## Chapter 12

# Sociophonetic analysis of mid front vowel production in Barcelona 

Annie Helms ${ }^{\text {a }}$<br>${ }^{\text {a }}$ University of California, Berkeley


#### Abstract

Studies aimed at observing Spanish contact-induced changes in Catalan among Spanish-Catalan bilinguals in Catalonia have evidenced both assimilation and dissimilation in the production of Catalan mid front vowels. However, the general lack of studies aimed at observing Catalan contact-induced changes in Spanish is an oversight of bidirectionality as an expectation of language contact. Accordingly, the present study uses both Catalan and Spanish mid front vowel production data from Barcelona to investigate the roles of age, gender, and language dominance in the processes of assimilation (Catalan cognate effects in Spanish) and dissimilation (distinctly produced cross-linguistic mid front vowel categories). While no cognate effect was observed in Spanish, younger speakers maintain less distinction between Catalan mid front vowel categories than older speakers, and females have less overlap between Spanish /e/ and Catalan $/ \varepsilon /$ than males. These results are consistent with a male-led change towards greater assimilation and more overlapping productions of all three mid front vowel categories. While cognitive factors such as language dominance and cognate status are central to models of bilingual phonetic representation, it is paramount to situate the bilingual individual within the context of the community and acknowledge the external social factors which also mediate variation in acquisition and production.


## 1 Introduction

In Barcelona, Spanish and Catalan have been in close contact for centuries and many instances of lexical and phonological imposition and borrowing have been recorded (Galindo i Solé 2003, 2006, Arnal 2011). However, the nearly exclusive focus on the variable acquisition of Catalan in Spain has furthered a longstanding asymmetry favoring the study of the Spanish-influence in Catalan over

Annie Helms. 2023. Sociophonetic analysis of mid front vowel production in Barcelona. In Barbara E. Bullock, Cinzia Russi \& Almeida Jacqueline Toribio (eds.), A half century of Romance linguistics: Selected proceedings of the 50th Linguistic Symposium on Romance Languages, 249268. Berlin: Language Science Press. DOI: 10.5281/zenodo. 7525114 @ ©

## Annie Helms

Catalan-influence in Spanish (Galindo i Solé 2003: 18). This trend in the literature, exemplified by Arnal (2011: 22) who states that "in the current situation of generalized bilingualism in Catalonia, the change caused by contact does not affect Spanish, but rather only affects Catalan", runs counter to the expectation of bidirectionality in language contact situations with widespread bilingualism (Thomason \& Kaufman 1988, Davidson 2020). Furthermore, the lack of cross-linguistic production data in the literature limits the conclusions that can be drawn about outcomes of language contact, especially regarding the processes of assimilation and dissimilation theorized by the Speech Learning Model (SLM; Flege 1995) and the revised Speech Learning Model (SLM-r; Flege \& Bohn 2021).

A variable studied among Spanish-Catalan bilinguals in which this asymmetry is often present is the production of mid vowels in Catalan. The mid front and mid back vowels of Catalan are contrasting, yielding minimal pairs (e.g. /net/ 'grandson' and /nct/ 'clean'; /os/ 'bear' and /os/ 'bone'). Regarding the mid front vowels, with which this study is concerned, Spanish/e/ is produced lower (higher F1) than the Catalan /e/, but higher (lower F1) and more fronted (higher F2) than the Catalan $/ \varepsilon /$ (Figure 1). However, the most significant sources of variability


Figure 1: Vowel Spaces of Spanish (Ladefoged \& Johnson 2015: 237) and Catalan (Carbonell \& Llisterri 1999: 62)
among the three vowel categories are found across F1, not F2 (Bosch \& RamonCasas 2011, Cortés et al. 2019, Recasens \& Espinosa 2006, Simonet 2011). Certain factors, such as cognate status, language dominance, and age have been linked to the variable production of these vowels cross-linguistically. Therefore, the production of these mid front vowels in bilingual settings have been the focus of many studies as they provide an opportunity to study outcomes of language contact, phonological representations of bilinguals, and the social factors that mediate these processes.

### 1.1 Models of phonological acquisition and representation

The Speech Learning Model (SLM; Flege 1995) and the Revised Speech Learning Model (SLM-r; Flege \& Bohn 2021) postulate that bilingual speakers do not maintain separate phonetic systems for each language, rather, the two systems co-exist in a mutual phonetic space and may influence one another. Whereas the SLM focuses on between-group differences and whether L2 speakers are able to produce target L1 categories, the SLM-r shifts focus to the individual and an individual's differentiation of L1 and L2 categories based on quantity and quality of input of L1 and L2.

### 1.1.1 Assimilatory outcomes

Both the SLM and SLM-r postulate that if phonetic differences between an L2 category and an L1 category are not perceived by a bilingual individual, the formation of a new category will be blocked. Blockage of category formation results in assimilation, by which the speaker may produce a composite L1-L2 category (Evans \& Iverson 2007, Kendall \& Fridland 2012). The acoustic properties of this composite category are "defined by the statistical regularities present in the combined distributions of the perceptually linked L1 and L2 sounds" (Flege \& Bohn 2021: 41). A number of factors, such as the quality and quantity of L1 and L2 input the bilingual receives in their lifetime (Flege 2002, Yeni-Komshian et al. 2000), individual cognitive differences (Lev-Ari \& Peperkamp 2013), or relative language activation (Grosjean 2001), will affect the overall acoustic profile of this composite sound. Additionally, assimilation may occur between two L2 categories when sufficient acoustic differences are not perceived or produced within the contrast.

### 1.1.2 Dissimilatory outcomes

Alternatively, the SLM and SLM-r describe the process of dissimilation, which may occur in the inventory of a bilingual that is able to perceive sufficient acoustic difference between the L1 and L2 categories, or between two L2 categories, thus preventing assimilation from occurring. In this case, the individual maintains distinct categories in each language. Categories may "deflect" one another to maintain contrast in the shared phonetic space (Baker \& Trofimovich 2005, Flege \& Bohn 2021), yielding categories that differ from those of a monolingual speaker. However, evidence of cross-linguistic dissimilation is less abundant in the literature as many studies report perception or production data from only one of the languages in question.

## Annie Helms

### 1.2 Cognate effects

Through a cognate effect, the phonology of cognate words in the non-target language may be activated during speech, potentially affecting the production of a word in the target language (Costa et al. 2000, Colomé \& Miozzo 2010). Assimilation may be realized through cognate effects, where the influence of the L1 on L2 category production is strengthened in words that are cognate between the L1 and L2. Amengual (2016a) provides evidence for a cognate effect in the production of Catalan mid back vowels in Mallorca, where productions of Catalan /o/ with incongruent Spanish cognates are raised, evidencing assimilation towards Spanish /o/.

### 1.3 Language dominance

Language dominance is a measure of linguistic history, linguistic attitudes, language proficiency, and language use, and is often a predictor of category production and perception of bilinguals (Birdsong et al. 2012). Under the SLM and SLMr (Flege 1995, Flege \& Bohn 2021), language dominance, through language exposure and experience, can predict whether assimilation or dissimilation may occur. Amengual (2016b) provides evidence for language dominance as a predictor of assimilation, where Spanish-dominant bilinguals assimilated both Catalan /e/ and Catalan $/ \varepsilon /$ to Spanish /e/. Alternatively, Bosch \& Ramon-Casas (2011) find that systematic and consistent exposure to Catalan (i.e., greater Catalan-dominance) contributes to the production of distinct Catalan mid front vowel categories by bilinguals raised in Spanish-dominant homes, thus providing evidence for dissimilation between two L2 categories. No cross-linguistic comparisons were performed in the study, thus it is unknown whether or not language dominance additionally contributed to cross-linguistic dissimilation between Catalan /e/ and Spanish /e/ among these bilinguals.

### 1.4 Age and gender

Within the variationist sociolinguistic framework, social factors such as age and gender are correlated with sound changes in progress. Under the apparent-time construct (Bailey 2004), generational sound change can be observed by comparing the speech of older speakers from that of younger speakers, where the speakers pertain to separate generations (Labov 1994: 45-46). Patterns of language use across gender often are consistent with the Gender Paradox, where "women conform more closely than men to sociolinguistic norms that are overtly prescribed, but conform less than men when they are not" (Labov 2001: 292-293).

Therefore, in both changes from above and below the level of conscious awareness, women tend to be the leaders of change. The combination of this principle with the apparent-time construct results in the methodological practice of treating younger women's speech patterns as suggestive of possible community-wide changes in progress (Labov 2001: 279).

In Barcelona, age is additionally a correlate to access to explicit language instruction as the Catalonian Linguistic Normalization Law of 1983 has yielded a generational divide between those that have and have not had access to Catalan instruction in school. Although most studies of Spanish and Catalan production in Catalonia focus on the effects of language dominance and other cognitive factors, Cortés et al. (2019) examine the language of preschool-aged children and their parents in three neighborhoods of Barcelona. They observe that the children's ability to produce Catalan mid vowels is most affected by the language environment (i.e., strength of Spanish-influence in neighborhood), whereas the production of adults is more affected by personal relationships and connections maintained in the present and the past. Despite the relative lack of research on the role of age and gender in vowel production in this bilingual community, but due to increasing immigration and the documented mid front vowel merger in progress in some areas in Barcelona (Mora \& Nadeu 2012), I predict that a potential change in progress, if observed, would be led by younger female speakers and be advancing in the direction of vowel assimilation.

### 1.5 The present study

The present study uses cross-linguistic production data to address the following research questions relating social factors to processes of assimilation and dissimilation. First, how do the factors of gender, age, and language dominance mediate a cognate effect from Catalan in the production of Spanish /e/, demonstrating assimilation? I hypothesize that a cognate effect will occur less among less Catalandominant speakers, less in females than males, and less in younger speakers despite exposure to new Catalan educational policies - as these speakers are less likely to have maintained the Catalan mid front vowel contrast, thereby inhibiting a cognate effect. Secondly, how do these social factors mediate the degree to which Spanish /e/ overlaps with each Catalan mid front vowel? I hypothesize that decreased Catalan-dominance will contribute to greater overlap (assimilation), and that female speakers and younger speakers will additionally produce categories with less acoustic distinction.

## 2 Methodology

### 2.1 Subject population

Seventeen participants were recruited with flyers posted at the University of Barcelona and were stratified according to age and gender. Two generations are represented, one group between 18-25 years old and the other between 40-65 years old. All participants are bilingual in Spanish and Catalan and have lived in Barcelona for the past 10 years. All participants were connected to the University, or had been connected in the past, which may yield a similar exposure to Catalan in a professional setting across participants.

Table 1: Number of participants and mean language dominance scores (with standard deviations) across the four social cells

| Participant Group | Count | Language Dominance |
| :--- | ---: | ---: |
| Younger Women | 6 | $+55.7(17.2)$ |
| Older Women | 4 | $+50.8(77.5)$ |
| Younger Men | 4 | $+1.0(47.0)$ |
| Older Men | 3 | $+93.9(10.4)$ |

Each participant was assigned a Catalan-Spanish dominance score (minimum: -218; maximum: +218) after completing the Bilingual Language Profile (Birdsong et al. 2012), where a more positive dominance score is correlated with greater dominance in Catalan and a more negative dominance score is correlated with greater dominance in Spanish. Table 1 shows the distribution of participants across the four participant groups, as well as the mean language dominance score for each group. The relatively low number of speakers that are more Spanishdominant $(\mathrm{n}=4)$ and the imbalance of their distribution across the four social cells prevent the use of a categorical dominance score without crossing language dominance with other social factors. Instead, dominance will remain a continuous factor in the present analysis and will not interact with either gender or age. Although the sample size is relatively small, 3-5 participants per cell is the statistical minimum to reflect group tendencies more than individual idiosyncrasies (Tagliamonte 2006: 31). At the time of data collection, no participant reported ever having any history of speech or hearing disorders.

### 2.2 Materials and procedure

All experimental sessions took place in an empty classroom at the University of Barcelona. Participants were asked if they would prefer to interact with the
researcher in Spanish or in Catalan. All communication, including the consent form, questionnaire, and sociolinguistic interview, were subsequently conducted in the preferred language. First, the participants were instructed to read and sign the consent form and complete the Bilingual Language Profile (Birdsong et al. 2012), adapted as a Qualtrics survey (Qualtrics 2005). Next, the participants engaged in a sociolinguistic interview (data not analyzed in the present study), followed by two elicited production tasks, the first in Spanish and the second in Catalan. The productions of these token stimuli were recorded using a Zoom H4N Multitrack Recorder and Comica Lavalier microphone.

The Spanish word list ${ }^{1}$ used in the elicited production task was stratified according to cognate status: 20 words have congruent Catalan cognates (e.g., Sp. conc[e]pto, Cat. conc[e]pte 'concept'); 20 words have incongruent Catalan cognates (e.g., Sp. inter[e]s, Cat. inter[ $\varepsilon$ ]s 'interest'); and 20 words have no Catalan cognate (e.g., Sp. mad[e]ra, Cat. fusta 'wood'). In order to determine the target vowel for each Catalan cognate, an online dictionary with transcriptions of the Barcelona variety of Catalan was consulted (Alcover \& Moll 2002). The Catalan word list consisted only of the 40 congruent and incongruent Catalan cognates. According to the online corpus NIM (Guasch et al. 2013), all words from the Spanish word list have a relative frequency of at least 10 parts per million ( ppm ), and all words from the Catalan word list have a relative frequency of at least 5 ppm . In each word list, all target vowels occur in stressed syllables. Additionally, Spanish words where /e/ is followed by a palatal consonant, or either an /x/ or an $/ \mathrm{r} /$, were excluded, as these segments either lower or raise the F1 of /e/ (Hualde 2013: 115). Before data collection began in Barcelona, four trained linguists who are native speakers of Catalan and/or Spanish participated in a pilot study. After the experiment, none of the participants were able to identify the sound of interest, so to reduce the duration of the experiment, neither word list included filler tokens. Each word list was randomized and all participants saw the same list orders appear on a tablet in the form of isolated words.

### 2.3 Acoustic analysis

A total of 1,020 Spanish mid front vowels ( 17 participants x 20 words x 3 cognate levels) and a total of 680 Catalan mid front vowels ( 17 participants x 20 words x 2 target vowels) were submitted to acoustic analysis. For the Spanish data, timealigned, word- and phoneme-segmented Praat TextGrid files were generated using Montreal Forced Aligner (McAuliffe et al. 2017) with a Spanish dictionary (Morgan 2017). The TextGrids were hand-corrected in Praat (Boersma \& Weenink

[^0]
## Annie Helms

2019), and a Praat script (Riebold 2013) was used to extract measurements for F1, F2, and F3 at the midpoint of each stressed /e/ phone marked in the TextGrid in order to minimize co-articulation effects upon the formant measurements. The same procedure was carried out for the 680 Catalan mid front vowels, and vowel categories were classified following the target vowels in the word list. The F1 and F2 measurements for all mid front vowels of Spanish and Catalan were normalized across vocal tract length, using the Lammert and Narayanan $\Delta \mathrm{F}$ normalization method (Johnson 2020), which can be calculated using only a subset of vowels from the acoustic space.

### 2.4 Statistical analysis

To analyze a possible mid front vowel merger in Catalan, a Pillai score was calculated for the two mid front Catalan vowels for each speaker in the data set and measures of F1 were submitted to a mixed effects linear regression model. Although normally both Pillai scores and measures of Euclidean distance are employed to analyze possible mergers, the two Catalan mid front vowels predominantly differ across F1, so the regression model of F1 measures provides roughly the same information as Euclidean distance. The Pillai score is a measure of the degree of overlap between vowel categories and is calculated for each speaker from multivariate analysis of variance (MANOVA) models fitted with F1 and F2 measurements by vowel category (Nycz \& Hall-Lew 2013). The Pillai scores from each speaker were calculated using a custom function and submitted to a fixed effects linear regression model using the glm() function in R (R Core Team 2018). The model includes a main effect of language dominance and a two-way interaction term of age and gender. The mixed effects linear regression model predicting F1 was built using the lmerTest package (Kuznetsova et al. 2017). This model serves as another indicator of a possible mid front vowel merger, and additionally provides information about variation occurring across this formant axis. The model contains a two-way interaction of language dominance and target catalan vowel, a three-way interaction of GENDER, AGE, and TARGET CATALAN vOWEL, and random intercepts of PARTICIPANT and TOKEN WORD.

In order to observe a possible cognate effect within productions of Spanish /e/, F1 measures were submitted to a mixed effects linear regression model. The model included a two-way interaction of language dominance and catalan COGNATE VOWEL, a three-way interaction between GENDER, AGE, and catalan COGNATE VOWEL, and random intercepts of TOKEN WORD and Participant. Additionally, to observe the impact of social factors on the degree of overlap between Spanish /e/ and Catalan /e/, and between Spanish/e/ and Catalan / $\varepsilon /$, Pillai
scores were calculated for each participant for each vowel category comparison. Scores were submitted to separate fixed effects linear regression models with the two-way interaction between AGE and GENDER and the main effect of language dominance. For these regression models, and all previous models, the emmeans package (Lenth 2021) was used to calculate Cohen's $d$ effect sizes for pairwise comparisons and to perform necessary post-hoc tests using a Tukey pairwise comparison. The heplots package (Fox et al. 2021) was used to calculate partial eta-squared $\left(\eta_{p}^{2}\right)$ effect sizes for fixed effects models, and the r2glmm package (Jaeger 2017) was used to calculate marginal R-squared $\left(R^{2}\right)$ effect sizes for mixed effects models.

## 3 Results

### 3.1 Catalan production

Table 2: Regression coefficients for fixed effects linear model predicting overlap between Catalan /e/ and Catalan $/ \varepsilon /$ (Pillai scores) across the two-way interaction of AGE and GENDER and the main effect of LANguage dominance (Dom.). The intercept is the overlap of older female speakers with a language dominance score of 0 .

|  | Estimate | Std. Error | $t$-value | $p$-value |  |
| :--- | ---: | :--- | ---: | ---: | :--- |
| (Intercept) | 0.3669835 | 0.0822150 | 4.464 | 0.000774 | $* * *$ |
| Younger | -0.1727733 | 0.0900524 | -1.919 | 0.079134 | $*$ |
| Male | -0.1587978 | 0.1126758 | -1.409 | 0.184121 |  |
| Dom. | 0.0003611 | 0.0008584 | 0.421 | 0.681425 |  |
| Younger: Male | 0.2068923 | 0.1626608 | 1.272 | 0.227497 |  |

To look for evidence of assimilation via a cognate effect, it must first be determined if $/ \mathrm{e} /$ and $/ \varepsilon /$ are produced as distinct vowel categories. An initial visual examination of the Catalan mid front vowels (Figure 2) suggests that older females produce more contrasting vowels than other participants. To investigate this observation further, Pillai scores and productions along F1 were analyzed. The coefficients of the regression model of individuals' Pillai scores (Table 2) indicate that neither age, gender, or LANGUAGE DOMINANCE significantly impact the degree of overlap between the two mid front vowel categories. However, the main effect of AGE is approaching significance, where younger speakers produce


Figure 2: Vowel space plot showing distribution of three vowel categories in acoustic space, and displayed across social factors of age and gender. Ellipses are drawn 1 SD from the mean. Formant values appear in normalized units, derived originally from raw hertz values.

Catalan mid front vowels with greater overlap than older speakers ( $\beta=-0.17, \eta_{p}^{2}$ $=0.565, p=0.079$ ).

Regression coefficients for the mixed effect linear regression model predicting Catalan F1 (Table 3) indicate significant main effects of language dominance, TARGET CATALAN VOWEL, GENDER, and a significant interaction of TARGET CATAlan vowel and age. A post-hoc Tukey pairwise comparison of target catalan vowel and AGE reveals that /e/ is produced significantly higher than $/ \varepsilon /$ for older speakers ( $\beta=0.0771, d=1.474, p<0.001$ ) and for younger speakers $(\beta=0.0433, d$ $=0.828, p<0.05)$. Additionally, while the acoustic distinction between the $/ \varepsilon /$ of older speakers and the $/ \varepsilon /$ of younger speakers is not significant, older speakers produce /e/ higher than younger speakers ( $\beta=0.0885, d=1.691, p<0.01$ ). Main effects of LANGUAGE DOMINANCE $\left(\beta=0.00059, R^{2}=0.047, p<0.05\right)$ and GENDER ( $\beta=-0.095, R^{2}=0.068, p<0.01$ ) indicate that mid front vowels are produced lower with increasing Catalan-dominance, and that male productions are higher
than female productions. Although these social factors influence Catalan production, they do not impact the degree of acoustic distinction between the two mid front vowel categories.

Table 3: Regression coefficients for mixed effects linear model predicting Catalan F1, with a three-way interaction of AGE, GENDER, and target vowel and a two-way interaction of language dominance (Dom.) and target vowel. The intercept is older, female speakers producing Catalan $/ \varepsilon /$ with a language dominance score of 0 .

|  | Estimate | Std. Error | $t$-value | $p$-value |  |
| :--- | ---: | ---: | ---: | ---: | :--- |
| (Intercept) | $5.092 \mathrm{e}-01$ | $2.220 \mathrm{e}-02$ | 22.934 | $<2 \mathrm{e}-16$ | $* * *$ |
| Dom. | $5.939 \mathrm{e}-04$ | $2.098 \mathrm{e}-04$ | 2.830 | 0.010688 | $*$ |
| /e/ | $-9.065 \mathrm{e}-02$ | $1.654 \mathrm{e}-02$ | -5.480 | $5.06 \mathrm{e}-07$ | $* * *$ |
| Male | $-9.454 \mathrm{e}-02$ | $2.754 \mathrm{e}-02$ | -3.433 | 0.002790 | $* *$ |
| Younger | $4.266 \mathrm{e}-02$ | $2.201 \mathrm{e}-02$ | 1.938 | 0.067620 | $*$ |
| Dom.: /e/ | $2.003 \mathrm{e}-05$ | $1.019 \mathrm{e}-04$ | 0.197 | 0.844256 |  |
| Male: Younger | $2.405 \mathrm{e}-02$ | $3.976 \mathrm{e}-02$ | 0.605 | 0.552480 |  |
| /e/: Male | $2.514 \mathrm{e}-02$ | $1.338 \mathrm{e}-02$ | 1.879 | 0.060672 | $*$ |
| /e/: Younger | $3.895 \mathrm{e}-02$ | $1.069 \mathrm{e}-02$ | 3.643 | 0.000292 | $* * *$ |
| /e/: Male: Younger | $-1.028 \mathrm{e}-02$ | $1.931 \mathrm{e}-02$ | -0.532 | 0.594572 |  |

### 3.2 Spanish production

From the Spanish word lists, F1 measures of Spanish /e/ were submitted to a mixed effects linear regression model and the output of the model appears in Table 4. The model output indicates a significant main effect of language domINANCE $\left(\beta=0.00052, R^{2}=0.040, p<0.01\right)$, where speakers that are more Catalandominant produce Spanish /e/ lower than speakers that are less Catalan-dominant. The model also reveals a significant main effect of AGE $\left(\beta=0.049, R^{2}\right.$ $=0.033, p<0.05$ ), where younger speakers have lower productions than older speakers. Additionally, the main effect of GENDER is approaching significance, where males produce /e/ higher than females ( $\beta=0.047, R^{2}=0.020, p=0.0524$ ). Importantly, there is no significant interaction containing levels of the factor CATALAN COGNATE VOWEL, which could indicate a cognate effect from Catalan in the production of Spanish/e/. Accordingly, it seems that these participants do not evidence assimilation via a cognate effect to Catalan mid front vowel categories in their production of Spanish /e/.

Table 4: Regression coefficients for mixed effects linear model predicting Spanish F1 with a three-way interaction of AGE, GENDER, and CATAlan cognate vowel ( $\mathrm{NC}=$ non-cognate) and a two-way interaction of language dominance (Dom.) and catalan cognate vowel. The intercept is older, female speakers with a language dominance score of 0 producing Spanish /e/ where the Catalan cognate vowel is $/ \varepsilon /$.

|  | Estimate | Std. Error | $t$-value | $p$-value |  |
| :--- | ---: | ---: | ---: | ---: | :--- |
| (Intercept) | $4.097 \mathrm{e}-01$ | $1.716 \mathrm{e}-02$ | 23.872 | $<2 \mathrm{e}-16$ | $* * *$ |
| Dom. | $5.186 \mathrm{e}-04$ | $1.737 \mathrm{e}-04$ | 2.986 | 0.0071 | $* *$ |
| /e/ | $-2.370 \mathrm{e}-03$ | $1.073 \mathrm{e}-02$ | -0.221 | 0.8253 |  |
| NC | $-5.334 \mathrm{e}-03$ | $1.073 \mathrm{e}-02$ | -0.497 | 0.6194 |  |
| Male | $-4.690 \mathrm{e}-02$ | $2.280 \mathrm{e}-02$ | -2.057 | 0.0524 | . |
| Younger | $4.903 \mathrm{e}-02$ | $1.822 \mathrm{e}-02$ | 2.691 | 0.0138 | $*$ |
| Dom.: /e/ | $-8.966 \mathrm{e}-05$ | $9.308 \mathrm{e}-05$ | -0.963 | 0.3357 |  |
| Dom.: NC | $9.668 \mathrm{e}-06$ | $9.308 \mathrm{e}-05$ | 0.104 | 0.9173 |  |
| Male: Younger | $3.510 \mathrm{e}-02$ | $3.291 \mathrm{e}-02$ | 1.067 | 0.2984 |  |
| le/: Male | $-2.240 \mathrm{e}-03$ | $1.222 \mathrm{e}-02$ | -0.183 | 0.8546 |  |
| NC: Male | $3.753 \mathrm{e}-03$ | $1.222 \mathrm{e}-02$ | 0.307 | 0.7588 |  |
| /e/: Younger | $1.125 \mathrm{e}-02$ | $9.765 \mathrm{e}-03$ | 1.152 | 0.2497 |  |
| NC: Younger | $4.063 \mathrm{e}-03$ | $9.765 \mathrm{e}-03$ | 0.416 | 0.6774 |  |
| le/: Male: Younger | $7.526 \mathrm{e}-03$ | $1.764 \mathrm{e}-02$ | 0.427 | 0.6697 |  |
| NC: Male: Younger | $1.032 \mathrm{e}-02$ | $1.764 \mathrm{e}-02$ | 0.585 | 0.5585 |  |

### 3.3 Cross-linguistic production

To observe how social factors mediate the degree of overlap between Spanish /e/ and each Catalan mid front vowel, Pillai scores measuring the degree of overlap between Spanish /e/ and each Catalan mid front vowel were calculated for each participant. The values were submitted to fixed effects linear regression models with main effects of language dominance and two-way interactions of age and Gender. The regression coefficients for the overlap between Spanish /e/ and Catalan /e/ appear in Table 5. The model output indicates that social factors do not impact the degree of overlap of these two categories. In other words, all participants regardless of age, gender, or language dominance, demonstrate a similar degree of overlap of Spanish /e/ and Catalan /e/. As a relative measure, Pillai scores do not convey a specific percentage of overlap in productions. However, as scores can range from 0 (merged) to 1 (unmerged), the fairly low $\beta$-coefficient of the intercept $(0.156)$ suggests that the categories have a considerable degree of overlap.

Table 5: Regression coefficients for the fixed linear effects model predicting the degree of overlap between Spanish /e/ and Catalan /e/ (Pillai scores) across the two-way interaction of AGE and GENDER and the main effect of language dominance (Dom.). The intercept is the overlap of productions by older, female speakers with a language dominance score of 0 .

|  | Estimate | Std. Error | $t$-value | $p$-value |  |
| :--- | :---: | :---: | :---: | :---: | :---: |
| (Intercept) | 0.1557052 | 0.0552384 | 2.819 | 0.0155 | $*$ |
| Younger | 0.0274665 | 0.0605042 | 0.454 | 0.6580 |  |
| Male | -0.0125400 | 0.0757043 | -0.166 | 0.8712 |  |
| Dom. | 0.0001259 | 0.0005767 | 0.218 | 0.8309 |  |
| Younger: Male | -0.1134477 | 0.1092881 | -1.038 | 0.3197 |  |

Table 6: Regression coefficients for the fixed linear effects model predicting the degree of overlap between Spanish /e/ and Catalan $/ \varepsilon /$ (Pillai scores) across the two-way interaction of AGE and GENDER and the main effect of language dominance (Dom.). The intercept is the overlap of productions by older, female speakers with a language dominance of 0 .

|  | Estimate | Std. Error | $t$-value | $p$-value |  |
| :--- | ---: | :---: | ---: | ---: | :--- |
| (Intercept) | 0.5244873 | 0.0604239 | 8.680 | $1.62 \mathrm{e}-06$ | $* * *$ |
| Younger | -0.1223421 | 0.0661840 | -1.849 | 0.0893 | $*$ |
| Male | -0.2384856 | 0.0828110 | -2.880 | 0.0138 | $*$ |
| Dom. | 0.0012737 | 0.0006309 | 2.019 | 0.0664 | $*$ |
| Younger: Male | 0.0144009 | 0.1195474 | 0.120 | 0.9061 |  |

The regression coefficients for the model predicting the overlap of Spanish $/ \mathrm{e} /$ and Catalan $/ \varepsilon /$ are shown in Table 6 . In this model output, the main effect of GENDER is significant ( $\beta=0.238, R^{2}=0.99, p<0.05$ ), indicating that male speakers produce the two categories with more overlap, relative to female speakers. Additionally, the factors of AGE and language dominance are approaching significance, where younger speakers would evidence more overlap than older speakers ( $\beta=0.122, R^{2}=0.99, p=0.0893$ ) and more Catalan-dominant speakers would evidence less overlap ( $\beta=0.0013, R^{2}=0.99, p=0.0664$ ). The $\beta$-coefficient of the intercept, compared with that of the previously analyzed model, suggests that these participants have greater overlap of Spanish /e/ and Catalan /e/, than of Spanish /e/ and Catalan $/ \varepsilon /$, an observation that is supported visually in Figure 2.

## 4 Discussion

In order to address the first research question, whether there is a cognate effect from Catalan in the production of Spanish /e/, the Catalan data were first analyzed. The results of this analysis indicate that the categories of Catalan /e/ and Catalan $/ \varepsilon /$ have not yet fully merged for these participants, as each category was produced significantly differently across F1 by both older speakers and younger speakers. However, the results of the regression model of individual Pillai scores indicate that the degree of category overlap may be increasing in apparent-time. Since significant acoustic distinction was observed, it could be possible to find cognate effects in the Spanish production data. However, the factor of catalan cognate vowel was not a significant main effect of F 1 of Spanish /e/, nor was it involved in any significant interactions, suggesting that there is no observable cognate effect, regardless of the social profile of the participants analyzed in this study.
The second research question regards the variation due to social factors in the degree of overlap between Spanish /e/ and Catalan /e/, and also between Spanish /e/ and Catalan $/ \varepsilon /$. While the overlap between Spanish /e/ and Catalan /e/ was not seen to be socially-mediated, GENDER was a significant predictor of the degree of overlap between Spanish /e/ and Catalan $/ \varepsilon /$, where the male participants tend to evidence more overlap between the two categories than the female participants. Contrary to the research hypothesis, neither age nor language dominance significantly impacted the degree of overlap, though both factors approached statistical significance.
Regarding the role of age, older speakers maintained less overlap of Catalan mid front categories and greater acoustic distinction across F1, thus demonstrating adherence to prescriptive norms, where Catalan /e/ and Catalan $/ \varepsilon /$ are two distinct phones. This result is in line with the research hypothesis, despite the educational changes that occurred in Catalonia following the Catalonian Linguistic Normalization law of 1983. Though exposure to Catalan instruction in school may be greater for younger speakers, the large number of immigrants that are L2 Catalan speakers attending public schools alongside native Catalan speakers may contribute to increased exposure to exemplars with a merger. Therefore, exposure to Catalan under the new educational policies does not deterministically cause younger speakers to fully adhere to prescriptive norms. These findings therefore support the SLM-r's assertion that "[q]uality of input has been largely ignored in L2 speech research [in favor of quantity] even though it may well determine the extent to which L2 learners differ from native speakers" (Flege \& Bohn 2021: 32).

With regard to the role of gender, the female participants had less overlap between Spanish /e/ and Catalan $/ \varepsilon /$ than the male participants, thus conforming more to the prescriptive norm than males. That gender did not also impact the overlap of Spanish /e/ and Catalan /e/ could be attributed to the lack of salience of this distinction; one participant of this study mentioned during the sociolinguistic interview that students are taught in school that Spanish /e/ and Catalan /e/ are the same sound. Though the factor of age was only marginally significant, a visual examination of Figure 2 suggests that the younger speakers tend to conflate all three vowel categories whereas the older speakers mainly conflate Spanish /e/ and Catalan /e/. With greater statistical power, perhaps age would surface as a significant predictor of overlap, in which case the data would be consistent with a male-led change in progress. Alternatively, the lack of a significant age effect, coupled with female speakers' greater adherence to the prescriptive norm than male speakers, is consistent with a case of stable variation. However, since prior studies (e.g., Mora \& Nadeu 2012) have documented the merger as recent and ongoing, I will proceed considering the merger as a possible change in progress.

Whereas females are often the leaders of community-wide change (Labov 2001), a male-led change could suggest that the social meaning indexed by a production of Spanish /e/ that is conflated with Catalan mid front vowels is stratified by gender. For example, it could be that this variant is an index of solidarity or a Catalan-identity marker (akin to lateral velarization and other phonetic phenomena; Davidson 2019) that is generally associated with males and does not provide social gain if used by females (Chappell 2016). Similarly, because bilinguals that use this variant would only have one mid front vowel category crosslinguistically, the variant could also index some trait associated with metropolitan bilingualism or a unique Barcelona identity that is a blend of Spanish and Catalan identities (Newman \& Trenchs-Parera 2015). Of course, these presently speculative accounts can be empirically tested in future perception research that aims to uncover the social meanings that listeners afford to the assimilation of Spanish /e/ and Catalan $/ \varepsilon /$.

Models of second language acquisition and representation, such as the SLM, SLM-r, PAM-L2, and L2LP, predict language dominance to be an important factor in category formation. Although language dominance did not surface as a significant predictor of production in this study, the participant group was considerably homogeneous in terms of dominance, where only 4 participants were scored as Spanish-dominant by the BLP. A larger data set with more variability in language dominance could be more revealing of the importance of input in language production. Of the three models, the SLM-r should make the most

## Annie Helms

relevant predictions for the data as the participants are early bilinguals rather than naïve listeners or adult learners and the bias towards Catalan-dominance in the data set means that the majority of participants are attempting to produce a single L2 category (Spanish /e/) rather than an L2 category contrast (Catalan $/ \mathrm{e} /$ and $/ \varepsilon /$ ). Under this theory, the quality and quantity of input that participants receive of Catalan /e/ does not allow for sufficient differentiation from Spanish /e/, yielding category assimilation in their production as was seen in the Pillai scores.

Based on SLM-r, any differences in production based on gender, ethnicity, age, or other social factors must necessarily be derived from differences in L2 input. However, the connection between variants and social meaning central to variationist sociolinguistic theory is not accounted for in the SLM-r. Chappell \& Kanwit (2021) attempt to reconcile this disconnect by proposing a unified framework of theoretical models of second language learning, exemplar theory, and indexical meaning to explain variable outcomes in second language perception. The present data also support this unified approach for production, where variable productions may be influenced by the mapping of social meaning onto said variants, in addition to the L2 input received. While cognitive factors such as language dominance and cognate status are central to models of bilingual phonetic representation, it is paramount to situate the bilingual individual within the context of the community and acknowledge the external social factors which also mediate variation in acquisition and production.

## 5 Conclusions

This study found evidence for age and gender as predictors of assimilation among productions of Spanish and Catalan mid front vowels. Younger speakers evidence less acoustic distinction between Catalan mid front vowels than older speakers, and males produce Spanish /e/ and Catalan $/ \varepsilon /$ with more overlap than females. These findings contribute a variationist sociolinguistic analysis to the literature of bilingual production of mid front vowels, demonstrating the importance of viewing models of category acquisition and phonetic representation through the lens of social factors, in conjunction with cognitive factors. Future studies of contact varieties of Spanish found in bilingual settings within Barcelona, both in the realms of production and perception, will further reveal the impact of contemporary Catalan linguistic policies, and the social meaning indexed by variation within the production of these (vocalic) and other linguistic variables.

## Acknowledgements

I would like to thank Justin Davidson, Ernesto Gutiérrez Topete, Ana Belén Redondo Campillos, Bernat Bardagil i Más, Yamel Nuñez, Antonio Torres Torres, the audience members at LSRL50, and two anonymous reviewers for their contributions and insightful feedback. All remaining errors are my own.

## References

Alcover, Antoni Maria \& Francesc de Borja Moll. 2002. Diccionari català-valenciàbalear. https://dcvb.iec.cat/.
Amengual, Mark. 2016a. Cross-linguistic influence in the bilingual mental lexicon: Evidence of cognate effects in the phonetic production and processing of a vowel contrast. Frontiers in Psychology 7. 617. DOI: 10.3389/fpsyg.2016.00617.
Amengual, Mark. 2016b. The perception and production of language-specific midvowel contrasts: Shifting the focus to the bilingual individual in early language input conditions. International fournal of Bilingualism 20(2). 133-152. DOI: 10. 1177/1367006914544988.
Arnal, Antoni. 2011. Linguistic changes in the Catalan spoken in Catalonia under new contact conditions. Journal of Language Contact 4(1). 5-25. DOI: 10.1163/ 187740911x558815.
Bailey, Guy. 2004. Real and apparent time. In J. K. Chambers, Peter Trudgill \& Natalie Schilling-Estes (eds.), The handbook of language variation and change, 312-332. London: Blackwell Publishing. DOI: 10.1002/9780470756591.ch12.
Baker, Wendy \& Pavel Trofimovich. 2005. Interaction of native- and secondlanguage vowel system(s) in early and late bilinguals. Language and Speech 48(1). 1-27. DOI: 10.1177/00238309050480010101.
Birdsong, David, Libby M. Gertken \& Mark Amengual. 2012. Bilingual language profile: An easy-to-use instrument to assess bilingualism. COERLL, University of Texas at Austin. https://sites.la.utexas.edu/bilingual/ (31 August, 2016).
Boersma, Paul \& David Weenink. 2019. Praat: Doing phonetics by computer.
Bosch, Laura \& Marta Ramon-Casas. 2011. Variability in vowel production by bilingual speakers: Can input properties hinder the early stabilization of contrastive categories? Journal of Phonetics 39(4). 514-526. DOI: 10.1016/j.wocn. 2011.02.001.

Carbonell, Joam \& Joaquim Llisterri. 1999. Catalan. In International Phonetic Association (ed.), Handbook of the International Phonetic Association, 53-56. Cambridge: Cambridge University Press.

## Annie Helms

Chappell, Whitney. 2016. On the social perception of intervocalic /s/ voicing in Costa Rican Spanish. Language Variation and Change 28(3). 357-378. DOI: 10. 1017/S0954394516000107.

Chappell, Whitney \& Matthew Kanwit. 2021. Do learners connect sociophonetic variation with regional and social characteristics? The case of L2 perception of Spanish aspiration. Studies in Second Language Acquisition 44(1). 1-25. DOI: 10.1017/S0272263121000115.

Colomé, Àngels \& Michele Miozzo. 2010. Which words are activated during bilingual word production? fournal of Experimental Psychology: Learning, Memory, and Cognition 36(1). 96-109. DOI: 10.1037/a0017677.
Cortés, Susana, Conxita Lleó \& Ariadna Benet. 2019. Weighing factors responsible for the production of the Catalan vowel $/ \varepsilon /$ versus /e/ contrast in three districts of Barcelona. International fournal of Bilingualism 23(6). 1264-1277. DOI: 10.1177/1367006918781058.

Costa, Albert, Alfonso Caramazza \& Nuria Sebastian-Galles. 2000. The cognate facilitation effect: Implications for models of lexical access. Journal of Experimental Psychology: Learning, Memory, and Cognition 26(5). 1283-1296. DOI: 10.1037/0278-7393.26.5.1283.

Davidson, Justin. 2019. Covert and overt attitudes towards Catalonian Spanish laterals and intervocalic fricatives. In Withney Chappell (ed.), Recent advances in the study of Spanish sociophonetic perception, 40-83. DOI: 10.1075/ihll.21. 03dav.
Davidson, Justin. 2020. Asymmetry and directionality in Catalan-Spanish contact: Intervocalic fricatives in Barcelona and Valencia. Languages 5(4). 60. DOI: 10.3390/languages5040060.

Evans, Bronwen G. \& Paul Iverson. 2007. Plasticity in vowel perception and production: A study of accent change in young adults. The fournal of the Acoustical Society of America 121(6). 3814-3826. DOI: 10.1121/1.2722209.
Flege, James E. 1995. Second language speech learning: Theory, findings, and problems. In Strange Winifred (ed.), Speech perception and linguistic experience: Issues in cross-language research, 233-277. Baltimore: York Press.
Flege, James E. 2002. Interactions between the native and second-language phonetic systems. In Petra Burmeister, Thorsten Piske \& Andreas Rhod (eds.), An integrated view of language development: Papers in honor of Henning Wode, 217244. Berlin: Wissenschaftlicher Verlag.

Flege, James E. \& Ocke-Schwen Bohn. 2021. The Revised Speech Learning Model (SLM-r). In Ratree Wayland (ed.), Second language speech learning: Theoretical and empirical progress, 3-83. Cambridge: Cambridge University Press. DOI: 10 1017/9781108886901.002.

Fox, John, Michael Friendly \& Georges Monette. 2021. heplots: Visualizing Tests in Multivariate Linear Models. R package version 1.3-8.
Galindo i Solé, Mireia. 2003. Language contact phenomena in Catalonia: The influence of Catalan in spoken Castilian. In Lotfi Sayahi (ed.), Selected proceedings of the first workshop on Spanish sociolinguistics, 18-29. Somerville: Cascadilla Proceedings Project.
Galindo i Solé, Mireia. 2006. Les llengües a l'hora del pati: Usos lingüístics en les converses dels infants de primària a Catalunya. Barcelona: Universitat de Barcelona. (Doctoral dissertation).
Grosjean, François. 2001. The bilingual's language modes. In Janet Nicol (ed.), One mind, two languages: Bilingual language processing, 1-22. London: Blackwell.
Guasch, Marc, Roger Boada, Pilar Ferré \& Rosa Sánchez-Casas. 2013. NIM: A Webbased Swiss army knife to select stimuli for psycholinguistic studies. Behavior Research Methods 45(3). 765-771. DOI: 10.3758/s13428-012-0296-8.
Hualde, José Ignacio. 2013. Los sonidos del español: Spanish language edition. Cambridge: Cambridge University Press. DOI: 10.1017/cbo9780511719943.
Jaeger, Byron. 2017. r2glmm: Computes R Squared for Mixed (Multilevel) Models. R package version 0.1.2.
Johnson, Keith. 2020. The $\Delta \mathrm{F}$ method of vocal tract length normalization for vowels. Laboratory Phonology: Journal of the Association for Laboratory Phonology 11(1). 10. DOI: 10.5334/labphon.196.
Kendall, Tyler \& Valerie Fridland. 2012. Variation in perception and production of mid front vowels in the US Southern vowel shift. Journal of Phonetics 40(2). 289-306. DOI: 10.1016/j.wocn.2011.12.002.
Kuznetsova, Alexandra, Per B. Brockhoff \& Rune H. B. Christensen. 2017. Lmertest package: Tests in linear mixed effects models. fournal of Statistical Software 82(13). 1-26. DOI: 10.18637/jss.v082.i13.
Labov, William. 1994. Principles of language change: Internal factors. Malden, MA: Blackwell.
Labov, William. 2001. Principles of language change: Social factors, vol. 2. Malden, MA: Blackwell.
Ladefoged, Peter F. \& Keith Johnson. 2015. A course in phonetics. Haidian: Peking University Press.
Lenth, Russell V. 2021. emmeans: Estimated Marginal Means, aka Least-Squares Means. R package version 1.6.0.
Lev-Ari, Shiri \& Sharon Peperkamp. 2013. Low inhibitory skill leads to non-native perception and production in bilinguals' native language. fournal of Phonetics 41(5). 320-331. DOI: 10.1016/j.wocn.2013.06.002.

## Annie Helms

McAuliffe, Michael, Michaela Socolof, Sarah Mihuc, Michael Wagner \& Morgan Sonderegger. 2017. Montreal forced aligner: Trainable text-speech alignment using Kaldi. In Francisco Lacerda (ed.), Interspeech, 498-502. Stockholm: ISCA. DOI: 10.21437/Interspeech.2017-1386.
Mora, Joan C \& Marianna Nadeu. 2012. L2 effects on the perception and production of a native vowel contrast in early bilinguals. International Journal of Bilingualism 16(4). 484-500. DOI: 10.1177/1367006911429518.
Morgan, John. 2017. Santiago Spanish lexicon: A pronouncing dictionary for the Spanish language. www.openslr.org/34/.
Newman, Michael \& Mireia Trenchs-Parera. 2015. Language policies, ideologies and attitudes in Catalonia. Part 1: Reversing language shift in the twentieth century. Language and Linguistics Compass 9(7). 285-294. DOI: 10.1111/lnc3 12141.

Nycz, Jennifer \& Lauren Hall-Lew. 2013. Best practices in measuring vowel merger. Proceedings of Meetings on Acoustics 20(1). 060008. DOI: 10.1121/1 4894063.

Qualtrics. 2005. Qualtrics survey software. https://www.qualtrics.com.
R Core Team. 2018. R: A Language and Environment for Statistical Computing. R Foundation for Statistical Computing. Vienna. https://www.R-project.org/.
Recasens, Daniel \& Aina Espinosa. 2006. Dispersion and variability of Catalan vowels. Speech Communication 48(6). 645-666. DOI: 10.1016/j.specom.2008.09 002.

Riebold, John. 2013. Vowel analyzer. Praat script.
Simonet, Miquel. 2011. Production of a Catalan-specific vowel contrast by early Spanish-Catalan bilinguals. Phonetica 68(1-2). 88-110. DOI: 10.1159/000328847.
Tagliamonte, Sali A. 2006. Analysing sociolinguistic variation. Cambridge: Cambridge University Press.
Thomason, Sara Grey \& Terrence Kaufman. 1988. Language contact, creolization and genetic linguistics. Berkeley, CA: University of California Press.
Yeni-Komshian, Grace H., James E. Flege \& Serena Liu. 2000. Pronunciation proficiency in the first and second languages of Korean-English bilinguals. Bilingualism: Language and cognition 3(2). 131-149. DOI: 10. 1017 / s1366728900000225.


[^0]:    ${ }^{1}$ Spanish and Catalan word lists available at https://anniehelms.github.io/lsrl50_supplemental/

