DISTRIBUTED RECURSIVE ESTIMATION UNDER HEAVY-TAIL 2 COMMUNICATION NOISE

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Abstract. We consider distributed recursive estimation of an unknown vector parameter 5 ${m heta}^* \in \mathbb{R}^M$ in the presence of impulsive communication noise. That is, we assume that inter-agent 6 communication is subject to an additive communication noise that may have heavy-tails or is con-7 8 taminated with outliers. To combat this effect, within the class of consensus+innovations distributed 9 estimators, we introduce for the first time a nonlinearity in the consensus update. We allow for a 10 general class of nonlinearities that subsumes, e.g., the sign function or component-wise saturation 11 function. For the general nonlinear estimator and a general class of additive communication noises that may have infinite moments of order higher than one - we establish almost sure (a.s.) convergence 12 13to the parameter θ^* . We further prove asymptotic normality and evaluate the corresponding asymp-14totic covariance. These results reveal interesting tradeoffs between the negative effect of "loss of 15 information" due to incorporation of the nonlinearity, and the positive effect of communication noise reduction. We also demonstrate and quantify benefits of introducing the nonlinearity in high-noise (low signal-to-noise ratio) and heavy-tail communication noise regimes. 17

18 Key words. Distributed inference; distributed estimation; recursive estimation; heavy-tail 19noise; consensus+innovations; stochastic approximation.

AMS subject classifications. 93E10, 93E35, 60G35, 94A13, 62M05 20

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21**1.** Introduction. We consider distributed inference in networked systems, whe-22 re each agent in a generic network continuously (over time instances t = 0, 1, ...,) makes noisy linear observations of an unknown vector parameter $\boldsymbol{\theta}^* \in \mathbb{R}^M$. Each 23 agent, at each time t, generates a local estimate of θ^* through the so-called consen-24 sus+innovations strategy, i.e., by 1) weight-averaging its current solution estimate 25with those of its neighbors, and 2) assimilating its new observation. 26

27In this paper, we are interested in consensus+innovations distributed estimation in the presence of an impulsive communication noise, e.g., when the communication 28 noise that corresponds to inter-neighbor communications is heavy-tailed or contami-29nated with outliers. It is highly relevant to consider impulsive communication noise 30 in many application scenarios. For example, edge devices in Internet of Things (IoT) systems or sensor networks can be subject to impulsive noise distributions that may 33 not have finite moments of order higher than one, e.g., [8, 32, 13, 37, 12, 9]. In this work, we allow the communication noise to be a zero-mean random variable that 34 may have infinite moments of order α , for any $\alpha > 1$. In particular, communication 35 noise may have an infinite variance. To the best of our knowledge, such scenarios 36 have not been studied in the past work, wherein communication noise in consensus+innovations inference is always assumed to have a finite moment of at least second 38 order (finite variance). Actually, as demonstrated ahead in the paper, existing con-39

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sensus+innovation estimators – that are always *linear* in the consensus update part – 40 41 can fail to converge under a heavy-tail communication noise. To combat the effect of the (impulsive or high variance) communication noise, we introduce for the first time 42 a general nonlinearity in the consensus update. More precisely, we apply a nonlinear 43operator (e.g., a sign function, a saturation-like function, or a sigmoid function) on 44 the difference between an agent's current iterate and a noisy version of its neighbor's 45 iterate, for every agent in the neighborhood set. We establish, under a general setting 46 for the nonlinearity and the additive communication noise, almost sure (a.s.) conver-47 gence of the nonlinear estimator to the true parameter θ^* . We also prove asymptotic 48normality and evaluate the corresponding asymptotic covariance in terms of the un-49derlying network topology, observation noise, communication noise, and the employed 50 51nonlinearity. The results reveal interesting interplay among these different problem dimensions. Most notably, we show that, provided that the nonlinearity has uniformly bounded outputs, the nonlinear estimator converges a.s. and achieves a finite asymp-53 totic covariance, even when the communication noise has no finite moments of order 54 α for any $\alpha > 1$. We then demonstrate that, in the same regime, the corresponding 56 linear consensus+innovations estimator has an infinite asymptotic covariance. We further provide several studies in the finite communication noise variance case that highlight the regimes where employing the nonlinearity strictly improves performance 58of consensus+innovations estimation over linear schemes. Typically, there is a threshold on the communication noise variance above which the nonlinear scheme achieves 60 a strictly better performance over a linear counterpart.

62 We now review existing literature to help us contrast our contributions with respect to existing work. There has been extensive work on consensus+innovation 63 distributed estimation, e.g., [17, 15, 16] and related distributed estimation methods, 64 e.g., [20, 22, 23, 27, 31, 24, 38]. For example, reference [17] derives distributed estima-65 tors for both linear and nonlinear observation models, and establishes a.s. convergence 66 and asymptotic normality of the methods under a general setting for inter-agent com-67 68 munication and observation noises. Specifically, their network model accounts for random link failures and dithered quantization, which, from the analysis perspective, 69 effectively translates into an additive communication noise. Reference [15] considers 70 consensus+innovations distributed estimation in the presence of random link fail-71ures without quantization or additive noise and develops estimators that are asymp-72totically efficient, i.e., that achieve the best achievable asymptotic covariance. The 73 74 authors of [16] propose adaptive asymptotically efficient estimators, wherein the innovation gains are adaptively learned during the algorithm progress. There have been 75several recent works that consider robust distributed estimation in the presence of 76impulsive observation (sensing) noise; see [26] for a very recent survey and the ref-77 78 erences therein. To develop robust estimators, various techniques have been utilized, including, e.g., distributed estimators based on Wilcoxon norm, e.g., [19], Huber loss, 79 e.g., [21], and mean error minimization, e.g., [36], and novel robust variants of gradi-80 ent descent [30]. Reference [1] also considers distributed recursive estimation in the 81 presence of heavy-tail (impulsive) sensing (observation) noise and develops a distrib-82 83 uted estimator that seeks the unknown parameter while at the same time identifying the optimal error nonlinearity. Reference [6] considers distributed estimation under 84 85 measurement attacks. In this setting, the authors develop a consensus+innovations estimator that employs a saturation nonlinearity in the *innovations update*. Refer-86 ences [1, 6] utilize nonlinearities in the *innovations update* to combat the *observation* 87 attacks or heavy-tail noise. This is in contrast with the current paper that employs a 88 general nonlinearity in the *consensus update* to combat the heavy-tail communication 89

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noise. Reference [7] (see also [35]) considers robust distributed estimation methods 90 91 based on adaptive subgradient projections. They are also not concerned with combating the effect of heavy-tail inter-agent communication noise. There have also been 92 several works on consensus+innovations and related distributed detection methods, 93 e.g., [25, 3, 2, 14]. In particular, reference [14] considers consensus+innovations 94 95 distributed detection in the presence of Gaussian additive communication noise. In summary, with respect to existing work on consensus+innovations distributed infer-96 ence, we employ for the first time a general nonlinearity in the consensus update, we 97 allow for the first time for heavy-tail additive communication noise, and establish for 98 the considered setting strong convergence guarantees, namely a.s. convergence and 99 asymptotic normality. 100

101 The idea of employing a nonlinearity into a "baseline" linear scheme has also been used in nonlinear versions of the standard average consensus algorithm, e.g., [18, 33, 9]. 102Average consensus is a distributed algorithm that compute a network-wide average 103 of scalar values, e.g. [5, 10, 11]. In more detail, the authors of [18] introduce a 104trigonometric nonlinearity into a standard linear consensus dynamics and show an 105improved dependence of the method on initial conditions. References [33, 9] employ a 106 107 general nonlinearity in the linear consensus dynamics and show that it improves the method's resilience to additive communication noise. The above works are different 108from ours as they focus on the average consensus problem, where the observations are 109 given to agents beforehand; the corresponding consensus algorithms hence involve only 110 a consensus step and not an innovation step in the iterative update rule. In contrast, 111 112we consider here the consensus+innovations framework, where new observations are 113 assimilated at each time instant (algorithm iteration). This technically leads to a very different analysis with respect to [18, 33, 9], and to qualitatively very different 114 results. For example, asymptotic performance of the nonlinear consensus+innovations 115estimators is determined by an interplay between the effects of network topology, 116 observation noise and communication noise; observation noise is a model dimension 117 118 not present in standard average consensus.

There have also been works that employ a specific nonlinearity in the consensus 119 update within distributed optimization problems. In this context, the authors of [34] 120 modify the linear consensus update by taking out from the averaging operation the 121maximal and minimal estimates among the estimates from all neighbors of an agent. 122Reference [4] employs the sign nonlinearity in the consensus update part for distrib-123124uted consensus optimization. The works [4, 34] contrast from ours in that they employ a specific nonlinearity, while we consider a general nonlinearity class. Furthermore, 125these works assume deterministic functions in the corresponding distributed consen-126 sus optimization problem, that effectively translates into having the observation data 127128 available beforehand. On the other hand, we consider a streaming data scenario that corresponds to the innovations update part in the algorithm we study. 129

Paper organization. Section 2 describes the distributed estimation model that we consider and presents the nonlinear consensus+innovations estimator that we propose. Section 3 explains our main results on the almost sure convergence and the asymptotic normality of the proposed distributed estimator. Section 4 provides several analytical and numerical examples that demonstrate benefits of the proposed nonlinear estimator over the linear counterpart in high and heavy-tail noise regimes. Finally, Section 5 concludes the paper.

137 **Notation**. We denote by \mathbb{R} the set of real numbers and by \mathbb{R}^m the *m*-dimensional 138 Euclidean real coordinate space. We use normal lower-case letters for scalars, lower 139 case boldface letters for vectors, and upper case boldface letters for matrices. Further,

to represent a vector $\mathbf{a} \in \mathbb{R}^m$ through its component, we write $\mathbf{a} = [\mathbf{a}_1, \mathbf{a}_2, ..., \mathbf{a}_m]^\top$ 140 and we denote by: \mathbf{a}_i or $[\mathbf{a}]_i$, as appropriate, the *i*-th element of vector \mathbf{a} ; \mathbf{A}_{ij} or 141 $[\mathbf{A}]_{ij}$, as appropriate, the entry in the *i*-th row and *j*-th column of a matrix $\mathbf{A}; \mathbf{A}^{\top}$ 142 the transpose of a matrix \mathbf{A} ; \otimes the Kronecker product of matrices. Further, we use 143either $\mathbf{a}^{\top}\mathbf{b}$ or $\langle \mathbf{a}, \mathbf{b} \rangle$ for the inner products of vectors \mathbf{a} and \mathbf{b} . Next, we let $\mathbf{I}, \mathbf{0}$, 144 and 1 be, respectively, the identity matrix, the zero vector, and the column vector 145 with unit entries. Further, $Diag(\mathbf{a})$ is the diagonal matrix whose diagonal entries are 146 the elements of vector **a**; $Tr(\mathbf{A})$ the trace of matrix **A**; **J** the $N \times N$ matrix **J** := 147 $(1/N)\mathbf{1}\mathbf{1}^{\mathsf{T}}$. When appropriate, we indicate the matrix or vector dimension through 148a subscript. Next, $\mathbf{A} \succ 0$ ($\mathbf{A} \succeq 0$) means that the symmetric matrix A is positive 149definite (respectively, positive semi-definite). We further denote by: $\|\cdot\| = \|\cdot\|_2$ the 150151 Euclidean (respectively, spectral) norm of its vector (respectively, matrix) argument; $\lambda_i(\cdot)$ the *i*-th smallest eigenvalue; g'(v) the derivative evaluated at v of a function 152 $g: \mathbb{R} \to \mathbb{R}; \nabla h(\mathbf{w})$ and $\nabla^2 h(\mathbf{w})$ the gradient and Hessian, respectively, evaluated at 153w of a function $h: \mathbb{R}^m \to \mathbb{R}, m > 1; \mathbb{P}(\mathcal{A})$ and $\mathbb{E}[u]$ the probability of an event \mathcal{A} and 154expectation of a random variable u, respectively; and by sign(a) the sign function, 155i.e., sign(a) = 1, for a > 0, sign(a) = -1, for a < 0, and sign(0) = 0. Finally, for two 156positive sequences η_n and χ_n , we have: $\eta_n = O(\chi_n)$ if $\limsup_{n \to \infty} \frac{\eta_n}{\chi_n} < \infty$. 157

2. Model and Algorithm. Subsection 2.1 explains the network and observation models that we assume. Subsection 2.2 presents the nonlinear consensus+innovations distributed estimator that we propose and states the technical assumptions needed for subsequent analysis presented in Section 3.

162 **2.1. Problem model.** Consider a network of N agents (sensors). Each agent i163 at each time t = 0, 1, ..., collects a linear transformation of the parameter of interest 164 $\boldsymbol{\theta}^* \in \mathbb{R}^M$, corrupted by noise, as follows:

 $z_i^t = \mathbf{h}_i^\top \boldsymbol{\theta}^* + n_i^t.$

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Here, $z_i^t \in \mathbb{R}$ is the observation, $\mathbf{h}_i \in \mathbb{R}^M$ is the deterministic, non-zero linear transformation vector and $n_i^t \in \mathbb{R}$ is a scalar zero-mean noise. The above update in (2.1) can be written in a compact form as follows:

 $\begin{array}{l} 170 \quad (2.2) \\ \end{array} \qquad \qquad \mathbf{z}^{t} = \mathbf{H} \left(\mathbf{1}_{N} \otimes \boldsymbol{\theta}^{*} \right) + \mathbf{n}^{t}. \end{array}$

Here, $\mathbf{z}^t = [z_1^t, z_2^t, ..., z_N^t]^\top \in \mathbb{R}^N$ is the observation vector. **H** is the $N \times (MN)$ matrix whose *i*-th row vector equals $[\mathbf{0}, ..., \mathbf{0}, \mathbf{h}_i^\top, \mathbf{0}, ..., \mathbf{0}] \in \mathbb{R}^{MN}$, where the *i*-th block of size M equals \mathbf{h}_i^\top , and the other M-size blocks are zero vectors; and $\mathbf{n}^t =$ $[n_1^t, n_2^t, ..., n_N^t]^\top \in \mathbb{R}^N$ is the noise vector at time t.

The agents constitute a network G = (V, E), where $V = \{1, ..., N\}$ is the set of agents, 176and E is the set of (undirected) inter-agent communication links (edges) $\{i, j\}$. For 177 future reference, introduce the $N \times N$ graph Laplacian matrix **L**, defined by $\mathbf{L} = \mathbf{D} - \mathbf{A}$, 178where **D** is the degree matrix and **A** is the adjacency matrix. That is, $\mathbf{D} = \text{Diag}(\{d_i\})$, 179where d_i is the degree (number of neighbors) of agent i, and A is a zero-one symmetric 180 matrix with zero diagonal, such that, for $i \neq j$, $\mathbf{A}_{ij} = 1$ if and only if $\{i, j\} \in E$. Also, 181 denote by Ω_i the set of neighbors of agent *i* (excluding *i*). For an undirected edge 182 $\{i, j\} \in E$, we denote by (i, j) the arc that points from j to i, and similarly, (j, i) is 183the arc that points from i to j. Following this convention, the communication noise 184injected when agent i communicates to agent i will be indexed by subscript ij (see 185ahead (2.3)). 186

187 **2.2.** Proposed algorithm and technical assumptions. The agents perform 188 an iterative consensus+innovations distributed algorithm to collaboratively estimate 189 the unknown vector parameter $\boldsymbol{\theta}^* \in \mathbb{R}^M$ in the presence of noisy communication links.

We assume that communication noise may be heavy-tailed, e.g., [8, 32, 13, 37, 12, 9]. 190

191To combat the heavy-tail communication noise, we introduce for the first time a

nonlinear consensus step in consensus+innovations-type methods. More precisely, 192

the proposed distributed estimator is as follows. At each time t = 0, 1, ..., each agent 193

i updates its estimate $\mathbf{x}_i^t \in \mathbb{R}^M$ of the parameter $\boldsymbol{\theta}^*$ in the following fashion: 194

195 (2.3)
$$\mathbf{x}_{i}^{t+1} = \mathbf{x}_{i}^{t} - \alpha_{t} \left(\frac{b}{a} \sum_{j \in \Omega_{i}} \Psi \left(\mathbf{x}_{i}^{t} - \mathbf{x}_{j}^{t} + \boldsymbol{\xi}_{ij}^{t} \right) - \mathbf{h}_{i} \left(z_{i}^{t} - \mathbf{h}_{i}^{\top} \mathbf{x}_{i}^{t} \right) \right)$$
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Here, $\alpha_t = a/(t+1)$ is a step-size, a, b > 0 are constants, $\boldsymbol{\xi}_{ij}^t \in \mathbb{R}^M$ is a zero-mean additive communication noise that models the imperfect communication from agent 197 198 j to agent i. Next, $\Psi : \mathbb{R}^M \to \mathbb{R}^M$ is a non-linear map that operates component-wise 199on any vector as follows: 200

$$\Psi(\mathbf{y}_1, \mathbf{y}_2, ..., \mathbf{y}_M) = [\Psi(\mathbf{y}_1), \Psi(\mathbf{y}_2), ..., \Psi(\mathbf{y}_M)]^\top,$$

where, abusing notation, $\Psi : \mathbb{R} \to \mathbb{R}$ is a component-wise non-linear function. With 203 algorithm (2.3), upon reception of the noisy version of agent j's parameter estimate $\widehat{\mathbf{x}}_{ij}^t = \mathbf{x}_j^t - \boldsymbol{\xi}_{ij}^t$, agent *i* applies the nonlinearity $\boldsymbol{\Psi} : \mathbb{R}^M \to \mathbb{R}^M$ on the consensus 204205contribution $(\mathbf{x}_i^t - \hat{\mathbf{x}}_{ij}^t)$. Intuitively, the role of Ψ is to combat the communication 206noise effect (e.g., truncate large values) while maintaining sufficient useful information flow. When in algorithm (2.3) we set $\Psi : \mathbb{R}^M \to \mathbb{R}^M$ to be the identity map, we 207208recover the \mathcal{LU} (linear estimator) in [17]. 209

For future reference, we write algorithm (2.3) in compact form.

Let $\mathbf{x}^t = [\mathbf{x}_1^t, \mathbf{x}_2^t, ..., \mathbf{x}_N^t]^\top \in \mathbb{R}^{MN}$. Furthermore, for $\mathbf{x} = [\mathbf{x}_1, \mathbf{x}_2, ..., \mathbf{x}_N]^\top \in \mathbb{R}^{MN}$ and $\boldsymbol{\xi} = [\boldsymbol{\xi}_1, \boldsymbol{\xi}_2, ..., \boldsymbol{\xi}_N]^\top \in \mathbb{R}^{MNN}$, where $\boldsymbol{\xi}_i = [\boldsymbol{\xi}_{i1}, \boldsymbol{\xi}_{i2}, ..., \boldsymbol{\xi}_{iN}]^\top \in \mathbb{R}^{MN}$ and $\boldsymbol{\xi}_{ij} = 0$ if 211212 $j \notin \Omega_i$, define $\mathbf{L}_{\Psi}(\mathbf{x}, \boldsymbol{\xi})$ by 213

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 $\begin{bmatrix} \vdots \\ \vdots \end{bmatrix}$ That is, the map $\mathbf{L}_{\Psi}(\mathbf{x}, \boldsymbol{\xi}) : \mathbb{R}^{MN} \times \mathbb{R}^{MNN} \to \mathbb{R}^{MN}$ stacks the *N* vectors of size *M*, $\sum_{j \in \Omega_i} \Psi(\mathbf{x}_i - \mathbf{x}_j + \boldsymbol{\xi}_{ij}), \ i = 1, 2, ..., N, \text{ one on top of another. Then, algorithm (2.3)}$ 216 217

219 (2.4)
$$\mathbf{x}^{t+1} = \mathbf{x}^t - \alpha_t \left(\frac{b}{a} \mathbf{L}_{\Psi}(\mathbf{x}^t, \boldsymbol{\xi}^t) - \mathbf{H}^\top \left(\mathbf{z}^t - \mathbf{H} \mathbf{x}^t \right) \right),$$

for $t = 0, 1, \dots$. 221

We make the following assumptions on the underlying network, non-linear map, ob-222 servation noise, and communication noise. The assumed nonlinearity class is similar 223224 to that in [29].

- Assumtion 2.1. Network model: 225
- Graph G = (V, E) is undirected, simple and static. 226

Assumption 2.2. Nonlinearity Ψ : 227

The non-linear function $\Psi : \mathbb{R} \to \mathbb{R}$ satisfies the following properties: 228

- 1. Function Ψ is odd, i.e., $\Psi(a) = -\Psi(-a)$, for any $a \in \mathbb{R}$; 229
- 2. $\Psi(a) > 0$, for any a > 0; 230
- 3. Function Ψ is a monotonically nondecreasing function; 231
- 4. Ψ is continuous, except possibly on a point set with Lebesque measure of 232 zero. Moreover, Ψ is piecewise differentiable. 233
- Also, $\Psi : \mathbb{R} \to \mathbb{R}$ satisfies one of the following two properties: 234

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235	5. $ \Psi(a) \leq c_1(1+ a)$, for any $a \in \mathbb{R}$, for some constant $c_1 > 0$;
236	5'. $ \Psi(a) \leq c_2$, for some constant $c_2 > 0$.
237	There are many interesting examples of nonlinearities that satisfy Assumption 2.2,
238	including, e.g., the following:
239	• (NL1) Sign function: $\Psi(a) = \operatorname{sign}(a);$
240	• (NL2) Saturation or clipping function: $\Psi(a) = a$, for $ a \le m$; and $\Psi(a) =$
241	$m \operatorname{sign}(a)$, for $ a > m$, for some constant $m > 0$;
242	• (NL3) Relay function with insensitivity zone: $\Psi(a) = 0$, for $ a \leq r$; and
243	$\Psi(a) = \operatorname{sign}(a)$, for $ a > r$, for some constant $r > 0$.
244	Assumtion 2.3. Observation model:
245	1. For each agent $i = 1,, N$, the observation noise sequence $\{n_i^t\}$ in (2.1), is
246	zero-mean and independent identically distributed (i.i.d.);
247	2. Random variables n_i^t and n_i^s are mutually independent whenever the tuple
248	(i,t) is different from (j,s) ;
249	3. Random variable n_i^t has a finite variance equal to σ_{obs}^2 , for any $t = 0, 1,$
250	and for any $i = 1,, N$;
251	4. The matrix $\sum_{i=1}^{N} \mathbf{h}_i \mathbf{h}_i^{\top}$ is invertible.
252	The condition 4 in Assumption 2.3 is a standard global observability assumption,
253	see. e.g. [17]; if it does not hold, then a central estimator that collects all observations
254	according to (2.1) for each $t = 0, 1,$ and for each $i = 1,, N$, is not able to provide
255	a consistent sequence of estimates over times $t = 0, 1,$
256	Assumtion 2.4. Communication noise:
257	1. Additive communication noise $\{\boldsymbol{\xi}_{ij}^t\}, \boldsymbol{\xi}_{ij}^t \in \mathbb{R}^M$ in (2.3), is i.i.d. in time t ,
258	independent of the observation noise family $\{n_i^t\}, i = 1,, N, t = 0, 1, and$
259	independent across different arcs (i, j) of graph G ;
260	2. Each random variable $[\boldsymbol{\xi}_{ij}^t]_{\ell}$, for each $t = 0, 1,$ for each arc (i, j) , for each
261	entry $\ell = 1,, M$, has the same cumulative distribution function Φ ;
262	3. The distribution function Φ is symmetric, i.e., for all $a \in \mathbb{R}$ we have that
263	$\Phi(a) = 1 - \Phi(-a)$, and has strictly positive second moment.
264	We assume that at least one of the conditions 4. or $4'$. below holds.
265	4. Function Ψ is strictly increasing (from Assumption 2.2) and functions Φ and
266	Ψ have a common growth point, i.e.,
267	$\Psi(a_0 + \varepsilon) \ge \Psi(a_0 - \varepsilon),$
268	$[\Phi_{ij}]_l(a_0 + \varepsilon) \ge [\Phi_{ij}]_l(a_0 - \varepsilon),$
270	for some $a_0 \in \mathbb{R}$ and all $\varepsilon > 0$;
271	4'. Distribution Φ has a pdf $p(u), p : \mathbb{R} \to \mathbb{R}$, that is strictly unimodal, i.e., there
272	holds $p(0) < +\infty$ and $p(u_1) < p(u_2)$ for $ u_1 > u_2 $;
273	5. There holds that $\int a d\Phi(a) < \infty$, and the communication noise is zero-mean,
274	i.e., $\int a d\Phi(a) = 0;$
275	6. If part 5 of Assumption 2.2 holds, then we additionally require that commu-
276	nication noise has a finite variance, i.e.:
	$\int \frac{1}{2} 4\pi \langle x \rangle$
277	$\int a^{2} d\Phi(a) < \infty;$
278	7. Distribution Φ has a well-defined pdf $p : \mathbb{R} \to \mathbb{R}$ in the vicinity of discontinuity
279	points of function $\Psi : \mathbb{R} \to \mathbb{R}$ from Assumption 2.2.
280	For notational simplicity and a clearer presentation, we assume that the com-
281	munication noise has the same distribution Φ across all arcs (i, j) such that $\{i, j\} \in$
282	E. We additionally assume that each element of communication noise vector $[\xi_{ij}^t]_{\ell}$,

 $\ell = 1, 2, ..., M$, has the same cumulative distribution function Ψ , and that $[\boldsymbol{\xi}_{ij}^t]_{\ell}$ and [$\boldsymbol{\xi}_{ij}^t]_s$ are mutually independent for $\ell \neq s$. Extensions to heterogeneous choices of nonlinearity Ψ across links and heterogeneous communication noises with mutually dependent $[\boldsymbol{\xi}_{ij}^t]_{\ell}$ and $[\boldsymbol{\xi}_{ij}^t]_s$ for $\ell \neq s$, are presented in Remark 1 in Section 3.1 (see also Supplementary material C). Similarly, we assume that the observation noise has the same variance across all agents *i*; analogous extensions to different agents' observation noise variances can be performed as well.

3. Main results. Subsection 3.1 states and proves almost sure convergence of the proposed nonlinear consensus+innovations distributed estimator in (2.3). Subsection 3.2 establishes asymptotic normality of the estimator and evaluates the corresponding asymptotic variance.

3.1. Almost sure convergence. We have the following Theorem.

THEOREM 3.1 (Almost sure convergence). Let Assumptions 2.1-2.4 hold. Then, for each agent i = 1, ..., N, the sequence of iterates $\{\mathbf{x}_i^t\}$ generated by algorithm (2.3) converges almost surely to the true vector parameter $\boldsymbol{\theta}^*$.

Theorem 3.1 establishes, for a nonlinearity Ψ with bounded outputs (e.g., the 298nonlinearities NL1-3 introduced in Section 2), almost sure convergence of the pro-299posed algorithm (2.3) under heavy-tail communication noise that may not have finite 300 301 moments of order greater than one. In contrast, it can be shown that the corresponding linear \mathcal{LU} scheme in [17] (obtained by taking Ψ to be the identity function in (2.3)) 302 generates a sequence of iterates with unbounded second moments for all t = 1, 2, ...303 (see Supplementary material B). The Theorem also establishes almost sure conver-304 gence of (2.3) for nonlinearities with unbounded outputs, more precisely, those that 305 satisfy part 5 of Assumption 2.2, when the communication noise has finite second 306 307 moment. As a special case, by taking Ψ to be the identity map, we recover for the letter case almost sure convergence of the linear estimator (the \mathcal{LU} algorithm) in [17]. 308 Setting up the proof. We next outline our strategy for proving Theorem 3.1. 309 We base our analysis on stochastic approximation arguments. More precisely, we 310 use Theorem 29 in [17] adapted from [28] (see also Theorem 3 in the supplementary 311 material) to establish a.s. convergence of \mathbf{x}^t to $\mathbf{1}_N \otimes \boldsymbol{\theta}^*$ by verifying assumptions 312 313 B1–B5 of Theorem 29 in [17].

The proof strategy is as follows. We first prove a.s. convergence of algorithm (2.3) for the case without communication noise, i.e., by setting $\boldsymbol{\xi}_{ij}^t \equiv 0$ in (2.3). In this setting, we first prove the result assuming a continuous function $\Psi : \mathbb{R} \to \mathbb{R}$. Then, we handle the case with discontinuous Ψ by additionally assuming that we can associate to $\Psi : \mathbb{R} \to \mathbb{R}$ a "lower bound" surrogate function $\underline{\Psi} : \mathbb{R} \to \mathbb{R}$ that is *continuous*, satisfies assumption 2.2, and the following holds:

320 (3.1)
$$|\Psi(a)| \ge |\Psi(a)|, \text{ for any } a \in \mathbb{R}.$$

This enables us to complete the proof for the noiseless case. To transition to the noisy communications case, a key argument is to consider an auxiliary function $\varphi : \mathbb{R} \to \mathbb{R}$, defined by

$$\begin{array}{l} 324\\ 325 \end{array} \quad (3.2) \qquad \qquad \varphi(a) = \int \Psi(a+w) d\Phi(w). \end{array}$$

Intuitively, $\varphi : \mathbb{R} \to \mathbb{R}$ is a convolution-like transformation of nonlinearity $\Psi : \mathbb{R} \to \mathbb{R}$, where the convolution is taken with respect to the communication noise cumulative distribution function Φ .

As we will demonstrate ahead, function $\varphi : \mathbb{R} \to \mathbb{R}$ in the noisy communications case effectively plays the role that function $\Psi : \mathbb{R} \to \mathbb{R}$ has in the noiseless case. Moreover,

function φ inherits all the key properties of function Ψ . More precisely, we exploit 331 the following Lemma in [29] (see Lemmas 1-6 in [29]). 332

LEMMA 3.2 ([29]). Consider function φ in (3.2), where function $\Psi : \mathbb{R} \to \mathbb{R}$, 333 satisfies Assumption 2.2. Then, the following holds: 334

1. φ is odd; 335

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2. If $|\Psi(\nu)| \leq c_1$, for any $\nu \in \mathbb{R}$, then $|\varphi(a)| \leq c'_2$, for any $a \in \mathbb{R}$, for some 337 $c_1' > 0;$

3. If $|\Psi(\nu)| \leq c_2(1+|\nu|)$, for any $\nu \in \mathbb{R}$, then $|\varphi(a)| \leq c'_2(1+|a|)$, for any $a \in \mathbb{R}$, 338 for some $c'_2 > 0$; 339

- 4. $\varphi(a)$ is monotonically nondecreasing; 340
- 5. $\varphi(a) > 0$, for any a > 0. 341
- 6. φ is continuous at zero; 342

7. φ is differentiable at zero, with a strictly positive derivative at zero, equal to: 343

(3.3)
$$\varphi'(0) = \sum_{i=1}^{s} \left(\Psi(\nu_i + 0) - \Psi(\nu_i - 0)\right) p(\nu_i) + \sum_{i=0}^{s} \int_{\nu_i}^{\nu_{i+1}} \Psi'(\nu) p(\nu) d\nu,$$

where $\nu_i, i = 1, ..., s$ are points of discontinuity of Ψ such that $\nu_0 = -\infty$ 345and $\nu_{s+1} = +\infty$, and we recall that p(u) is the pdf of distribution Φ (see 346 Assumption 2.2). 347

Lemma 3.2 allows that the treatment of the noisy case becomes completely analogous 348 to the noiseless case, by replacing function Ψ with φ . Finally, to address the case 349 when φ may not be continuous over \mathbb{R} , we make use of the following Lemma that is 350 a trivial corollary of Lemma 3.2. 351

352 LEMMA 3.3. Consider φ in (3.2). Then, there exists a positive constant ξ such that $|\varphi(a)| \ge \frac{1}{2}\varphi'(0) |a|$, for $|a| \le \xi$. 353

Lemma 3.3 allows us to define a continuous function $\underline{\varphi} : \mathbb{R} \mapsto \mathbb{R}$, $\underline{\varphi}(a) = \begin{cases} \frac{1}{2}\varphi'(0) a &, & |a| \leq \xi \\ \xi \operatorname{sign}(a) &, & \operatorname{else} \end{cases},$ 354

that satisfies Assumption 2.2 and obeys the property 357

(3.4) $|\varphi(a)| \geq |\varphi(a)|$, for any $a \in \mathbb{R}$. 358

Function φ will then clearly play the role of function $\underline{\Psi}$ in (3.1) in the noiseless case. 359 We are now ready to prove Theorem 3.1. 360

Proof. (Proof of Theorem 3.1) 361

Step 1: No communication noise. We start the proof by verifying conditions 362 B1–B5 of Theorem 29 in [17] for the case without communication noise. We use the 363 following Lyapunov function $V: \mathbb{R}^{MN} \to \mathbb{R}, V(\mathbf{x}) = ||\mathbf{x} - \mathbf{1}_N \otimes \boldsymbol{\theta}^*||^2$. For this, only 364 365 for condition B3, we need to analyze separately the case with continuous Ψ and the case when Ψ may not be continuous. Also, it can be shown that (2.3) can be put in 366 the form required by Theorem 29 in [17] (see also (36) in the supplementary material) 367 by letting 368

369 (3.5)
$$\mathbf{r}(\mathbf{x}) = -\mathbf{H}^{\top}\mathbf{H}\left(\mathbf{x} - \mathbf{1}_N \otimes \boldsymbol{\theta}^*\right) - \frac{b}{a}\mathbf{L}_{\boldsymbol{\Psi}}(\mathbf{x}, \mathbf{0}),$$

370 (3.6)
$$\boldsymbol{\gamma}(t+1,\mathbf{x},\omega) = \mathbf{H}^{\top}\mathbf{n}^{t},$$

where ω denotes an element of the underlying probability space. 372

Consider the filtration \mathcal{F}_t , t = 1, 2, ..., where \mathcal{F}_t is the σ - algebra generated by $\{\mathbf{n}^s\}_{s=0}^{t-1}$. 373 Denote by $(\Omega, \mathcal{F}, \mathbb{P})$ the underlying probability space that generates random vectors

- 374
- \mathbf{n}^t , t = 0, 1, 2, ..., and by $\omega \in \Omega$ its arbitrary element. Clearly, for each t, function 375 $\gamma(t+1,\cdot,\cdot)$ is $\mathcal{B}^{MN} \otimes \mathcal{F}$ measurable, where \mathcal{B}^{MN} is the Borel sigma algebra on \mathbb{R}^{MN} . 376

Also, $\mathbf{r}(\cdot)$ is \mathcal{B}^{MN} measurable. Hence, condition B1 holds. Further, the family of random vectors $\boldsymbol{\gamma}(t+1, \mathbf{x}, \omega)$ is \mathcal{F}_t measurable, zero-mean and independent of \mathcal{F}_{t-1} . Thus, condition B2 holds.

We now inspect condition B3. Assume first that function $\Psi : \mathbb{R} \to \mathbb{R}$ is continuous. The gradient of V equals $\nabla V(\mathbf{x}) = 2(\mathbf{x} - \mathbf{1}_N \otimes \boldsymbol{\theta}^*)$. Clearly, function $V(\cdot)$ is twice continuously differentiable and has uniformly bounded second order partial derivatives. We consider

$$S = \sup_{\substack{\|\mathbf{x}-\mathbf{1}_N \otimes \boldsymbol{\theta}^*\| \in (\epsilon, 1/\epsilon)}} \langle \mathbf{r}(\mathbf{x}), \nabla V(\mathbf{x}) \rangle,$$

We will show that S < 0, thus verifying condition B3. We have, for any $\mathbf{x} \in \mathbb{R}^{MN}$:

387
$$\langle \mathbf{r}(\mathbf{x}), \nabla V(\mathbf{x}) \rangle = -2 \left(\mathbf{x} - \mathbf{1}_N \otimes \boldsymbol{\theta}^* \right)^\top \left(\mathbf{H}^\top \mathbf{H} \left(\mathbf{x} - \mathbf{1}_N \otimes \boldsymbol{\theta}^* \right) + \frac{b}{a} \mathbf{L}_{\Psi}(\mathbf{x}) \right)$$

388 (3.8)
$$= -2\underbrace{\left(\left(\mathbf{x} - \mathbf{1}_N \otimes \boldsymbol{\theta}^*\right)^\top \mathbf{H}^\top \mathbf{H} \left(\mathbf{x} - \mathbf{1}_N \otimes \boldsymbol{\theta}^*\right)\right)}_{T_1(\mathbf{x})} - 2\frac{b}{a}\underbrace{\left(\mathbf{x} - \mathbf{1}_N \otimes \boldsymbol{\theta}^*\right)^\top \mathbf{L}_{\Psi}(\mathbf{x})}_{T_2(\mathbf{x})}.$$

Clearly $T_1 = T_1(\mathbf{x}) \ge 0$. We will also show that $T_2 = T_2(\mathbf{x}) \ge 0$. Utilizing the fact that, $\Psi(\cdot)$ is an odd function, we have that,

392 (3.9)
$$T_2 = \sum_{\{i,j\}\in E, i < j} (\mathbf{x}_i - \mathbf{x}_j)^\top \Psi (\mathbf{x}_i - \mathbf{x}_j) \ge 0,$$

as for $\mathbf{g} = (\mathbf{x}_i - \mathbf{x}_j)$, we have that,

395 (3.10)
$$\left(\mathbf{x}_{i} - \mathbf{x}_{j}\right)^{\top} \boldsymbol{\Psi} \left(\mathbf{x}_{i} - \mathbf{x}_{j}\right) = \sum_{\ell=1}^{M} \mathbf{g}_{\ell} \boldsymbol{\Psi} \left(\mathbf{g}_{\ell}\right) \ge 0,$$

397 because \mathbf{g}_{ℓ} and $\Psi(\mathbf{g}_{\ell})$ have the same sign, by Assumption 2.2. Therefore,

$$\langle \mathbf{r}(\mathbf{x}), \nabla V(\mathbf{x}) \rangle = -2 T_1 - 2 \frac{b}{a} T_2 \le 0,$$

 $\begin{array}{ll} 399\\ 400 & \text{for any } \mathbf{x} \in \mathbb{R}^{MN}. \end{array}$

398

401 We will further show that S in (3.7) is strictly less than 0. First, consider the set 402 $C = \{ \mathbf{x} \in \mathbb{R}^{MN} : \|\mathbf{x} - \mathbf{1}_N \otimes \boldsymbol{\theta}^*\| \in [\epsilon, 1/\epsilon] \}$. Note that set C is nonempty and compact. 403 Clearly, we have that:

404 (3.11)
$$S \leq S_{\mathcal{C}} := \sup_{\mathbf{x} \in \mathcal{C}} \langle \mathbf{r}(\mathbf{x}), \nabla V(\mathbf{x}) \rangle$$

It is thus sufficient to show that $S_{\mathcal{C}} < 0$. Suppose the contrary is true, i.e., suppose 405that $S_{\mathcal{C}} = 0$. As set \mathcal{C} is compact and function $\mathbf{x} \mapsto \langle \mathbf{r}(\mathbf{x}), \nabla V(\mathbf{x}) \rangle$ is continuous, by the 406Weierstrass theorem, we have that $S_{\mathcal{C}} = 0$ is equivalent to having $\langle \mathbf{r}(\mathbf{x}^{\bullet}), \nabla V(\mathbf{x}^{\bullet}) \rangle = 0$, 407 for some point $\mathbf{x}^{\bullet} \in \mathcal{C}$. In this case, \mathbf{x}^{\bullet} has to be of the form, $\mathbf{x}^{\bullet} = \mathbf{1}_N \otimes \mathbf{m}$, where $\mathbf{m} \in \mathcal{C}$. 408 \mathbb{R}^{M} . As otherwise, we would have that, T_{2} is strictly positive. But then, we have, $T_{1} =$ 409 $\left(\left(\mathbf{x}^{\bullet} - \mathbf{1}_N \otimes \boldsymbol{\theta}^*\right)^{\top} \mathbf{H}^{\top} \mathbf{H} \left(\mathbf{x}^{\bullet} - \mathbf{1}_N \otimes \boldsymbol{\theta}^*\right)\right) = \left(\mathbf{m} - \boldsymbol{\theta}^*\right)^{\top} \left(\sum_{i=1}^N \mathbf{h}_i \mathbf{h}_i^{\top}\right) \left(\mathbf{m} - \boldsymbol{\theta}^*\right) > 0,$ 410 which is a contradiction in view of (3.7). Hence, we conclude that, for a continuous 411 function Ψ , it holds that S < 0, and that condition B3 holds, i.e., 412 $\sup_{\|\mathbf{x}-\mathbf{1}_N\otimes\boldsymbol{\theta}^*\|\in(\epsilon,1/\epsilon)}\langle \mathbf{r}(\mathbf{x}),\nabla V(\mathbf{x})\rangle<0.$ 413

⁴¹⁴ $\|\mathbf{x}-\mathbf{I}_N \otimes \boldsymbol{\theta}^*\| \in (\epsilon, 1/\epsilon)$ ⁴¹⁵ Now, we verify condition B3 for function Ψ that is not continuous but to which we can ⁴¹⁶ associate function $\underline{\Psi}$ that obeys Assumption 2.2 and for which condition (3.1) holds. ⁴¹⁷ Then, the verification of condition B3 follows analogously to the case with continuous ⁴¹⁸ Ψ by replacing T_2 in (3.9) with the following lower bound of T_2

410
$$(2,10)$$
 T T $\sum_{i=1}^{n} (1, -i)^{T} \mathbf{\Psi}(1, -i)$

419 (3.12)
$$\underline{T}_2 = \sum_{\{i,j\}\in E, i< j} (\mathbf{x}_i - \mathbf{x}_j)^{\top} \underline{\Psi} (\mathbf{x}_i - \mathbf{x}_j),$$

420 where $\underline{\Psi}(\mathbf{a}) = [\underline{\Psi}(\mathbf{a}_1), \underline{\Psi}(\mathbf{a}_2), ..., \underline{\Psi}(\mathbf{a}_M)]^{\top}$. Hence, condition B3 is verified.

We next verify condition B4. Recalling the definition of $\mathbf{r}(\mathbf{x})$ in (3.5), we have, 421 $\parallel \mathbf{T}_{\mathbf{r}}$ $\rightarrow \parallel 2$

$$\|\mathbf{r}(\mathbf{x})\|^{2} \leq c_{3}V(\mathbf{x}) + c_{4}\|\mathbf{\Psi}(\mathbf{x})\|^{2}$$

where $c_3 = 2a^2 \|\mathbf{H}^{\top}\mathbf{H}\|^2$ and $c_4 = 2b^2 \|\mathbf{L}\|^2$. 424We also have that. 425

426
$$\|\Psi(\mathbf{x})\| \le c_5 \sum_{\{i,j\}\in E} (|\mathbf{x}_i - \boldsymbol{\theta}^*| + |\mathbf{x}_j - \boldsymbol{\theta}^*|) + c_6 \le c_7 \|\mathbf{x} - \mathbf{1}_N \otimes \boldsymbol{\theta}^*\| + c_6,$$

- for some positive constants c_5, c_6, c_7 . Therefore, we have (3.14) $\|\Psi(\mathbf{x})\|^2 \leq 2c_7 V(\mathbf{x}) + 2c_8^2$, 428
- 430

for some positive constant c_8 . 431

Thus, we have that 432

$$\left\|\mathbf{r}(\mathbf{x})\right\|^2 \le c_9 V(\mathbf{x}) + c_{10},$$

form some positive constants c_9, c_{10} . Recall $\gamma(t+1, \mathbf{x}, \omega)$ in (3.6). Using the bound-435edness of the second moment of the observation noise, we finally have that, 436

$$\|\mathbf{r}(\mathbf{x})\|^2 + \mathbb{E}\left[\|\boldsymbol{\gamma}(t+1,\mathbf{x}^t,\omega)\|^2\right] \le c_{11}\left(V(\mathbf{x})+1\right)$$

for some positive constant c_{11} . Hence, condition B4 is satisfied. Finally, condition B5 439 clearly holds. Therefore, we conclude that $\mathbf{x}^t \to \mathbf{1}_N \otimes \boldsymbol{\theta}^*$, almost surely. 440

Step 2: The case with communication noise. We proceed by considering algo-441 rithm (2.3) under communication noise. 442

We clarify the steps needed to transition from the noiseless to the noisy case. If we 443 444 write

448

 $437 \\ 438$

$$\Psi(\mathbf{x}_i^t - \mathbf{x}_j^t + \boldsymbol{\xi}_{ij}^t) = arphi(\mathbf{x}_i^t - \mathbf{x}_j^t) + \eta_{ij}^t,$$

where $\boldsymbol{\eta}_{ij}^t = \left[\boldsymbol{\Psi}(\mathbf{x}_i^t - \mathbf{x}_j^t + \boldsymbol{\xi}_{ij}^t) - \boldsymbol{\varphi}(\mathbf{x}_i^t - \mathbf{x}_j^t) \right]$ and $\boldsymbol{\varphi} : \mathbb{R}^M \to \mathbb{R}^M$ is component-wise map 447 defined as $\varphi(\mathbf{x}_1, \mathbf{x}_2, ..., \mathbf{x}_M) = [\varphi(\mathbf{x}_1), \varphi(\mathbf{x}_2), ..., \varphi(\mathbf{x}_M)]^\top$. We will see that quantity 448 $\boldsymbol{\eta}_{ij}^t$ is a key ingredient of $\boldsymbol{\gamma}(t+1,\mathbf{x},\omega)$ in Theorem 29 in [17] (see also Theorem 3 in 449the supplementary material). 450

The algorithm (2.3) can be written in compact form: 451

$$\begin{array}{l} 452 \quad (3.15) \quad \mathbf{x}^{t+1} = \mathbf{x}^t - \alpha_t \left(\frac{b}{a} \mathbf{L}_{\varphi}(\mathbf{x}^t) - \mathbf{H}^T(\mathbf{z}^t - \mathbf{H}\mathbf{x}^t) + \frac{b}{a} \boldsymbol{\eta}^t \right). \end{array}$$

454Here,

455 (3.16)
$$\mathbf{L}_{\boldsymbol{\varphi}}(\mathbf{x}^t) = \begin{bmatrix} \vdots \\ \sum_{j \in \Omega_i} \boldsymbol{\varphi}(\mathbf{x}_i^t - \mathbf{x}_j^t) \\ \vdots \end{bmatrix} \in \mathbb{R}^{MN}, \quad \boldsymbol{\eta}^t = \begin{bmatrix} \vdots \\ \sum_{j \in \Omega_i} \boldsymbol{\eta}_{ij}^t \\ \vdots \end{bmatrix} \in \mathbb{R}^{MN},$$

456

where the $M \times 1$ blocks $\sum_{j \in \Omega_i} \varphi(\mathbf{x}_i^t - \mathbf{x}_j^t)$ and $\sum_{j \in \Omega_i} \eta_{ij}^t$ are stacked one on top of another 457458 for j = 1, ..., N.

The differences of (3.15) with respect to the case without additive communication 459460

noise are that \mathbf{L}_{φ} replaces \mathbf{L}_{Ψ} and the term $\frac{b}{a} \alpha_t \boldsymbol{\eta}^t$ is added. We define the Lyapunov function $V : \mathbb{R}^{MN} \to \mathbb{R}$, and quantities $\mathbf{r}_{\varphi}(x)$ and $\boldsymbol{\gamma}_{\varphi}(t, \mathbf{x}, \omega)$ 461 as follows: 462

463 (3.17)
$$V(\mathbf{x}) = ||\mathbf{x} - \mathbf{1}_N \otimes \boldsymbol{\theta}^*||^2$$

464 (3.18)
$$\mathbf{r}_{\varphi}(\mathbf{x}) = -\mathbf{H}^T \mathbf{H}(\mathbf{x} - \mathbf{1}_N \otimes \boldsymbol{\theta}^*) - \frac{b}{a} \mathbf{L}_{\varphi}(\mathbf{x}),$$

 $\boldsymbol{\gamma}_{\boldsymbol{\varphi}}(t+1,\mathbf{x},\omega) = \mathbf{H}^{\top}\mathbf{n}^t - \frac{b}{a}\boldsymbol{\eta}^t,$ (3.19)465 466

Now, make the following identification with respect to the transition from the noiseless 467

to the noisy case. Quantity $\mathbf{H}^{\top}\mathbf{n}^{t}$ in the noiseless case is replaced with quantity 468

469 $\mathbf{H}^{\top}\mathbf{n}^{t} - \frac{b}{a}\boldsymbol{\eta}^{t}$ in the noisy case. The map $\mathbf{L}_{\Psi}(\cdot, \mathbf{0}) : \mathbb{R}^{MN} \to \mathbb{R}^{MN}$ in (2.4) is replaced 470 with the map $\mathbf{L}_{\varphi} : \mathbb{R}^{MN} \to \mathbb{R}^{MN}$ given in (3.16).

The proof proceeds analogously by again verifying Assumptions B1-B5. We only 471 clarify the differences in verifying these conditions with respect to the noiseless case. 472The filtration \mathcal{F}_t is replaced with the filtration \mathcal{G}_t , t = 1, 2, ..., which is generated 473not only by $\{\mathbf{n}^s\}_{s=0}^{t-1}$ but also by $\{\boldsymbol{\xi}_{ij}^s\}_{s=0}^{t-1}$ for $(i,j) \in E$. Clearly, for each t, function 474 $\gamma_{\boldsymbol{\varphi}}(t+1;\cdot;\cdot)$ is $\mathcal{B}^{MN}\otimes\mathcal{F}$ measurable. Also, $\mathbf{r}_{\boldsymbol{\varphi}}(\cdot)$ is \mathcal{B}^{MN} measurable. Hence, condition 475B1 holds. Further, the family of random vectors $\gamma_{\omega}(t+1, \mathbf{x}, \omega)$ is \mathcal{F}_t measurable, 476zero-mean and independent of \mathcal{F}_{t-1} . Thus, condition B2 holds. As function φ is 477 odd, non-decreasing, strictly positive for its positive arguments, and has a positive 478 derivative at zero by Lemma 3.2, condition B3 is derived analogously to the noiseless 479case. Conditions B4 and B5 hold analogously to the noiseless case. Thus, the result 480is verified. 481 Γ

482 **Remark 1:** Theorem 3.1 continues to hold under the following generalizations:

- A different nonlinear function $\Psi_{ij,\ell} : \mathbb{R} \to \mathbb{R}$ is assigned to each arc (i, j) and to each element $\ell = 1, ..., M$ of the communication noise $[\boldsymbol{\xi}_{ij}^t]_{\ell}$. Each function $\Psi_{ij,\ell}$ obeys Assumption 2.2.
- The observation noise $\sigma_{obs,i}^2$ is different for each agent i = 1, 2, ..., N.
- The communication noise $\boldsymbol{\xi}_{ij}^t$ has the joint cumulative distribution function 488 $\boldsymbol{\Phi}_{ij}$ such that:
- 489

490

$$\int_{\mathbf{a}\in\mathbb{R}^M} \|\mathbf{a}\| d\Phi_{ij}(\mathbf{a}) < \infty, \quad \int_{\mathbf{a}\in\mathbb{R}^M} \mathbf{a} d\Phi_{ij}(\mathbf{a}) = 0,$$

491 and $\mathbf{\Phi}_{ij}(\mathbf{a}) = 1 - \mathbf{\Phi}_{ij}(-\mathbf{a})$, for all $\mathbf{a} \in \mathbb{R}^M$.

492 All the remaining assumptions in 2.1-2.4 continue to hold.

Note that the above means that the communication noise $\boldsymbol{\xi}_{ij}^t$ may have mutually dependent elements $[\boldsymbol{\xi}_{ij}^t]_{\ell}$, for $\ell = 1, ..., M$.

For the above generalization, it can be shown that Theorem 3.1 continues to hold (see Supplementary material C).

497 **3.2.** Asymptotic normality. We now present our results on asymptotic nor-498 mality of estimator (2.3).

THEOREM 3.4 (Asymptotic normality). Let Assumptions 2.1 – 2.4 hold. Consider algorithm (2.3) with step-size $\alpha_t = a/(t+1)$, t = 0, 1, ..., a > 0. Then, the normalized sequence of iterates $\{\sqrt{t+1}(\mathbf{x}^t - \mathbf{1}_N \otimes \boldsymbol{\theta}^*)\}$ converges in distribution to a zero-mean multivariate normal random vector, i.e., the following holds:

583
$$\sqrt{t+1}(\mathbf{x}^t - \mathbf{1}_N \otimes \boldsymbol{\theta}^*) \Rightarrow \mathcal{N}(\mathbf{0}, \mathbf{S}),$$

505 where the asymptotic covariance matrix \mathbf{S} equals:

506 (3.20)
$$\mathbf{S} = a^2 \int_{0}^{\infty} e^{\mathbf{\Sigma} v} \mathbf{S}_0 e^{\mathbf{\Sigma}^{\top} v} dv$$

507 Here, $\mathbf{S}_0 = \sigma_{obs}^2 \mathbf{H}^\top \mathbf{H}^+ \frac{b^2}{a^2} \sigma^2 \operatorname{Diag}\left(\{d_i \mathbf{I}_M\}\right)$, where we recall that d_i is the degree of 508 agent i; $\sigma^2 = \int |\Psi(w)|^2 d\Phi(w)$ is the effective communication noise variance after 509 passing through the nonlinearity Ψ ; we recall the observation matrix \mathbf{H} in (2.2); the 510 observation noise variance σ_{obs}^2 in (2.1); function ϕ in (3.2); and $\Sigma = \frac{1}{2}\mathbf{I} - a(\mathbf{H}^\top \mathbf{H} + \frac{b}{a}\varphi'(0)(\mathbf{L}\otimes\mathbf{I}_M))$, where a is taken large enough such that matrix Σ is stable (i.e., real 512 parts of Σ 's eigenvalues are negative).

513 Theorem 3.4 shows that, for the communication noise with finite variance and un-

bounded nonlinearities that satisfy part 5 of Assumption 2.2, the variance with the proposed nonlinear estimator (2.3) decays (in the weak convergence sense) at (the best achievable) rate O(1/t). In particular, by taking Ψ to be the identity function, we recover the asymptotic normality result in [17] of the corresponding linear estimator (the \mathcal{LU} scheme in [17]). Note also that the asymptotic variance expression in (3.20)

for the indentity function $\Psi(a) = a$ coincides with that in [17] for the \mathcal{LU} scheme.

- Theorem 3.4 further demonstrates that, even under a heavy-tailed communication 520 noise (with unbounded variance) with a bounded nonlinearity (e.g., nonlinearities
- NL1-3 in Section 2), the variance with algorithm (2.3) still decays at rate O(1/t). In contrast, the corresponding linear scheme (obtained by taking Ψ in (2.3) to be the
- identity function) generates a sequence with unbounded variances for each t = 1, 2, ...

More precisely, we then have that $E[||\mathbf{x}^t - \mathbf{1}_N \otimes \boldsymbol{\theta}^*||^2] = \infty$, for any t = 1, 2, ... (see Supplementary material C).

527 Theorem 3.4 explicitly quantifies asymptotic variance of (2.3). This also allows, in

the finite communication noise regime, to compare the nonlinear versus the linear scheme (when both schemes achieve a finite asymptotic variance). See Subsection 4.1 for details.

531 Theorem 3.4 also reveals an interesting tradeoff when including the nonlinearity Ψ into the consensus update. On the one hand, nonlinearity makes a beneficial effect in 532that the communication noise plays the role only through the effective variance $\sigma^2 =$ 533 $\int |\Psi(w)|^2 d\Phi(w)$. In contrast, with the linear scheme, σ^2 is replaced with $\int w^2 d\Phi(w)$ 534that is infinite under a heavy tail setting. On the other hand, the nonlinearity Ψ 536 makes a negative effect in that it "reduces quality" of matrix Σ through the quantity $\varphi'(0)$ that is typically less than one with a nonlinear scheme and equal to one with the 537 linear scheme. Clearly, the tradeoff goes in favor of the nonlinear scheme in the heavy 538 tail setting (finite variance with the nonlinear estimator versus infinite variance with the linear estimator). In the finite communication noise variance setting, the nonlinear 540scheme typically improves performance under a sufficiently low communication signal 541542 to noise ratio (SNR); see also Subsection 4.1. We are now ready to prove Theorem 3.4. 543

544 *Proof.* (Proof of Theorem 3.4) We establish asymptotic normality by verifying 545 assumptions C1-C5 of Theorem 29 in [17] (see also Theorem 3 in the supplementary 546 material). Firstly, we show that condition C1 hold. Since function φ is differentiable 547 at zero, we have that

548 (3.21)
$$\varphi(a) = \varphi(0) + \varphi'(0)a + \Delta(a) = \varphi'(0)a + \Delta(a),$$

where for the function $\Delta : \mathbb{R} \to \mathbb{R}$, we have that $\lim_{a\to 0} \frac{\Delta(a)}{a} = 0$. Hence, the function $\mathbf{r}_{\varphi}(\mathbf{x})$ admits representation as in Theorem 29 of [17] (see also (37) of Theorem 3 in the supplementary material), with matrix

$$\mathbf{B} = -\mathbf{H}^T \mathbf{H} - \frac{b}{a} \varphi'(0) \big[\mathbf{L} \otimes \mathbf{I}_M \big],$$

and function $\boldsymbol{\delta} : \mathbb{R}^{MN} \to \mathbb{R}^{MN}$, given with $\boldsymbol{\delta}(\mathbf{x}) = -\frac{b}{a} \mathbf{L}_{\boldsymbol{\Delta}}(\mathbf{x})$. Here, function $\mathbf{L}_{\boldsymbol{\Delta}}(\mathbf{x}) :$ 556 $\mathbb{R}^{MN} \to \mathbb{R}^{MN}$ is defined by

557
$$\mathbf{L}_{\Delta}(\mathbf{x}) = \begin{bmatrix} \vdots \\ \sum_{j \in \Omega_i} \Delta(\mathbf{x}_i - \mathbf{x}_j) \\ \vdots \end{bmatrix}$$

558

where function $\boldsymbol{\Delta} : \mathbb{R}^M \to \mathbb{R}^M$ is defined by (3.21), $\boldsymbol{\Delta}(\mathbf{y}_1, \mathbf{y}_1, ..., \mathbf{y}_M) = [\Delta(\mathbf{y}_1), \Delta(\mathbf{y}_2), ..., \Delta(\mathbf{y}_M)]^\top$, 560 $\mathbf{y} \in \mathbb{R}^M$. 561 Condition C2 trivially holds, if we use that $\alpha_t = \frac{a}{t+1}$. Furthermore, $\Sigma = a\mathbf{B} + \frac{1}{2}\mathbf{I}$ is 562 stable if *a* is large enough, because matrix $-\mathbf{B}$ is positive definite (see [17]). Thus, 563 condition C3 also holds.

For $\mathbf{A}(t, \mathbf{x}) = \mathbb{E}[\boldsymbol{\gamma}_{\boldsymbol{\varphi}}(t+1, \mathbf{x}, \omega) \boldsymbol{\gamma}_{\boldsymbol{\varphi}}^{\top}(t+1, \mathbf{x}, \omega)]$, using the Lebesgue's dominated convergence theorem, it can be shown that

$$\lim_{t \to \infty, \mathbf{x} \to \boldsymbol{\theta}^*} \mathbf{A}(t, \mathbf{x}) = \sigma_{\text{obs}}^2 \mathbf{H}^\top \mathbf{H} + \sigma^2 \operatorname{Diag}\left(\{d_i \, \mathbf{I}_M\}\right)$$

Therefore, condition C4 also holds. It remains to verify condition C5. Recall quantity $\gamma_{\varphi}(t+1, \mathbf{x}, \omega)$ in (3.19). Note that this condition is equivalent to saying that the family of random variables $\{\|\gamma_{\varphi}(t+1, \mathbf{x}, \omega)\|^2\}_{t=0,1,...,\|\mathbf{x}-\theta^{\star}\|<\epsilon}$ is uniformly integrable. If the condition 5 in Assumption 2.2 holds (the case with finite communication noise variance and the nonlinearity with unbounded outputs), then: (3.22) $\|\gamma_{\varphi}(t+1, \mathbf{x}, \omega)\|^2 \leq c_{12} + c_{13} \|\mathbf{n}^t\|^2 + c_{14} \|\boldsymbol{\eta}^t\|^2$,

574 for some positive constants c_{12}, c_{13}, c_{14} .

575 Consider the family $\{\widetilde{\mathbf{g}}(t+1,\mathbf{x},\omega)\}_{t=0,1,\dots,\|\mathbf{x}-\boldsymbol{\theta}^{\star}\|<\epsilon}$, with

576 (3.23) $\widetilde{\mathbf{g}}(t+1,\mathbf{x},\omega) = c_{12} + c_{13} \|\mathbf{n}^t\|^2 + c_{14} \|\boldsymbol{\eta}^t\|^2.$

577 Clearly, $\tilde{\mathbf{g}}(t+1, x, \omega)$ is integrable, for any t = 0, 1, ..., for any $\epsilon > 0$, due to 578 the finite second moment of sensing and observation noises. The family { $\tilde{\mathbf{g}}(t + 1, x, \omega)$ } $_{t=0,1,...,\|\mathbf{x}-\boldsymbol{\theta}^{\star}\|<\epsilon}$ is i.i.d. and hence it is uniformly integrable. The family 580 { $\|\boldsymbol{\gamma}_{\boldsymbol{\varphi}}(t+1,x,\omega)\|^2$ } $_{t=0,1,...,\|\mathbf{x}-\boldsymbol{\theta}^{\star}\|<\epsilon}$ is dominated by { $\tilde{\mathbf{g}}(t+1,x,\omega)$ } $_{t=0,1,...,\|\mathbf{x}-\boldsymbol{\theta}^{\star}\|<\epsilon}$ 581 that is uniformly integrable, and hence { $\|\boldsymbol{\gamma}_{\boldsymbol{\varphi}}(t+1,x,\omega)\|^2$ } $_{t=0,1,...,\|\mathbf{x}-x^{\star}\|<\epsilon}$ is also uni-582 formly integrable. An analogous argument can be applied if condition 5' in Assump-583 tion 2.2 holds (bounded nonlinearity, communication noise with infinite variance). 584 Hence, condition C5 holds; thus, the result.

4. Analytical and numerical examples. Subsection 4.1 provides analytical examples, and Subsection 4.2 provides simulation examples, that illustrate the main results presented in Section 3.

4.1. Analytical examples. We provide several analytical examples that illustrate Theorem 3.4. The examples demonstrate that, in the considered setting, the proposed nonlinear method in (2.3) achieves a lower asymptotic variance than the corresponding linear scheme, for a low SNR regime, i.e., for the case when the communication noise variance is above a threshold. We also consider optimization of the nonlinearity Ψ for a given nonlinearity class; more precisely, for the given analytical example, we consider optimization of parameter *B* for the NL2 nonlinearity class in Section 2.

Example 1: We follow a setup similar to [17], but we consider the nonlinear consensus+innovations scheme in (2.3), with the non-linear operator $\Psi : \mathbb{R} \to \mathbb{R}$ of the following form (the NL2 nonlinearity):

599 (4.1)
$$\Psi(w) = \begin{cases} w , & |w| \le B \\ +B , & w > B \\ -B , & w < B \end{cases},$$

600

685

for some parameter B > 0. Notice that letting $B \to \infty$ in (4.1) leads to the linear consensus+innovations \mathcal{LU} scheme in [17].

Each agent *i* observes a scalar parameter $\theta^* \in \mathbb{R}$ according to:

$$z_i(t) = h\theta^* + n_i^t,$$

where $h \neq 0$ and n_i^t is i.i.d. in time and across sensors with variance σ_{obs}^2 and zero mean. Communication noise is i.i.d. across arcs and in time and is independent of $\{n_i^t\}$, for all i = 1, 2, ..., N. Assume that the communication noise has a probability

distribution function f(w) that is strictly positive in the vicinity of zero. Denote the 609 eigenvalues of **L** by $0 = \lambda_1 < \lambda_2 \leq \cdots \leq \lambda_N$. Let the graph be regular, for simplicity, 610 with degree d. Using Theorem 3.4, we have that the asymptotic covariance matrix 611 612 equals:

$$\mathbf{S} = a^2 \int\limits_0^\infty e^{\mathbf{\Sigma} v} \mathbf{S}_0 e^{\mathbf{\Sigma} v} dv.$$

Here, $\mathbf{S}_0 = \left(h^2 \sigma_{\text{obs}}^2 + \frac{b^2}{a^2} d\sigma^2\right) \mathbf{I}$; also, recall $\sigma^2 = \int_{-\infty}^{\infty} |\Psi(w)|^2 f(w) dw$, the effective 615 communication noise per link. We assume that f(w) has a zero mean and variance 616

 $\sigma_{\text{comm}}^2 = \int_{-\infty}^{\infty} w^2 f(w) dw$ that is finite. Also, 617

$$\Sigma = \frac{1}{2}\mathbf{I} - a\left(h^{2}\mathbf{I} + \frac{b}{a}\varphi'(0)\mathbf{L}\right)$$

where φ is given in (3.2). For the nonlinearity considered here, we have that 620

621
$$\sigma^2 = 2 \int_{0}^{+B} w^2 f(w) dw + B^2 \left(1 - 2 \int_{0}^{+B} f(w) dw \right),$$
$$+B$$

622
$$\varphi'(0) = 2 \int_{0}^{\infty} f(w) dw.$$

Denote by $\sigma_B^2 = \frac{1}{N} \operatorname{Tr}(\mathbf{S})$ the average per-agent asymptotic variance. Analogously to (76)-(86) in [17], for $a > \frac{1}{2h^2}$ we get: 624625

$$\sigma_B^{26} = \frac{a^2 h^2 \sigma_{\text{obs}}^2 + b^2 d\sigma^2}{N \left(2ah^2 - 1\right)} + \frac{a^2 h^2 \sigma_{\text{obs}}^2 + b^2 d\sigma^2}{N} \sum_{i=2}^N \frac{1}{2b\lambda_i \varphi'(0) + (2ah^2 - 1)}$$

We next analyze the values of σ_B^2 as $B \to 0$ and $B \to +\infty$. For $B \to 0$, we have that $\sigma^2 \to 0$, $\varphi'(0) \to 0$ and 628

629

$$\sigma_B^2 \to \frac{a^2 h^2 \sigma_{\rm obs}^2}{2ah^2 - 1} =: \sigma_0^2$$

That is, when $B \to 0$, we effectively have the case that each agent is working in 632 isolation, hence not seeing the effect of the communication noise. 633

For $B \to +\infty$, we have that $\varphi'(0) \to 1$, $\sigma^2 \to \sigma^2_{\text{comm}}$ and 634

$$\begin{array}{ccc} {}^{635} & \sigma_B^2 \to \frac{a^2 h^2 \sigma_{\rm obs}^2 + b^2 d \sigma_{\rm comm}^2}{N \left(2ah^2 - 1\right)} + \frac{a^2 h^2 \sigma_{\rm obs}^2 + b^2 d \sigma_{\rm comm}^2}{N} \sum_{i=2}^N \frac{1}{2b\lambda_i + (2ah^2 - 1)} =: \sigma_\infty^2. \end{array}$$

This is the asymptotic variance of the linear \mathcal{LU} scheme in [17]. Note that, for any 637

638 set of values of system parameters and any
$$a > \frac{1}{2h^2}$$
 and $b > 0$, there holds that
638 (4.2) $\sigma_{\infty}^2 > \sigma_0^2$

641

for a sufficiently large σ_{comm}^2 . Assume from now on that (4.2) holds. It can be shown that there exists an optimal B, i.e., there exists B^* such that $B^* \in (0, +\infty)$ and $\inf_{B \in (0, +\infty)} \sigma_B^2 = \sigma_{B^*}^2$ (see Supple-642 643

mentary material D). 644

Note that the above analysis generalizes also to the case when 645

646
$$\sigma_{\text{comm}}^2 = \int_{-\infty}^{+\infty} w^2 f(w) dw = +\infty,$$

630

i.e., when the noise variance is $+\infty$. In this case, we have that $\sigma_{\infty}^2 = +\infty$, for the linear scheme and $\sigma_0^2 = \frac{a^2 h^2 \sigma_{obs}^2}{2ah^2 - 1}$ for the isolation scheme. It can be shown that $\inf_{B \in (0, +\infty)} \sigma_B^2$ is achieved at some $B^* \in (0, +\infty)$ (see Supplementary material D).

In order to demonstrate the results above, we minimize σ_B^2 and calculate B^* for a specific numerical example (see Figure 1a). We consider a sensor (agents) network with N = 8 agents, where the underlying topology is given by a regular graph with degree d = 3. We set innovation and consensus constants as a = b = 1, the observation parameter h = 1, and the true parameter $\theta^* = 1$. The observation noise for each sensor's measurements is standard normal, and the communication noise for each communication link has the following pdf

658 (4.3)
$$f(w) = \frac{\beta - 1}{2(1 + |w|)^{\beta}}$$

with $\beta = 2.05$. (This pdf's distribution has the infinite variance.) Figure 1b shows 660 performance of the nonlinear consensus+innovations estimator (2.3) in terms of the 661 estimated per-sensor mean squared error (MSE) across iterations, for the optimal B^* 662 and for some sub-optimal choices of B, obtained through a Monte Carlo simulation. 663 664 We can see that the scheme with B^* performs better than for the considered suboptimal choices of B. Figure 1c shows that Monte Carlo estimate of the per-agent 665 asymptotic variance, i.e., $\hat{S} = \frac{1}{N} \| \mathbf{x}^t - \mathbf{1}_N \otimes \boldsymbol{\theta}^* \|^2 t$ matches well the corresponding 666 theoretical value as per Theorem 3.4. 667



FIG. 1. (a) Per-agent asymptotic variance σ_B^2 versus B for the nonlinear consensus+ innovations estimator and the NL2 nonlinearity. (b) Monte Carlo-estimated per-sensor MSE error on logarithmic scale for the nonlinear consensus+innovations estimator with the NL2 nonlinearity for different choices of B. (c) Monte Carlo estimate of the per-agent asymptotic variance, and the corresponding theoretical value as per Theorem 3.4.

Example 2: We consider the same network and sensing models as in Example 1 and the heavy-tail communication noise distribution in (4.3). Furthermore, we assume that $\Psi(w) = \operatorname{sign}(w)$ (the NL3 nonlinearity). For the \mathcal{LU} scheme, it can be shown that (see Supplementary material E):

672
$$\sigma^2 = \sigma_{\rm comm}^2 = \frac{2}{(\beta - 3)(\beta - 2)},$$

$$\varphi'(0) = 1.$$

It can be shown here that the average per-agent asymptotic variance $\sigma_{\rm L}^2 = \frac{1}{N} \operatorname{Tr}(\mathbf{S})$ for the \mathcal{LU} scheme is equal to

677 (4.4)
$$\sigma_{\rm L}^2 = \begin{cases} \infty , & 2 < \beta \le 3, \\ \frac{a^2 h^2 \sigma_{\rm obs}^2 + b^2 d\sigma^2}{N(2ah^2 - 1)} + \frac{a^2 h^2 \sigma_{\rm obs}^2 + b^2 d\sigma^2}{N} \sum_{i=2}^N \frac{1}{2b\lambda_i + (2ah^2 - 1)} , & \beta > 3. \end{cases}$$

679 For $\beta > 3$, quantity $\sigma_{\rm L}^2$ can be written as

680 (4.5)
$$\sigma_{\rm L}^2 = A_{\rm L} + B_{\rm L} \frac{1}{(\beta - 3)(\beta - 2)},$$

682 where

16

$$A_{\rm L} = \frac{a^2 h^2 \sigma_{\rm obs}^2}{N(2ah^2 - 1)} + \frac{a^2 h^2 \sigma_{\rm obs}^2}{N} \sum_{i=2}^N \frac{1}{2b\lambda_i + (2ah^2 - 1)}$$

683

$$B_{\rm L} = 2\left(\frac{b^2d}{N(2ah^2 - 1)} + \frac{b^2d}{N}\sum_{i=2}^{N}\frac{1}{2b\lambda_i + (2ah^2 - 1)}\right)$$

685 We next consider the nonlinear consensus+innovations scheme with the nonlinearity 686 $\Psi(w) = \operatorname{sign} w$. We have that 687

$$\sigma^{2} = 1,$$

$$\varphi(a) = 2 \int_{0}^{a} f(w) dw,$$

689

$$\varphi(a) = 2\int_{0}^{a} f(a) da$$

690 which means that $\varphi'(a) = 2f(a)$ and $\varphi'(0) = 2f(0) = (\beta - 1)$. Hence, we have that 691 the average per-agent asymptotic variance for the nonlinear scheme $\sigma_{\rm NL}^2 = \frac{1}{N} \operatorname{Tr}(S)$ 692 693 is given by:

694 (4.6)
$$\sigma_{\rm NL}^2 = \frac{a^2 h^2 \sigma_{\rm obs}^2 + b^2 d\sigma^2}{N \left(2ah^2 - 1\right)} + \frac{a^2 h^2 \sigma_{\rm obs}^2 + b^2 d\sigma^2}{N} \sum_{i=2}^N \frac{1}{4b\lambda_i f(0) + (2ah^2 - 1)},$$

which can be written in the form 696

697 (4.7)
$$\sigma_{\rm NL}^2 = A_{\rm NL} + B_{\rm NL} \frac{P_{N-2}(\beta)}{\prod_{i=2}^{N} (\beta - \beta_i)},$$

where 699

700

703704

$$A_{
m NL} = rac{a^2 h^2 \sigma_{
m obs}^2 + b^2 d}{N \left(2 a h^2 - 1
ight)}$$

701
$$B_{\rm NL} = \frac{a^2 h^2 \sigma_{\rm obs}^2 + b^2 d}{N \prod_{i=2}^{N} 2b\lambda_i},$$

702
$$P_{N-2}(\beta) = \sum_{\substack{i=2\\j\neq i}}^{N} \prod_{\substack{j=2\\j\neq i}}^{N} 2b\lambda_j(\beta - \beta_j).$$

$$\beta_i = 1 - \frac{2ah^2 - 1}{2b\lambda_i}, \quad i = 2, ..., N$$

We next compare the average per-agent asymptotic variances for the linear consen-705sus+innovations scheme and the nonlinear consensus+innovations scheme. From (4.4)706it is obvious that $\sigma_{\rm NL}^2 < \sigma_{\rm L}^2$ for $\beta \in (2,3]$. For $\beta > 3$, if $A_{\rm L} \gg A_{\rm NL}$ (see Supplementary material E), the linear scheme is worse than the nonlinear scheme for all $\beta > 3$. 707 708It is obvious that $\sigma_{\rm L}^2$ decreases on interval $(3, \infty)$ and $\sigma_{\rm NL}^2$ decreases on the interval (β_m, ∞) , where $\beta_m = \max_{i=2,\dots,N} \beta_i < 1$ is closest β_i to 1. Function $\sigma_{\rm L}^2 = \sigma_{\rm L}^2(\beta)$ has an 709710 asymptote at $\beta = 3$, and function $\sigma_{\rm NL}^2 = \sigma_{\rm NL}^2(\beta)$ at $\beta = \beta_m$, where $\beta_m < 3$, also, 711

 $A_{\rm L}$ and $A_{\rm NL}$ are horizontal asymptotes for $\sigma_{\rm L}^2$ and $\sigma_{\rm NL}^2$, respectively. Therefore, if $A_{\rm L}$ is much larger than $A_{\rm NL}$, $\sigma_{\rm L}^2$ is above $\sigma_{\rm NL}^2$ for all $\beta > 3$. Moreover, if $A_{\rm L} < A_{\rm NL}$ there exists $\beta^* > 3$ such that the average per-agent asymptotic variance is still better 712 713 714 for the nonlinear than for the linear scheme for $\beta \in (2, \beta^*]$. Defining $k = \frac{\sigma_L^2}{\sigma_{xy}^2}$, it is 715possible to show that $k \to \infty$ as $\beta \to 3$, and $k \to \frac{A_{\rm L}}{A_{\rm NL}}$ as $\beta \to \infty$. Therefore, if $A_{\rm L} < A_{\rm NL}$, there exists β^* such that $\sigma_{\rm NL}^2 < \sigma_{\rm L}^2$ for all $\beta \in (2, \beta^*)$. In other words, 716 717 there exists a threshold value $\beta^* > 3$, such that the nonlinear scheme outperforms the 718linear scheme for the "heavy-tail regime" $\beta \in (2, \beta^*)$, and the linear scheme performs 719 better for $\beta > \beta^*$. To summarize, in Example 2, depending on sensing and network 720 parameters, it holds that either the nonlinear scheme outperforms the linear one for 721 all β , or there exists a threshold value β^* such that the nonlinear scheme is better 722 than the linear one for $\beta \in (2, \beta^*)$. Figure 2 shows the ratio $k = \frac{\sigma_L^2}{\sigma_{NL}^2}$ versus β for 723 the same sensing and network parameters as in Example 1. As it can be seen, there 724 exists a threshold β^* , that here approximately equals $\beta^* = 3.9$, such that k > 1 for 725 $\beta \in (2, \beta^*)$. On the other hand, for $\beta > \beta^*$, the ratio becomes smaller than one, which 726 means that for the given numerical parameters, the linear scheme performs better for 727 728 $\beta > \beta^*$. This is in accordance with the analysis that we provided above.



FIG. 2. Ratio $k = \frac{\sigma_{\rm L}^2}{\sigma_{\rm NL}^2}$ versus β for Example 2.

4.2. Simulation examples. In this section, we illustrate the performance of the proposed nonlinear consensus+innovations estimator for two different choices of the non-linear operator Ψ . For both nonlinearity choices, our method is compared with the corresponding linear consensus+innovations estimator \mathcal{LU} in [17], when the communication noise has probability distribution function given by (4.3).

We consider a sensor network with N = 40 agents. The underlying topology is an 734 instance of a random geometric graph. We use the same initialization $\mathbf{x}^0 = \mathbf{0}$ and same 735 step sizes $\alpha_t = \frac{1}{t+1}, a = 1, b = 1$, for both the linear and the nonlinear estimators. 736 Also, we assume that the observation noise is normally distributed, i.e., $n_i^t \sim \mathcal{N}(0, 1)$, 737 for each t, for each i. The true parameter $\theta^* \in \mathbb{R}^{10}$ is generated randomly, where 738 the entries of θ^* are drawn mutually independently from the uniform distribution 739 on [-10,10]. The observation vectors $\mathbf{h}_i \in \mathbb{R}^{10}$ are also generated at random, for 740 which the condition 4 of Assumption 2.3 is true. We use the communication noise 741 742 pdf in (4.3) with $\beta = 2.05$. Note that, in this case, the communication noise has an infinite variance. 743

Figure 4 compares the linear \mathcal{LU} estimator in [17] with the nonlinear estimator (2.3) with $\Psi(w)$ given in (4.1) for B = 5. Figure 3 shows the comparison between \mathcal{LU} and [17] with $\Psi(w) = \operatorname{sign}(w)$. Both Figures show the iteration counter t at the x-axis and a Monte-Carlo estimate of the average mean square error (MSE) across agents on the *y*-axis. We can see that, as predicted by our theory, the nonlinear estimator, for both nonlinearity choices, persistently decreases MSE along iterations, despite the fact that the communication noise has an infinite variance. At the same

751 time, \mathcal{LU} fails to produce a useful estimation result.



FIG. 3. Monte-Carlo average per-agent MSE estimate versus iteration counter on logarithmic scale for the proposed nonlinear estimator (2.3) with the nonlinearity in (4.1) for B = 5 and the linear LU scheme in [17].



FIG. 4. Monte-Carlo average per-agent MSE estimate versus iteration counter on logarithmic scale for the proposed nonlinear estimator (2.3) with the nonlinearity $\Psi(w) = \text{sign}(w)$ and the linear $\mathcal{L}\mathcal{U}$ scheme in [17].

5. Conclusion. We studied consensus+innovations distributed estimation in the 752presence of impulsive, heavy-tail communication noise. To combat the impulsive 753communication noise, we introduce for the first time a general nonlinearity in the 754755consensus update for consensus+innovations distributed estimation. We establish almost sure convergence of the nonlinear consensus+innovations estimator to the true 756 parameter, prove its asymptotic normality, and explicitly evaluate the corresponding 757 asymptotic variance. We compare the proposed nonlinear estimator with conventional 758 759 consensus+innovation estimators that utilize linear consensus update. Analytical and

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760 numerical examples demonstrate significant gains of introducing consensus nonlinear-

761 ity in low SNR (high communication noise) regimes. Most notably, we demonstrate

that, when the communication noise has infinite variance, the proposed nonlinear con-

real sensus+innovations estimator is strongly consistent (converges almost surely), while the corresponding linear counterpart provides a sequence of estimators with infinite

765 variance.

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