplementary figures and tables for ease of reference.





Figure S23 Elevation map of the study area. Data source: Jarvis et al. (2008)

Table S1 General settings for the time window analyses

Reference day of the year:-Spring:1 June-Autumn:31 DecemberMinimum time window length:14 P_c^* limit:< 0.3</td>

* Bailey & van de Pol (2016); van de Pol et al. (2016); Haest et al. (2018)
Table S2 Settings for the time window analysis, specific to each of the analysed weather variables

Weather variable*	Operation on each time window	Maximum time window length	Period of the year analysed (approximate)	Number of windows analysed per grid cell	
Temperature	mean	365	Spring: 2 June to 1 June	62481	
Precipitation	sum	505	Autumn: 1 January to 31 December	02401	
Wind - 925-hPa pressure level - coming from Bracken Cave					
Wind - 925-hPa pressure level - going to Bracken Cave					
Wind - 850-hPa pressure level - coming from Bracken Cave	1 6 1				
Wind - 850-hPa pressure level - going to Bracken Cave	number of days		Spring: 1 December to 1 June		
Wind - 700-hPa pressure level - coming from Bracken Cave		182		14706	
Wind - 700-hPa pressure level - going to Bracken Cave			Autumn: 2 July to 31 December		
Wind - 925-hPa pressure level - wind assistance					
Wind - 850-hPa pressure level - wind assistance	mean				
Wind - 700-hPa pressure level - wind assistance					

*For definitions of how the weather variables were calculated, see the "Materials and Methods - Weather data" section.

Table S3 Overview of the relative variable importance of the candidate weather signals for spring migration phenology, calculated using three different methods: (a) the sum of the multi-model AICc weights across all the possible models with maximum four independent variables (Burnham & Anderson, 2002); (b) the boruta method (Kursa & Rudnicki, 2010); and (c) the game-theory-based LMG metric for variance decomposition in linear models (Grömping, 2006, 2015). This table includes only the candidate weather signals that were not removed due to high collinearity with another better performing candidate signal, low performance compared to an intercept-only model, or because of relative variable importances using the boruta method that are likely obtained by chance. The ID values are used consistently throughout all supplementary figures and tables for ease of reference. WindGT: Wind going in the direction of Bracken Cave; WindCF: Wind coming from the direction of Bracken Cave.

						Relative Variable Importance		Ranks of Relative Variable Importance						
		∆AICc	ΔAICe											
		(compared	(compared to	Window	Window									
		to trend	intercept-only	Open	Close	Model				Model			Mean	Rank of
Climate variable	ID	model)	model)	Date	Date	weights	LMG	Boruta	Mean	weights	LMG	Boruta	Rank	the mean
WindGT – 925-hPa	S12	-17.29	-19.35	13 Feb	13 Mar	0.25	0.09	0.10	0.15	14	12	14	13.3	14
WindGT – 850-hPa	S17	-18.51	-23.11	05 Dec	25 Dec	0.25	0.11	0.08	0.15	13	14	9	12	13
WindGT - 850-hPa	S16	-9.77	-12.33	23 Feb	23 Mar	0.20	0.08	0.07	0.12	12	10	8	10	12
WindGT-850-hPa	S20	-14.15	-19.02	04 Dec	19 Jan	0.09	0.10	0.10	0.10	11	13	12	12	11
WindGT - 700-hPa	S27	-17.54	-16.75	17 Dec	31 Dec	0.04	0.08	0.10	0.07	9	11	13	11	10
Wind Assistance - 850-hPa	S37	-14.15	-16.01	18 Dec	09 Jan	0.05	0.07	0.07	0.07	10	9	7	8.7	9
WindCF - 700-hPa	S21	-14.48	-15.25	04 Dec	24 Dec	0.03	0.07	0.09	0.06	7	7	11	8.3	8
WindGT - 700-hPa	S32	-16.64	-15.83	17 Feb	07 Mar	0.03	0.07	0.06	0.06	8	8	6	7.3	7
Wind Assistance – 925-hPa	S34	-13.70	-11.36	06 Dec	09 Feb	0.00	0.05	0.09	0.05	2	2	10	4.7	6
Wind Assistance - 700-hPa	S43	-15.60	-14.42	15 Feb	05 Mar	0.01	0.06	0.06	0.04	6	6	5	5.7	5
WindGT-700-hPa	S28	-19.72	-13.36	04 Mar	25 Mar	0.01	0.06	0.04	0.04	4	5	1	3.3	4
WindGT - 925-hPa	S10	-12.02	-10.89	15 Mar	10 May	0.01	0.06	0.04	0.04	5	4	2	3.7	3
WindCF - 700-hPa	S25	-13.33	-9.25	20 Apr	27 May	0.01	0.05	0.05	0.03	3	1	4	2.7	2
Precipitation	S5	-10.30	-12.89	27 Aug	25 Feb	0.00	0.05	0.05	0.03	1	3	3	2.3	1

Table S4 Overview of the relative variable importance of the candidate weather signals for autumn migration phenology, calculated using three different methods: (a) the sum of the multi-model AICc weights across all the possible models with maximum four independent variables (Burnham & Anderson, 2002); (b) the boruta method (Kursa & Rudnicki, 2010); and (c) the game-theory-based LMG metric for variance decomposition in linear models (Grömping, 2006, 2015). This table includes only the candidate weather signals that were not removed due to high collinearity with another better performing candidate signal, low performance compared to an intercept-only model, or because of relative variable importances using the boruta method that are likely obtained by chance. The ID values are used consistently throughout all supplementary figures and tables for ease of reference. WindGT: Wind going in the direction of Bracken Cave; WindCF: Wind coming from the direction of Bracken Cave.

Δι		ΔAICc	ΔAICc			Relativ	Relative Variable Importance				Ranks of Relative Variable Importance				
		(compared	(compared	Window	Window										
		to trend	to intercept-	Open	Close	Model				Model			Mean	Rank of	
Climate variable	ID	model)	only model)	Date	Date	weights	LMG	Boruta	Mean	weights	LMG	Boruta	Rank	the mean	
Precipitation	A11	-22.27	-19.41	23-Feb	14-Mar	0.24	0.09	0.09	0.14	14	12	11	12.3	14	
Precipitation	A9	-16.46	-14.50	03-Aug	26-Dec	0.22	0.08	0.09	0.13	13	10	12	11.7	13	
WindCF – 700-hPa	A32	-17.05	-17.49	03-Nov	28-Nov	0.16	0.09	0.11	0.12	12	13	14	13	12	
Wind Assistance – 925-hPa	A47	-12.03	-11.79	09-Nov	01-Dec	0.05	0.07	0.11	0.08	7	8	13	9.3	11	
WindCF - 700-hPa	A33	-21.97	-22.51	23-Sep	30-Oct	0.05	0.10	0.07	0.07	10	14	9	11	10	
WindCF - 700-hPa	A36	-15.81	-14.40	10-Sep	01-Dec	0.07	0.07	0.06	0.07	11	7	6	8	9	
Wind Assistance - 700-hPa	A51	-19.57	-15.33	31-Aug	14-Oct	0.05	0.08	0.06	0.06	9	11	3	7.7	8	
Precipitation	A7	-12.44	-12.16	01-Oct	18-Dec	0.05	0.08	0.06	0.06	8	9	5	7.3	7	
WindGT – 700-hPa	A43	-14.59	-12.14	22-Jul	31-Aug	0.02	0.07	0.07	0.05	5	6	10	7	6	
WindGT – 925-hPa	A21	-9.95	-9.63	28-Oct	18-Nov	0.05	0.06	0.05	0.05	6	5	1	4	5	
WindGT – 700-hPa	A42	-15.92	-14.60	22-Jul	08-Aug	0.00	0.06	0.07	0.04	1	4	8	4.3	4	
Wind Assistance - 700-hPa	A52	-9.48	-10.02	10-Oct	29-Nov	0.02	0.05	0.06	0.04	4	2	4	3.3	3	
WindGT – 700-hPa	A44	-8.04	-8.59	29-Aug	28-Sep	0.01	0.05	0.06	0.04	3	1	7	3.7	2	
Wind Assistance – 925-hPa	A46	-8.33	-8.82	04-Sep	23-Sep	0.00	0.06	0.06	0.04	2	3	2	2.3	1	

Table S5 Contribution of the effect of each weather variable on mean passage date (MPD) in spring and autumn to the overall trend in MPD over the period 1995-2017. Negative values are in italics. IDs are identical to those in Table S3 and S4, and Figure S1 to S22. Coef.: coefficient; SE: standard error. The ID values are used consistently throughout all supplementary figures and tables for ease of reference.

			<u>ƏMPD</u>		dClin	nate	$\frac{\partial MPD}{\partial Climate} \times$	dClimate	Total change	
Season	Climate variable	ID	Coef	Coof SF		me SE	Coef	SE	_ Over 25 years	
Scason		1D ~ ()		SE .	0.001.	SE	0001.	SE	(uays)	
Spring	tailwind 925-hPa	S12	-2.20	0.45	0.09	0.07	-0.21	0.16	-4.81	
	tailwind 850-hPa	S17	-2.06	0.36	0.24	0.07	-0.50	0.17	-11.54	
Autumn	precipitation	A9	0.06	0.01	-5.77	2.70	-0.34	0.18	-7.88	
	precipitation	A11	-0.41	0.07	-0.43	0.60	0.18	0.25	4.08	
	headwind 700-hPa	A32	-2.46	0.75	0.06	0.06	-0.14	0.16	-3.24	

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