DISTRIBUTED RECURSIVE ESTIMATION UNDER HEAVY-TAIL 2 COMMUNICATION NOISE

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Abstract. We consider distributed recursive estimation of an unknown vector parameter $\theta^* \in \mathbb{R}^M$ in the presence of impulsive communication noise. That is, we assume that inter-agent communication is subject to an additive communication noise that may have heavy-tails or is con- taminated with outliers. To combat this effect, within the class of consensus+innovations distributed estimators, we introduce for the first time a nonlinearity in the consensus update. We allow for a general class of nonlinearities that subsumes, e.g., the sign function or component-wise saturation function. For the general nonlinear estimator and a general class of additive communication noises – that may have infinite moments of order higher than one – we establish almost sure (a.s.) convergence 13 to the parameter θ^* . We further prove asymptotic normality and evaluate the corresponding asymp- totic covariance. These results reveal interesting tradeoffs between the negative effect of "loss of information" due to incorporation of the nonlinearity, and the positive effect of communication noise reduction. We also demonstrate and quantify benefits of introducing the nonlinearity in high-noise (low signal-to-noise ratio) and heavy-tail communication noise regimes.

 Key words. Distributed inference; distributed estimation; recursive estimation; heavy-tail noise; consensus+innovations; stochastic approximation.

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1. Introduction. We consider distributed inference in networked systems, whe-22 re each agent in a generic network continuously (over time instances $t = 0, 1, \ldots$) 23 makes noisy linear observations of an unknown vector parameter $\boldsymbol{\theta}^* \in \mathbb{R}^M$. Each 24 agent, at each time t, generates a local estimate of θ^* through the so-called consen- sus+innovations strategy, i.e., by 1) weight-averaging its current solution estimate with those of its neighbors, and 2) assimilating its new observation.

 In this paper, we are interested in consensus+innovations distributed estimation in the presence of an impulsive communication noise, e.g., when the communication noise that corresponds to inter-neighbor communications is heavy-tailed or contami- nated with outliers. It is highly relevant to consider impulsive communication noise in many application scenarios. For example, edge devices in Internet of Things (IoT) systems or sensor networks can be subject to impulsive noise distributions that may not have finite moments of order higher than one, e.g., [\[8,](#page-18-0) [32,](#page-19-0) [13,](#page-18-1) [37,](#page-19-1) [12,](#page-18-2) [9\]](#page-18-3). In this work, we allow the communication noise to be a zero-mean random variable that 35 may have infinite moments of order α , for any $\alpha > 1$. In particular, communication noise may have an infinite variance. To the best of our knowledge, such scenarios have not been studied in the past work, wherein communication noise in consen- sus+innovations inference is always assumed to have a finite moment of at least second order (finite variance). Actually, as demonstrated ahead in the paper, existing con-

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40 sensus+innovation estimators – that are always *linear* in the consensus update part – can fail to converge under a heavy-tail communication noise. To combat the effect of the (impulsive or high variance) communication noise, we introduce for the first time a general nonlinearity in the consensus update. More precisely, we apply a nonlinear operator (e.g., a sign function, a saturation-like function, or a sigmoid function) on the difference between an agent's current iterate and a noisy version of its neighbor's iterate, for every agent in the neighborhood set. We establish, under a general setting for the nonlinearity and the additive communication noise, almost sure (a.s.) conver-48 gence of the nonlinear estimator to the true parameter $\boldsymbol{\theta}^*$. We also prove asymptotic normality and evaluate the corresponding asymptotic covariance in terms of the un- derlying network topology, observation noise, communication noise, and the employed nonlinearity. The results reveal interesting interplay among these different problem dimensions. Most notably, we show that, provided that the nonlinearity has uniformly bounded outputs, the nonlinear estimator converges a.s. and achieves a finite asymp- totic covariance, even when the communication noise has no finite moments of order α for any $\alpha > 1$. We then demonstrate that, in the same regime, the corresponding linear consensus+innovations estimator has an infinite asymptotic covariance. We further provide several studies in the finite communication noise variance case that highlight the regimes where employing the nonlinearity strictly improves performance of consensus+innovations estimation over linear schemes. Typically, there is a thresh- old on the communication noise variance above which the nonlinear scheme achieves a strictly better performance over a linear counterpart.

 We now review existing literature to help us contrast our contributions with re- spect to existing work. There has been extensive work on consensus+innovation distributed estimation, e.g., [\[17,](#page-19-2) [15,](#page-18-4) [16\]](#page-18-5) and related distributed estimation methods, e.g., [\[20,](#page-19-3) [22,](#page-19-4) [23,](#page-19-5) [27,](#page-19-6) [31,](#page-19-7) [24,](#page-19-8) [38\]](#page-19-9). For example, reference [\[17\]](#page-19-2) derives distributed estima- tors for both linear and nonlinear observation models, and establishes a.s. convergence and asymptotic normality of the methods under a general setting for inter-agent com- munication and observation noises. Specifically, their network model accounts for random link failures and dithered quantization, which, from the analysis perspective, effectively translates into an additive communication noise. Reference [\[15\]](#page-18-4) considers consensus+innovations distributed estimation in the presence of random link fail- ures without quantization or additive noise and develops estimators that are asymp- totically efficient, i.e., that achieve the best achievable asymptotic covariance. The authors of [\[16\]](#page-18-5) propose adaptive asymptotically efficient estimators, wherein the inno- vation gains are adaptively learned during the algorithm progress. There have been several recent works that consider robust distributed estimation in the presence of impulsive observation (sensing) noise; see [\[26\]](#page-19-10) for a very recent survey and the ref- erences therein. To develop robust estimators, various techniques have been utilized, including, e.g., distributed estimators based on Wilcoxon norm, e.g., [\[19\]](#page-19-11), Huber loss, e.g., [\[21\]](#page-19-12), and mean error minimization, e.g., [\[36\]](#page-19-13), and novel robust variants of gradi- ent descent [\[30\]](#page-19-14). Reference [\[1\]](#page-18-6) also considers distributed recursive estimation in the presence of heavy-tail (impulsive) sensing (observation) noiseand develops a distrib- uted estimator that seeks the unknown parameter while at the same time identifying the optimal error nonlinearity. Reference [\[6\]](#page-18-7) considers distributed estimation under measurement attacks. In this setting, the authors develop a consensus+innovations 86 estimator that employs a saturation nonlinearity in the *innovations update*. Refer-87 ences $\begin{bmatrix}1, 6\end{bmatrix}$ utilize nonlinearities in the *innovations update* to combat the *observation* 88 attacks or heavy-tail noise. This is in contrast with the current paper that employs a general nonlinearity in the consensus update to combat the heavy-tail communication

 noise. Reference [\[7\]](#page-18-8) (see also [\[35\]](#page-19-15)) considers robust distributed estimation methods based on adaptive subgradient projections. They are also not concerned with com-

 bating the effect of heavy-tail inter-agent communication noise. There have also been several works on consensus+innovations and related distributed detection methods, e.g., [\[25,](#page-19-16) [3,](#page-18-9) [2,](#page-18-10) [14\]](#page-18-11) . In particular, reference [\[14\]](#page-18-11) considers consensus+innovations distributed detection in the presence of Gaussian additive communication noise. In summary, with respect to existing work on consensus+innovations distributed infer- ence, we employ for the first time a general nonlinearity in the consensus update, we allow for the first time for heavy-tail additive communication noise, and establish for the considered setting strong convergence guarantees, namely a.s. convergence and asymptotic normality.

 The idea of employing a nonlinearity into a "baseline" linear scheme has also been used in nonlinear versions of the standard average consensus algorithm, e.g., [\[18,](#page-19-17) [33,](#page-19-18) [9\]](#page-18-3). Average consensus is a distributed algorithm that compute a network-wide average of scalar values, e.g. [\[5,](#page-18-12) [10,](#page-18-13) [11\]](#page-18-14). In more detail, the authors of [\[18\]](#page-19-17) introduce a trigonometric nonlinearity into a standard linear consensus dynamics and show an improved dependence of the method on initial conditions. References [\[33,](#page-19-18) [9\]](#page-18-3) employ a general nonlinearity in the linear consensus dynamics and show that it improves the method's resilience to additive communication noise. The above works are different from ours as they focus on the average consensus problem, where the observations are given to agents beforehand; the corresponding consensus algorithms hence involve only a consensus step and not an innovation step in the iterative update rule. In contrast, we consider here the consensus+innovations framework, where new observations are assimilated at each time instant (algorithm iteration). This technically leads to a very different analysis with respect to [\[18,](#page-19-17) [33,](#page-19-18) [9\]](#page-18-3), and to qualitatively very different results. For example, asymptotic performance of the nonlinear consensus+innovations estimators is determined by an interplay between the effects of network topology, observation noise and communication noise; observation noise is a model dimension not present in standard average consensus.

 There have also been works that employ a specific nonlinearity in the consensus update within distributed optimization problems. In this context, the authors of [\[34\]](#page-19-19) modify the linear consensus update by taking out from the averaging operation the maximal and minimal estimates among the estimates from all neighbors of an agent. Reference [\[4\]](#page-18-15) employs the sign nonlinearity in the consensus update part for distrib- uted consensus optimization. The works [\[4,](#page-18-15) [34\]](#page-19-19) contrast from ours in that they employ a specific nonlinearity, while we consider a general nonlinearity class. Furthermore, these works assume deterministic functions in the corresponding distributed consen- sus optimization problem, that effectively translates into having the observation data available beforehand. On the other hand, we consider a streaming data scenario that corresponds to the innovations update part in the algorithm we study.

Paper organization. Section [2](#page-3-0) describes the distributed estimation model that we consider and presents the nonlinear consensus+innovations estimator that we pro- pose. Section [3](#page-6-0) explains our main results on the almost sure convergence and the asymptotic normality of the proposed distributed estimator. Section [4](#page-12-0) provides sev- eral analytical and numerical examples that demonstrate benefits of the proposed nonlinear estimator over the linear counterpart in high and heavy-tail noise regimes. Finally, Section [5](#page-17-0) concludes the paper.

137 Notation. We denote by $\mathbb R$ the set of real numbers and by $\mathbb R^m$ the *m*-dimensional Euclidean real coordinate space. We use normal lower-case letters for scalars, lower case boldface letters for vectors, and upper case boldface letters for matrices. Further,

140 to represent a vector $\mathbf{a} \in \mathbb{R}^m$ through its component, we write $\mathbf{a} = [\mathbf{a}_1, \mathbf{a}_2, ..., \mathbf{a}_m]^\top$ 141 and we denote by: \mathbf{a}_i or $[\mathbf{a}]_i$, as appropriate, the *i*-th element of vector **a**; \mathbf{A}_{ij} or $[{\bf A}]_{ij}$, as appropriate, the entry in the *i*-th row and *j*-th column of a matrix ${\bf A}$; ${\bf A}^{\top}$ 143 the transpose of a matrix \mathbf{A} ; ⊗ the Kronecker product of matrices. Further, we use 144 either $\mathbf{a}^\top \mathbf{b}$ or $\langle \mathbf{a}, \mathbf{b} \rangle$ for the inner products of vectors **a** and **b**. Next, we let **I**, **0**, 145 and 1 be, respectively, the identity matrix, the zero vector, and the column vector 146 with unit entries. Further, Diag(a) is the diagonal matrix whose diagonal entries are 147 the elements of vector a; Tr(A) the trace of matrix A; J the $N \times N$ matrix J := $(1/N)$ 11[⊤]. When appropriate, we indicate the matrix or vector dimension through 149 a subscript. Next, $\mathbf{A} \succ 0$ ($\mathbf{A} \succ 0$) means that the symmetric matrix A is positive 150 definite (respectively, positive semi-definite). We further denote by: $\|\cdot\| = \|\cdot\|_2$ the 151 Euclidean (respectively, spectral) norm of its vector (respectively, matrix) argument; 152 $\lambda_i(\cdot)$ the *i*-th smallest eigenvalue; $g'(v)$ the derivative evaluated at v of a function 153 $g: \mathbb{R} \to \mathbb{R}; \nabla h(\mathbf{w})$ and $\nabla^2 h(\mathbf{w})$ the gradient and Hessian, respectively, evaluated at 154 w of a function $h : \mathbb{R}^m \to \mathbb{R}$, $m > 1$; $\mathbb{P}(\mathcal{A})$ and $\mathbb{E}[u]$ the probability of an event A and 155 expectation of a random variable u, respectively; and by $sign(a)$ the sign function, 156 i.e., $sign(a) = 1$, for $a > 0$, $sign(a) = -1$, for $a < 0$, and $sign(0) = 0$. Finally, for two 157 positive sequences η_n and χ_n , we have: $\eta_n = O(\chi_n)$ if $\limsup_{n \to \infty} \frac{\eta_n}{\chi_n} < \infty$.

 2. Model and Algorithm. Subsection [2.1](#page-3-1) explains the network and observa- tion models that we assume. Subsection [2.2](#page-3-2) presents the nonlinear consensus+inno- vations distributed estimator that we propose and states the technical assumptions needed for subsequent analysis presented in Section [3.](#page-6-0)

162 **2.1. Problem model.** Consider a network of N agents (sensors). Each agent i 163 at each time $t = 0, 1, \ldots$, collects a linear transformation of the parameter of interest 164 $\theta^* \in \mathbb{R}^M$, corrupted by noise, as follows:

 $z_i^t = \mathbf{h}_i^{\top} \boldsymbol{\theta}^* + n_i^t.$ 166

167 Here, $z_i^t \in \mathbb{R}$ is the observation, $\mathbf{h}_i \in \mathbb{R}^M$ is the deterministic, non-zero linear trans-168 formation vector and $n_i^t \in \mathbb{R}$ is a scalar zero-mean noise. The above update in (2.1) 169 can be written in a compact form as follows:

 $\mathbf{z}^t = \mathbf{H}\left(\mathbf{1}_N\otimes\boldsymbol{\theta}^*\right) + \mathbf{n}^t$ $\mathbf{z}^t = \mathbf{H}\left(\mathbf{1}_N\otimes\boldsymbol{\theta}^*\right) + \mathbf{n}^t.$ 171

172 Here, $\mathbf{z}^t = [z_1^t, z_2^t, ..., z_N^t]^\top \in \mathbb{R}^N$ is the observation vector. **H** is the $N \times (MN)$ 173 matrix whose *i*-th row vector equals $[\mathbf{0},...,\mathbf{0},\mathbf{h}_i^{\top},\mathbf{0},...,\mathbf{0}] \in \mathbb{R}^{MN}$, where the *i*-th 174 block of size M equals \mathbf{h}_i^{\top} , and the other M-size blocks are zero vectors; and \mathbf{n}^t = 175 $[n_1^t, n_2^t, ..., n_N^t]^\top \in \mathbb{R}^N$ is the noise vector at time t.

176 The agents constitute a network $G = (V, E)$, where $V = \{1, ..., N\}$ is the set of agents, 177 and E is the set of (undirected) inter-agent communication links (edges) $\{i, j\}$. For 178 future reference, introduce the $N \times N$ graph Laplacian matrix **L**, defined by **L** = **D**-**A**, 179 where **D** is the degree matrix and **A** is the adjacency matrix. That is, $\mathbf{D} = \text{Diag}(\{d_i\}),$ 180 where d_i is the degree (number of neighbors) of agent i, and **A** is a zero-one symmetric 181 matrix with zero diagonal, such that, for $i \neq j$, $\mathbf{A}_{ij} = 1$ if and only if $\{i, j\} \in E$. Also, 182 denote by Ω_i the set of neighbors of agent i (excluding i). For an undirected edge 183 $\{i, j\} \in E$, we denote by (i, j) the arc that points from j to i, and similarly, (j, i) is 184 the arc that points from i to j. Following this convention, the communication noise 185 injected when agent j communicates to agent i will be indexed by subscript ij (see 186 ahead [\(2.3\)](#page-4-0)).

187 2.2. Proposed algorithm and technical assumptions. The agents perform 188 an iterative consensus+innovations distributed algorithm to collaboratively estimate 189 the unknown vector parameter $\boldsymbol{\theta}^* \in \mathbb{R}^M$ in the presence of noisy communication links.

190 We assume that communication noise may be heavy-tailed, e.g., [\[8,](#page-18-0) [32,](#page-19-0) [13,](#page-18-1) [37,](#page-19-1) [12,](#page-18-2) [9\]](#page-18-3).

- 191 To combat the heavy-tail communication noise, we introduce for the first time a
- 192 nonlinear consensus step in consensus+innovations-type methods. More precisely,
- 193 the proposed distributed estimator is as follows. At each time $t = 0, 1, \dots$, each agent
- 194 *i* updates its estimate $\mathbf{x}_i^t \in \mathbb{R}^M$ of the parameter $\boldsymbol{\theta}^*$ in the following fashion:

195
$$
\mathbf{x}_i^{t+1} = \mathbf{x}_i^t - \alpha_t \left(\frac{b}{a} \sum_{j \in \Omega_i} \Psi \left(\mathbf{x}_i^t - \mathbf{x}_j^t + \boldsymbol{\xi}_{ij}^t \right) - \mathbf{h}_i \left(z_i^t - \mathbf{h}_i^\top \mathbf{x}_i^t \right) \right).
$$

197 Here, $\alpha_t = a/(t+1)$ is a step-size, $a, b > 0$ are constants, $\xi_{ij}^t \in \mathbb{R}^M$ is a zero-mean additive communication noise that models the imperfect communication from agent *j* to agent *i*. Next, $\Psi : \mathbb{R}^M \to \mathbb{R}^M$ is a non-linear map that operates component-wise on any vector as follows:

$$
\mathbf{\Psi}(\mathbf{y}_1, \mathbf{y}_2, ..., \mathbf{y}_M) = [\Psi(\mathbf{y}_1), \Psi(\mathbf{y}_2), ..., \Psi(\mathbf{y}_M)]^\top,
$$

203 where, abusing notation, $\Psi : \mathbb{R} \to \mathbb{R}$ is a component-wise non-linear function. With 204 algorithm (2.3) , upon reception of the noisy version of agent j's parameter estimate 205 $\hat{\mathbf{x}}_{ij}^t = \mathbf{x}_j^t - \boldsymbol{\xi}_{ij}^t$, agent i applies the nonlinearity $\Psi : \mathbb{R}^M \to \mathbb{R}^M$ on the consensus 206 contribution $(\mathbf{x}_i^t - \hat{\mathbf{x}}_{ij}^t)$. Intuitively, the role of Ψ is to combat the communication 207 noise effect (e.g., truncate large values) while maintaining sufficient useful information 208 flow. When in algorithm [\(2.3\)](#page-4-0) we set $\Psi : \mathbb{R}^M \to \mathbb{R}^M$ to be the identity map, we 209 recover the $\mathcal{L} \mathcal{U}$ (linear estimator) in [\[17\]](#page-19-2).

210 For future reference, we write algorithm [\(2.3\)](#page-4-0) in compact form.

211 Let $\mathbf{x}^t = [\mathbf{x}_1^t, \mathbf{x}_2^t, ..., \mathbf{x}_N^t]^\top \in \mathbb{R}^{MN}$. Furthermore, for $\mathbf{x} = [\mathbf{x}_1, \mathbf{x}_2, ..., \mathbf{x}_N]^\top \in \mathbb{R}^{MN}$ and 212 $\boldsymbol{\xi} = [\boldsymbol{\xi}_1, \boldsymbol{\xi}_2, ..., \boldsymbol{\xi}_N]^{\top} \in \mathbb{R}^{MNN}$, where $\boldsymbol{\xi}_i = [\boldsymbol{\xi}_{i1}, \boldsymbol{\xi}_{i2}, ..., \boldsymbol{\xi}_{iN}]^{\top} \in \mathbb{R}^{MN}$ and $\boldsymbol{\xi}_{ij} = 0$ if 213 $j \notin \Omega_i$, define $\mathbf{L}_{\mathbf{\Psi}}(\mathbf{x}, \boldsymbol{\xi})$ by

$$
\mathbf{L}_{\mathbf{\Psi}}(\mathbf{x}, \boldsymbol{\xi}) = \begin{bmatrix} \vdots \\ \sum_{j \in \Omega_i} \mathbf{\Psi}(\mathbf{x}_i - \mathbf{x}_j + \boldsymbol{\xi}_{ij}) \\ \vdots \end{bmatrix}.
$$

215

216 That is, the map $\mathbf{L}_{\mathbf{\Psi}}(\mathbf{x}, \boldsymbol{\xi}) : \mathbb{R}^{MN} \times \mathbb{R}^{MNN} \to \mathbb{R}^{MN}$ P stacks the N vectors of size M , 217 $\sum_{j \in \Omega_i} \Psi(\mathbf{x}_i - \mathbf{x}_j + \boldsymbol{\xi}_{ij}), i = 1, 2, ..., N$, one on top of another. Then, algorithm [\(2.3\)](#page-4-0)

218 can be written as:

$$
\mathbf{x}^{t+1} = \mathbf{x}^t - \alpha_t \left(\frac{b}{a} \mathbf{L}_{\mathbf{\Psi}}(\mathbf{x}^t, \boldsymbol{\xi}^t) - \mathbf{H}^\top \left(\mathbf{z}^t - \mathbf{H} \mathbf{x}^t \right) \right),\tag{2.4}
$$

221 for $t = 0, 1, ...$.

222 We make the following assumptions on the underlying network, non-linear map, ob-223 servation noise, and communication noise. The assumed nonlinearity class is similar 224 to that in [\[29\]](#page-19-20).

225 Assumtion 2.1. Network model:

226 Graph $G = (V, E)$ is undirected, simple and static.

227 Assumtion 2.2. Nonlinearity Ψ :

228 The non-linear function $\Psi : \mathbb{R} \to \mathbb{R}$ satisfies the following properties:

- 229 1. Function Ψ is odd, i.e., $\Psi(a) = -\Psi(-a)$, for any $a \in \mathbb{R}$;
- 230 2. $\Psi(a) > 0$, for any $a > 0$;
- 231 3. Function Ψ is a monotonically nondecreasing function;
- 232 4. Ψ is continuous, except possibly on a point set with Lebesque measure of 233 zero. Moreover, Ψ is piecewise differentiable.
- 234 Also, $\Psi : \mathbb{R} \to \mathbb{R}$ satisfies one of the following two properties:

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 $\ell = 1, 2, ..., M$, has the same cumulative distribution function Ψ , and that $[\xi_{ij}^t]_\ell$ and $[\xi_{ij}^t]_s$ are mutually independent for $\ell \neq s$. Extensions to heterogeneous choices of 285 nonlinearity Ψ across links and heterogeneous communication noises with mutually 286 dependent $[\xi_{ij}^t]_\ell$ and $[\xi_{ij}^t]_s$ for $\ell \neq s$, are presented in Remark 1 in Section [3.1](#page-6-1) (see also Supplementary material C). Similarly, we assume that the observation noise has the same variance across all agents i; analogous extensions to different agents' observation noise variances can be performed as well.

 3. Main results. Subsection [3.1](#page-6-1) states and proves almost sure convergence of the proposed nonlinear consensus+innovations distributed estimator in [\(2.3\)](#page-4-0). Sub- section [3.2](#page-10-0) establishes asymptotic normality of the estimator and evaluates the cor-responding asymptotic variance.

294 3.1. Almost sure convergence. We have the following Theorem.

295 THEOREM 3.1 (Almost sure convergence). Let Assumptions [2.1-](#page-4-2)[2.4](#page-5-1) hold. Then, 296 for each agent $i = 1, ..., N$, the sequence of iterates $\{x_i^t\}$ generated by algorithm (2.3) 297 converges almost surely to the true vector parameter $\boldsymbol{\theta}^*$.

298 Theorem [3.1](#page-6-2) establishes, for a nonlinearity Ψ with bounded outputs (e.g., the nonlinearities NL1-3 introduced in Section 2), almost sure convergence of the pro- posed algorithm [\(2.3\)](#page-4-0) under heavy-tail communication noise that may not have finite moments of order greater than one. In contrast, it can be shown that the correspond-302 ing linear $\mathcal{L}U$ scheme in [\[17\]](#page-19-2) (obtained by taking Ψ to be the identity function in [\(2.3\)](#page-4-0)) 303 generates a sequence of iterates with unbounded second moments for all $t = 1, 2, ...$ (see Supplementary material B). The Theorem also establishes almost sure conver- gence of [\(2.3\)](#page-4-0) for nonlinearities with unbounded outputs, more precisely, those that satisfy part 5 of Assumption [2.2,](#page-4-1) when the communication noise has finite second 307 moment. As a special case, by taking Ψ to be the identity map, we recover for the 308 letter case almost sure convergence of the linear estimator (the $\mathcal{L}U$ algorithm) in [\[17\]](#page-19-2). Setting up the proof. We next outline our strategy for proving Theorem [3.1.](#page-6-2) We base our analysis on stochastic approximation arguments. More precisely, we use Theorem 29 in [\[17\]](#page-19-2) adapted from [\[28\]](#page-19-21) (see also Theorem 3 in the supplementary 312 material) to establish a.s. convergence of x^t to $1_N \otimes \theta^*$ by verifying assumptions B1–B5 of Theorem 29 in [\[17\]](#page-19-2).

314 The proof strategy is as follows. We first prove a.s. convergence of algorithm [\(2.3\)](#page-4-0) 315 for the case without communication noise, i.e., by setting $\xi_{ij}^t \equiv 0$ in [\(2.3\)](#page-4-0). In this 316 setting, we first prove the result assuming a continuous function $\Psi : \mathbb{R} \mapsto \mathbb{R}$. Then, we 317 handle the case with discontinuous Ψ by additionally assuming that we can associate 318 to $\Psi : \mathbb{R} \mapsto \mathbb{R}$ a "lower bound" surrogate function $\Psi : \mathbb{R} \mapsto \mathbb{R}$ that is *continuous*, 319 satisfies assumption [2.2,](#page-4-1) and the following holds:

320 (3.1)
$$
|\Psi(a)| \geq |\Psi(a)|, \text{ for any } a \in \mathbb{R}.
$$

321 This enables us to complete the proof for the noiseless case. To transition to the noisy 322 communications case, a key argument is to consider an auxiliary function $\varphi : \mathbb{R} \to \mathbb{R}$, 323 defined by

324 (3.2)
$$
\varphi(a) = \int \Psi(a+w) d\Phi(w).
$$

326 Intuitively, $\varphi : \mathbb{R} \to \mathbb{R}$ is a convolution-like transformation of nonlinearity $\Psi : \mathbb{R} \to \mathbb{R}$, 327 where the convolution is taken with respect to the communication noise cumulative 328 distribution function Φ.

329 As we will demonstrate ahead, function $\varphi : \mathbb{R} \to \mathbb{R}$ in the noisy communications case

330 effectively plays the role that function $\Psi : \mathbb{R} \to \mathbb{R}$ has in the noiseless case. Moreover,

331 function φ inherits all the key properties of function Ψ . More precisely, we exploit 332 the following Lemma in [\[29\]](#page-19-20) (see Lemmas 1-6 in [\[29\]](#page-19-20)).

333 LEMMA 3.2 ([\[29\]](#page-19-20)). Consider function φ in [\(3.2\)](#page-6-3), where function $\Psi : \mathbb{R} \to \mathbb{R}$, 334 satisfies Assumption [2.2.](#page-4-1) Then, the following holds:

335 1. φ is odd;

336 2. If $|\Psi(\nu)| \leq c_1$, for any $\nu \in \mathbb{R}$, then $|\varphi(a)| \leq c'_2$, for any $a \in \mathbb{R}$, for some 337 $c'_1 > 0;$

338 **3.** $I_f^f |\Psi(\nu)| \le c_2(1+|\nu|)$, for any $\nu \in \mathbb{R}$, then $|\varphi(a)| \le c'_2(1+|a|)$, for any $a \in \mathbb{R}$, 339 *for some c*'₂ > 0;

- 340 4. $\varphi(a)$ is monotonically nondecreasing;
- 341 5. $\varphi(a) > 0$, for any $a > 0$.
- 342 6. φ is continuous at zero;

343 7. φ is differentiable at zero, with a strictly positive derivative at zero, equal to:

$$
344 \qquad (3.3) \quad \varphi'(0) = \sum_{i=1}^{s} \left(\Psi(\nu_i + 0) - \Psi(\nu_i - 0) \right) p(\nu_i) + \sum_{i=0}^{s} \int_{\nu_i}^{\nu_{i+1}} \Psi'(\nu) p(\nu) d\nu,
$$

345 where $\nu_i, i = 1, ..., s$ are points of discontinuity of Ψ such that $\nu_0 = -\infty$ 346 and $\nu_{s+1} = +\infty$, and we recall that $p(u)$ is the pdf of distribution Φ (see 347 Assumption [2.2\)](#page-4-1).

348 Lemma [3.2](#page-7-0) allows that the treatment of the noisy case becomes completely analogous 349 to the noiseless case, by replacing function Ψ with φ . Finally, to address the case 350 when φ may not be continuous over R, we make use of the following Lemma that is 351 a trivial corollary of Lemma [3.2.](#page-7-0)

352 LEMMA 3.3. Consider φ in [\(3.2\)](#page-6-3). Then, there exists a positive constant ξ such 353 $that |\varphi(a)| \geq \frac{1}{2}\varphi'(0) |a|, for |a| \leq \xi.$

354 Lemma [3.3](#page-7-1) allows us to define a continuous function $\varphi : \mathbb{R} \mapsto \mathbb{R}$,

$$
\underline{\varphi}(a) = \begin{cases} \frac{1}{2}\varphi'(0) a & , |a| \leq \xi \\ \xi \operatorname{sign}(a) & , \quad \text{else} \end{cases},
$$

356 357 that satisfies Assumption [2.2](#page-4-1) and obeys the property:

358 (3.4) $|\varphi(a)| \ge |\varphi(a)|$, for any $a \in \mathbb{R}$.

359 Function φ will then clearly play the role of function Ψ in [\(3.1\)](#page-6-4) in the noiseless case. 360 We are now ready to prove Theorem [3.1.](#page-6-2)

361 Proof. (Proof of Theorem [3.1\)](#page-6-2)

362 Step 1: No communication noise. We start the proof by verifying conditions 363 B1–B5 of Theorem 29 in [\[17\]](#page-19-2) for the case without communication noise. We use the 364 following Lyapunov function $V : \mathbb{R}^{MN} \to \mathbb{R}, V(\mathbf{x}) = ||\mathbf{x} - \mathbf{1}_N \otimes \boldsymbol{\theta}^*||^2$. For this, only 365 for condition B3, we need to analyze separately the case with continuous Ψ and the 366 case when Ψ may not be continuous. Also, it can be shown that (2.3) can be put in 367 the form required by Theorem 29 in [\[17\]](#page-19-2) (see also (36) in the supplementary material) 368 by letting

369 (3.5)
$$
\mathbf{r}(\mathbf{x}) = -\mathbf{H}^\top \mathbf{H} (\mathbf{x} - \mathbf{1}_N \otimes \boldsymbol{\theta}^*) - \frac{b}{a} \mathbf{L}_{\boldsymbol{\Psi}} (\mathbf{x}, \mathbf{0}),
$$

$$
\begin{aligned}\n\frac{370}{7} \quad (3.6) \quad \gamma(t+1, \mathbf{x}, \omega) &= \mathbf{H}^\top \mathbf{n}^t,\n\end{aligned}
$$

372 where ω denotes an element of the underlying probability space.

373 Consider the filtration \mathcal{F}_t , $t = 1, 2, ...,$ where \mathcal{F}_t is the σ - algebra generated by $\{\mathbf{n}^s\}_{s=0}^{t-1}$.

- 374 Denote by $(\Omega, \mathcal{F}, \mathbb{P})$ the underlying probability space that generates random vectors
- 375 \mathbf{n}^t , $t = 0, 1, 2, \dots$, and by $\omega \in \Omega$ its arbitrary element. Clearly, for each t, function 376 $\gamma(t+1,\cdot,\cdot)$ is $\mathcal{B}^{MN}\otimes\mathcal{F}$ measurable, where \mathcal{B}^{MN} is the Borel sigma algebra on \mathbb{R}^{MN} .

377 Also, $\mathbf{r}(\cdot)$ is \mathcal{B}^{MN} measurable. Hence, condition B1 holds. Further, the family of 378 random vectors $\gamma(t + 1, \mathbf{x}, \omega)$ is \mathcal{F}_t measurable, zero-mean and independent of \mathcal{F}_{t-1} . 379 Thus, condition B2 holds.

380 We now inspect condition B3. Assume first that function $\Psi : \mathbb{R} \mapsto \mathbb{R}$ is contin-381 uous. The gradient of V equals $\nabla V(\mathbf{x}) = 2(\mathbf{x} - \mathbf{1}_N \otimes \boldsymbol{\theta}^*)$. Clearly, function $V(\cdot)$ 382 is twice continuously differentiable and has uniformly bounded second order partial 383 derivatives. We consider

384 (3.7)
$$
S = \sup_{\|\mathbf{x}-\mathbf{1}_N\otimes\boldsymbol{\theta}^*\| \in (\epsilon,1/\epsilon)} \langle \mathbf{r}(\mathbf{x}), \nabla V(\mathbf{x}) \rangle,
$$

386 We will show that $S < 0$, thus verifying condition B3. We have, for any $\mathbf{x} \in \mathbb{R}^{MN}$:

$$
387 \qquad \langle \mathbf{r}(\mathbf{x}), \nabla V(\mathbf{x}) \rangle = -2 (\mathbf{x} - \mathbf{1}_N \otimes \boldsymbol{\theta}^*)^\top \left(\mathbf{H}^\top \mathbf{H} (\mathbf{x} - \mathbf{1}_N \otimes \boldsymbol{\theta}^*) + \frac{b}{a} \mathbf{L}_{\boldsymbol{\Psi}}(\mathbf{x}) \right)
$$

$$
\begin{array}{ll}\n\text{388} & (3.8) & = -2 \underbrace{\left(\left(\mathbf{x} - \mathbf{1}_N \otimes \boldsymbol{\theta}^* \right)^{\top} \mathbf{H}^{\top} \mathbf{H} \left(\mathbf{x} - \mathbf{1}_N \otimes \boldsymbol{\theta}^* \right) \right)}_{T_1(\mathbf{x})} - 2 \frac{b}{a} \underbrace{\left(\mathbf{x} - \mathbf{1}_N \otimes \boldsymbol{\theta}^* \right)^{\top} \mathbf{L}_{\boldsymbol{\Psi}}(\mathbf{x})}_{T_2(\mathbf{x})}.\n\end{array}
$$

390 Clearly $T_1 = T_1(\mathbf{x}) \geq 0$. We will also show that $T_2 = T_2(\mathbf{x}) \geq 0$. Utilizing the fact 391 that, $\Psi(\cdot)$ is an odd function, we have that,

$$
T_2 = \sum_{\{i,j\} \in E, i < j} (\mathbf{x}_i - \mathbf{x}_j)^\top \, \Psi \, (\mathbf{x}_i - \mathbf{x}_j) \geq 0,
$$
\n393

394 as for $\mathbf{g} = (\mathbf{x}_i - \mathbf{x}_j)$, we have that,

$$
395 \quad (3.10) \qquad \qquad (\mathbf{x}_i - \mathbf{x}_j)^{\top} \mathbf{\Psi} (\mathbf{x}_i - \mathbf{x}_j) = \sum_{\ell=1}^{M} \mathbf{g}_{\ell} \mathbf{\Psi} (\mathbf{g}_{\ell}) \geq 0,
$$

397 because \mathbf{g}_{ℓ} and $\Psi(\mathbf{g}_{\ell})$ have the same sign, by Assumption [2.2.](#page-4-1) Therefore,

$$
\langle \mathbf{r}(\mathbf{x}), \nabla V(\mathbf{x}) \rangle = -2 T_1 - 2 \frac{b}{a} T_2 \le 0,
$$

400 for any $\mathbf{x} \in \mathbb{R}^{MN}$.

399

 401 We will further show that S in (3.7) is strictly less than 0. First, consider the set 402 $\mathcal{C} = \{ \mathbf{x} \in \mathbb{R}^{MN} : ||\mathbf{x} - \mathbf{1}_N \otimes \boldsymbol{\theta}^*|| \in [\epsilon, 1/\epsilon] \}.$ Note that set \mathcal{C} is nonempty and compact. 403 Clearly, we have that:

404 (3.11)
$$
S \leq S_{\mathcal{C}} := \sup_{\mathbf{x} \in \mathcal{C}} \langle \mathbf{r}(\mathbf{x}), \nabla V(\mathbf{x}) \rangle,
$$

405 It is thus sufficient to show that $S_{\mathcal{C}} < 0$. Suppose the contrary is true, i.e., suppose 406 that $S_{\mathcal{C}} = 0$. As set C is compact and function $\mathbf{x} \mapsto \langle \mathbf{r}(\mathbf{x}), \nabla V(\mathbf{x}) \rangle$ is continuous, by the 407 Weierstrass theorem, we have that $S_{\mathcal{C}} = 0$ is equivalent to having $\langle \mathbf{r}(\mathbf{x}^{\bullet}), \nabla V(\mathbf{x}^{\bullet}) \rangle = 0$, 408 for some point $\mathbf{x}^{\bullet} \in \mathcal{C}$. In this case, \mathbf{x}^{\bullet} has to be of the form, $\mathbf{x}^{\bullet} = \mathbf{1}_N \otimes \mathbf{m}$, where $\mathbf{m} \in \mathcal{C}$. ⁴⁰⁹ R $\sqrt{ }$ ^M. As otherwise, we would have that, T_2 is strictly positive. But then, we have, $T_1 =$ $\left\{ \mathbf{x}^{\bullet}-\mathbf{1}_{N}\otimes \boldsymbol{\theta}^{*}\right\} ^{\top}\mathbf{H}^{\top}\mathbf{H}\left(\mathbf{x}^{\bullet}-\mathbf{1}_{N}\otimes \boldsymbol{\theta}^{*}\right)\right\} =\left(\mathbf{m}-\boldsymbol{\theta}^{\star}\right)^{\top}\left(\sum_{i=1}^{N}\mathbf{h}_{i}\mathbf{h}_{i}^{\top}\right)\left(\mathbf{m}-\boldsymbol{\theta}^{\star}\right)>0,$ 411 which is a contradiction in view of (3.7) . Hence, we conclude that, for a continuous 412 function Ψ , it holds that $S < 0$, and that condition B3 holds, i.e., sup $\|\mathbf{x}-\mathbf{1}_N\!\otimes\! \boldsymbol{\theta}^*\| \!\in\! (\epsilon,1/\epsilon)$ 413 $\langle \mathbf{r}(\mathbf{x}), \nabla V(\mathbf{x}) \rangle < 0.$ 414

 Now, we verify condition B3 for function Ψ that is not continuous but to which we can 416 associate function Ψ that obeys Assumption [2.2](#page-4-1) and for which condition [\(3.1\)](#page-6-4) holds. Then, the verification of condition B3 follows analogously to the case with continuous Ψ by replacing T_2 in [\(3.9\)](#page-8-1) with the following lower bound of T_2

419 (3.12)
$$
\underline{T}_2 = \sum_{\{i,j\} \in E, i < j} (\mathbf{x}_i - \mathbf{x}_j)^\top \underline{\Psi} (\mathbf{x}_i - \mathbf{x}_j),
$$

420 where $\underline{\Psi}(\mathbf{a}) = [\underline{\Psi}(\mathbf{a}_1), \underline{\Psi}(\mathbf{a}_2), ..., \underline{\Psi}(\mathbf{a}_M)]^\top$. Hence, condition B3 is verified.

421 We next verify condition B4. Recalling the definition of $r(x)$ in [\(3.5\)](#page-7-2), we have, $||\cdot|| \times ||^2$ >

$$
\lim_{423} (3.13) \t ||\mathbf{r}(\mathbf{x})||^2 \le c_3 V(\mathbf{x}) + c_4 ||\mathbf{\Psi}(\mathbf{x})||^2,
$$

where $c_3 = 2a^2 ||\mathbf{H}^\top \mathbf{H}||$ 424 where $c_3 = 2a^2 ||\mathbf{H}^\top \mathbf{H}||^2$ and $c_4 = 2b^2 ||\mathbf{L}||^2$. 425 We also have that,

426
$$
\|\mathbf{\Psi}(\mathbf{x})\| \leq c_5 \sum_{\{i,j\} \in E} (|\mathbf{x}_i - \boldsymbol{\theta}^*| + |\mathbf{x}_j - \boldsymbol{\theta}^*|) + c_6 \leq c_7 \|\mathbf{x} - \mathbf{1}_N \otimes \boldsymbol{\theta}^*\| + c_6,
$$

428 for some positive constants c_5, c_6, c_7 . Therefore, we have

439 (3.14) $\|\Psi(\mathbf{x})\|^2 \le 2c_7V(\mathbf{x}) + 2c_8^2,$ 430

 4

431 for some positive constant c_8 .

432 Thus, we have that,

$$
\|\mathbf{r}(\mathbf{x})\|^2 \le c_9 V(\mathbf{x}) + c_{10},
$$

435 form some positive constants c_9, c_{10} . Recall $\gamma(t + 1, \mathbf{x}, \omega)$ in [\(3.6\)](#page-7-3). Using the bound-436 edness of the second moment of the observation noise, we finally have that,

$$
\|\mathbf{r}(\mathbf{x})\|^2 + \mathbb{E}\left[\left\|\boldsymbol{\gamma}(t+1,\mathbf{x}^t,\omega)\right\|^2\right] \leq c_{11}\left(V(\mathbf{x})+1\right),\,
$$

 $\begin{array}{c} 437 \\ 438 \\ 439 \end{array}$ for some positive constant c_{11} . Hence, condition B4 is satisfied. Finally, condition B5 440 clearly holds. Therefore, we conclude that $\mathbf{x}^t \to \mathbf{1}_N \otimes \boldsymbol{\theta}^*$, almost surely.

441 Step 2: The case with communication noise. We proceed by considering algo-442 rithm [\(2.3\)](#page-4-0) under communication noise.

443 We clarify the steps needed to transition from the noiseless to the noisy case. If we 444 write

446

434

$$
\Psi(\mathbf{x}_i^t - \mathbf{x}_j^t + \boldsymbol{\xi}_{ij}^t) = \varphi(\mathbf{x}_i^t - \mathbf{x}_j^t) + \boldsymbol{\eta}_{ij}^t,
$$

447 where $\boldsymbol{\eta}_{ij}^t = \left[\boldsymbol{\Psi}(\mathbf{x}_i^t\!-\!\mathbf{x}_j^t\!+\!\boldsymbol{\xi}_{ij}^t)\!-\!\boldsymbol{\varphi}(\mathbf{x}_i^t\!-\!\mathbf{x}_j^t)\right]$ and $\boldsymbol{\varphi}: \mathbb{R}^M \to \mathbb{R}^M$ is component-wise map 448 defined as $\boldsymbol{\varphi}(\mathbf{x}_1, \mathbf{x}_2, ..., \mathbf{x}_M) = [\varphi(\mathbf{x}_1), \varphi(\mathbf{x}_2), ..., \varphi(\mathbf{x}_M)]^\top$. We will see that quantity 449 η_{ij}^t is a key ingredient of $\gamma(t+1, \mathbf{x}, \omega)$ in Theorem 29 in [\[17\]](#page-19-2) (see also Theorem 3 in 450 the supplementary material).

451 The algorithm [\(2.3\)](#page-4-0) can be written in compact form:

$$
\mathbf{x}^{t+1} = \mathbf{x}^t - \alpha_t \left(\frac{b}{a} \mathbf{L}_{\varphi}(\mathbf{x}^t) - \mathbf{H}^T(\mathbf{z}^t - \mathbf{H}\mathbf{x}^t) + \frac{b}{a} \boldsymbol{\eta}^t \right).
$$

454 Here,

455 (3.16)
$$
\mathbf{L}_{\boldsymbol{\varphi}}(\mathbf{x}^{t}) = \begin{bmatrix} \vdots \\ \sum_{j \in \Omega_{i}} \boldsymbol{\varphi}(\mathbf{x}_{i}^{t} - \mathbf{x}_{j}^{t}) \\ \vdots \end{bmatrix} \in \mathbb{R}^{MN}, \quad \boldsymbol{\eta}^{t} = \begin{bmatrix} \vdots \\ \sum_{j \in \Omega_{i}} \boldsymbol{\eta}_{ij}^{t} \\ \vdots \end{bmatrix} \in \mathbb{R}^{MN},
$$

456

where the $M \times 1$ blocks \sum $j \in \Omega_i$ $\varphi(\mathbf{x}_i^t - \mathbf{x}_j^t)$ and \sum $j \in \Omega_i$ 457 where the $M \times 1$ blocks $\sum \varphi(\mathbf{x}_i^t - \mathbf{x}_j^t)$ and $\sum \eta_{ij}^t$ are stacked one on top of another 458 for $j = 1, ..., N$.

459 The differences of [\(3.15\)](#page-9-0) with respect to the case without additive communication 460 noise are that \mathbf{L}_{φ} replaces \mathbf{L}_{Ψ} and the term $\frac{b}{a} \alpha_t \eta^t$ is added.

461 We define the Lyapunov function $V : \mathbb{R}^{MN} \to \mathbb{R}$, and quantities $\mathbf{r}_{\varphi}(x)$ and $\gamma_{\varphi}(t, \mathbf{x}, \omega)$ 462 as follows:

$$
V(\mathbf{x}) = ||\mathbf{x} - \mathbf{1}_N \otimes \boldsymbol{\theta}^*||^2
$$

464 (3.18)
$$
\mathbf{r}_{\varphi}(\mathbf{x}) = -\mathbf{H}^T \mathbf{H}(\mathbf{x} - \mathbf{1}_N \otimes \boldsymbol{\theta}^*) - \frac{b}{a} \mathbf{L}_{\varphi}(\mathbf{x}),
$$

 $\boldsymbol{\gamma}_{\boldsymbol{\varphi}}(t+1, \mathbf{x}, \omega) = \mathbf{H}^{\top} \mathbf{n}^t - \frac{b}{a}$ $\gamma_{\varphi}(t+1, \mathbf{x}, \omega) = \mathbf{H}^{\top} \mathbf{n}^{t} - \frac{\partial}{\partial \theta} \boldsymbol{\eta}^{t},$
466 (3.19) 466

467 Now, make the following identification with respect to the transition from the noiseless 468 to the noisy case. Quantity $H^T n^t$ in the noiseless case is replaced with quantity

469 $\mathbf{H}^{\top} \mathbf{n}^t - \frac{b}{a} \boldsymbol{\eta}^t$ in the noisy case. The map $\mathbf{L}_{\Psi}(\cdot, \mathbf{0}) : \mathbb{R}^{MN} \to \mathbb{R}^{MN}$ in [\(2.4\)](#page-4-3) is replaced 470 with the map $\mathbf{L}_{\varphi}: \mathbb{R}^{MN} \to \mathbb{R}^{MN}$ given in [\(3.16\)](#page-9-1).

471 The proof proceeds analogously by again verifying Assumptions B1–B5. We only 472 clarify the differences in verifying these conditions with respect to the noiseless case. 473 The filtration \mathcal{F}_t is replaced with the filtration \mathcal{G}_t , $t = 1, 2, ...,$, which is generated 474 not only by $\{\mathbf{n}^{s}\}_{s=0}^{t-1}$ but also by $\{\boldsymbol{\xi}_{ij}^{s}\}_{s=0}^{t-1}$ for $(i, j) \in E$. Clearly, for each t, function 475 $\gamma_{\varphi}(t+1;\cdot;\cdot)$ is $\mathcal{B}^{MN}\otimes\mathcal{F}$ measurable. Also, $\mathbf{r}_{\varphi}(\cdot)$ is \mathcal{B}^{MN} measurable. Hence, condition 476 B1 holds. Further, the family of random vectors $\gamma_{\varphi}(t+1, \mathbf{x}, \omega)$ is \mathcal{F}_t measurable, 477 zero-mean and independent of \mathcal{F}_{t-1} . Thus, condition B2 holds. As function φ is 478 odd, non-decreasing, strictly positive for its positive arguments, and has a positive 479 derivative at zero by Lemma [3.2,](#page-7-0) condition B3 is derived analogously to the noiseless 480 case. Conditions B4 and B5 hold analogously to the noiseless case. Thus, the result 481 is verified. Г

482 Remark 1: Theorem [3.1](#page-6-2) continues to hold under the following generalizations:

- 483 A different nonlinear function $\Psi_{i,i,\ell} : \mathbb{R} \to \mathbb{R}$ is assigned to each arc (i, j) and 484 to each element $\ell = 1, ..., M$ of the communication noise $[\xi_{ij}^t]_\ell$. Each function 485 $\Psi_{ij,\ell}$ obeys Assumption [2.2.](#page-4-1)
- 486 The observation noise $\sigma_{obs,i}^2$ is different for each agent $i = 1, 2, ..., N$.
- 487 The communication noise ξ_{ij}^t has the joint cumulative distribution function 488 Φ_{ij} such that:
-

490

489
$$
\int_{\mathbf{a}\in\mathbb{R}^M} \|\mathbf{a}\|d\Phi_{ij}(\mathbf{a}) < \infty, \quad \int_{\mathbf{a}\in\mathbb{R}^M} \mathbf{a}d\Phi_{ij}(\mathbf{a}) = 0,
$$

491 and $\mathbf{\Phi}_{ij}(\mathbf{a}) = 1 - \mathbf{\Phi}_{ij}(-\mathbf{a})$, for all $\mathbf{a} \in \mathbb{R}^M$.

492 All the remaining assumptions in [2.1-](#page-4-2)[2.4](#page-5-1) continue to hold.

493 Note that the above means that the communication noise ξ_{ij}^t may have mutually 494 dependent elements $[\xi_{ij}^t]_\ell$, for $\ell = 1, ..., M$.

495 For the above generalization, it can be shown that Theorem [3.1](#page-6-2) continues to hold (see 496 Supplementary material C).

497 3.2. Asymptotic normality. We now present our results on asymptotic nor-498 mality of estimator [\(2.3\)](#page-4-0).

⁴⁹⁹ Theorem 3.4 (Asymptotic normality). Let Assumptions [2](#page-4-2).1 − [2](#page-5-1).4 hold. Con-500 sider algorithm [\(2.3\)](#page-4-0) with step-size $\alpha_t = a/(t+1), t = 0, 1, ..., a > 0$. Then, the 501 normalized sequence of iterates $\{\sqrt{t+1}(\mathbf{x}^t - \mathbf{1}_N \otimes \boldsymbol{\theta}^*)\}$ converges in distribution to a 502 zero-mean multivariate normal random vector, i.e., the following holds:

$$
\mathfrak{H}_{4}^{3} \qquad \qquad \mathcal{N}t+1(\mathbf{x}^{t}-\mathbf{1}_{N}\otimes\boldsymbol{\theta}^{*})\Rightarrow\mathcal{N}(\mathbf{0},\mathbf{S}),
$$

505 where the asymptotic covariance matrix S equals:

506 (3.20)
$$
\mathbf{S} = a^2 \int_0^\infty e^{\mathbf{\Sigma} v} \mathbf{S}_0 e^{\mathbf{\Sigma}^\top v} dv.
$$

507 Here, $\mathbf{S}_0 = \sigma_{obs}^2 \mathbf{H}^\top \mathbf{H} + \frac{b^2}{g^2} \sigma^2 \text{Diag}(\{d_i \mathbf{I}_M\})$, where we recall that d_i is the degree of 508 agent i; $\sigma^2 = \int |\Psi(w)|^2 d\Phi(w)$ is the effective communication noise variance after 509 passing through the nonlinearity Ψ ; we recall the observation matrix **H** in [\(2.2\)](#page-3-4); the 510 observation noise variance σ_{obs}^2 in [\(2.1\)](#page-3-3); function ϕ in [\(3.2\)](#page-6-3); and $\Sigma = \frac{1}{2}\mathbf{I} - a(\mathbf{H}^\top \mathbf{H} +$ 511 $\frac{b}{a}\varphi'(0)(\mathbf{L}\otimes\mathbf{I}_M)$, where a is taken large enough such that matrix Σ is stable (i.e., real 512 parts of Σ 's eigenvalues are negative).

513 Theorem [3.4](#page-10-1) shows that, for the communication noise with finite variance and un-

514 bounded nonlinearities that satisfy part 5 of Assumption [2.2,](#page-4-1) the variance with the 515 proposed nonlinear estimator [\(2.3\)](#page-4-0) decays (in the weak convergence sense) at (the best 516 achievable) rate $O(1/t)$. In particular, by taking Ψ to be the identity function, we 517 recover the asymptotic normality result in [\[17\]](#page-19-2) of the corresponding linear estimator 518 (the $\mathcal{L}U$ scheme in [\[17\]](#page-19-2)). Note also that the asymptotic variance expression in [\(3.20\)](#page-10-2)

519 for the indentity function $\Psi(a) = a$ coincides with that in [\[17\]](#page-19-2) for the $\mathcal{L}U$ scheme.

- 520 Theorem [3.4](#page-10-1) further demonstrates that, even under a heavy-tailed communication 521 noise (with unbounded variance) with a bounded nonlinearity (e.g., nonlinearities
- 522 NL1-3 in Section 2), the variance with algorithm (2.3) still decays at rate $O(1/t)$. In
- 523 contrast, the corresponding linear scheme (obtained by taking Ψ in [\(2.3\)](#page-4-0) to be the
- 524 identity function) generates a sequence with unbounded variances for each $t = 1, 2, ...$

525 More precisely, we then have that $E[||\mathbf{x}^t - \mathbf{1}_N \otimes \boldsymbol{\theta}^*||^2] = \infty$, for any $t = 1, 2, ...$ (see

526 Supplementary material C).

 Theorem [3.4](#page-10-1) explicitly quantifies asymptotic variance of [\(2.3\)](#page-4-0). This also allows, in the finite communication noise regime, to compare the nonlinear versus the linear scheme (when both schemes achieve a finite asymptotic variance). See Subsection [4.1](#page-12-1) for details.

 Theorem [3.4](#page-10-1) also reveals an interesting tradeoff when including the nonlinearity Ψ into the consensus update. On the one hand, nonlinearity makes a beneficial effect in that the communication noise plays the role only through the effective variance σ^2 = $\int |\Psi(w)|^2 d\Phi(w)$. In contrast, with the linear scheme, σ^2 is replaced with $\int w^2 d\Phi(w)$ that is infinite under a heavy tail setting. On the other hand, the nonlinearity Ψ 536 makes a negative effect in that it "reduces quality" of matrix Σ through the quantity $537 \quad \varphi'(0)$ that is typically less than one with a nonlinear scheme and equal to one with the linear scheme. Clearly, the tradeoff goes in favor of the nonlinear scheme in the heavy tail setting (finite variance with the nonlinear estimator versus infinite variance with the linear estimator). In the finite communication noise variance setting, the nonlinear scheme typically improves performance under a sufficiently low communication signal to noise ratio (SNR); see also Subsection [4.1.](#page-12-1) We are now ready to prove Theorem 543 [3.4.](#page-10-1)

 Proof. (Proof of Theorem [3.4\)](#page-10-1) We establish asymptotic normality by verifying assumptions C1-C5 of Theorem 29 in [\[17\]](#page-19-2) (see also Theorem 3 in the supplementary 546 material). Firstly, we show that condition C1 hold. Since function φ is differentiable at zero, we have that

$$
\mathbf{H}_{48}^4 \quad (3.21) \qquad \qquad \varphi(a) = \varphi(0) + \varphi'(0)a + \Delta(a) = \varphi'(0)a + \Delta(a),
$$

where for the function $\Delta : \mathbb{R} \to \mathbb{R}$, we have that $\lim_{a \to 0} \frac{\Delta(a)}{a} = 0$. Hence, the function 550 $\mathfrak{r}_{\varphi}(\mathbf{x})$ admits representation as in Theorem 29 of [\[17\]](#page-19-2) (see also (37) of Theorem 3 in 552 the supplementary material), with matrix

$$
\mathbf{B} = -\mathbf{H}^T \mathbf{H} - \frac{b}{a} \varphi'(0) \big[\mathbf{L} \otimes \mathbf{I}_M \big],
$$
⁵⁵³

and function $\delta : \mathbb{R}^{MN} \to \mathbb{R}^{MN}$, given with $\delta(\mathbf{x}) = -\frac{b}{a} \mathbf{L}_{\Delta}(\mathbf{x})$. Here, function $\mathbf{L}_{\Delta}(\mathbf{x})$: 555 556 $\mathbb{R}^{MN} \to \mathbb{R}^{MN}$ is defined by

$$
\mathbf{L}_{\mathbf{\Delta}}(\mathbf{x}) = \begin{bmatrix} \vdots \\ \sum_{j \in \Omega_i} \mathbf{\Delta}(\mathbf{x}_i - \mathbf{x}_j) \\ \vdots \end{bmatrix},
$$
558

558

where function $\Delta : \mathbb{R}^M \to \mathbb{R}^M$ is defined by $(3.21), \Delta(\mathbf{y}_1, \mathbf{y}_1, ..., \mathbf{y}_M) = [\Delta(\mathbf{y}_1), \Delta(\mathbf{y}_2), ..., \Delta(\mathbf{y}_M)]^\top$ $(3.21), \Delta(\mathbf{y}_1, \mathbf{y}_1, ..., \mathbf{y}_M) = [\Delta(\mathbf{y}_1), \Delta(\mathbf{y}_2), ..., \Delta(\mathbf{y}_M)]^\top$, 560 $\mathbf{y} \in \mathbb{R}^M$.

h

561 Condition C2 trivially holds, if we use that $\alpha_t = \frac{a}{t+1}$. Furthermore, $\Sigma = a\mathbf{B} + \frac{1}{2}\mathbf{I}$ is 562 stable if a is large enough, because matrix $-\mathbf{B}$ is positive definite (see [\[17\]](#page-19-2)). Thus, 563 condition C3 also holds.

564 For $\mathbf{A}(t, \mathbf{x}) = \mathbb{E}[\gamma_{\varphi}(t+1, \mathbf{x}, \omega)\gamma_{\varphi}^{\top}(t+1, \mathbf{x}, \omega)],$ using the Lebesgue's dominated 565 convergence theorem, it can be shown that

$$
\frac{56}{5}
$$

$$
\lim_{t \to \infty, \mathbf{x} \to \boldsymbol{\theta}^*} \mathbf{A}(t, \mathbf{x}) = \sigma_{\text{obs}}^2 \mathbf{H}^\top \mathbf{H} + \sigma^2 \operatorname{Diag}(\{d_i \mathbf{I}_M\}).
$$

567 Therefore, condition C4 also holds. It remains to verify condition C5. Recall quantity 569 $\gamma_{\varphi}(t+1, \mathbf{x}, \omega)$ in [\(3.19\)](#page-9-2). Note that this condition is equivalent to saying that the family 570 of random variables $\{\|\gamma_{\varphi}(t+1,\mathbf{x},\omega)\|^2\}_{t=0,1,\ldots, \|\mathbf{x}-\theta^*\|<\epsilon}$ is uniformly integrable. If 571 the condition 5 in Assumption [2.2](#page-4-1) holds (the case with finite communication noise 572 variance and the nonlinearity with unbounded outputs), then:

573 (3.22) $\|\gamma_{\varphi}(t+1,x,\omega)\|^2 \leq c_{12} + c_{13} \|\mathbf{n}^t\|^2 + c_{14} \|\boldsymbol{\eta}^t\|^2,$

574 for some positive constants c_{12}, c_{13}, c_{14} .

575 Consider the family $\{\widetilde{\mathbf{g}}(t+1, \mathbf{x}, \omega)\}_{t=0,1,\dots, \|\mathbf{x}-\theta^*\|<\epsilon}$, with

(3.23) $\widetilde{\mathbf{g}}(t+1, \mathbf{x}, \omega) = c_{12} + c_{13} \|\mathbf{n}\|$ 576 (3.23) $\widetilde{\mathbf{g}}(t+1,\mathbf{x},\omega) = c_{12} + c_{13} \|\mathbf{n}^t\|^2 + c_{14} \|\boldsymbol{n}^t\|^2.$

577 Clearly, $\tilde{\mathbf{g}}(t + 1, x, \omega)$ is integrable, for any $t = 0, 1, \dots$, for any $\epsilon > 0$, due to the finite second moment of sensing and observation poises. The family $\tilde{\mathbf{g}}(t + 1)$ 578 the finite second moment of sensing and observation noises. The family $\{\tilde{\mathbf{g}}(t + 579 - 1, x, \omega)\}_{t=0,1}$ = $\mathbf{g}(\mathbf{x} + \mathbf{g})$ = $\mathbf{g}(\mathbf{x} + \mathbf{g})$ = $\mathbf{g}(\mathbf{x} + \mathbf{g})$ = $\mathbf{g}(\mathbf{x} + \mathbf{g})$ = $\mathbf{g}(\mathbf{x} + \mathbf$ $\{1, x, \omega\}\}_{t=0,1,\ldots,\|\mathbf{x}-\mathbf{\theta}^*\|<\epsilon}$ is i.i.d. and hence it is uniformly integrable. The family 580 $\{\|\gamma_{\varphi}(t+1,x,\omega)\|^2\}_{t=0,1,\ldots,\|\mathbf{x}-\theta^*\|<\epsilon}$ is dominated by $\{\widetilde{\mathbf{g}}(t+1,x,\omega)\}_{t=0,1,\ldots,\|\mathbf{x}-\theta^*\|<\epsilon}$ 581 that is uniformly integrable, and hence $\{\|\gamma_{\varphi}(t+1,x,\omega)\|^2\}_{t=0,1,\ldots,\|x-x^*\|<\epsilon}$ is also uni-582 formly integrable. An analogous argument can be applied if condition $5'$ in Assump-583 tion [2.2](#page-4-1) holds (bounded nonlinearity, communication noise with infinite variance). 584 Hence, condition C5 holds; thus, the result. \Box

585 4. Analytical and numerical examples. Subsection [4.1](#page-12-1) provides analytical 586 examples, and Subsection [4.2](#page-16-0) provides simulation examples, that illustrate the main 587 results presented in Section [3.](#page-6-0)

 4.1. Analytical examples. We provide several analytical examples that illus- trate Theorem [3.4.](#page-10-1) The examples demonstrate that, in the considered setting, the proposed nonlinear method in [\(2.3\)](#page-4-0) achieves a lower asymptotic variance than the corresponding linear scheme, for a low SNR regime, i.e., for the case when the com- munication noise variance is above a threshold. We also consider optimization of the 593 nonlinearity Ψ for a given nonlinearity class; more precisely, for the given analytical example, we consider optimization of parameter B for the NL2 nonlinearity class in Section [2.](#page-3-0)

596 Example 1: We follow a setup similar to [\[17\]](#page-19-2), but we consider the nonlinear con-597 sensus+innovations scheme in [\(2.3\)](#page-4-0), with the non-linear operator $\Psi : \mathbb{R} \to \mathbb{R}$ of the 598 following form (the NL2 nonlinearity):

599 (4.1)
$$
\Psi(w) = \begin{cases} w, & |w| \leq B \\ +B, & w > B \\ -B, & w < B \end{cases}
$$

600

605

601 for some parameter $B > 0$. Notice that letting $B \to \infty$ in [\(4.1\)](#page-12-2) leads to the linear 602 consensus+innovations $\mathcal{L} \mathcal{U}$ scheme in [\[17\]](#page-19-2).

603 Each agent i observes a scalar parameter $\theta^* \in \mathbb{R}$ according to:

$$
\mathcal{E}_i(t) = h\theta^* + n_i^t,
$$

606 where $h \neq 0$ and n_i^t is i.i.d. in time and across sensors with variance σ_{obs}^2 and zero 607 mean. Communication noise is i.i.d. across arcs and in time and is independent of 608 $\{n_i^t\}$, for all $i = 1, 2, ..., N$. Assume that the communication noise has a probability

609 distribution function $f(w)$ that is strictly positive in the vicinity of zero. Denote the 610 eigenvalues of **L** by $0 = \lambda_1 < \lambda_2 \leq \cdots \leq \lambda_N$. Let the graph be regular, for simplicity, 611 with degree d. Using Theorem [3.4,](#page-10-1) we have that the asymptotic covariance matrix 612 equals:

$$
f_{\rm{max}}
$$

613
$$
\mathbf{S} = a^2 \int_0^\infty e^{\mathbf{\Sigma} v} \mathbf{S}_0 e^{\mathbf{\Sigma} v} dv.
$$

614 Here, $\mathbf{S}_0 = \left(h^2 \sigma_{\text{obs}}^2 + \frac{b^2}{a^2} d\sigma^2\right) \mathbf{I}$; also, recall $\sigma^2 = \int_{0}^{\infty}$ 615 Here, $\mathbf{S}_0 = \left(h^2 \sigma_{\text{obs}}^2 + \frac{b^2}{\sigma^2} d\sigma^2\right) \mathbf{I}$; also, recall $\sigma^2 = \int |\Psi(w)|^2 f(w) dw$, the effective

−∞ 616 communication noise per link. We assume that $f(w)$ has a zero mean and variance $\sigma_{\text{comm}}^2 = \int_{0}^{\infty}$ −∞ 617 $\sigma_{\text{comm}}^2 = \int w^2 f(w) dw$ that is finite. Also,

$$
\Sigma = \frac{1}{2}\mathbf{I} - a\left(h^2\mathbf{I} + \frac{b}{a}\varphi'(0)\mathbf{L}\right),\,
$$
619

620 where φ is given in [\(3.2\)](#page-6-3). For the nonlinearity considered here, we have that

621
$$
\sigma^{2} = 2 \int_{0}^{+B} w^{2} f(w) dw + B^{2} \left(1 - 2 \int_{0}^{+B} f(w) dw \right),
$$

$$
+ B
$$

622
$$
\varphi'(0) = 2 \int_{0}^{1} f(w) dw.
$$

624 Denote by $\sigma_B^2 = \frac{1}{N} \text{Tr}(\mathbf{S})$ the average per-agent asymptotic variance. Analogously to 625 (76)-(86) in [\[17\]](#page-19-2), for $a > \frac{1}{2h^2}$ we get:

626
$$
\sigma_B^2 = \frac{a^2 h^2 \sigma_{\text{obs}}^2 + b^2 d\sigma^2}{N (2ah^2 - 1)} + \frac{a^2 h^2 \sigma_{\text{obs}}^2 + b^2 d\sigma^2}{N} \sum_{i=2}^N \frac{1}{2b\lambda_i \varphi'(0) + (2ah^2 - 1)}
$$

628 We next analyze the values of σ_B^2 as $B \to 0$ and $B \to +\infty$.

629 For $B \to 0$, we have that $\sigma^2 \to 0$, $\varphi'(0) \to 0$ and

630
$$
\sigma_B^2 \to \frac{a^2 h^2 \sigma_{\text{obs}}^2}{2ah^2 - 1} =: \sigma_0^2.
$$

632 That is, when $B \to 0$, we effectively have the case that each agent is working in 633 isolation, hence not seeing the effect of the communication noise.

634 For $B \to +\infty$, we have that $\varphi'(0) \to 1$, $\sigma^2 \to \sigma_{\text{comm}}^2$ and

635
$$
\sigma_B^2 \to \frac{a^2 h^2 \sigma_{\text{obs}}^2 + b^2 d \sigma_{\text{comm}}^2}{N (2ah^2 - 1)} + \frac{a^2 h^2 \sigma_{\text{obs}}^2 + b^2 d \sigma_{\text{comm}}^2}{N} \sum_{i=2}^N \frac{1}{2b\lambda_i + (2ah^2 - 1)} =: \sigma_{\infty}^2.
$$
637 This is the asymptotic variance of the linear $\mathcal{L} \mathcal{U}$ scheme in [17]. Note that, for any

638 set of values of system parameters and any
$$
a > \frac{1}{2h^2}
$$
 and $b > 0$, there holds that
\n $6\frac{3}{40}(4.2)$ $\sigma_{\infty}^2 > \sigma_0^2$

641 for a sufficiently large σ_{comm}^2 .

642 Assume from now on that [\(4.2\)](#page-13-0) holds. It can be shown that there exists an optimal 643 B, i.e., there exists B^* such that $B^* \in (0, +\infty)$ and $\inf_{B \in (0, +\infty)} \sigma_B^2 = \sigma_{B^*}^2$ (see Supple-

- 644 mentary material D).
- 645 Note that the above analysis generalizes also to the case when

646
$$
\sigma_{\text{comm}}^2 = \int_{-\infty}^{+\infty} w^2 f(w) dw = +\infty,
$$
647

631

648 i.e., when the noise variance is $+\infty$. In this case, we have that $\sigma_{\infty}^2 = +\infty$, for the 649 linear scheme and $\sigma_0^2 = \frac{a^2 h^2 \sigma_{\text{obs}}^2}{2ah^2-1}$ for the isolation scheme. It can be shown that 650 $\inf_{B\in(0,+\infty)} \sigma_B^2$ is achieved at some $B^* \in (0,+\infty)$ (see Supplementary material D).

651 In order to demonstrate the results above, we minimize σ_B^2 and calculate B^* for a 652 specific numerical example (see Figure [1a\)](#page-14-0). We consider a sensor (agents) network 653 with $N = 8$ agents, where the underlying topology is given by a regular graph with 654 degree $d = 3$. We set innovation and consensus constants as $a = b = 1$, the observation 655 • parameter $h = 1$, and the true parameter $\theta^* = 1$. The observation noise for each 656 sensor's measurements is standard normal, and the communication noise for each 657 communication link has the following pdf

658 (4.3)
$$
f(w) = \frac{\beta - 1}{2(1 + |w|)^{\beta}},
$$

660 with $\beta = 2.05$. (This pdf's distribution has the infinite variance.) Figure [1b](#page-14-1) shows 661 performance of the nonlinear consensus+innovations estimator [\(2.3\)](#page-4-0) in terms of the estimated per-sensor mean squared error (MSE) across iterations, for the optimal B^* 662 663 and for some sub-optimal choices of B, obtained through a Monte Carlo simulation. 664 We can see that the scheme with B^* performs better than for the considered sub-665 optimal choices of B. Figure [1c](#page-14-2) shows that Monte Carlo estimate of the per-agent 666 asymptotic variance, i.e., $\hat{S} = \frac{1}{N} ||\mathbf{x}^t - \mathbf{1}_N \otimes \boldsymbol{\theta}^*||^2 t$ matches well the corresponding 667 theoretical value as per Theorem 3.4.

FIG. 1. (a) Per-agent asymptotic variance σ_B^2 versus B for the nonlinear consensus + innovations estimator and the NL2 nonlinearity. (b) Monte Carlo-estimated per-sensor MSE error on logarithmic scale for the nonlinear consensus+innovations estimator with the NL2 nonlinearity for different choices of B. (c) Monte Carlo estimate of the per-agent asymptotic variance, and the corresponding theoretical value as per Theorem [3.4.](#page-10-1)

 Example 2: We consider the same network and sensing models as in Example 1 and the heavy-tail communication noise distribution in [\(4.3\)](#page-14-3). Furthermore, we assume 670 that $\Psi(w) = \text{sign}(w)$ (the NL3 nonlinearity). For the $\mathcal{L}U$ scheme, it can be shown that (see Supplementary material E):

672
$$
\sigma^2 = \sigma_{\text{comm}}^2 = \frac{2}{(\beta - 3)(\beta - 2)},
$$

$$
6\bar{7}2\qquad \qquad \varphi'(0) = 1.
$$

675 It can be shown here that the average per-agent asymptotic variance $\sigma_{\rm L}^2 = \frac{1}{N} \text{Tr}(\mathbf{S})$ 676 for the $\mathcal{L}U$ scheme is equal to

677
$$
(4.4) \quad \sigma_{\rm L}^2 = \begin{cases} \infty & , \quad 2 < \beta \le 3, \\ \frac{a^2 h^2 \sigma_{\rm obs}^2 + b^2 d \sigma^2}{N (2 a h^2 - 1)} + \frac{a^2 h^2 \sigma_{\rm obs}^2 + b^2 d \sigma^2}{N} \sum_{i=2}^N \frac{1}{2 b \lambda_i + (2 a h^2 - 1)}, & , \quad \beta > 3. \end{cases}
$$

679 For $\beta > 3$, quantity $\sigma_{\rm L}^2$ can be written as

$$
\sigma_{\mathcal{L}}^2 = A_{\mathcal{L}} + B_{\mathcal{L}} \frac{1}{(\beta - 3)(\beta - 2)},
$$
\n(88)

682 where

$$
A_{\rm L} = \frac{a^2 h^2 \sigma_{\rm obs}^2}{N(2ah^2 - 1)} + \frac{a^2 h^2 \sigma_{\rm obs}^2}{N} \sum_{i=2}^{N} \frac{1}{2b\lambda_i + (2ah^2 - 1)},
$$

$$
B_{\rm L} = 2\left(\frac{b^2d}{N(2ah^2 - 1)} + \frac{b^2d}{N}\sum_{i=2}^{N}\frac{1}{2b\lambda_i + (2ah^2 - 1)}\right).
$$

685

686 We next consider the nonlinear consensus+innovations scheme with the nonlinearity 687 $\Psi(w) = \text{sign } w$. We have that

$$
\sigma^2 = 1,
$$

$$
\varphi(a) = 2 \int_{0}^{a} f(w) dw,
$$

690 691 which means that $\varphi'(a) = 2f(a)$ and $\varphi'(0) = 2f(0) = (\beta - 1)$. Hence, we have that 692 the average per-agent asymptotic variance for the nonlinear scheme $\sigma_{NL}^2 = \frac{1}{N} \text{Tr}(S)$ 693 is given by:

$$
694 \quad (4.6) \qquad \sigma_{\rm NL}^2 = \frac{a^2 h^2 \sigma_{\rm obs}^2 + b^2 d \sigma^2}{N \left(2 a h^2 - 1\right)} + \frac{a^2 h^2 \sigma_{\rm obs}^2 + b^2 d \sigma^2}{N} \sum_{i=2}^N \frac{1}{4b\lambda_i f(0) + (2ah^2 - 1)},
$$

696 which can be written in the form

697 (4.7)
$$
\sigma_{\rm NL}^2 = A_{\rm NL} + B_{\rm NL} \frac{P_{N-2}(\beta)}{\prod_{i=2}^N (\beta - \beta_i)},
$$
698

699 where

$$
A_{\rm NL} = \frac{a^2 h^2 \sigma_{\rm obs}^2 + b^2 d}{N (2ah^2 - 1)},
$$

$$
B_{\rm NL} = \frac{a^2 h^2 \sigma_{\rm obs}^2 + b^2 d}{N \prod_{i=2}^{N} 2b\lambda_i},
$$

$$
P_{N-2}(\beta) = \sum_{i=2}^{N} \prod_{\substack{j=2 \ j \neq i}}^{N} 2b\lambda_j(\beta - \beta_j).
$$

 $\beta_i = 1 - \frac{2ab}{2b\lambda_i}, \quad i = 2, ..., N.$

704 705 We next compare the average per-agent asymptotic variances for the linear consen-706 sus+innovations scheme and the nonlinear consensus+innovations scheme. From [\(4.4\)](#page-15-0) 707 it is obvious that $\sigma_{\text{NL}}^2 < \sigma_{\text{L}}^2$ for $\beta \in (2,3]$. For $\beta > 3$, if $A_{\text{L}} \gg A_{\text{NL}}$ (see Supplemen-708 tary material E), the linear scheme is worse than the nonlinear scheme for all $\beta > 3$. 709 It is obvious that $\sigma_{\rm L}^2$ decreases on interval $(3,\infty)$ and $\sigma_{\rm NL}^2$ decreases on the interval 710 (β_m, ∞) , where $\beta_m = \max_{i=2,\dots,N} \beta_i < 1$ is closest β_i to 1. Function $\sigma_L^2 = \sigma_L^2(\beta)$ has an 711 asymptote at $\beta = 3$, and function $\sigma_{NL}^2 = \sigma_{NL}^2(\beta)$ at $\beta = \beta_m$, where $\beta_m < 3$, also,

 $\beta_i = 1 - \frac{2ah^2 - 1}{2b}$

712 A_{L} and A_{NL} are horizontal asymptotes for σ_{L}^2 and σ_{NL}^2 , respectively. Therefore, if 713 A_{L} is much larger than A_{NL} , σ_{L}^2 is above σ_{NL}^2 for all $\beta > 3$. Moreover, if $A_{\text{L}} < A_{\text{NL}}$ 714 there exists $\beta^* > 3$ such that the average per-agent asymptotic variance is still better 715 for the nonlinear than for the linear scheme for $\beta \in (2, \beta^*]$. Defining $k = \frac{\sigma_{\rm L}^2}{\sigma_{\rm NL}^2}$, it is 716 possible to show that $k \to \infty$ as $\beta \to 3$, and $k \to \frac{A_{\text{L}}}{A_{\text{NL}}}$ as $\beta \to \infty$. Therefore, if 717 $A_{\text{L}} < A_{\text{NL}}$, there exists β^* such that $\sigma_{\text{NL}}^2 < \sigma_{\text{L}}^2$ for all $\beta \in (2, \beta^*)$. In other words, 718 there exists a threshold value $\beta^* > 3$, such that the nonlinear scheme outperforms the 719 linear scheme for the "heavy-tail regime" $\beta \in (2, \beta^*)$, and the linear scheme performs 720 better for $\beta > \beta^*$. To summarize, in Example 2, depending on sensing and network 721 parameters, it holds that either the nonlinear scheme outperforms the linear one for 722 all β , or there exists a threshold value β^* such that the nonlinear scheme is better 7[2](#page-16-1)3 than the linear one for $\beta \in (2, \beta^*)$. Figure 2 shows the ratio $k = \frac{\sigma_{\rm L}^2}{\sigma_{\rm NL}^2}$ versus β for 724 the same sensing and network parameters as in Example 1. As it can be seen, there 725 exists a threshold β^* , that here approximately equals $\beta^* = 3.9$, such that $k > 1$ for 726 $\beta \in (2,\beta^*)$. On the other hand, for $\beta > \beta^*$, the ratio becomes smaller than one, which 727 means that for the given numerical parameters, the linear scheme performs better for 728 $\beta > \beta^*$. This is in accordance with the analysis that we provided above.

FIG. 2. Ratio $k = \frac{\sigma_{\rm L}^2}{\sigma_{\rm NL}^2}$ versus β for Example 2.

 4.2. Simulation examples. In this section, we illustrate the performance of the proposed nonlinear consensus+innovations estimator for two different choices of the non-linear operator Ψ. For both nonlinearity choices, our method is compared 732 with the corresponding linear consensus+innovations estimator $\mathcal{L} \mathcal{U}$ in [\[17\]](#page-19-2), when the communication noise has probability distribution function given by [\(4.3\)](#page-14-3).

734 We consider a sensor network with $N = 40$ agents. The underlying topology is an instance of a random geometric graph. We use the same initialization $x^0 = 0$ and same 736 step sizes $\alpha_t = \frac{1}{t+1}, a = 1, b = 1$, for both the linear and the nonlinear estimators. 737 Also, we assume that the observation noise is normally distributed, i.e., $n_i^t \sim \mathcal{N}(0, 1)$, 738 for each t, for each i. The true parameter $\boldsymbol{\theta}^* \in \mathbb{R}^{10}$ is generated randomly, where 739 the entries of θ^* are drawn mutually independently from the uniform distribution 740 on [-10,10]. The observation vectors $\mathbf{h}_i \in \mathbb{R}^{10}$ are also generated at random, for 741 which the condition 4 of Assumption [2.3](#page-5-0) is true. We use the communication noise 742 pdf in (4.3) with $\beta = 2.05$. Note that, in this case, the communication noise has an 743 infinite variance.

 744 744 744 Figure 4 compares the linear $\mathcal{L}U$ estimator in [\[17\]](#page-19-2) with the nonlinear estima-745 tor [\(2.3\)](#page-4-0) with $\Psi(w)$ given in [\(4.1\)](#page-12-2) for $B = 5$. Figure [3](#page-17-2) shows the comparison between 746 LU and [\[17\]](#page-19-2) with $\Psi(w) = \text{sign}(w)$. Both Figures show the iteration counter t at the x-axis and a Monte-Carlo estimate of the average mean square error (MSE) across agents on the y-axis. We can see that, as predicted by our theory, the nonlinear estimator, for both nonlinearity choices, persistently decreases MSE along iterations, despite the fact that the communication noise has an infinite variance. At the same

time, $\mathcal{L}U$ fails to produce a useful estimation result.

Fig. 3. Monte-Carlo average per-agent MSE estimate versus iteration counter on logarithmic scale for the proposed nonlinear estimator [\(2.3\)](#page-4-0) with the nonlinearity in [\(4.1\)](#page-12-2) for $B = 5$ and the linear $\mathcal{L} \mathcal{U}$ scheme in [\[17\]](#page-19-2).

Fig. 4. Monte-Carlo average per-agent MSE estimate versus iteration counter on logarithmic scale for the proposed nonlinear estimator [\(2.3\)](#page-4-0) with the nonlinearity $\Psi(w) = \text{sign}(w)$ and the linear $\mathcal{L} \mathcal{U}$ scheme in [\[17\]](#page-19-2).

 5. Conclusion. We studied consensus+innovations distributed estimation in the presence of impulsive, heavy-tail communication noise. To combat the impulsive communication noise, we introduce for the first time a general nonlinearity in the consensus update for consensus+innovations distributed estimation. We establish al- most sure convergence of the nonlinear consensus+innovations estimator to the true parameter, prove its asymptotic normality, and explicitly evaluate the corresponding asymptotic variance. We compare the proposed nonlinear estimator with conventional consensus+innovation estimators that utilize linear consensus update. Analytical and

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numerical examples demonstrate significant gains of introducing consensus nonlinear-

ity in low SNR (high communication noise) regimes. Most notably, we demonstrate

that, when the communication noise has infinite variance, the proposed nonlinear con-

 sensus+innovations estimator is strongly consistent (converges almost surely), while the corresponding linear counterpart provides a sequence of estimators with infinite

variance.

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REFERENCES

- [1] S. Al-Sayed, A. M. Zoubir, and A. H. Sayed, Robust distributed estimation by networked agents, IEEE Transactions on Signal Processing, 65 (2017), pp. 3909–3921.
- [2] D. Bajovic, D. Jakovetic, J. M. Moura, J. Xavier, and B. Sinopoli, Large deviations per- formance of consensus+ innovations distributed detection with non-gaussian observations, IEEE Transactions on Signal Processing, 60 (2012), pp. 5987–6002.
- [3] D. Bajovic, D. Jakovetic, J. Xavier, B. Sinopoli, and J. Moura, Distributed detection via gaussian running consensus: Large deviations asymptotic analysis, Signal Processing, IEEE Transactions on, 59 (2011), pp. 4381 – 4396.
- 782 [4] W. BEN-AMEUR, P. BIANCHI, AND J. JAKUBOWICZ, Robust distributed consensus using total variation, IEEE Transactions on Automatic Control, 61 (2016), pp. 1550–1564.
- [5] M. Cao, A. Morse, and B. Anderson, Reaching a consensus in a dynamically changing en- vironment: A graphical approach, SIAM J. Control and Optimization, 47 (2008), pp. 575– 600.
- [6] Y. Chen, S. Kar, and J. Moura, Resilient distributed field estimation, SIAM Journal on Control and Optimization, 58 (2020), pp. 1429–1456.
- 789 [7] S. CHOUVARDAS, K. SLAVAKIS, AND S. THEODORIDIS, Adaptive robust distributed learning in diffusion sensor networks, Signal Processing, IEEE Transactions on, 59 (2011), pp. 4692 – 4707.
- [8] L. Clavier, T. Pedersen, I. Larrad, M. Lauridsen, and M. Egan, Experimental evidence 793 for heavy tailed interference in the IoT, IEEE Communications Letters, 25 (2021), pp. 692– 695.
- 795 [9] S. DASARATHAN, C. TEPEDELENLIOĞLU, M. K. BANAVAR, AND A. SPANIAS, Robust consensus in the presence of impulsive channel noise, IEEE Transactions on Signal Processing, 63 (2015), pp. 2118–2129.
- [10] F. Fagnani and S. Zampieri, Average consensus with packet drop communication, in Proceed-ings of the 45th IEEE Conference on Decision and Control, 2006, pp. 1007–1012.
- 800 [11] M. HUANG AND J. MANTON, Coordination and consensus of networked agents with noisy mea- surements: Stochastic algorithms and asymptotic behavior, SIAM J. Control and Opti-mization, 48 (2009), pp. 134–161.
- 803 [12] B. HUGHES, Alpha-stable models of multiuser interference, in 2000 IEEE International Sympo-sium on Information Theory (Cat. No.00CH37060), 2000, pp. 383–.
- 805 [13] J. ILOW AND D. HATZINAKOS, Analytic alpha-stable noise modeling in a poisson field of in- terferers or scatterers, Signal Processing, IEEE Transactions on, 46 (1998), pp. 1601 – 1611.
- 808 [14] D. JAKOVETIC, J. M. F. MOURA, AND J. XAVIER, *Distributed detection over noisy networks:*
809 Large deviations analysis, IEEE Transactions on Signal Processing, 60 (2012), pp. 4306– Large deviations analysis, IEEE Transactions on Signal Processing, 60 (2012), pp. 4306– 4320.
- 811 [15] S. KAR AND J. MOURA, Asymptotically efficient distributed estimation with exponential family statistics, IEEE Transactions on Information Theory, 60 (2014), pp. 4811–4831.
- [16] S. Kar, J. Moura, and H. V. Poor, Distributed linear parameter estimation: Asymptoti-814 cally efficient adaptive strategies, SIAM Journal on Control and Optimization, 51 (2013), pp. 2200–2229.

- [17] S. Kar, J. M. F. Moura, and K. Ramanan, Distributed parameter estimation in sensor networks: Nonlinear observation models and imperfect communication, IEEE Transactions on Information Theory, 58 (2012), pp. 3575–3605.
- [18] U. A. Khan, S. Kar, and J. M. F. Moura, Distributed average consensus: Beyond the realm 820 of linearity, in 2009 Conference Record of the Forty-Third Asilomar Conference on Signals, Systems and Computers, 2009, pp. 1337–1342.
- 822 [19] S. KUMAR, U. K. SAHOO, A. K. SAHOO, AND D. P. ACHARYA, Diffusion minimum-wilcoxon- norm over distributed adaptive networks: Formulation and performance analysis, Digital Signal Processing, 51 (2016), pp. 156–169.
- 825 [20] A. LALITHA, T. JAVIDI, AND A. D. SARWATE, Social learning and distributed hypothesis testing, IEEE Transactions on Information Theory, 64 (2018), pp. 6161–6179.
- 827 [21] Z. Li AND S. GUAN, Diffusion normalized huber adaptive filtering algorithm, Journal of the Franklin Institute, 355 (2018), pp. 3812–3825.
- [22] Q. Liu and A. Ihler, Distributed estimation, information loss and exponential families, 2014. 830 [23] C. LOPES AND A. SAYED, Diffusion least-mean squares over adaptive networks: Formulation 831 and performance analysis, IEEE Transactions on Signal Processing, 56 (2008), pp. 3122– 3136.
- 833 [24] G. MATEOS, I. SCHIZAS, AND G. GIANNAKIS, *Distributed recursive least-squares for consensus-* based in-network adaptive estimation, Signal Processing, IEEE Transactions on, 57 (2009), pp. 4583 – 4588.
- 836 [25] V. MATTA, P. BRACA, S. MARANO, AND A. H. SAYED, Diffusion-based adaptive distributed 837 detection: Steady-state performance in the slow adaptation regime, IEEE Transactions on
838 Information Theory, 62 (2016), pp. 4710–4732. Information Theory, 62 (2016), pp. 4710–4732.
- 839 [26] S. MODALAVALASA, U. SAHOO, A. SAHOO, AND S. BARAHA, A review of robust distributed esti-mation strategies over wireless sensor networks, Signal Processing, 188 (2021), p. 108150.
- 841 [27] A. NEDIC, A. OLSHEVSKY, AND C. A. URIBE, Nonasymptotic convergence rates for cooperative learning over time-varying directed graphs, in 2015 American Control Conference (ACC), **IEEE**, 2015, pp. 5884–5889.
- 844 [28] M. B. Nevel'SON AND R. Z. HAS' MINSKII, Stochastic approximation and recursive estimation, vol. 47, American Mathematical Soc., 1976.
- 846 [29] B. POLYAK AND Y. TSYPKIN, Adaptive estimation algorithms: Convergence, optimality, stabil-*ity*, Automation and Remote Control, 1979 (1979).
- [30] A. Prasad, A. S. Suggala, S. Balakrishnan, and P. Ravikumar, Robust estimation via robust gradient estimation, Journal of the Royal Statistical Society: Series B (Statistical Methodology), 82 (2020), pp. 601–627.
- 851 [31] S. RAM, V. VEERAVALLI, AND A. NEDIC, *Distributed and Recursive Parameter Estimation*, Springer Science & Business Media, 2009, pp. 17–38.
- 853 [32] B. SELIM, M. S. ALAM, V. CARVALHO, G. KADDOUM, AND B. L. AGBA, Noma-based iot net- works: Impulsive noise effects and mitigation, IEEE Communications Magazine, 58 (2020), pp. 69–75.
- [33] S. Stankovic, M. Beko, and M. Stankovic, A robust consensus seeking algorithm, in IEEE EUROCON 2019-18th International Conference on Smart Technologies, 2019, pp. 1–6.
- [34] S. Sundaram and B. Gharesifard, Consensus-based distributed optimization with malicious nodes, in 2015 53rd Annual Allerton Conference on Communication, Control, and Com-puting (Allerton), 2015, pp. 244–249.
- 861 [35] S. THEODORIDIS, K. SLAVAKIS, AND I. YAMADA, Adaptive learning in a world of projections, Signal Processing Magazine, IEEE, 28 (2011), pp. 97 – 123.
- [36] F. Wen, Diffusion least mean p-power algorithms for distributed estimation in alpha-stable noise environments, Electronics Letters, 49 (2013).
- [37] X. Yang and A. Petropulu, Co-channel interference modeling and analysis in a poisson 866 field of interferers in wireless communications, IEEE Transactions on Signal Processing, 51 (2003), pp. 64–76.
- 868 [38] X. ZHAO, S.-Y. TU, AND A. H. SAYED, *Diffusion adaptation over networks under imperfect information exchange and non-stationary data*, IEEE Transactions on Signal Processing, 60 (2012), pp. 3460–3475.

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