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Abstract

 The methylation status of the O(6)-methylguanine-DNA methyltransferase (*MGMT*) gene is an important predictive biomarker for benefit from alkylating agent therapy in glioblastoma. Our model MGMT-STP27 allows prediction of the methylation status of the *MGMT* promoter using data from the HumanMethylationBeadChip (Illumina, HM-27K and HM-450K) that is publically available for many cancer datasets. Here we present investigations addressing the impact of the context of genetic and epigenetic alterations and tumor type on the classification, report on technical aspects, such as robustness of cut-off definition and preprocessing of the data. The association between gene copy number variation (CNV), predicted *MGMT* methylation and *MGMT* expression revealed a gene dosage effect on *MGMT* expression in lower grade glioma (WHO grade II/III) that in contrast to glioblastoma usually carry two copies of chromosome 10 on which *MGMT* resides (10q26.3). This implies some *MGMT* expression, potentially conferring residual repair function blunting the therapeutic effect of alkylating agents. A sensitivity analyses corroborated the performance of the original cut-off for various optimization criteria and for most data preprocessing methods. Finally, we propose a R package mgmtstp27 that allows prediction of the methylation status of the *MGMT* promoter and calculation of appropriate confidence and/or prediction intervals. Overall the MGMT-STP27 is a robust model for *MGMT* classification that is independent of tumor type, and is adapted for single sample prediction.

Introduction

 Large scale analyses of the methylome of gliomas have provided relevant insights into tumor biology and cell of origin that has important implications for tumor classification and choice 47 of therapy $^{1, 2}$ $^{1, 2}$ $^{1, 2}$ $^{1, 2}$. The DNA methylation status of the promoter of the O(6)-methylguanine-DNA methyltransferase (*MGMT)* gene that encodes a DNA repair protein is the most important predictive factor for benefit from alkylating agents such as temozolomide in glioblastoma (GBM) $3-6$. However, in anaplastic and low grade glioma a prognostic versus a predictive 51 value is more controversial $6-9$. A principle difference between GBM and lower grade glioma (WHO grade II and III) is the high frequency of mutations in the isocitrate dehydrogenase (IDH) genes 1 or 2 in lower grade glioma that is mechanistically linked with the development 54 of a CpG island methylator phenotype (CIMP+)^{[10](#page-18-4)}. In glioma CIMP is almost invariably associated with *MGMT* promoter methylation regardless of tumor grade as we have reported 56 . previously 11 11 11 . This raises the question whether the mechanistic underpinnings of CIMP may lead to functionally relevant differences in the methylation pattern affecting epigenetic silencing of the *MGMT* gene. It has been shown that DNA hypermethylation in CIMP results from inhibition of α-ketoglutarate-dependent dioxygenases such as the epigenetic modifier TET2, by high concentrations of the oncometabolite 2-hydroxyglutarate produced by the 61 neomorphic enzymatic function of the IDH1 and 2 mutants $^{10, 12, 13}$ $^{10, 12, 13}$ $^{10, 12, 13}$ $^{10, 12, 13}$ $^{10, 12, 13}$. Furthermore, loss of 1 copy of chromosome 10, home of *MGMT* (10q26), is a hallmark of primary GBM (>80%), while it is a rare event in lower grade glioma. Hence in *MGMT* methylated lower grade gliomas *MGMT* could be transcribed from the second potentially intact strand.

 Genome-wide DNA methylation data on human methylation 27K (HM-27K) or 450K (HM- 450K) BeadChips have become publically available for large datasets of glioma. This data can be used to determine the *MGMT* methylation status using our previously developed 68 logistic regression model, MGMT-STP27 11 11 11 . The input into the model are measures of 2 key

 CpG probes located in the *MGMT* promoter that we identified to be functionally highly relevant and which are available on both versions of the chip. The model was trained with a dataset of 63 GBM from homogenously treated patients, for which the *MGMT* methylation status was previously shown to be predictive for outcome, based on classification by methylation-specific PCR (MSP). The MGMT-STP27 model provided good classification properties and prognostic value (kappa=0.85; logrank p<0.001), and has been successfully validated in independent datasets including clinical trials, by us and other groups $2, 9, 11, 14, 15$ $2, 9, 11, 14, 15$ $2, 9, 11, 14, 15$ $2, 9, 11, 14, 15$ $2, 9, 11, 14, 15$. The original preprocessing procedure was based on the conversion of the Red/Green channel from the Illumina methylation array into the methylation signal, without using any normalization. However, the rising interest into epigenetics has stimulated development of methods to analyze DNA methylation data including numerous procedures for normalization 80 and bias correction $16-19$. Triche et al. 17 listed no fewer than seven methods to correct background such as substraction of fifth percentile of negative control distribution (Illumina procedure) and normal-exponential deconvolution (Noob). The use of one of these new procedures may modify the estimation of signal intensities in ways that affect the suitability of the parameters in the current MGMT-STP27 model thereby impacting classification.

 The aim of the present study was to determine the impact of methodological/computational procedures, sample type (frozen versus formalin fixed paraffin embedded, FFPE), and biological context [CIMP, gene copy number alterations (CNA), tumor type] on the evaluation of the *MGMT* status using the MGMT-STP27 method. The functional validity of the classification model, including the previously established cut-off, is tested across tumor grades, CIMP-status, and extended to non-brain tumor entities. This includes the investigation of the spatial correlations of CpG-methylation and *MGMT* expression that informs on the functionality of the methylation to actually impact *MGMT* expression and thereby indicating the potential of the tumor cells for DNA repair. The simultaneous effects of CIMP, promoter

 methylation and gene dosage on *MGMT* expression are evaluated. To complete the sensitivity analysis for the model MGMT-STP27, we investigate how our classifier can be affected by different background and normalization procedures for data from the HM-27K and HM-450K platforms. Finally, we provide a R package called "mgmtstp27" [\(https://github.com/badozor/mgmtstp27\)](https://github.com/badozor/mgmtstp27) that allows easy computation of MGMT-STP27 classification for individual samples, and includes new features such as the calculation of the confidence intervals of the *MGMT* methylation scores (*MGMT* methylation probability), comparison of the score distribution of external datasets with the training set, and quality control.

Materials and methods

Datasets

 Clinical information and DNA methylation data (HM-27K and 450K) from 7 publically available glioma data-sets (761 individuals, 119 WHO grade II, 258 WHO grade III and 384 GBM) were used for this study. The first, originally used as the training set, contained DNA methylation profiles and expression data for 63 GBM tissues from 59 patients treated within 110 clinical trials and five non-tumoral brain tissues (epilepsy surgery) (M-GBM) $^{11, 20, 21}$ $^{11, 20, 21}$ $^{11, 20, 21}$ $^{11, 20, 21}$ $^{11, 20, 21}$. The external datasets used are VB-Glioma-III, from patients treated within a clinical trial (n= 110 112 glioma grade III)⁹[;](#page-18-7) T-Glioma-II/III (29 WHO grade II, 42 grade III)^{[10](#page-18-4)}; and the following datasets from The Cancer Genome Atlas (TCGA, http://cancergenome.nih.gov; [https://tcga-](https://tcga-data.nci.nih.gov/tcga/) [data.nci.nih.gov/tcga/\)](https://tcga-data.nci.nih.gov/tcga/): TCGA-GBM-27, TCGA-GBM-450 (n= 321 GBM) and TCGA- Glioma-II/III (n=197; 90 WHO grade II, 106 WHO grade III, n=1, unspecified grade; website 116 http://cancergenome.nih.gov/) $22-24$. Three additional TCGA datasets for non-brain tumors comprise colon adenocarcinoma (TCGA-COAD, n= 227), breast cancer (TCGA-BRCA, n=

 305, randomly selected from a set of 642 samples), head and neck squamous cell carcinoma (TCGA-HNSC, n=442), and lung squamous cell carcinoma (TCGA-LUSC, n=328). The dbGaP accession number to the specific version of the TCGA data set is phs000178.v9.p8. The datasets and their accession numbers, including their corresponding expression datasets, are described in detail in the Supplemental Table S1. The clinical and molecular baseline description for the glioma datasets is summarized in Supplemental Table S2.

Procedures for preprocessing and *MGMT* **promoter methylation prediction**

 The pipeline for computation of the *MGMT* classification is summarized in Supplemental Figure S1. The prediction of the DNA methylation status of *MGMT* promoter requires the conversion of the Red/Green channel information derived from the Illumina methylation array into signals for methylated and unmethylated, respectively, without normalization. .The M-130 values 25 25 25 (log2-ratio of methylated and unmethylated intensities corrected by an offset equal to 1,) for the methylation probes of interest located in the *MGMT* promoter, cg12434587 and cg12981137 (location see Figure 1) were used as input into the logistic regression model 133 (MGMT-STP27) to predict the methylation status of the $MGMT$ gene 11 11 11 . The calculation of the confidence intervals for the logistic regression model is described ^{[26](#page-19-9)}. The *MGMT* score was obtained by logit-transformation of the probability that the *MGMT* promoter is methylated to obtain a quasi-normal score. The predicted values (probabilities and *MGMT* score), confidence intervals, and *MGMT* classification can be directly obtained by the function MGMTpredict from the R package mgmtstp27 [\(https://github.com/badozor/mgmtstp27\)](https://github.com/badozor/mgmtstp27).

 The effect of normalization and preprocessing of the HM-450K data on the prediction of the *MGMT* status was tested for five additional procedures and compared to the original (raw) 142 preprocessing used for developing the method : control normalization which requires the selection of a reference array (Genome Studio), preprocessing including only background correction, quantile normalization of the separated unmethylated and methylated signals, 145 Subset-quantile within array normalization (SWAN) procedure and Noob normalization, including background correction based on normal-exponential deconvolution with dye-bias 147 correction .

Preprocessing for determination of gene copy number alterations from HM-450K and HM-27K

 Gene copy number alterations (CNA) were calculated basically according to the procedure 152 described by Feber et al ^{[19](#page-19-10)} and adapted for the HM-27k platform and Genome Studio output. 153 As proposed for Illumina Infinium Whole-genome SNP data , the quantile normalization was performed individually for each sample using intensity for unmethylated and methylated signals. The combined intensities for methylated and unmethylated (total intensity, T) was calculated from the normalized intensities. Because matched reference samples were not available, the value *log2(R)* was defined as the difference of intensity between samples and a synthetic reference corresponding to the median profile from a reference dataset containing 159 eight non-tumor brain samples from the TCGA database and M-GBM 11 11 11 .

$$
log2(R) = log2(T_{observed} + 1) - log2(T_{reference} + 1)
$$

 An additional smoothing procedure was applied to remove the wave bias for more accurate 161 breakpoint detection in profiles . The unmethylated and methylated intensities from chemistry II (see Illumina technical sheet; [http://www.illumina.com/content/dam/illumina-](http://www.illumina.com/content/dam/illumina-marketing/documents/products/datasheets/datasheet_humanmethylation450.pdf)[marketing/documents/products/datasheets/datasheet_humanmethylation450.pdf\)](http://www.illumina.com/content/dam/illumina-marketing/documents/products/datasheets/datasheet_humanmethylation450.pdf) were

 corrected by a scaling factor method to reduce the chemistry type-bias before the computation of the total intensity. As indicated above, probes with non-significant p-values (typically $166 \rightarrow 0.01$) were excluded from our analysis when raw data served as input.

Determination of gene copy alteration state

169 For determination of CNA the R package CGHcall 29 29 29 was used that performs circular binary 170 segmentation (CBS)^{[30](#page-19-14)} starting with normalized $log2(R)$ values for each sample. Afterwards, 171 each probe (CpG) was classified by a mixture model 29 29 29 into five classes: amplified, gained, normal, deleted and homozygously deleted. For genomic region (or gene), the CNA events were detected in using copy number probe means (CpGs) contained in the selected region (e.g. chromosomal arms 1p and 19q, region of 10q26.3).

Statistical Analysis

 CIMP positive tumors were identified using unsupervised clustering methods (Ward's 178 algorithm with Euclidean distance) as previously reported . The relationships between categorical variables were assessed by Chi-squared tests with *p* values computed by Monte 180 Carlo simulation, because cell counts were expected to be less than five .

 The classical two-way ANOVA is replaced by Monte-Carlo version to test the effects of CNA and DNA methylation on expression of *MGMT* based on F-statistics (two-way ANOVA-like 183 approach) $32, 33$ $32, 33$ $32, 33$, this method is more robust for the unbalanced data and non-normal assumption for the distribution of the data.

 Evaluation of cut-off robustness, including determination of optimal values and performances was tested for six criteria (cost functions) using the training dataset (M-GBM) for which 187 classification by MSP is also available, which served as gold standard : maximization of 188 sensitivity and specificity, $MaxSpSE^{34}$ $MaxSpSE^{34}$ $MaxSpSE^{34}$; maximization of the product of sensitivity and specificity, *MaxProdSpSe*³⁵; equality (balance) of sensitivity and specificity, *SpEqualSe*^{[36](#page-20-4)}; 190 maximization of the Youden's index ³⁷;maximization of the accuracy, *MaxEfficiency* ^{[38](#page-20-6)}; and 191 maximization of the Kappa index, $MaxKappa^{39}$ $MaxKappa^{39}$ $MaxKappa^{39}$. The optimal values and performances were provided by the R packages OptimalCutpoints^{[40](#page-20-8)} and epiR. The statistical tests, analyses and graphical representations were performed using R-3.2.0.

Results

Epigenetic context of *MGMT* **promoter methylation and expression of** *MGMT*

 The fact that almost all CIMP+ glioma are predicted to have a methylated *MGMT* status using the MGMT-STP27 model $9, 11, 15$ $9, 11, 15$ $9, 11, 15$ raised the question whether the functional correlation of the pattern of *MGMT* promoter methylation and *MGMT* expression is similar between CIMP+ and CIMP- glioma and thus the prediction model remains valid. The spatial pattern of the correlations between methylation of the 19 individual CpGs (7 for 27K) interrogated in the *MGMT* promoter region and *MGMT* expression is displayed separately for CIMP+ and CIMP- gliomas across tumor grades (WHO II, III, IV) (Figure 1). It was similar between CIMP+ and CIMP- gliomas, and across tumor grades. As previously observed, CpG methylation close to the initiation start site (ISS) displayed little correlation with expression. Methylation at the two CpGs (cg12434587 and cg12981137) comprised in the MGMT-STP27 model consistently exhibited substantial negative correlation with expression of *MGMT,* with maximal values close to -0.5, regardless of glioma subtype, CIMP-status, and tumor grade

 (Figure 1). The pattern was also very similar in colon adenocarcinoma (TCGA-COAD), head and neck cancer (TCGA-HNSC), and lung squamous cell carcinoma (TCGA-LUSC), but not in breast cancer (TCGA-BRCA) (Supplemental Figure S2). In the latter, correlation between expression and methylation is very weak. However *MGMT* methylation is rare (see below).

 The distribution of the *MGMT* score (logit-transformed probability of methylation) revealed bimodal distributions for all glioma subtypes clearly separating methylated from unmethylated (Figure 2, CIMP+ and CIMP- cases are visualized separately) and were almost superimposable onto the original GBM training set (M-GBM). Similar bimodal distributions were obtained for TCGA-COAD, TCGA-HNSC and TCGA-LUSC, while TCGA-BRCA basically only displays a peak for *MGM*T unmethylated tumors (Figure 3). The original cut- off, based on the maximized sum of sensitivity and specificity of the training cohort (M- GBM) was located at the nadir (lowest point between two populations) of the density plots in all glioma subpopulations, and including other tumor types, hence efficiently differentiating *MGMT* unmethylated and methylated (Figure 2 & 3). The majority of CIMP+ samples were *MGMT* methylated across all glioma datasets (Figure 2). Of note, samples with codeletion of 1p/19q were without exception *MGMT* methylated and displayed a high *MGMT* score confirmed in other datasets by other groups using MGMT-STP27 $^{14, 15}$ $^{14, 15}$ $^{14, 15}$ $^{14, 15}$ $^{14, 15}$. The calculated proportions of *MGMT* methylation were 36.6% in TCGA-COAD, 31.2% for TCGA-HNSC, 16.2% in TCGA-LUSC, and 4.3 % in the TCGA-BRCA population (Figure 3) in line with the 228 literature . A meta-analysis based on 13 colon cancer studies using different technologies 229 and comprising 2772 cases $42-53$ revealed 37% (Supplemental Figure S3) that is in good agreement with the *MGMT* methylation proportion detected by MGMT-STP27 model in TCGA-COAD.

Robustness of the cut-off to varying optimization criteria

 The assessment of cut-off robustness was conducted to determine how the definition of cut- off points would influence the dichotomization into unmethylated and methylated subgroups using the M-GBM dataset for which *MGMT* classification based on MSP is available. Six criteria (cost functions, see methods) were used to determine the optimal cut-off. Four yielded the same cut-off as obtained originally for the MGMT-STP27 model (0.358, Table 1). A different cut-off of 0.405 was obtained by two of the procedures (Table 1) that balance the errors among false positives (FP) and false negatives (FN) (as previously defined based on 241 MSP)^{[11](#page-18-5)}. The use of this cut-off value reduced the sensitivity by 6%, but only slightly 242 improved the specificity $\langle 2\%, \rangle$, while it had minor impact on the rate of good classification accuracy (Table 1). When testing the second cut-off (0.405) on the 788 glioma samples, we only identified five discrepancies, two for the training dataset (M-GBM), two for the TCGA- Glioma-II/III dataset and one for the T-Glioma-II/III dataset. No discrepancy was observed for TCGA-GBM-27, TCGA-GBM-450, and VB-Glioma-III datasets.

Association of CNA at the *MGMT* **Locus and CIMP status on Expression of** *MGMT*

 Loss of the chromosomal region comprising the *MGMT* gene (10q26) is common in GBM (>80%) as opposed to lower grade glioma. We assessed, whether there is a statistical relation (an "effect") between gene dosage, methylation, and expression of the *MGMT* gene using an additive model. Promoter methylation significantly affected *MGMT* expression in all glioma subtypes and grades (Table 2). Loss of 10q26 had a significant effect on expression in the lower grade glioma populations (p-value=0.003, T-Glioma-II/III; p-value=0.001, TCGA- Glioma-II/III; Table 2), while the effect was not significant in GBM (p-value=0.692, TCGA-GBM-450; p-value=0.848, TCGA-GBM-27; p-value=0.544, M-GBM; Table 2, Figure 4). In the other cancer types, we observed that promoter methylation was significantly associated with *MGMT* expression (p-value=0.001, TCGA-COAD; p-value=0.001, TCGA-HNSC; p- value=0.001 TCGA-LUSC; Table 2, Supplemental Figure S4). No significant associations were detected between 10q26.3 deletion and *MGMT* expression, but such deletion events were rare in TCGA-LUSC (4%), TCGA-COAD (2%) and TCGA-HNSC (2%) datasets that can affect the robustness of the statistical tests (Table 2).

263 The interaction between deletion and methylation was not significant (p=0.196, Monte-Carlo ANOVA with 999 permutations) in the TCGA-Glioma-II/III dataset, suggesting an additive effect. The other datasets could not be analyzed because the distributions of patients in each cross-category were highly unbalanced, in particular due to the high frequency of loss of one copy of chromosome 10 in GBM that harbors *MGMT* (10q26) that can reduce the power of the statistical tests. Further, the CIMP status did not significantly affect the expression of the *MGMT* gene (Supplemental Table S3 and Supplemental Figure S5) in the LGG populations and it was not reasonably testable in the GBM populations considering the very low frequency of this event (7%, Supplemental Table S2).

Effect of tumor matrix (frozen versus FFPE)

 The beadchip platform can be used for frozen and with the addition of a restoration step also for formalin fixed paraffin embedded (FFPE) samples. Here we tested whether datasets originating from different sample matrices can be combined. The VB-Glioma-III dataset, containing 51 frozen samples and 59 FFPE samples, was analyzed (Supplemental Table S1). The distributions of the *MGMT* scores calculated for FFPE and frozen samples, respectively, were not significantly different (p=0.253, Kolmogorov-Smirnov test, Supplemental Figure S6). Furthermore, the original cut-off of 0.3582 efficiently differentiated the unmethylated and methylated *MGMT* promoters for FFPE tissues. Hence, the two datasets were combined for the present study.

Effect of data preprocessing

 The datasets M-GBM and TCGA-GBM-450 were used to compare five normalization and preprocessing procedures for HM-450K with the original (raw) preprocessing used to build the model MGMT-STP27 (Figure 5, Supplemental Figure S7, Supplemental Table S4). The control normalization and preprocessing including only background correction lead to a slight underestimation of the methylation probabilities compared to the standard procedure. However, we only observed three (2.5%) differently reclassified samples for TCGA-GBM- 450 (Figure S7) and four (5.9 %) for the training dataset, M-GBM (Figure 5). The background correction based on normal-exponential deconvolution (Noob) (Supplemental Table S4) similarly underestimated the methylation probabilities. Five and four samples were misclassified for TCGA-GBM-450 and M-GBM, respectively. In contrast, the SWAN normalization resulted in a slight overestimation of the methylation probabilities. Five (4.1%) and one (1.5%) reclassified samples were detected for TCGA-GBM-450 and M-GBM, respectively (Supplemental Table S4). In contrast, the concordance between the initial classification and outputs resulting from a procedure using quantile normalization separately on each signal was extremely low (Figure 5C and Supplemental Figure S7C), indicating incompatibility between this procedure and the current MGMT-STP27 default parameters.

 For the HM-27K platform, we investigated the cohort of 241 TCGA GBM samples (TCGA- GBM-27) and compared the *MGMT* scores obtained with raw data (TCGA level 1) and already preprocessed data including Noob background correction (Level 2, preprocessed data) (Supplemental Figure S7F and G, Supplemental Table S4). The methylation probabilities

 trended to be underestimated for data from Level 2 (Supplemental Figure S7G), with 9 (3.7%) 306 misclassified samples in comparison with the original results . The use of Level 1 (raw) data provided similar predictions as originally determined.

 In spite of a moderate bias for probability estimation, the final *MGMT* classification was robust for both Infinium platforms, except for quantile normalization. The effect of data preprocessing on classification was limited. The strong bimodal distribution of the *MGMT* scores and the low proportion of samples contained in the intermediate probability range [0.3; 0.7] favor this robust behavior.

Discussion

 In the present study we tested the robustness of the MGMT-STP27 model to predict the *MGMT* methylation status. Considerations included biological effects, such as the context of pathogenetic and epigenetic alterations of the tumors analyzed. On the other hand we investigated technical issues, ranging from impact of tissue matrix to preprocessing of the data and cut-off definitions.

 First, we demonstrated that the functional relationship, corresponding to the pattern of the spatial correlation between methylation and expression was preserved across glioma subtypes, WHO grade and CIMP-status, and was also valid in other tumor types. The probes of the two CpGs used in the MGMT-STP27 model displayed a strong negative correlation between methylation and expression in all datasets. Clear bimodal distributions of the *MGMT* scores allowing classification into methylated and unmethylated samples was conserved across all datasets. The original cut-off used for dichotomization was located at the nadir of the distributions in all datasets analyzed including the non-glioma tumor cohorts. The robustness

 of the original cut-off was further confirmed by comparing different procedures of cut-off optimization that had little effect on classification.

 An essential issue for any model is the estimation of the uncertainty related to the prediction. The computation of the confidence intervals as proposed in the new R package mgmtstp27 permits evaluation of the pertinence and quality of the classification for a new sample as we have reported previously 11 . The implemented quality control procedures allow visualization of multiple or single sample predictions in comparison to the training set (Figure 6). The confidence intervals on the methylation status probability are important to assess the confidence in the classification, particularly useful when the prediction is close to the cut-off. This is clinically relevant in particular when deciding not to give TMZ, e.g in clinical trials 338 where patients are selected according to their *MGMT* status , or to use TMZ as monotherapy, as recommended for elderly patients whose GBM is *MGMT* methylated ^{[4,](#page-18-8) [55](#page-21-0)}. In other tumor types, like metastatic colon cancer, alkylating agents may be a treatment option among 341 others ^{[56](#page-21-1)}, and only patients with a higher *MGMT* score may be considered.

 A significant effect of gene dosage on *MGMT* expression was observed in LGG that usually have two gene copies in contrast to GBM. This may indicate that not both copies are methylated, which cannot be distinguished by the assay, potentially yielding some expression conferring residual repair function in these tumors. In other words, residual *MGMT*-related resistance to TMZ may not be excluded in LGG, even when they are classified methylated. In GBM the effect of gene dosage was not statistically evaluable due to the characteristic high frequency of loss of one copy of chromosome 10, home of *MGMT*. In contrast, no effect on expression was observed for CIMP in LGG, while it was not testable in GBM. However, it is of note that the *MGMT* status in LGG is not independent of CIMP due to the nested relationship.

 The effect of preprocessing on the classification was relatively moderate for the tested scenarios, except for quantile normalization that is clearly not suitable. For the other methods, the effect on classification was minor due to the strong bimodal distribution with few samples close to the cut-off. Additionally, the classification robustness can be explained by the limited difference of the probe specific bias in M-values among background correction methods for 357 Infinium chemistry type I probes . This corroborates our previous results 11 showing that the M-value distributions of the two selected probes from the training dataset (M-GBM) and TCGA-GBM-27 were not significantly different.

 A major constraint for direct inter-study prediction are normalization procedures, such as quantile methods, as they can be affected by biological differences in the sample populations across studies and by study design (e.g. presence or absences of control or non-tumor samples, overrepresentation of subgroups). Testing of five preprocessing/normalizing procedures revealed that quantile normalization was clearly not compatible with MGMT- STP27, while for the other four only moderate differences were observed. Unless the compatibility is tested, we recommend to use the raw data (format IDAT), and convert the Red/Green channel from the Illumina methylation array into methylation signal, without using any normalization. This avoids potential dataset dependent biases associated with normalization procedures and allows for single sample prediction that is an essential r_{eq} requirement for clinical utility 57 . In practice, functions such as preprocessRaw or 371 methylumIDAT from the R packages minfi 58 58 58 and methylumi 59 59 59 offer appropriate solutions to import and to preprocess the raw HM-450K and HM-27K data.

 Overall the MGMT-STP27 is a robust model for classification of samples into *MGMT* methylated and unmethylated that is independent on glioma subtype, is adapted for single sample prediction, and is also valid in other tumor types.

Note Added in Proof

 The new Infinium MethylationEPIC BeadChip (850K) proposed by Illumina contains both probes used in the model MGMT-STP27. The annotations (eg, chemistry type and probe location) suggest that our model can be extended to this new platform.

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581 ^{*}See methods for explication of Criterion: *Youden*, maximization of Youden's index; *MaxEfficiency*, maximization of accuracy; *MaxKappa*, maximization of Kappa index; *MaxProdSpSe*, maximization of product of sensitivity and specificity; *SpEqualSe*, equality (balance) of sensitivity and specificity; *MaxSpSE*: maximization of sensitivity and specificty † ⁵⁸⁵ The maximization of the sum of specificity and sensitivity used for developing MGMT-STP27^{[11](#page-18-5)} was identical to the maximization of Youden's index.

587 Abbreviations: FP, false positives; FN, false negatives; prev meth, prevalence of methylation;

588 sens, sensitivity; spec, specificity; diag acc; diagnostic accuracy; Youden, Youden index

590 **Table 2** Effects of CNA and DNA methylation on expression of *MGMT* in Glioma and Non-

591 Glioma tumors.

592 CNA 10q26.3 very common event, unbalanced data!

† 593 10q26.3 loss very rare event, unbalanced data!

[#] simulated p-values estimated by Monte-Carlo procedures (999 permutations); significant p-595 values are indicated in bold.

Bady et al. J Mol Diagn *18*, 350-61, 2016

Figure Legends

 Figure 1. Spatial correlation between *MGMT* expression and CpG methylation in the *MGMT* promoter. The correlation between the Infinium probes, in the *MGMT* promoter (genome assemble 37, hg19) present on the 450K and the 27K, respectively, and expression of *MGMT* is displayed for 5 glioma datasets (AFFYmetrix probe, ; RNA sequencing for TCGA-Glioma II/III). The black, green and red line correspond to the correlation for all samples, CIMP- and CIMP+ populations respectively. The CpG island located in the *MGMT* promoter region is illustrated with a green bar, and the location of the two Inifinium HM-450K/27K probes used in the model MGMT-STP27 are indicated with dark blue marks, and the transcription start site (TSS) with an arrow.

 Figure 2. Distribution of the *MGMT* scores in glioma grade II-IV stratified by CIMP-status. The density plots of the *MGMT* scores, corresponding to the logit-transformed probabilities (*MGMT* score) that the *MGMT* promoter is methylated, are shown for the LGG (grade II and III) and GBM (grade IV) populations. The smoothened lines are provided by kernel density estimate, and indicate in green grade IV (GBM), in red grade III, and in blue for grade II glioma. The vertical dotted lines identify the position of the cut-off used to classify in into methylated and unmethylated *MGMT* promoter status.

 Figure 3. Distribution of *MGMT* score for non-Glioma datasets from TCGA. The score corresponds to the logit-transformed probabilities that *MGMT* promoter is methylated. The black smoothened line is provided by kernel density estimate. The vertical dotted line 620 identifies the position of the cut-off used to determinate the *MGMT* promoter state ^{[11](#page-18-5)}. The

 proportion of *MGMT* methylation for head and neck cancer (TCGA-HNSC) is 138/442 (31.2%, 95% confidence interval [CI, 26.9-35.8%]), 53/328 (16% [CI, 12.3-20.6%]) for lung squamous cell carcinoma (TCGA-LUSC), 13/305 (4.3% [CI, 2.3-7.2%]) for breast carcinoma (TCGA-BRCA), and 83/227 (36.6% [CI, 3.0-4.3]) for colon adenocarcinoma (TCGA-COAD).

 Figure 4. Boxplot representation of *MGMT* expression in function of CNA and *MGMT* methylation status in glioma grade II to IV. For each dataset the number of samples for each subpopulation is provided next to the box. Subpopulations with deletions at 10q26.3 (del) are indicated in white, the ones with normal copy number (no-del) in black. *MGMT* methylated, M; *MGMT* unmethylated, U.

 Figure 5. Effect of data preprocessing procedures on *MGMT* classification. Paired comparisons of the probabilities of *MGMT* promoter methylation (MGMT-STP27) between preprocessing procedures for the M-GBM dataset. Five preprocessing procedures for the HM- 450K platform were compared with the initial procedure used to build the model MGMT- STP27. The outputs from recommended preprocessing were compared with (A) outputs from the Illumina-like procedure based on control normalization (a reference sample was used during the normalization step), (B) preprocessing with Illumina-like background correction only, (C) quantile normalization, (D) SWAN normalization, and (E) Noob normalization. Each dataset contained exactly the same samples. The grey dashed lines identify the original 642 cut-off of 0.3582. The straight, dashed black line corresponds to the equation $y=x$ and the grey line to the loess regression, respectively. The proportions of good classification (diagnostic accuracy, DA) are provided for the original cut-off on each panel.

 Figure 6. Quality control visualization for multi-sample and single sample predictions from R package mgmtstp27. The M-values of the two probes *cg12434587* and *cg12981137* are illustrated in (A) for multi-sample predictions and (D) for single sample prediction. The inertia ellipses identify the training dataset and the dots correspond to the location of the new sample prediction. The red and blue colors visualize methylated and unmethylated status, respectively. (B) illustrates the comparison of the *MGMT* score distribution of a new multi- sample dataset (black curve) with the training dataset (M-GBM, green curve, histogram). For single sample prediction, the new sample is indicated by the black vertical line (E). The multi- sample predictions (*MGMT* score and Probabilities) for the dataset TCGA-GBM-27 (black points and lines) associated with their prediction intervals (grey polygons) are shown in (C). The prediction for the sample TCGA-02-0057 from the dataset TCGA-GBM-27 is indicated in (F) associated with the prediction interval. As reference, the green curve and grey polygons correspond to the prediction and confidence intervals for the training dataset (M-GBM).

Multi-sample predictions (dataset TCGA-GBM-27)

Supplementary Figures ¬ Tables

Sensitivity analysis of the MGMT-STP27 model and impact of genetic/epigenetic context to predict the *MGMT* **methylation status in gliomas and other tumors**

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Running head: Sensitivity analysis MGMT-STP27

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Legends Supplementary Figures

Figure S1. Pipeline for computation of *MGMT* classification using the R package mgmtstp27. The R package minfi and methylumi can be used to import and to preprocess raw data. The prediction of the DNA methylations status of *MGMT* promoter requires preprocessed intensities for the signals for unmethylated and methylated as initially proposed for HM-27k in Illumina Genome Studio software in 2009-2011 and originally used in TCGA database. For raw HM-450K data, this operation was performed by the function preprocessRaw from R package minfi. When the raw IDAT format was not available, we assumed an adequate normalization procedure.

Figure S2. Spatial correlation between *MGMT* expression and CpG methylation in the *MGMT* promoter for Non-Glioma Tumors from TCGA. The correlation between expression and DNA methylation for the Infinium HM-450K probes in *MGMT* promoter (genome assemble 37, hg19) is given for TCGA-COAD, TCGA-BRCA, TCGA-HNSC and TCGA-LUSC datasets. The green rectangle corresponds to the CpG island located in the *MGMT* promoter region and the two dark blue rectangles identify the location of the two Inifinium HM-450K/27K probes used in the model MGMT-STP27.

Figure S3. Forest plot of the meta-analysis for the proportion of *MGMT* methylation in colon cancer. The calculation of an overall proportion of *MGMT* methylation from 13 studies (2779 patients). This analysis used logit transformation and inverse variance method. DerSimonian-Laird estimate was used in the random effects model and Clopper-Pearson intervals were given for *MGMT* proportion in each study ('exact' binomial interval).

Figure S4. Boxplot representation of *MGMT* expression in function of CNA and *MGMT* methylation status in non-Glioma datasets from TCGA (TCGA-COAD, TCGA-BRCA,TCGA-HNSC and TCGA-LUSC). For each dataset the number of samples for each subpopulation is provided next to the box. Subpopulations with deletions at 10q26.3, del; subpopulations with normal copy number, no-del; *MGMT* methylated, M; *MGMT* unmethylated, U.

Figure S5. Boxplot representation of *MGMT* expression in function of CIMP status and *MGMT* methylation status in glioma grade II to IV. The number of samples for each subpopulation is provided next to the box for each dataset. The combined effect of the two variables CIMP status and *MGMT* methylation status on the expression of *MGMT* was not efficiently testable because the data was strongly unbalanced. Presence of CIMP, CIMP+; absence of CIMP, CIMP-; *MGMT* methylated, M; *MGMT* unmethylated, U.

Figure S6. Comparison of *MGMT* score distributions (logit-transformed probability) among FFPE and Frozen Tissues from VB-Glioma-III dataset. The *MGMT* score distributions were represented by histogram for frozen tissue (A, n=51), for FFPE tissue (B, n=59) and for aggregated data (C, n=110). The dotted, dashed and solid red curves correspond to kernel density estimates for frozen tissues, FFPE tissues and all samples. The vertical dashed black line identifies the position of the cutoff used to determinate the *MGMT* promoter state (0.3582). The QQ-plot representation (D) compares the *MGMT* score distributions from Frozen and FFPE data (VB-Glioma-II/III). The distributions were compared by Smirnov-Kolmogorov tests (D=0.187, p-value=0.253). The solid red line corresponds to line of equation $y=x$.

Figure S7. Effect of preprocessing procedures on *MGMT* classification. Paired comparison of the probabilities that *MGMT* promoter was methylated to evaluate the effect of preprocessing procedure for TCGA datasets (TCGA-GBM-450, TCGA-Glioma-II/III). Five preprocessing procedures for the HM-450K platform were compared with the initial procedure used to build the model MGMT-STP27. The outputs from recommended preprocessing were compared with outputs from (A) Illumina-like procedure based on control normalization (a reference sample was used during the normalization step), (B) preprocessing with Illumina-like background correction only, (C) quantile normalization, (D) SWAN normalization and (E) Noob normalization. Each dataset contained exactly the same samples. The predictions from the level 1 (F) and level 2 (G) for HM-27k data from TCGA GBM database were compared with outputs of the originally calculated probabilities 11 . The grey dashed lines identify the original cut-off of 0.3582. The straight, dashed black line corresponds to the equation $y=x$ and the grey line to the loess regression, respectively. The proportions of good classification (diagnostic accuracy, DA) are provided for the original cut-off on each figure.

Meta-analysis for *MGMT* **methylation proportion in colon cancer**

0.2 0.3 0.4 0.5 0.6 *MGMT* methylation proportion

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 0.2 $\overline{0}$

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0.0 0.2 0.4 0.6 0.8 1.0

probabilities from background correction only (level2)

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0.0 0.2 0.4 0.6 0.8 1.0

probabilities from raw preprocessing (level1)

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Table S1. Description of datasets

† Accession number: Gene Expression Omnibus, [www.ncbi.nlm.nih.gov/](http://www.ncbi.nlm.nih.gov/geo/)**geo**/ *;*The Cancer Genome Atlas (TCGA),<https://tcga-data.nci.nih.gov/tcga/>

Table S2. Description of the main clinical and molecular variables of the Glioma datasets (WHO grade II, III and IV).

* The proportions were associated with their exact binomial confidence intervals at 95%.

† The age was encoded in three categories: young for age ≤ 40 , middle for age > 40 and ≤ 60 and for age > 60. ‡ one missing value

Dataset (N)		Type Variables	% (N)	F-statistic	[†] Pvalue
M-GBM (59)	GBM	MGMT meth	55.93 (33)	10.933	0.003
		CIMP+	6.78(4)	0.232	0.627
TCGA-GBM-27 (212)	GBM	MGMT meth	50.94 (108)	141.068	0.001
		CIMP+	8.02(17)	2.154	0.145
TCGA-GBM-450 (67)	GBM	MGMT meth	43.28 (29)	8.103	0.008
		້CIMP+	5.97(4)	0.529	0.46
TCGA-Glioma-II/III (195)	LGG	MGMT meth	84.62 (165)	19.114	0.001
		CIMP+	81.54 (159)	0.002	0.97
T-Glioma-II/III (48)	LGG	MGMT meth	85.42 (41)	9.374	0.005
		$CIMP+$	75 (36)	0.002	0.97

Table S3. Effects of CIMP and DNA methylation status on expression of *MGMT*.

* CIMP+ very rare event, unbalanced data!

† simulated p-values estimated by Monte-Carlo procedures (999 permutations)

Table S4. Description of preprocessing and normalization procedures for HM-27K and HM-450K.

