# The Detection of Micro-Calcifications in Mammographic Images Using High Dimensional Features

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Abstract

This paper examines techniques for the efficient use of high dimensional feature sets in the detection of micro-calcifications in mammograms. The paper focuses on techniques for dimensionality reduction and discriminant analysis. The paper examines the use of principal components and Fisher's linear discriminant for dimensionality reduction along with parametric and nonparametric statistical techniques for discriminant analysis.

# 1: Introduction

It is estimated that there are approximately 150,000 new cases of breast cancer diagnosed annually[1]. Besides the large toll in human suffering that this represents, the diagnosis and treatment of these cases is a burden on the health care system. As with many medical conditions a favorable prognosis for the patient is based on the earliest possible detection of the abnormality.

There has been a great deal of recent interest in using computer techniques in image processing and pattern recognition to help the radiologist's diagnostic efforts[2]. This work, along with work in medical expert systems, has come under the general heading of computer aided diagnosis (CAD). Much of the CAD work seeks to help the radiologist by drawing attention to regions of

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the image that the computer has identified as suspicious. This would be an extremely valuable tool to aid in the timely and accurate detection of anomolous tissue. CAD techniques have been applied to such diverse areas as liver ultrasound[3], chest radiography[4], and x-ray mammography[5].

One of the early indicators of breast cancer is the radiographic presence of clusters of malignant microcalcifications. In many cases these clusters are located in the midst of dense parenchymal tissue and it takes very careful inspection by the radiologist to locate them. Their prominent role in diagnosis and difficult detectability make their CAD an important goal.

The next section focuses on our techniques used to extract features from the microcalcifications. This is followed by a brief discussion of the Fisher linear discriminant and principal component analysis approaches to dimensionality reduction. Next we turn our attention to a few words on parametric and nonparametric Bayesian based classification techniques. The fruits of the application of these techniques to the detection problem at hand is presented in the results section and we conclude the paper with a few words on future directions.

## 2: Feature Extraction

The first step in the feature extraction process is the identification of regions of interest from the image. It is expected that these extracted regions contain both true groups of malignant microcalcifications and groups of benign microcalcifications or other anatomical features. It is the function of the classification portion of the system to sort out the true malignancies from the false positives, which will be discussed below. The segmentation is done by thresholding a difference image which is created by subtracting from every pixel the average of a 15x15 pixel (1.13mm) window surrounding it. Region growing is performed on all the pixels left in order to group them into candidate objects for input into the feature extraction system. This segmentation routine can produce hundreds of candidate objects due to its sensitivity. To reduce the number of candidate objects only 3% of the total number of segmented objects with the highest contrast (with a minimum of 100) are retained. The twenty-one extracted features along with references to where they first appeared in the literature are listed in Table 1. A designator of "Standard" indicates a standard image processing technique.

Ref. Number	Name	Description
[6]	Contrast 1	Avg. Obj Avg. Bck.
[6]	Contrast 2	Avg. Obj. / Avg. Bck.
[7]	Contrast 3	Max. in Obj Avg. Bck.
Standard	Smoothness B	Computed on Background
[8]	Area	Number of Pixels
[8]	Smoothness 1	RMS Noise Fluctuations in Object
[7]	Smoothness 2	RMS Noise Fluctuations in 51x51 windowed BCK.
[9]	Edge Strength	Avg. of Robert's Gradient for Perimeter Pixels

#### **Table 1: Microcalcification Features.**

Ref.	Name	Description
Number		
Standard	Gradient 1	Avg. of 3x3 Sobel Gradient
		for Perimeter Pixels
[10]	Moment F1	About Centroid
[10]	Moment F2	About Centroid
[10]	Moment F3	About Centroid
[10]	Moment (F3-F1)	About Centroid
[10]	Compactness	P <sup>2</sup> /A, P=Perimeter, A=Area
Standard	Shape 1	Rmax/Rmin
[8]	Shape 2	Std. Dev. from Perimeter Pix-
		els to centroid
Standard	Shape 3	Area/Rmax
[8]	Average Gray	Computed on Object
[10]	Gray Level Moment	Computed on Perimeter Pix-
	1	els
[10]	Gray Level Moment	Computed on Object Pixels
	2	
Standard	Average Edge Step	Computed on Perimeter Pix-
		els

### **Table 1: Microcalcification Features.**

The reader is referred to Woods et al[11], [12] for a more complete description of the feature extraction procedure.

### **3: Dimensionality Reduction**

Given our set of extracted features in R<sup>21</sup>, Bellman's curse of dimensionality [13] dictates that we project our features to as low a dimensional space as possible before we parametrically or nonparametrically model their underlying density. There are many techniques for dimensionality reduction. Since our ultimate goal is classification, it is beneficial to consider those transformations that enhance the separation between the probability density function of the features computed on the malignant microcalcifications and the one computed on the other nonmalignant tissue.

The two techniques for dimensionality reduction that were used here are the Fishers linear discriminant (FLD) [14] and principal component analysis (PCA) [13]. The FLD procedure seeks the best projection from the feature space to  $R^1$ , i. e. it seeks the best line to project upon. In this case optimality is measured by J(w), the ratio of between-class scatter, SB, to within-class scatter, S<sub>W</sub>, given by

$$J(\vec{w}) = \frac{\vec{w}^{t} S_{B} \vec{w}}{\vec{w}^{t} S_{w} \vec{w}}.$$
 (1)

 $\overline{w} \, S_W w$ In eq. 1  $S_B = (\vec{m}_1 - \vec{m}_2) (\vec{m}_1 - \vec{m}_2)^t$  where  $\vec{m}_1$  is the mean vector for class 1, and  $\vec{m}_2$  is the mean for class 2. The within class scatter matrix  $S_W = S_1 + S_2$ , where  $S_i = \sum_{\substack{k \in X_i \\ k \in X_i}} (\hat{x} - \vec{m}_i) (\hat{x} - \vec{m}_i)^t$ .

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In PCA one seeks to rotate the original data axes so that they align with the eigenvectors of the covariance matrix  $\Sigma$ . In this manner we seek a lower dimensional space that accounts for the variance of the features. In our case, in order to put the PCA approach on equal footing with the FLD we seek the best 1-d subspace to project upon. In order to account for the largest variance in the features we project upon the eigenvector corresponding to the largest eigenvalue.

### 4: Discriminant Analysis

Given the projected features in  $\mathbb{R}^1$  we proceed with a Bayesian approach to the classification problem. We seek parametric or nonparametric models of the class-conditional

probability densities  $p(x|c_i)$ . Once we have modeled these class-conditional or state-conditional probability densities we employ a standard likelihood ratio hypothesis testing procedure[15] to determine the class of an unknown observation.

In the simplest approach each of the two class-conditional probability density functions is modeled as  $N(\mu_i, \Sigma_p)$ , where  $\mu_i$  is the mean for the i-th class and  $\Sigma_p$  is the pooled covariance matrix. This procedure is known in the literature as a linear classifier. If we allow each class-conditional density to have its own covariance matrix then the resulting classifier is known as a quadratic classifier. The reader is referred to Duda and Hart [14] for a good treatment of these topics.

If on the other hand there is reason to believe that the underlying class-conditional densities are inherently nonnormal, then it is beneficial to employ a nonparametric estimator. One of the simplest such procedures is the kernel estimator[16]. In this case given  $X=\{x_i\}_{i=1}^n$  where  $x_i$ -i.i.d f(x), we model the unknown distribution as

$$\hat{f}(x) = \frac{1}{nh} \sum_{i=1}^{n} g\left(\frac{x - x_i}{h}\right),$$
(2)

where the smoothing function g(z) is chosen to be N(0,1), and h is referred to as the kernel bandwidth. This procedure has not only the benefits of computational simplicity, but it possesses powerful asymptotic properties that provide it robustness to the nonnormality of the class-conditional probability density functions. In our analysis h was chosen optimal based on the normal assumption[16].

Given models of the class conditional probability densities, one is interested in assessing the performance of the classifier. One approach to this task is through the production of receiver operator characteristic (ROC) curves. These curves typically display some function of the probability of misclassifying a target (microcalcification in our case) as a function of false alarm rate. The probability of the first type of error is designated as  $\alpha$  and  $(1-\alpha)$  is often referred to in the medical community as the sensitivity. A false alarm of course is when the classifier incorrectly identifies normal tissue as tumorous. In order to produce the most accurate estimation of the ROC curve of our classifiers we employ a variant of the "leave one out" procedure[17]. In this procedure the features from a given image are left out and the FLD and PCA projections along with the classconditional probability densities are built on the features from the remaining images. The projections are applied to the features in the image that was left out and the constructed class-conditional probability densities are used to build the likelihood ratios that are eventually used to determine class membership at each point. This procedure is repeated using the other images. At the end of the procedure the full set of likelihood ratios is used to perform the ROC analysis.

### 5: Results

The 24 images in the study were digitized at 70 micron resolution with 8 bit gray levels. The location of the malignant microcalcifications in the image was determined by having the groups of microcalcifications first identified by a radiologist and then the subsequent exact location of each microcalcification labeled by a trained technician. After passing the images through the segmentation portion of the algorithm there were 5730 nonmalignant regions and 542 malignant regions produced. The twenty-one features were computed for each of these regions.

Let's first examine the structure of the data for the two classes. Figure 1 represents kernel estimator approximations of the probability densities for each of the two classes' feature sets in FLD space. The normal appearance of the densities in Figure 1 is to be contrasted with the kernel estimated densities in PCA space shown in Figure 2.

Figures 3, 4, and 5 show the ROC curves for the linear, quadratic, and kernel based classifiers for the FLD and PCA features. The curves plot  $\alpha$  as a function of false alarm rate. In each case the classifiers perform better using the FLD feature then the PCA one. It is also interesting to note that the performance of the three classifiers is virtually identical. To provide a figure for comparison we note that at a false alarm rate of 10% the classifiers possesses an  $\alpha$  level of 20% accross the set of images. This value is in the same ballpark as some of our previous work utilizing seven dimensional features[11].

# 6: Conclusions

The preliminary results put forth in this paper indicate the superiority of the Fisher linear discriminate transformation over projection on the strongest principal component as a means for dimensionality reduction prior to Bayesian based discriminant analysis. These results are based on one area of application and future work is needed to determine if this trend will hold in general. With the curse of dimensionality in mind it is also important to reemphasize the comparable performance of the classifiers using the one dimensional features and our previous higher dimensional work. Future work will focus on "optimal" projections to  $R^2$  and  $R^3$  along with the application of our semi-parametric adaptive mixtures density estimator[18].



Figure 1. Kernel estimates for the microcalcification and nonmicrocalcification probability densities in Fisher linear discriminant space.

Figure 2. Kernel estimates for the microcalcification and nonmicrocalcification probability densities in the first principal component space.



Figure 3. Leave one image out ROC curve for the linear classifier on the FLD and PC features.





0.4

0.2

\* = PC + = FLD KE

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