

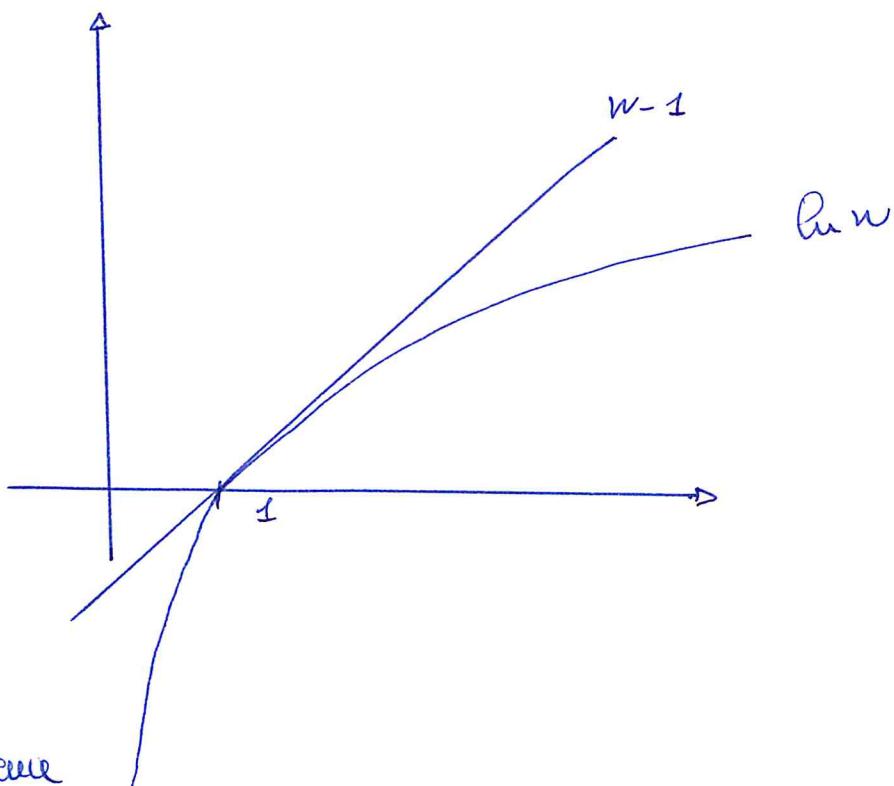
Proofs For TIC

06/06/13

if and only if

⊕

#1 Prove that $\rho_n(n) \leq n-1$, with $\iff n=1$.

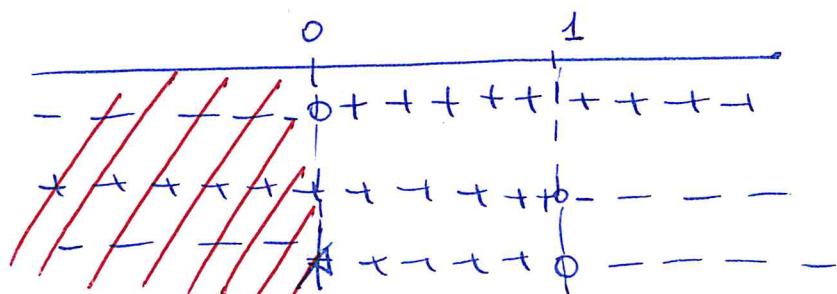


Studying the difference

? $\rho_n(n) - n + 1$

$$\frac{d\rho_n(n)}{dn} = \frac{d(\rho_n(n) - n + 1)}{dn} = \frac{1}{n} - 1 = \frac{1-n}{n}$$

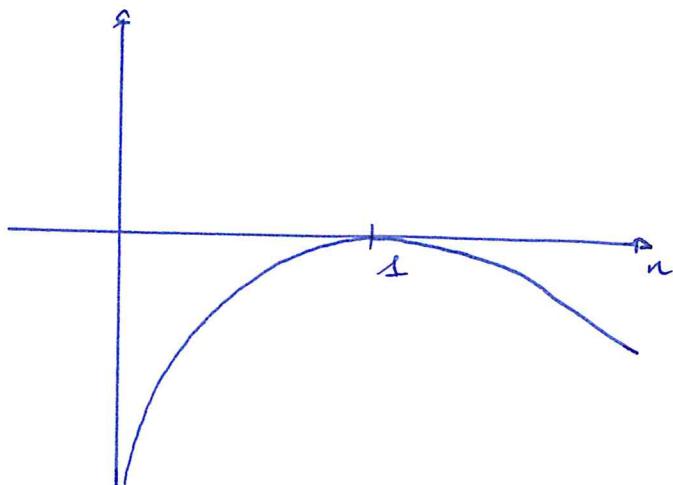
$$1-n \geq 0 \quad n \leq 1 \quad ; \quad n \neq 0$$



$\rho_n(n)$
 $n \in \mathbb{R}^+$

$$\frac{d^2 f(n)}{dn^2} = -\frac{1}{n^2} \quad -\frac{1}{n^2} > 0 \quad \forall n \in \mathbb{R} \Rightarrow \text{max!}$$

\Rightarrow The maximum is in $n=1$



$$\ln n - n + 1 @ n=1 \Rightarrow 0 - 1 + 1 = 0 \Rightarrow \ln n - n + 1 @ n=1$$

Information

$$I(n) = \log \frac{1}{p(n)}$$

Info given from the occurrence
of the event n .

If $\log = \log_2 \Rightarrow I(n) [\text{Sh}]$ Shannon.

Prop.

w_i, w_j independent,

$$p(w_i, w_j) = p(w_i)p(w_j) \Rightarrow I(w_i, w_j) = \log \left(\frac{1}{p(w_i)} \cdot \frac{1}{p(w_j)} \right) = I(w_i) + I(w_j)$$

Entropy

$$H(X) = \sum_n p(n) \log\left(\frac{1}{p(n)}\right) = \mathbb{E}[I(x)] [Sh]$$

If $p(n) \sim U[\text{all}]$

$$P(n_i) = \frac{1}{L}$$

$$H(P(n_i)) = \frac{1}{L} \sum_{i=1}^{N=L} \log\left(\frac{1}{L}\right) = -\frac{1}{L} \log \frac{1}{L} = \underline{\log L}$$

- $0 \leq H(X) \leq \log L$

$$H(n) \geq 0$$

Hp: $P(n) \geq 0$, $-\log(P(n))$ given $0 \leq P(n) \leq 1 \Rightarrow \geq 0$

$$\Rightarrow H(n) \geq 0$$

$$H(n) \leq \log_2 L$$

$$\begin{aligned} H(n) - \log_2 L &= -\sum_n p(n) \log_2(p(n)) - \sum_n p(n) \log_2 L \\ &= \sum_n p(n) \log_2\left(\frac{1}{p(n)L}\right) \end{aligned}$$

Given $\ln n \leq n-1 \Rightarrow \log_2 n \leq \log_2 e(n-1)$

$$\begin{aligned} &\leq \sum_n p(n) \log_2\left(\frac{1}{p(n)L} - 1\right) = \log_2 e \sum_n \left(\frac{1}{L} - p(n)\right) \\ &= \log_2 e (1-1) = 0 \end{aligned}$$

$$\Rightarrow H(n) \leq \log_2 L \Rightarrow 0 \leq H(n) \leq \log_2 L$$

The entropy is never < 0 !

Given x, y

- $H(X) = \sum_n p(n) \log \left(\frac{1}{p(n)} \right)$
- $H(Y) = \sum_y p(y) \log \left(\frac{1}{p(y)} \right)$
- $H(X, Y) = \sum_n \sum_y p(n, y) \log \frac{1}{p(n, y)}$
- $H(X|Y=y_i) = \sum_n p(n|Y=y_i) \log \frac{1}{p(n|Y=y_i)}$
- $H(X|Y) = \mathbb{E}[H(X|Y=y_i)] = \sum_y \sum_n p(y) p(n|Y=y_i) \log \frac{1}{p(n|Y=y_i)}$
 $= \sum_y \sum_n p(n, y) \log \frac{1}{p(n|Y=y_i)}$
- $H(X, Y) = H(X|Y) + H(Y)$
 $p(n, y) = p(n|Y=y_i) p(y)$
 $\sum_n \sum_y p(n|Y=y_i) p(y) \log \frac{1}{p(n|Y=y_i) p(y)} = 1$
 $= \sum_n \sum_y p(n|Y=y_i) p(y) \log \frac{1}{p(n|Y=y_i)} + \sum_n \sum_y p(n|Y=y_i) p(y) \log \frac{1}{p(y)}$
 $\sum_y \sum_n p(y) p(n|Y=y_i) \log \frac{1}{p(n|Y=y_i)} + \sum_y p(y) \log \frac{1}{p(y)} = H(X|Y) + H(Y)$

$$H(X|Y) \leq H(X)$$

$$H(X|Y) - H(X) = \sum_n \sum_y p(y) p(n|Y=y_i) \log \frac{1}{p(n|Y=y_i)} - \sum_n p(n) \cancel{\log} \frac{1}{p(n)}$$

$$= \sum_n \sum_y p(n,y) \left(\log \frac{1}{p(n|Y=y_i)} \right) - \sum_n \sum_y p(n,y) \log \frac{1}{p(n)}$$

$$= \sum_n \sum_y p(n,y) \left(\log \frac{1}{p(n|Y=y_i)} + \log p(n) \right)$$

$$= \sum_n \sum_y p(n,y) \left(\log \frac{p(y)}{p(n,y)} + \log p(n) \right)$$

--- es

$$\text{Where } p(n|Y=y_i) \cdot p(y) = p(n,y)$$

$$\Rightarrow p(n|Y=y_i) = \frac{p(n,y)}{p(y)}$$

$$= \sum_n \sum_y p(n,y) \log \frac{p(y)p(n)}{p(n,y)} \quad \text{es}$$

$$\leq \sum_n \sum_y p(n,y) \log_2 e \left(\frac{p(y)p(n)}{p(n,y)} - 1 \right)$$

$$\cancel{\log_2} \leq \left(p(n,y) + p(y)p(n) \right) = \log_2 e (1-1) = 0$$

$$H(X|Y) - H(X) \leq 0$$

$$H(X|Y) \leq H(X)$$

• Chain Rule

$$H(X, Y, Z) = H(X|Y, Z) + H(Y, Z) = H(X|Y, Z) + H(Y|Z) + H(Z)$$

$$H(X_1, X_2, \dots, X_n) = \sum_{i=1}^n H(X_i | X_{i-1}, \dots, X_1)$$

$$H(X_1, X_2, \dots, X_n) \leq \sum_{i=1}^n H(X_i)$$

$$H(X_k | X_{k-1}, \dots, X_1) \leq H(X_k)$$

\Rightarrow for the chain rule $\Rightarrow H(X_1, X_2, \dots, X_n) \leq \sum_{i=1}^n H(X_i)$

* Discrete Memory-Less Source (DMS)



rand. procen discrete values,
time discrete, ergodic
(memory less & stationary)

$$X_i \in \mathcal{X} = \{\tilde{w}_1, \tilde{w}_2, \dots, \tilde{w}_m\}$$

$$P(X_{n1}=w_1, X_{n2}=w_2, X_{n3}=w_3, \dots) = \prod_{i=1}^m P(X_{ni}=w_i)$$

Every symbol is independent.

Entropy of DMS

$$H(X) = H(X_m) \quad \left[\begin{array}{c} \text{Sh} \\ \text{symbol} \end{array} \right]$$

The same memory less channel has each time the same information

Kullback-Leibler Distance

Given $\underline{a} = (a_1, a_2, a_3, \dots, a_n)$ and \underline{b} ,

let $\sum a_i = 1$ and $\sum b_i = 1$

$$\sum_i a_i \log \frac{a_i}{b_i} \geq 0$$

$$\begin{aligned} \sum a_i \log \frac{a_i}{b_i} &\Rightarrow -\sum a_i \log \frac{b_i}{a_i} \leq \log e \sum a_i \left(\frac{b_i}{a_i} - 1 \right) \\ &= \log e \sum (b_i - a_i) = \log e (\sum b_i - \sum a_i) = 0 \end{aligned}$$

$$\sum a_i \log \frac{b_i}{a_i} \leq 0$$

↓

$$\sum a_i \log \frac{a_i}{b_i} \geq 0$$

It can be extended to the continuous case

$$\int a(u) \log \frac{a(u)}{b(u)} du \geq 0$$

Let X_1, X_2, \dots be stationary processes.

$$P\left\{X_1=w_1, X_2=w_2, \dots, X_M=w_M\right\} = P\left\{X_{l+\ell}=w_\ell, X_{l+2+\ell}=w_2, \dots, X_{n+\ell}=w_n\right\}$$

$$H(X_n) = H(X)$$

If it's possible to look at more symbols

$$H(X_1, X_2, X_3, X_4, \dots, X_k)$$

\Rightarrow The mean info for every symbol is

$$\frac{H(X_1, X_2, \dots, X_k)}{k}$$

and

$$\frac{H(X_1, X_2, \dots, X_k)}{k} \leq \frac{\sum_{i=1}^k H(X_i)}{k} = H(X).$$

The ENTROPY OF THE SOURCE or SOURCE ENTROPY is

$$H_\infty = \lim_{k \rightarrow \infty} \frac{H(X_1, X_2, \dots, X_k)}{k}$$

If the process is stationary, this is equal to

$$H'_\infty = \lim_{n \rightarrow \infty} H(X_n | X_{n-1}, \dots, X_1)$$

- $H'_\infty(x) = H_\infty(x)$ if stationary

$$\underbrace{H(X_{k+1} | X_k, \dots, X_1)}_{a_{k+1}} \leq H(X_{k+1} | X_k, \dots, X_2) = \underbrace{H(X_k | X_{k-1}, \dots, X_1)}_{a_k}$$

$a_{k+1} \leq a_k \Rightarrow$ The series is decreasing \Rightarrow
It converges

$$H_\infty(x) = \lim_{K \rightarrow \infty} \frac{H(X_K, X_{K-1}, \dots, X_1)}{K} = \lim_{K \rightarrow \infty} \frac{\sum_{i=1}^K H(X_i | X_{i-1}, \dots, X_1)}{K}$$

Zero Mean

$$a_n \xrightarrow{n \rightarrow \infty} a$$

$$\frac{1}{N} \sum_{i=1}^N a_i = b_n \quad b_n \xrightarrow{n \rightarrow \infty} a$$

$$= \lim_{K \rightarrow \infty} \frac{1}{K} \sum_{i=1}^K \underbrace{H(X_i | X_{i-1}, \dots, X_1)}_{a_i} = H_\infty(x)$$

Given that $b_i \xrightarrow{i \rightarrow \infty} H_\infty(x)$, for zero the $a_i \xrightarrow{i \rightarrow \infty} H_\infty(x)$.

But $a_i \xrightarrow{i \rightarrow \infty} = H'_\infty(x)$

$$\Rightarrow \boxed{H'_\infty(x) = H_\infty(x)}$$

SOURCE ENCODING AND DATA COMPRESSION



If the symbols are equiprobable and $\log_2 L = H_s(x)$
 The info should be preserved, so

$$H_s(x) \cdot B_s = \log_2 L \cdot B_s'$$

$$\Rightarrow \frac{H_s(x)}{\log_2 L} = \frac{B_s'}{B_s} \leq 1 \quad \text{= if mol. and uniform.}$$

- Non singular code

$$C(w_i) \neq C(w_j) \quad \forall i \neq j$$

- Uniquely decodable code

$$C(w_1), C(w_2), \dots \rightarrow w_1, w_2, w_3, \dots$$

each codeword brings to the original code.

- Prefix code (instant code)

There is not any codeword that is prefix to another one.

- Average code length

$$L(C) = E[\ell(x)] = \sum_n p(n) \ell(n) \quad \left[\begin{array}{c} \text{bit} \\ \text{source sym} \end{array} \right]$$

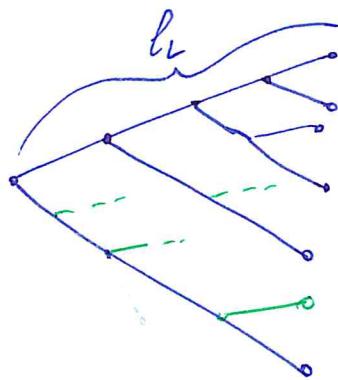
For any prefix code there is a binary tree.

- Kraft's Inequality

For any prefix code with $l_1 \leq l_2 \leq l_3 \leq \dots \leq l_L$ codeword length

$$\sum_{i=1}^L 2^{-l_i} \leq 1$$

Furthermore, for any $l_1 \leq l_2 \leq l_3 \leq \dots \leq l_L$ that satisfies Kraft there is a prefix code.



The total nodes are 2^{l_L} . If I remove from l_1 , ~~the tree~~, the nodes will be $\sum_{i=1}^L 2^{l_L - l_i} \leq 2^{l_L}$

This means $\sum_{i=1}^L 2^{-l_i} \leq 1$

The same reasoning could be done from the end.

- McMillan: The Kraft's Inequality is valid for any uniquely decodable codes.

Source Encoding Theorem : SHANNON 1948

Any code applied to a stationary source with X_n and entropy $H(X)$ has to obey

$$H(X) \leq E[\bar{e}(n)] \leq H(X) + 1$$

with more symbols, the mean entropy "

$$\frac{H(X)}{K} \leq \frac{E[\bar{e}(n)]}{K} \leq \frac{H(X)}{K} + \frac{1}{K}$$

$$\text{if } K \rightarrow \infty \Rightarrow E[\bar{e}(n)] = H(X)$$

• D.M

$$H(X) \leq E[\bar{e}(n)]$$

$$\begin{aligned} H(X) - E[\bar{e}(n)] &= \sum_n p(n) \log \frac{1}{p(n)} - \sum_n p(n) \bar{e}(n) \\ &= \sum_n p(n) \left(\log \frac{1}{p(n)} - \log_2 2^{\bar{e}(n)} \right) = \sum_n p(n) \log \left(\frac{1}{p(n) 2^{\bar{e}(n)}} \right) \\ &\leq \log_2 e \sum_n p(n) \left(\frac{1}{p(n) 2^{\bar{e}(n)}} - 1 \right) = \log_2 e \left(\sum_n \frac{1}{2^{\bar{e}(n)}} - \sum_n p(n) \right) \\ &= \log_2 e \left(\underbrace{\sum_n 2^{-\bar{e}(n)}}_{\leq 1} - \underbrace{\sum_n p(n)}_{1} \right) \leq 0 \end{aligned}$$

$$\underline{E[\bar{e}(n)] \geq H(n)}$$

$$H(X) + 1 \leq E[\ell(u)]$$

To show is used the Shannon-Fano code

$$\ell(u) = \sup \left[\log_2(p(u)) \right]$$

$$\underbrace{\log_2(p(u))}_{\Downarrow} \leq \ell(u) < \log_2(p(u)) + 1$$

$$2^{\log_2(p(u))} \leq 2^{\ell(u)} \quad \sum_n p(n) \leq \sum_n 2^{\ell(u)}$$

$$1 \leq \sum_n 2^{\ell(u)} \quad \Rightarrow \sum_n 2^{-\ell(u)} \leq 1$$

It satisfies the rule, so it is a possible code.

$$\log_2(p(u)) \leq \ell(u) < \log_2(p(u)) + 1$$

$$\sum_n p(n) \log_2(p(u)) \leq \sum_n p(n) \log_2(p(u)) < \sum_n p(u) \log_2(p(u)) + \sum_n p(u)$$

$$H(x) \leq E[\ell(u)] < H(x) + 1$$

HUFFMAN (1952)

3 steps

- Order the probabilities in a decreasing order.
- Sum the two least prob and create a new node
- Repeat and redo the steps

Block Encoding

$$K \rightarrow \infty \quad \frac{H(X)}{K} \leq \frac{\mathbb{E}[l(n)]}{K} < \frac{H(X) + 1}{K}$$

→ 0

$$\Rightarrow H(X) = \mathbb{E}[l(n)]$$

Typical Sequences

DMS

$$X = X_1, X_2, \dots, X_n \quad \text{iid} \quad x_i \in \mathcal{W} = \{\tilde{w}_1, \dots, \tilde{w}_L\}$$

With $n \gg 1$

$$X = \underbrace{w_1 w_1 \dots w_1}_{n p(w_1)} \underbrace{w_2 \dots}_{n p(w_2)} \underbrace{w_3 \dots}_{n p(w_3)} \dots$$

For the large number's law, $n p(w_i)$ times that occurrence

$$P(X) = n p(w_1)^{n p(w_1)} p(w_2)^{n p(w_2)} \dots = \prod_{i=1}^L p(w_i)^{n p(w_i)}$$

$$\log_2 P(X) = \sum_i n p(w_i) \log(p(w_i)) = -n H(X)$$

$$\log_2 P(X) = -nH(X)$$

$$P(X) = 2^{-nH(X)}$$

$$\text{Number of typical sequences} \Rightarrow \frac{1}{P(X)} = 2^{nH(X)}$$

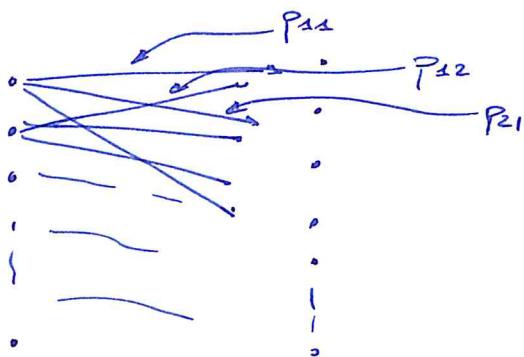
Transmission channels and capacity

DMC, Discrete Memory-Less Channel

DMS, Discrete Memory-Less Source

The DMC channel is defined by its transition matrix \underline{P}

$$\underline{P} = \{P_{ij}\} = \Pr\{Y = y_j | X = x_i\}$$



Let recall

$$H(X), H(Y) \quad \text{and} \quad H(X|Y) \leq H(X) = \text{.p.f. indip.}$$

The Mutual Information is defined as

$$I(X;Y) = H(X) - H(X|Y) \quad [\text{Sh sym}]$$

It is the reduction of the uncertainty after the observations of the output Y .

$$0 \leq I(X;Y) \leq H(X)$$

$$I(X;Y) = I(Y;X) = H(X) - H(X|Y) = H(Y) - H(Y|X)$$

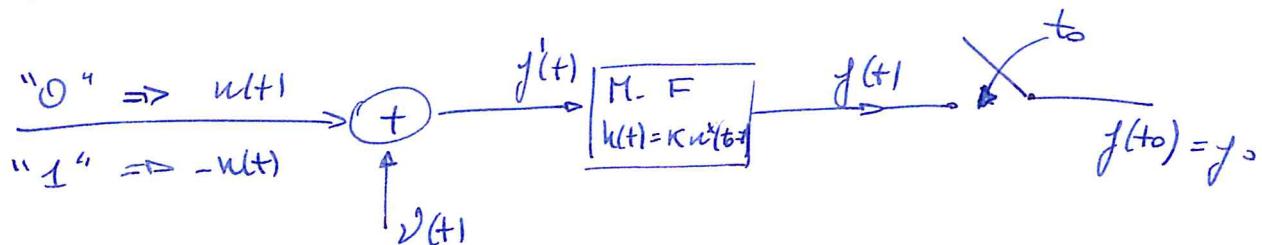
$$\begin{aligned} H(X) - H(X|Y) &= H(X) - H(Y|X) - H(X) + H(Y) \\ &= H(Y) - H(Y|X) \end{aligned}$$

Channel Capacity.

$$C = \max I(X;Y)$$

$$\left[\frac{S_y}{S_{\text{sym}}} \right] = \left[\frac{S_y}{S_{\text{chan. use}}} \right]$$

Antipodal - Patched Filter



$$v(t) = \text{AWGN}$$

$$f G_D(f)$$

$$h(t) = K w^*(t_0 - t) \quad \text{MATCHED FILTER}$$

$$y(t) = \underbrace{h(t) \otimes v(t)}_{y_R(t)} + \underbrace{h(t) \otimes w(t)}_{y_S(t)}$$

$$y_S(t) = \int_{\mathbb{R}} n(\xi) h(t - \xi) d\xi = \int_{\mathbb{R}} n(\xi) K w^*(t_0 - (t - \xi)) d\xi =$$

$$y_s(t_0) = \int_{-\infty}^{\rho} n(\xi) K n^*(t_0 - t_0 + \xi) d\xi = K \int_{\mathbb{R}} n(\xi) n^*(\xi) d\xi = K \underbrace{\int_{\mathbb{R}} |n(\xi)|^2 d\xi}_{E_x}$$

$$= K E_x$$

$$y_u(t) = \int_{\mathbb{R}} h_v(f) |H(f)|^2 df = \frac{N_0}{2} \int_{\mathbb{R}} |H(f)|^2 df \stackrel{\text{Parseval}}{=} \frac{N_0}{2} \int_{\mathbb{R}} |K(t)|^2 dt$$

$$= \frac{N_0}{2} \int_{\mathbb{R}} K^2 |n(t_0 - t)|^2 dt = \frac{N_0 k^2}{2} \int_{\mathbb{R}} |n(t)|^2 dt = \frac{N_0 k^2}{2} E_n$$

That is the power of the noise. At this point it is not still white but "colored" by the filter. Anyway it is still Gaussian with $\mu = 0 \Rightarrow y_u \sim \mathcal{N}(0, \sigma^2 = \frac{N_0 k^2 E_n}{2})$

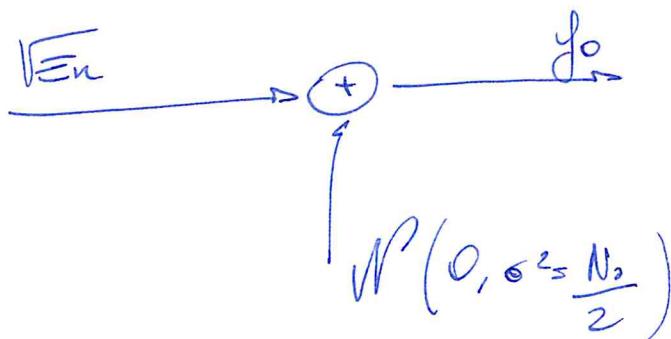
The total output signal

$$y_o = \mathcal{N}\left(\pm \sqrt{E_n}, \sigma^2 = \frac{N_0 k^2 E_x}{2}\right)$$

If we choose $K = \frac{1}{\sqrt{E_x}}$ it follows

$$\underline{y_o = \mathcal{N}\left(\pm \sqrt{E_n}, \sigma^2 = \frac{N_0}{2}\right)}$$

That is equivalent to write

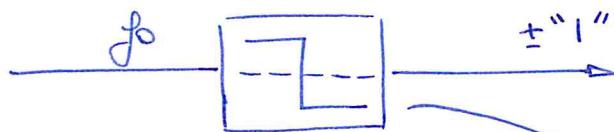


The equivalent time-discrete is referred to as $s(t)$.

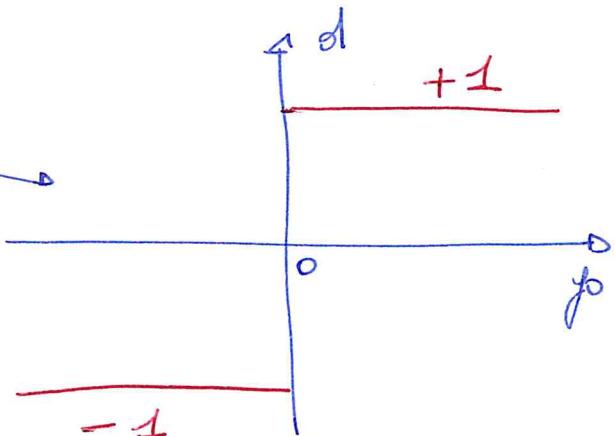
That was "Antipodal Transmission over the AWGN channel, matched filter".

Symbol recognition

- Hard Decision (HD)



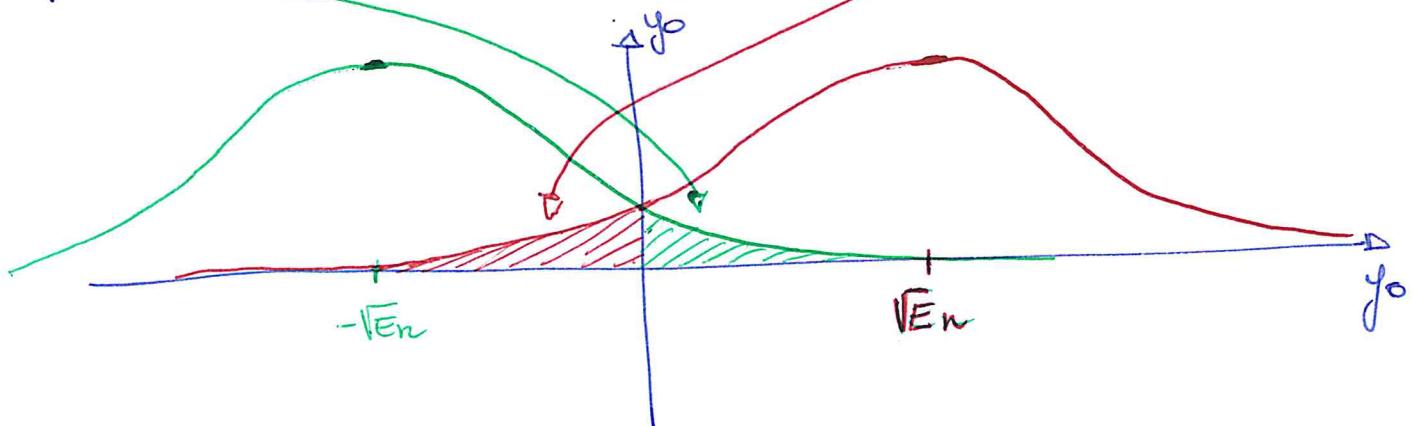
$$y_0 \sim \mathcal{N}(\pm \sqrt{E_n}, \sigma^2 = \frac{N_0}{2})$$



Since the signal is symmetric

$$P(d=1 | -n(+)) = P(y > 0 | -n(+)) = P(d=-1 | n(+)) = P(y < 0 | n(+))$$

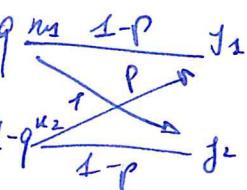
$$= P$$



$$P_e = P_e = \frac{1}{2} \operatorname{erfc} \frac{\sqrt{E_n}}{\sqrt{2} \frac{N_0}{2}} = \frac{1}{2} \operatorname{erfc} \sqrt{\frac{E_n}{N_0}} = P$$

- Being symmetric, it is possible to consider a BSC channel with error prob. p .

Since $C_{BSC} = 1 - H(p)$ where $H(p)$ is the entropy of a Bernoulli's Distribution with parameter "P"



$$I(X;Y) = H(Y) - H(Y|X)$$

$$H(Y|X) :$$

$$H(Y|X=n_1) = \mathcal{H}(\rho)$$

$$H(Y|X=x_2) = \mathcal{H}(\rho)$$

$$\begin{aligned} H(Y|X) &= p(u_1) \mathcal{H}(P) + p(u_2) \mathcal{H}(P) = \mathcal{H}(P)(p(u_1) + p(u_2)) \\ &= \mathcal{H}(P) \end{aligned}$$

$H(Y)$:

$$P(Y_1) = q(1-p) + (1-q)p = q - qp + p - qp = q + p - 2qp$$

$$P(Y_2) = (1-q)(1-p) + q \cdot p = 1 - q - p + qp = 1 - q - p + 2q$$

$\{ P(Y_1) + P(Y_2) = 1 \}$

$$\Rightarrow H(Y) = \mathcal{H}(q+p-2qp)$$

Since it is symmetric and $H(Y)=1$, it is easy to think that the user capacity is with

$$\Rightarrow \underline{C = 1 - \mathcal{H}(\rho)}$$

To use more time the BNC does not increase the capacity

$$I(X^{(n)}; Y^{(n)}) \leq nC$$

$$I(X^{(n)}; Y^{(n)}) = H(Y^{(n)}) - H(Y^{(n)}|X^{(n)})$$

$$H(Y^{(n)}) \leq \sum_{i=1}^n H(Y_i)$$

$$H(Y^{(n)}|X^{(n)}) = \sum_{i=1}^n H(Y_i|Y_{i-1}, \dots, Y_1, X_n, X_{n-1}, \dots, X_1)$$

Since it is memory less:

$$= \sum_{i=1}^n H(Y_i|X_i)$$

$$\Rightarrow I(X^{(n)}; Y^{(n)}) \leq \sum_{i=1}^n H(Y_i) - \sum_{i=1}^n H(Y_i|X_i) = \sum_{i=1}^n H(Y_i) - H(Y_i|X_i)$$

$$\approx nI(X; Y)$$

$$I(X^{(n)}; Y^{(n)}) \leq nI(X; Y) = nC$$

\vdots Channel Coding

$$\text{Code Rate } R_c = \frac{\log_2 L}{n}$$

Before the ch. encoder the channel cap. is

$$\frac{H_{\infty}(W)}{T}$$

After the encoder is

$$\frac{H_{\infty}(W)}{T_c}$$

This capacity has to be the same

$$\frac{H_{\infty}(w)}{T} = \frac{H_{\infty}(n)}{T_c}$$

$$\frac{H_{\infty}(w)}{H_{\infty}(n)} = \frac{T}{T_c} \quad \frac{H_{\infty}(w)}{C} \leq \frac{T}{T_c} \quad \text{where } H_{\infty}(n) \leq C$$

$$\frac{T_c}{T} \leq \frac{C}{H_{\infty}(w)}$$

if $H_{\infty}(w) = \log_2 L$

$$\frac{T_c}{T} \leq \frac{C}{\log_2 L} \quad \frac{1}{m} \leq \frac{C}{\log_2 L} \quad m \geq \frac{\log_2 L}{C}$$

$$C \geq \frac{\log_2 L}{m}$$

The code rate should be less than the channel capacity.

CHANNEL CODING THEOREM, Shannon 1948

-) $H_E > 0, R_c < C$

there are codes that, using an high m , allows the error probability ~~to be~~ to be under ϵ

-) Every code with $\epsilon \rightarrow 0$ ($m \rightarrow \infty$) has to have $R_c < C$.

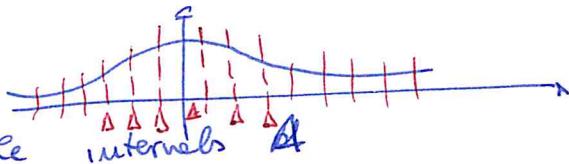
R.V.s continues

We define the Differential Entropy "

$$h(X) = \int_{\mathbb{R}} p(n) \log \frac{1}{p(n)} dn$$

It comes from:

Given a p.d.f.



We take some little intervals Δ

$$P(n_i) = \int_{n_i}^{n_i + \Delta} p(n) dn = p(n_i) \Delta$$

(there is a point inside the interval that multiplied for Δ gives $P(n_i)$)

$$H(X) = \sum_n P(n_i) \log \left(\frac{1}{p(n_i)} \right) = \sum_n p(n_i) \Delta \log \left(\frac{1}{p(n_i) \Delta} \right)$$

$$= \underbrace{\sum_n p(n_i) \log \left(\frac{1}{p(n_i)} \right) \Delta}_{\approx 1} + \underbrace{\sum_n p(n_i) \log \left(\frac{1}{\Delta} \right) \Delta}_{\approx 1}$$

$$\Delta \rightarrow 0 \Rightarrow \Delta = dn \quad \sum_n \Rightarrow \int_{\mathbb{R}}$$

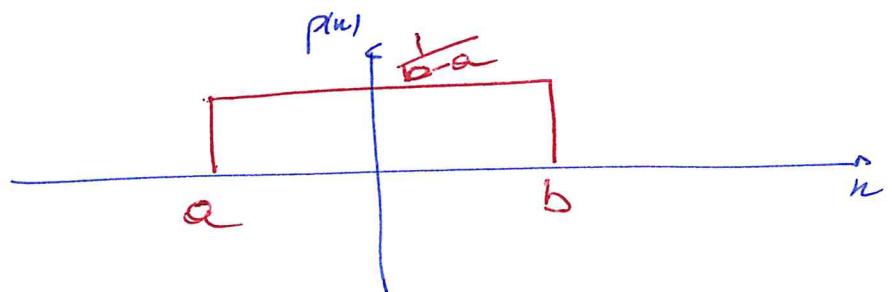
$$\int_{\mathbb{R}} p(n) \log \left(\frac{1}{p(n)} \right) dn + \log \left(\frac{1}{\Delta} \right) \approx \infty$$

This is interesting if the entropy is needed to calculate the entropy or something differential. It is not a measure of the uncertainty, since it should be ∞ .

$$I(X;Y) = h(Y) - h(Y|X)$$

- Uniform distribution

$$X \sim U[a, b]$$



$$h_u(x) = \int_{\mathbb{R}} p(u) \log \frac{1}{p(u)} du = \int_a^b \frac{1}{b-a} \log(b-a) du$$

$$= \underline{\log(b-a)}$$

It could be even < 0!

- The uniform distr. has the highest $h(x)$ for every limited interval distr.

$$h_u(u) - h(u)$$

$$\log(b-a) - \int_{\mathbb{R}} p(u) \log \left(\frac{1}{p(u)} \right) du = \int_{\mathbb{R}} p(u) \log(b-a) - \int_{\mathbb{R}} p(u) \log \left(\frac{1}{p(u)} \right) du$$

$$(b-a) = \frac{1}{p(u)}$$

$$= \int_{\mathbb{R}} p(u) \log \left(\frac{p(u)}{p(u)} \right) \geq 0 \Rightarrow h_u(u) \geq h(u)$$

$$X \sim N(\mu, \sigma^2)$$

$$p(n) = \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{(n-\mu)^2}{2\sigma^2}}$$

$$h(x) = \int_{\mathbb{R}} p(n) \log \frac{1}{p(n)} dn = \int_{\mathbb{R}} p(n) \log \sqrt{2\pi\sigma^2} dn + \int_{\mathbb{R}} p(n) \log \left(+ \frac{(n-\mu)^2}{2\sigma^2} \right) dn$$

$$= \log \sqrt{2\pi\sigma^2} + \underbrace{\frac{1}{2\sigma^2} \int_{\mathbb{R}} p(n) (n-\mu)^2 dn}_{\sigma^2} \cdot \text{Lage} = \log \sqrt{2\pi\sigma^2} + \frac{1}{2} \text{Lage}$$

$$= \log \sqrt{2\pi\sigma^2} + \log \sqrt{e} = \underline{\log \sqrt{2\pi\sigma^2 e}}$$

The Gaussian distr. is the one with the highest $h(x)$ for the non limited intervals distr.

$$h_a(x) - h(x) = \log \sqrt{2\pi\sigma^2 e} - \int_{\mathbb{R}} p(n) \log \frac{1}{p(n)} dn$$

$$= \int_{\mathbb{R}} p(n) \log \sqrt{2\pi\sigma^2 e} + \int_{\mathbb{R}} p(n) \log p(n) dn$$

$$= \int_{\mathbb{R}} p(n) \log \frac{1}{p(n)} + \int_{\mathbb{R}} p(n) \log p(n) dn$$

To demonstrate take $\int_{\mathbb{R}} p(n) \log \frac{1}{p(n)}$ and you will come to

$$= \int_{\mathbb{R}} p(n) \sqrt{2\pi\sigma^2 e} dn$$

$$= \int p(n) \log \left(\frac{p(n)}{p_{\text{a}}(n)} \right) dn \geq 0$$

$$h_u(n) \geq h(n)$$

• Exponentiel dn

$$p(n) = \begin{cases} ae^{-bn} & n \geq 0 \\ 0 & n < 0 \end{cases}$$

Normalisieren

$$\int_R ae^{-bn} dn = 1 \quad \left[\frac{ae^{-bn}}{-b} \right]^{+\infty} = + \frac{a}{b} = 1 \quad a=b$$

$$p(n) = \begin{cases} ae^{-an} & n \geq 0 \\ 0 & n < 0 \end{cases}$$

$$h_{\text{exp}}(x) = \int_R p(n) \log \frac{1}{p(n)} dn = a \int e^{-an} \log(ae^{-an}) dn$$

$$= a \underbrace{\int_{-1}^1 e^{-an} \log a dn}_{=1} + a \int e^{-an} \log(e^{-an}) dn$$

$$= \log a + a^2 \int e^{-an} n \log e dn$$

$$= \log a + a \underbrace{\int_{\mu}^{a e^{-\mu}} n d n}_{\text{in}} \cdot \log e$$

$$= \log a + \mu \log e$$

The expm. has the highest $h(x)$ of every in illuminated distn.

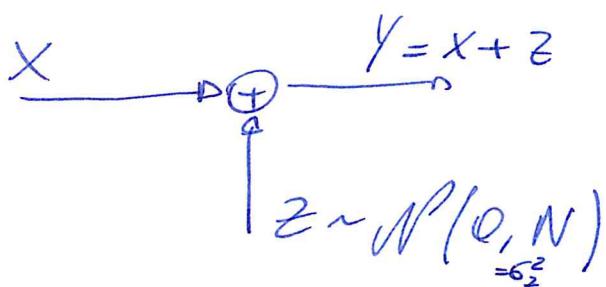
$$h_{\text{exp}}(x) = \log a + \mu \log e + \int p(n) \log p(n) d n$$

$$\int_{\mathbb{R}} p(n) \left(\frac{p(n)}{p_{\text{exp}}(n)} \right)^{\log(ae^{\mu n})} \frac{d n}{p_{\text{exp}}(n)} \geq 0$$

$h_{\text{exp}}(x) \geq h(x)$

AIXGN Channel capacity

- Soft Decision



Suppose $\mathbb{E}[X^2] = S$

$h(Y|X) = \log \sqrt{2\pi S e^2}$ Because fixed X the only variable is Z

$h(Y)$:

$$I(X; Y) = h(Y) - h(Y|X)$$

to maximize, since $h(Y|X)$ is maximum $h(Y)$ should be maximum and should be Gaussian

$$C = \log \sqrt{2\pi \sigma_y^2 e} - \log \sqrt{2\pi \sigma_z^2 e} = \frac{1}{2} \log \frac{\sigma_y^2}{\sigma_z^2} = \frac{1}{2} \log \frac{\sigma_u^2 + \sigma_e^2}{\sigma_z^2}$$

$$= \frac{1}{2} \log \left(1 + \frac{\sigma_u^2}{\sigma_z^2} \right) \xrightarrow{\text{X.W.M.} = 0} \frac{1}{2} \log \left(1 + \frac{S}{N} \right)$$

- Time Continuous, band-limited. AWGN

If the signal $u(t)$ has a band B , it has to be sampled with $\geq 2B$.

In a time T , there will be $2BT$ samples.

Let us this result and think every sample as a discrete sample.

$$C_i = \frac{1}{2} \log \left(1 + \frac{S}{N} \right) \Rightarrow$$

$$C = \frac{2BT}{2} \log \left(1 + \frac{S}{2N_0B} \right) = BT \log \left(1 + \frac{S}{N_0B} \right)$$

If time normalized

$$\frac{C}{T} = B \log \left(1 + \frac{S}{N_0B} \right)$$

Hartley-Shannon Formula

That is the maximum theoretically achievable for an AWGN channel.

If we define $S_r = \begin{bmatrix} Sh \\ \text{sec} \end{bmatrix}$ Bit rate at the end user

$$E_{sh} = \frac{S}{S_r}$$

$$M = \frac{S_r}{B}$$

$$S_r \leq C$$

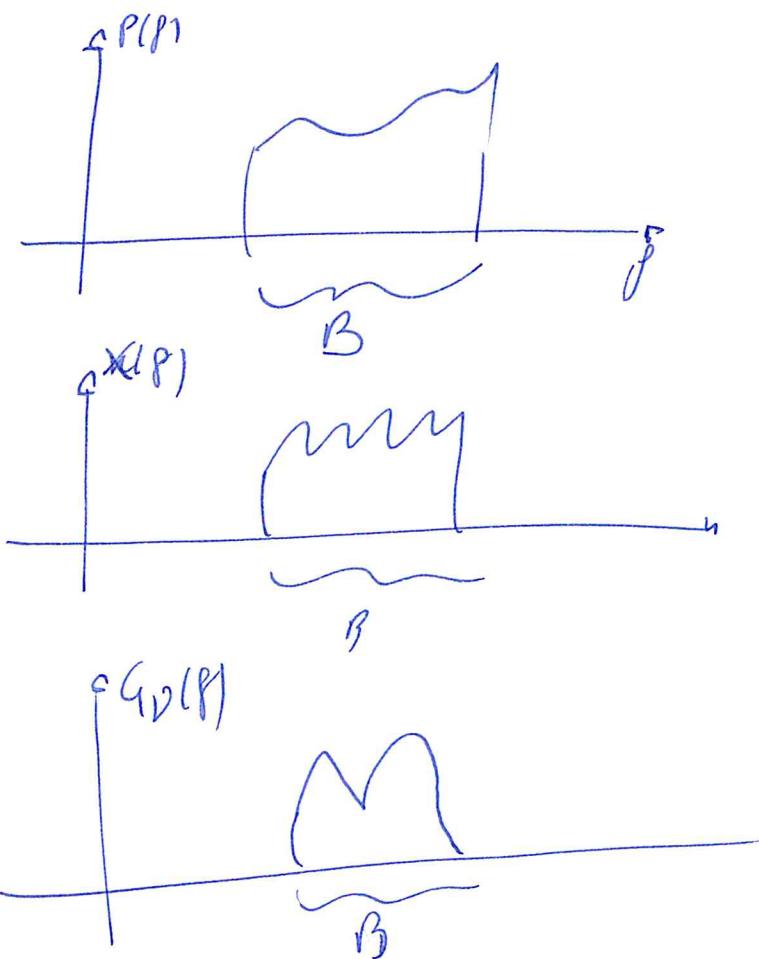
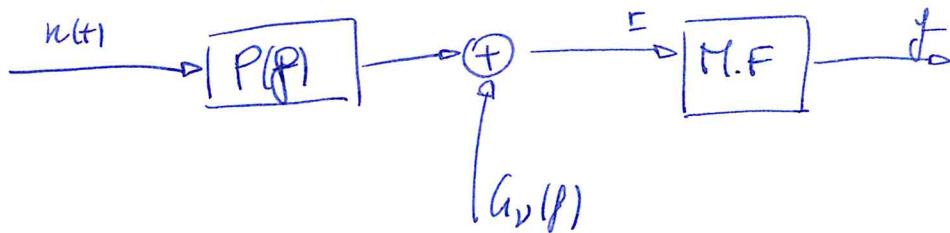
$$S_r \leq \frac{B}{2} \log \left(1 + \frac{S}{N_0B} \cdot \frac{Sh}{S_r} M \right) \Rightarrow \frac{S_r}{B} \leq \log \left(1 + \frac{M E_{sh}}{N_0} \right)$$

$$\gamma \leq \log \left(1 + \gamma \frac{E_{sh}}{N_0} \right)$$

$$2^M \leq 1 + \gamma \frac{E_{sh}}{N_0} \quad \frac{2^M - 1}{\gamma} \leq \frac{E_{sh}}{N_0}$$

This has to be verified for every transmission on an AWGN channel

- Now AWGN channel



Let think using an asymptotic approach:
We divide the channel in subchannels with bandwidth Δf

$$C_i = \Delta f \log \left(1 + \frac{G_{n,i}(f_i) \cdot |P(f_i)|^2 \cdot 2\Delta f}{G_{D,i}(f_i) \cdot 2\Delta f} \right)$$

$$= \Delta f \log \left(1 + \frac{G_{n,i}(f_i) \cdot |P(f_i)|^2}{G_{D,i}(f_i)} \right)$$

Now $\Delta f \rightarrow 0$ and sum every contribute

$$C = \int \log \left(1 + \frac{G_n(f) |P(f)|^2}{G_D(f)} \right) df$$

To know how should be $G_n(f)$ let try to maximize for $G_n(f)$.
Obviously there is a constraint.

$$S = 2 \int_B G_n(f) df$$

To minimize with a constraint, we use the Lagrange Multipliers.

$$\frac{\partial}{\partial G_n(f)} \left[\int \left(\log \left(1 + \frac{G_n(f) |P(f)|^2}{G_D(f)} \right) - \lambda G_n(f) \right) df \right] = 0$$

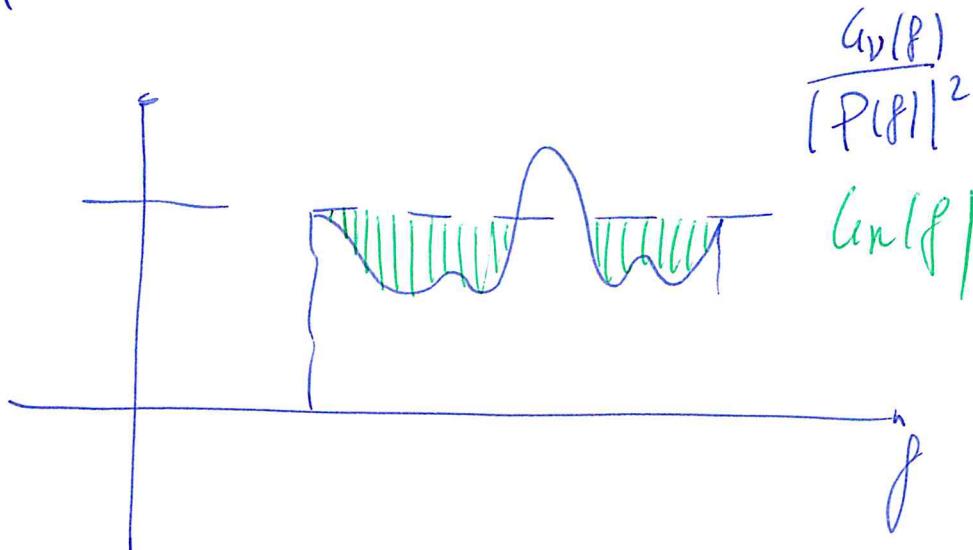
$$\frac{1}{1 + \frac{G_n(f) |P(f)|^2}{G_D(f)}} \cdot \frac{|P(f)|^2}{G_D(f)} - \lambda = 0$$

$$\frac{G_D(f)}{G_D(f) + G_n(f) |P(f)|^2} \cdot \frac{|P(f)|^2}{G_n(f)} = \lambda \quad \lambda = \frac{|P(f)|^2}{G_D(f) + G_n(f) |P(f)|^2}$$

$$\frac{1}{\lambda} = \frac{G_V(f) + G_W(f) |P(f)|^2}{|P(f)|^2} = \frac{G_V(f)}{|P(f)|^2} + G_W(f)$$

$$\begin{cases} \frac{G_V(f)}{|P(f)|^2} + G_W(f) = \text{const.} : G_W(f) > 0 \\ 0 \end{cases} \quad \text{otherwise}$$

$$\begin{cases} G_W(f) = \text{const} - \frac{G_V(f)}{|P(f)|^2} & \text{if } G_V(f) > 0 \\ 0 & \text{otherwise} \end{cases}$$



- AWGN soft Decision

Mixture Distr. see notes.

ERROR CORRECTION CODES

S = avg. power in t_n

B_r = bit rate [$\frac{\text{bit}}{\text{sec}}$]

B_{rc} = coded bit rate [$\frac{\text{bit coded}}{\text{sec}}$]

$$R_c = \frac{B_r}{B_{rc}} \quad E_b = \frac{S}{B_r} \quad E_{b,c} = \frac{S}{B_{rc}}$$

The error probability with the BSC channel HD

$$P_e = \frac{1}{2} \operatorname{erfc} \sqrt{\frac{E_b}{N_0}}$$

$$P_c = \frac{1}{2} \operatorname{erfc} \sqrt{\frac{E_b R_c}{N_0}}$$

\Rightarrow for the repetition

$$(3,1) \quad P_e < P_c.$$

It's not convenient to use!

- Weight of c_i = # of non zero elements of c_i
- Hamming distance = # of different elements between two codes c_i, c_j
- Linear code

if $\alpha c_i + \beta c_j = c_m$ where c_m is a codeword

the sum of two codewords is a codeword

A linear code has ever the $\underline{c} = \underline{0}$ element.

- For a binary code $d_H(c_i, c_j) = \underline{s}_i + \underline{s}_j = \underline{s}_i - \underline{s}_j$

- GENERATOR MATRIX (for linear codes).

$$\text{Coden } \underline{u}_M = \{u_{11}, u_{21}, u_{31}, \dots, u_{M1}\} \quad \underline{c}_M = \{c_{1M}, c_{2M}, c_{3M}, \dots, c_{MM}\}$$

$$\underline{c}_M = \{g_{11}u_{1M}, g_{12}u_{2M}, \dots, g_{1M}u_{M1}\}$$

It is possible to find a matrix \underline{G}

That $\underline{C}_m = \underline{U}_m \underline{G}$

$$\underline{G} = \begin{pmatrix} g_{11} & g_{12} & \cdots & g_{1n} \\ g_{21} & & & \\ \vdots & & & \\ g_{k1} & \cdots & g_{kn} \end{pmatrix} \quad \underline{G} \in M_{k \times n}$$

- Two codes are equivalents if it is possible to find the other matrix from the first one, using row and column operations.

- Systematic Codes.

$$\underline{G} = [\underline{\text{Id}} \mid \underline{P}]$$

The $\underline{C}_m = \underline{U}_m \underline{P}$

- Codewords

Each linear combination of every \underline{G} rows is a codeword.

- Dual Code

Let take a orthogonal vector \underline{H}

$$\underline{G} \underline{H}^T = 0 \quad \underline{H} \text{ is orthogonal to every codeword.}$$

The code generated from \underline{H} is said dual.

Take a \underline{G} for a systematic code

$$\underline{G} = [\underline{\text{Id}} \mid \underline{P}] \Rightarrow \underline{H} [\underline{P}^T \mid \underline{\text{Id}}]$$

Proof

$$\underline{G} \underline{H}^T = [\underline{\text{Id}} \mid \underline{P}] \begin{bmatrix} -\underline{P} \\ \underline{\text{Id}} \end{bmatrix} = -\underline{P} + \underline{P} = \underline{0}$$

• Hamming Code (7,4)

$$\underline{H} = \begin{pmatrix} 1 & 1 & 1 & 0 & 1 & 0 & 0 \\ 0 & 1 & 1 & 1 & 0 & 1 & 0 \\ 1 & 1 & 0 & 1 & 0 & 0 & 1 \end{pmatrix}$$

In the columns there are all combinations of three bits except 0.

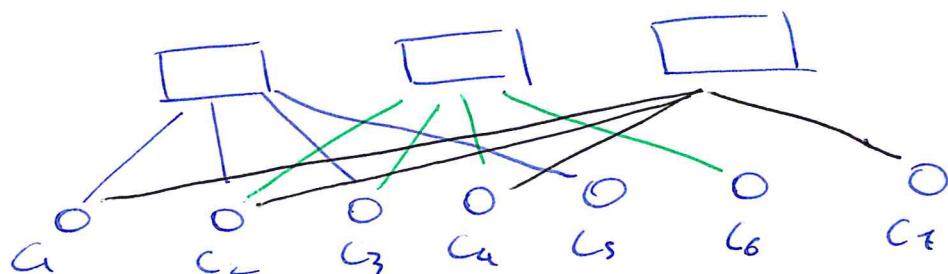


$$\underline{G} = \begin{pmatrix} 1 & 0 & 0 & 0 & 1 & 0 & 1 \\ 0 & 1 & 0 & 0 & 1 & 1 & 1 \\ 0 & 0 & 1 & 0 & 1 & 1 & 0 \\ 0 & 0 & 0 & 1 & 0 & 1 & 1 \end{pmatrix}$$

• Parity Equations

$$\underline{H}\underline{C}^T = 0 \quad \rightarrow \quad \begin{cases} c_1 + c_2 + c_3 + c_5 = 0 \\ c_2 + c_3 + c_4 + c_6 = 0 \\ c_1 + c_2 + c_4 + c_7 = 0 \end{cases}$$

This could be even represented with a bipartite graph



It is eq. to $\underline{H}\underline{C}^T = 0$

Detection theory



I want to obtain $\hat{\underline{c}}_m$ from y

$$P_e = \Pr\{\hat{\underline{c}}_m \neq \underline{c}_m\}$$

$P(\underline{c}_m)$ = A-PRIORI PROBABILITY

$P(\underline{c}_m|y)$ = A-POSTERIORI PROBABILITY

The perfect detection is done choosing the Maximum A Posteriori

MAP

$$\hat{\underline{c}}_m = \operatorname{argmax}_{\underline{c}_m \in \mathcal{C}} \{ P(\underline{c}_m|y) \}$$

This could be even written as:

$$P(\underline{c}_m|y) = \frac{P(y|\underline{c}_m) P(\underline{c}_m)}{P(y)}$$

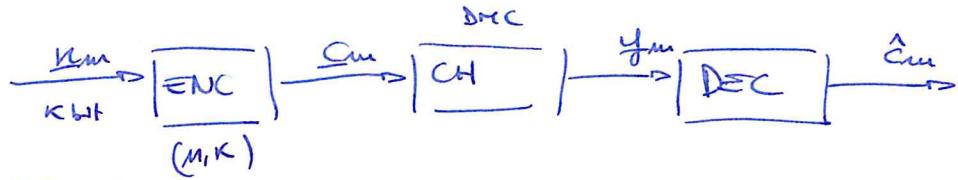
then

$$\hat{\underline{c}}_m = \operatorname{argmax}_{\underline{c}_m \in \mathcal{C}} \left\{ \frac{P(y|\underline{c}_m) P(\underline{c}_m)}{P(y)} \right\} = \operatorname{argmax}_{\underline{c}_m \in \mathcal{C}} \left\{ P(y|\underline{c}_m) P(\underline{c}_m) \right\}$$

If $P(\underline{c}_m)$ is equal for each codeword, so the code has the same A-PRIORI PROB.

$$\text{MAP} \Leftrightarrow \text{ML} \Rightarrow \hat{\underline{c}}_m = \operatorname{argmax}_{\underline{c}_m \in \mathcal{C}} \{ P(y|\underline{c}_m) \}$$

OPTIMUM DECODING FOR BLOCK CODES OVER BSC



$$\text{Hyp: } P(\underline{k}_m) = \frac{1}{2^K} \underline{k}_m \quad \Rightarrow \quad P(\underline{c}_m) = \frac{1}{2^k} \underline{c}_m$$

It is possible to use the ML

$$\hat{c}_m = \arg \max_{c_m \in \mathcal{C}} P(y_m | c_m)$$

Since it is DNC

$$P(y_m | c_m) = \prod_{i=1}^n P(y_{mi} | c_{mi}) \quad \text{as} \quad - - - - - \frac{-(y_{mi} - c_{mi})^2}{2\sigma^2} - - - - - \text{as}$$

$$P(y_{mi} | c_{mi}) = \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{(y_{mi} - c_{mi})^2}{2\sigma^2}}$$

Since we want a BSC, the prop. would be

$$P(c_m = \hat{c}_m) = 1 - P(c_m \neq \hat{c}_m) = P$$

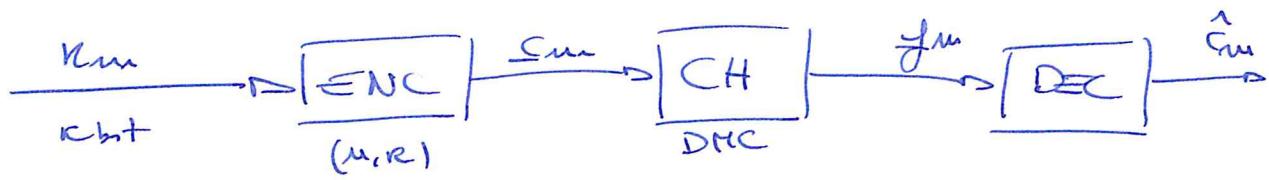
$$P(y_{mi} | c_{mi}) = (1-p) \left(\frac{p}{1-p} \right)^{d_H(y_{mi}, c_{mi})}$$

$$\Rightarrow P(y_m | c_m) = (1-p)^n \left(\frac{p}{1-p} \right)^{\sum_{i=1}^n d_H(y_{mi}, c_{mi})} \quad 0 \leq p \leq 1$$

$$\hat{c}_m = \arg \min_{c_m \in \mathcal{C}} \left\{ d_H(y_m, c_m) \right\}$$

MINIMUM HAMMING DISTANCE DECODER

AIX/GN Channel: Soft Decision



$$P(\underline{k}_m) = \frac{1}{Z^k} \Rightarrow P(\underline{c}_m) = \frac{1}{Z^n}$$

ML $\hat{\underline{c}}_m = \underset{\underline{c}_m \in \mathcal{C}}{\operatorname{argmax}} P(y_m | \underline{c}_m)$

Memoryless $\Rightarrow P(y_m | \underline{c}_m) = \prod_{i=1}^n P(y_{mi} | c_{mi})$

$$P(y_{mi} | c_{mi}) = \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{(y_{mi} - c_{mi})^2}{2\sigma^2}}$$

$$P(y_m | \underline{c}_m) = \left(\frac{1}{\sqrt{2\pi\sigma^2}} \right)^n e^{-\frac{1}{2\sigma^2} \cdot \sum (y_{mi} - c_{mi})^2}$$

$$\hat{\underline{c}}_m = \underset{\underline{c}_m \in \mathcal{C}}{\operatorname{argmax}} e^{-\sum_i (y_{mi} - c_{mi})^2}$$

$$\hat{\underline{c}}_m = \underset{\underline{c}_m \in \mathcal{C}}{\operatorname{argmin}} \sum_i (y_{mi} - c_{mi})^2 = \underset{\underline{c}_m \in \mathcal{C}}{\operatorname{argmin}} d_E^2(y_m, \underline{c}_m) = \underset{\underline{c}_m \in \mathcal{C}}{\operatorname{argmin}} d_E(y_m, \underline{c}_m)$$

MINIMUM EUCLIDEAN DISTANCE DECODER

This could be even seen as:

$$\sum_i (y_{mi} - c_{mi})^2 = \underbrace{\sum_i (y_{mi})^2}_{\text{const.}} + \underbrace{\sum_i (c_{mi})^2}_{\text{code symbol energy}} - 2 \sum_i (y_{mi} c_{mi})$$

\Leftrightarrow is equal to minimize $E_c - 2 \sum_i (y_{mi} c_{mi})$

correlation

SYNDROME AND STANDARD ARRAY DECODING

$$\underline{c}^H = 0 \quad \underline{y}^H = (\underline{c} + \underline{e})^H$$

\underline{e} = vector with 1 in the place where the error occurred.

$$(\underline{c} + \underline{e})^H = \cancel{\underline{c}^H} + \underline{e}^H = \underline{e}^H = \underline{s} \quad \text{SYNDROME}$$

Then we construct a table with a syndrome for each error pattern. It is kept only the minimum weight error word for each syndrome.

The correction is done by the sum of \underline{e} to \underline{y} .

The std. array impl. has a pattern of each error with each codeword (coset), so to decode it is only a look-up in the std. array.

$$B/R = \frac{\# \text{ERROR}}{\# \text{BITS}} \quad (\text{NUMBER OF WRONG BITS OVER THE TOTAL})$$

$$LER = \frac{\# \text{CODE ERRORS}}{\# \text{CODES}} \quad (\text{NUMBER OF WRONG WORDS OVER THE TOTAL})$$

• CYCLIC LINEAR BLOCK CODES

given $\underline{c} = (c_{n-1}, c_{n-2}, \dots, c_1, c_0)$ is a codeword

if the code is cyclic $\underline{c}' = (c_{n-2}, c_{n-3}, \dots, c_0, c_1, c_{n-1})$ is also a codeword

They are interesting because they could be converted to a polynomial

$$C(D) = c_{n-1}D^{n-1} + c_{n-2}D^{n-2} + \dots + c_1D + c_0$$

$$\bullet C''(D) = D \cdot C(D) \bmod (D^m - 1)$$

$$C(D) = c_{m-1} D^{m-1} + c_{m-2} D^{m-2} + \dots + c_1 D + c_0$$

$$DC(D) = c_{m-1} D^m + c_{m-2} D^{m-1} + \dots + c_1 D^2 + c_0 D$$

$$C'''(D) = c_{m-2} D^{m-1} + \dots + c_1 D^2 + c_0 D + c_{m-1}$$

$$DC(D) = c_{m-1} D^m + c_{m-2} D^{m-1} + \dots + c_1 D^2 + c_0 D + c_{m-1} - c_{m-1}$$

$$= c_{m-1}(D^m - 1) + C''(D)$$

$$DC(D) = \underbrace{c_{m-1}}_Q \underbrace{(D^{m-1})}_D + \underbrace{C''(D)}_R !$$

$$DC(D) = Q(D) D(D) + R(D)$$

∴ generally

$$C^{(m)}(D) = D^m C(D) \bmod (D^m - 1)$$

• Minimum grade Codeword

$$g(D) = D^{m-k} + \dots + 1$$

• It is unique : if I sum a smaller word, I could get something shorter \Rightarrow impossible.

• It has to have the 1 in the end, otherwise I can shift and make it shorter.

Shifting that word I can obtain another codeword

$$C^{(m)} = D^m g(D) \quad \text{if } m < k-1$$

For the others I can:

Multiply for a greater D^m w.r.t. $k-1$ and calculate the remainder, otherwise I can combine linearly two words obtained with such

$$C(D) = Q(D) g(D)$$

for $m-1$ $\geq k-1$ g is $n-k$

a) If C is a cyclic code, its generator $g(D)$ divides D^{m-1}

$$C(D) = Q(D) g(D)$$

$$C^{(1)}(D) = \tilde{Q}(D) g(D)$$

$$C^{(1)}(D) = D C(D) + c_{m-1} (D^{m-1})$$

$$\tilde{Q}(D) g(D) = D Q(D) g(D) + c_{m-1} (D^{m-1})$$

$$(\tilde{Q}(D) - D Q(D)) g(D) = c_{m-1} (D^{m-1})$$

with $c_{m-1} = 1$

$$\underline{D^{m-1} = \tilde{Q}(D) g(D)}$$
 so it divides D^{m-1}

b) If $g(D)$ divides D^{m-1} , it can generate a cyclic code

$$D^{m-1} = \tilde{Q}(D) g(D)$$

$$C(D) = Q(D) g(D)$$

$$C^{(1)}(D) = c_{m-1} (D^{m-1}) + C(D) D$$

$$C^{(1)}(D) = D \cdot Q(D) g(D) + \tilde{Q}(D) g(D) c_{m-1}$$

$$C^{(1)}(D) = g(D) (D Q(D) + \tilde{Q}(D) c_{m-1}) = g(D) \bar{Q}(D)$$

GENERATOR MATRIX

It is composed by every translation of $g(D)$ with $m < k-1$
and these are put as the rows.

• Systematic cyclic codes

Note

$D^{m-k} n(D)$ is $n(D)$ translated ~~to~~ of $m-k$ positions

$$\begin{array}{c|cccc|c|c|c|c} \hline & 1 & 0 & 0 & 0 & 0 & | & 1 & 1 \\ \hline & & & & & & \diagdown & & \\ & & & & & & k & & \\ \hline & t & & & & & m-k & & \\ \hline & 1 & | & 1 & - & - & | & 1 & 0 0 0 0 \\ \hline \end{array}$$

$$D^{m-k} n(D) : g(D) = Q(D) + R(D)$$

$$D^{m-k} n(D) = Q(D) g(D) + \underbrace{R(D)}_{k-1}$$

$$\underbrace{(D^{m-k} n(D) + R(D))}_{\text{gr. } m-1} = \underbrace{Q(D) f(D)}_{\text{gr. } n-1}$$

This is code!

And it is systematic!

To obtain systematic code

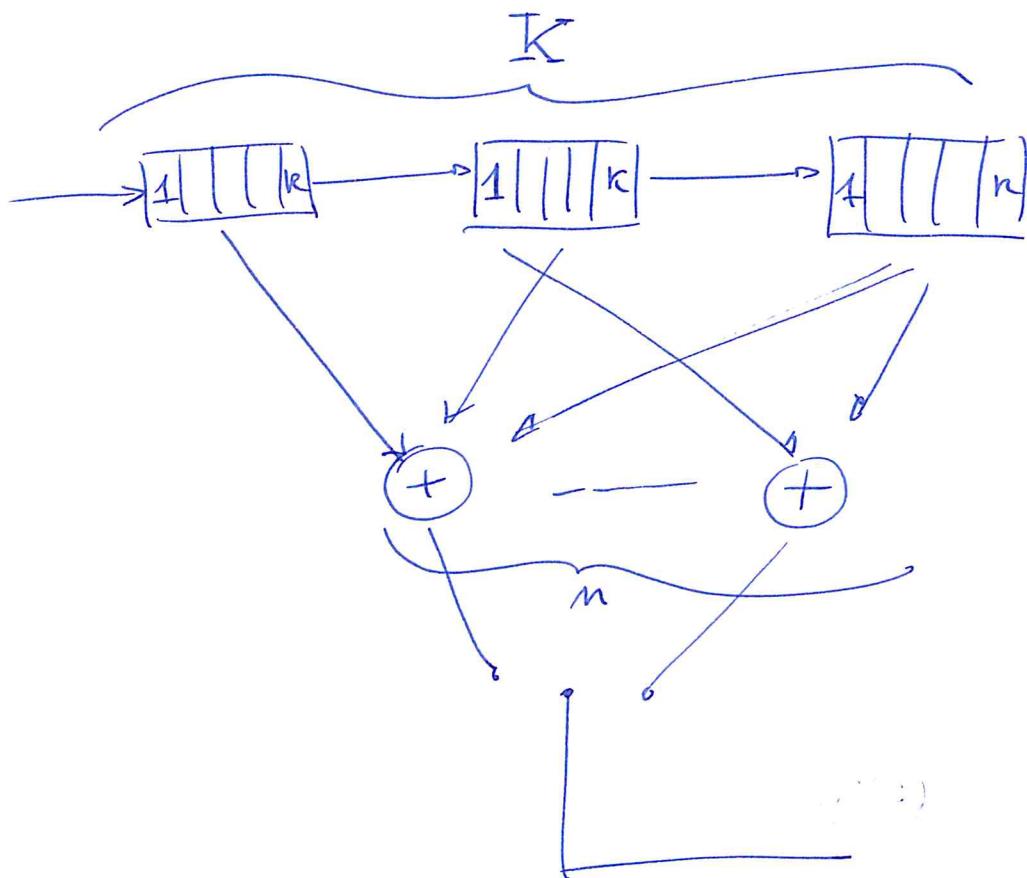
- Shift $m-k$ positions the input word
- Calculate the $D^{m-k} n(D) \bmod g(D)$
- Add this to the previous shifted code
- Synchronize correction

$$r(D) = C(D) + e(D)$$

$$= e(D) \bmod g(D)$$

$$r(D) \bmod g(D) = \underline{C(D) \bmod g(D) + e(D) \bmod g(D)}$$

Convolutional Codes:



It can be defined by a generator g , that could be a polynomial or defined in octal.

It is 1 where there is a connection to K^{th} register.
 K has to be 1.

These encoder/decoder has a memory, so it has a STATE.

It is possible to create a STATE DIAGRAM or the EQUIVALENT TRELLIS DIAGRAM. Each line in the TRELLIS is a state.

On the connection of the states there are usually the output:
 If $R_c=2$ two output and the line (different colors etc) should show the input bit.

To decode:

- create the trellis and then calculate the length of every branch.
- Viterbi Algorithm.
 - Calculate the length of every branch to every state and keep only the shortest for each ~~last~~ state (survivor)
 - Go ahead of one bit.

The original algorithm ends only when the trellis ends.

It is possible to have a really long code and to say that every bit that is 5-6 bit before the current node in the Trellis is right.

This is usually right.

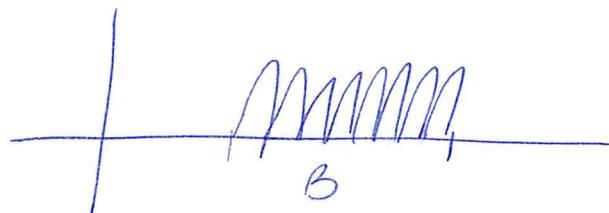
(I mean S-6 K) Truncated V.A.

The distances could be calculated using dt orde.

d_{free} is the min distance to return to the initial state.

The conv. codes makes burst errors, so it could be great to concatenate a Reed-Solomon or BCH to this one.

OFDM



We define a symbol as (or complex envelope)

$$a_0 = \sum_{i=0}^{n-1} a_i g(t - nT_s)$$

Then the ~~over~~ shifts the freq.

$$i(t) = \sum_n a_n e^{j2\pi f_n t}$$

$$f_m = \Delta f \cdot m$$

It is orthogonal if

$$\frac{1}{T_s} \int_0^{T_s} e^{j2\pi f_m t} e^{-j2\pi f_l t} dt = \frac{1}{T_s} \int_0^{T_s} e^{j2\pi f_l (m-l)t} dt \quad \text{if } T_s = \frac{1}{\Delta f} \Rightarrow = \delta_{m-l}$$

If the freq. is shifted by $\frac{1}{T_s}$ \Rightarrow the symbols are orthogonal.