

Article

Repeatability of taste recognition threshold measurements with *QUEST* and *quick Yes-No*

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Abstract: Taste perception, though vital for nutrient sensing, has long been overlooked in psychological and clinical sensory assessments. This can, at least in part, be attributed to challenges associated with the handling of liquid, perishable stimuli, but also with scarce efforts to optimize testing procedures to be more time-efficient. We have previously introduced an adaptive, QUEST-based procedure to measure taste sensitivity thresholds that can be quicker than other existing approaches, yet similarly reliable [1]. Despite its advantages, this procedure lacks experimental control of false alarms (i.e., response bias) and psychometric function slope. Variations of these parameters, however, may influence the threshold estimate likewise, raising the question as to whether a procedure assessing threshold, false-alarm rate, and slope simultaneously might be able to produce threshold estimates with higher repeatability, i.e., smaller variation between repeated measurements of the same participant. Here, we compared the performance of QUEST with a method that allows measurement of false-alarm rates and slopes, *quick Yes-No* (qYN), in a test-retest design for citric acid, sodium chloride, quinine hydrochloride, and sucrose recognition thresholds using complementary measures of repeatability, namely test-retest correlations and coefficients of repeatability [2]. Both threshold procedures yielded largely overlapping thresholds with good repeatability between measurements.

Keywords: sensitivity; taste; threshold; staircase; QUEST; bias; quick yes-no

1. Introduction

The ability to taste is undoubtedly crucial for nutrient sensing. Yet, the ability to taste is scarcely assessed in large cohorts and precise and "practical" methods for the measurement of taste sensitivity are needed to get a better understanding of the extent to which taste shapes preferences and eating behavior.

The precise measurement of sensory sensitivity – with typically limited measurement duration – has been a challenge in all senses. The relatively quick adaptation in gustation, however, increases the data collection burden enormously, as it requires long inter-stimulus intervals that cannot simply be countered with a reduction of the number of experimental trials. Using a Bayesian adaptive testing framework [3], we have previously shown that taste sensitivity can be precisely and reliably measured in a fraction of the time needed with conventional, non-adaptive methods [1,4].

Bayesian adaptive methods like QUEST [3] typically produce estimates of psychometric function parameters quicker than conventional staircase procedures, especially if the number of possible stimuli is large. This is mainly due to two features specific to Bayesian methods: firstly, they can incorporate prior knowledge by assigning a probability to each individual parameter value; and secondly, they incorporate the entire response history of a participant to predict the "true" parameter value and select

the next stimulus. Accordingly, on every trial, stimuli are selected such that the expected knowledge gain about the parameter (e.g., threshold) is maximized. This may lead to relatively strong intensity changes from one trial to the next, especially at the beginning of an experimental run, and allows the procedure to converge within fewer trials compared to a staircase.

When measuring thresholds, participants' responses not only depend on their sensory sensitivity, but are also influenced by mental processes. Specifically, whether a participant will report the presence of a stimulus is influenced by their cognitive *response criterion*: if a *liberal* criterion is adopted, even weak, non-obvious stimuli will be reported; if the criterion, however, is *strict*, only stimuli that the participant believes are sufficiently strong are going to be reported. Generally, a liberal criterion leads to a higher false-alarm rate (FAR), which describes the proportion of "yes" responses to a blank, while a stricter criterion reduces the frequency of false alarms [5,6].

Although QUEST was designed for alternative forced-choice tasks, which are commonly thought to control the response criterion (but c.f. [7] for criticism of this view), the simplicity of the yes-no experiment (see [7] for a systematic review of adaptive procedures) has been a strong motivator to explore the suitability of QUEST in a yes-no design, and it has yielded good performance in the chemosensory domain [1,4,8]. In these studies, participants were instructed to be "conservative" in their response behavior, in an attempt to keep false-alarm rates low and constant across sessions. This is imperative because QUEST can only estimate a single parameter, such that when the threshold is to be estimated, all other parameters defining the psychometric function, like FAR, slope, and the lapse rate (i.e. the proportion of "no" responses to high-intensity, supra-threshold stimuli) need to be set to a fixed value *a priori*. While the lapse rate can safely be assumed to be low in taste threshold testing, provided that sufficiently long inter-stimulus intervals are used and participants thoroughly rinse their mouth between trials, FAR may vary between repeated measures despite our instructions. Slope determines how well participants can detect intensity differences between stimuli, and as such can also serve as an (implicit) measure of threshold reliability. Just like thresholds differ between participants and, obviously, tastants, it is known from other sensory systems that "slopes are different for different stimuli, and this can lead to misleading results if slope is ignored." [7]

Here, we set out to address these concerns, and tested whether FAR and slope indeed vary between tastants and sessions, and whether the measurement of these parameters helps improve the repeatability of taste threshold measurements. To this end, we derived a test-retest design in which we employed two Bayesian methods, the previously used QUEST, which only measures threshold, and *quick Yes-No* (qYN, [9]), which assesses the sensory threshold, FAR, and psychometric function slope.

2. Materials and Methods

2.1. Participants

41 participants (34 women; mean age $M = 30.1$, standard deviation $SD = 11.4$, range 18–64 years) participated in the study and received compensatory payment. Their weight was within the normal range according to the body mass index (BMI; $M = 22.5$, $SD = 2.6$). 4 participants were smokers. Participants exhibited an eating style, measured with the Dutch Eating Behavior Questionnaire [10], that conforms well with recent norm data [11]; the average scores were 2.17 ($SD = 0.72$) for emotional, 2.95 ($SD = 9.79$) for external, and 2.49 ($SD = 0.77$) for restrained eating. The questionnaire data of one participant (female, 27 years old) is missing due to a technical fault; only her age, sex, and threshold data is reported here.

Exclusion criteria were: self-reported taste and smell disorders, smoking, current or recent oral, nasal or sinus infections, pregnancy, recent (during the last 6 months) childbirth, metabolic disease, and recent (during the last 3 months) weight change exceeding 10 kg. The study conformed to the revised Declaration of Helsinki and was approved by the ethical board of the German Society of Psychology (DGPs).

2.2. Procedure

2.2.1. Experimental sessions

Participants were invited to four experimental sessions that lasted 1 hour each: a Test and a Retest session for each of two threshold algorithms. To ensure similar testing conditions across sessions, participants were instructed to refrain from eating and drinking anything but water 30 min before visiting the laboratory. Further, the sessions were scheduled at approximately the same time of day, and within 10 days (inter-session interval: $M = 2.3$, $SD = 2.0$, range 1–10 days).

At the beginning of the first session, participants completed a screening questionnaire, the Dutch Eating Behavior Questionnaire, and rated how much they liked and how often they consumed salty, sour, sweet, and bitter foods. Following the ratings and in each subsequent session, taste recognition thresholds for citric acid (sour), sodium chloride (salty), quinine hydrochloride (bitter), and sucrose (sweet) were measured using either of two algorithms, QUEST or qYN, described below. The order of tastants was balanced across participants and kept constant for Test and Retest within each participant. The order of algorithms was balanced across participants.

2.2.2. Taste Ratings

As a measure of taste preference, participants rated how much they liked salty, sour, sweet, and bitter foods and beverages on separate, horizontal Visual Analogue Scales (VAS) anchored with 1 (not at all) and 5 (extremely). They furthermore provided the frequency at which they typically consume salty, sour, sweet, and bitter foods and beverages on a scale with 7 options: daily, 4–6 times per week, 2–3 times per week, once per week, 2–3 times per month, once per month, less than once per month. Specifically, they provided ratings for the following items: sweet sour, and bitter beverages, sweet, sour, and bitter fruits and vegetables, sweet cake/candy, salty snacks as well as added salt. Rating for each taste quality were averaged for further analysis. The ratings were assessed in paper–pencil format.

2.2.3. Taste Stimuli

Tastants were prepared by diluting prototypical chemicals that are known to elicit a clear taste perception in deionized (DI) water: citric acid (sour; molar mass $M = 192.12 \text{ g mol}^{-1}$), sodium chloride (salty; $M = 58.44 \text{ g mol}^{-1}$), quinine hydrochloride (bitter; $M = 396.91 \text{ g mol}^{-1}$), and sucrose (sweet; $M = 342.30 \text{ g mol}^{-1}$). All chemicals were produced by Sigma-Aldrich and purchased from Merck KGaA, Darmstadt, Germany.

Based on previous studies [1,4], we used the following sets of different concentrations that were equidistantly spaced on a decadic logarithmic grid for each tastant: citric acid, 0.015 mM to 46.846 mM (14 \log_{10} steps; step width: 0.269); sodium chloride, 0.342 mM to 342.231 mM (12 \log_{10} steps; step width: 0.273); quinine hydrochloride, 0.077×10^{-3} mM to 3.131 mM (21 \log_{10} steps; step width: 0.230); sucrose, 0.073 mM to 584.283 mM (14 \log_{10} steps; step width: 0.300).

Taste solutions were stored refrigerated at 4 °C for a maximum duration of seven days in glass bottles. During testing, they were sprayed manually by the experimenter to the anterior half of the tongue using a conventional spray head that released approx. 0.2 mL.

2.2.4. Taste recognition psychometric functions

The QUEST implementation we used assumes a psychometric function in which the proportion of "yes" responses to a tastant concentration, c , is given by

$$\Psi_{\text{yes}}(c) = \lambda\gamma + (1 - \lambda)[1 - (1 - \gamma)\exp(-10^{\beta(c+\tau)})].$$

Here, τ is the threshold parameter, i.e., the concentration at the perceptual threshold, which was to be estimated during the experiment. The parametrization was identical to the one used by

[1]: We defined "threshold" as the concentration with an expected proportion of 80% "yes" responses ($\Psi_{\text{yes}}(c) = 0.80$); the prior probabilities for the threshold parameter were given by a normal distribution with a standard deviation of 20, centered on the starting concentration of the respective tastant (see the *Taste threshold stimulus selection* section below). All other parameters were fixed *a priori*: slope, β , to 3.5; and both the false-alarm and lapse rate, γ and λ , to 0.01. Internally, QUEST works with an abstract "intensity grid", whose granularity we set to 0.01 as well.

In quick Yes-No [9], the psychometric function describing the proportion of "yes" responses for a given concentration is

$$\Psi_{\text{yes}}(c) = \epsilon + (1 - \epsilon) [1 - \Phi(\lambda - d'(c))]$$

with Φ being the cumulative normal distribution with a mean of 0 and a standard deviation of 1. ϵ is the lapse rate, which describes how frequently stimuli of a high intensity are not recognized. The decision criterion, λ , determines the FAR. Ψ_{yes} depends on the sensitivity function, d' , given by

$$d'(c) = \frac{\beta(c/\tau)^\gamma}{\sqrt{(\beta^2 - 1) + (c/\tau)^{2\gamma}}}.$$

τ is the "threshold intensity", which here is defined as the tastant concentration corresponding to a sensitivity of $d' = 1$. β defines the upper asymptote and γ the slope. In this study, we estimated threshold, τ ; slope, γ ; and decision criterion, λ . Consequently, the parameter space was a three-dimensional a grid. Different ranges of τ were used for each tastant. To achieve a finer granularity, additional (virtual) concentration steps were inserted halfway between the existing (physical) concentration steps, producing 27 values for citric acid and sucrose, 23 for sodium chloride, and 35 for quinine hydrochloride. For γ , we used 10 values in the interval [0.5, 3.0], evenly spaced on a decadic logarithmic grid; and for λ , we used 8 evenly spaced values in the interval [0.75, 2.50], corresponding to FARs of [22.7%, < 1.0%]. We assumed no prior knowledge regarding the "true" parameter values, and, hence, used an "uninformative" prior that assigned the same probability to all possible parameter value combinations. β was fixed at 5.0 [9], and ϵ was set to 0.

2.2.5. Taste recognition stimulus selection

The concentration presented in the first trial for each tastant was predefined such that it would be supra-threshold for most participants, in order to familiarize them with the particular tastant as testing commenced (citric acid: 7.328 mM; sodium chloride: 97.469 mM; quinine hydrochloride: 0.077 mM; sucrose: 73.509 mM). For subsequent trials, the QUEST and qYN procedures proposed the stimulus concentration to present based on response behavior in all previous trials. While QUEST aims to place stimuli at threshold concentration, qYN – trying to estimate *three* parameters at once – typically also suggests to present stimulus concentrations slightly above and below threshold to measure slope, and at very low concentrations to measure estimate the FAR.

As both algorithms internally worked with smaller, virtual concentration steps, they would sometimes propose concentrations that were not physically available. In this case, our computer program selected the concentration closest to the proposed one, and informed the algorithm about the actually used concentration. In QUEST, we added an additional rule: whenever the algorithm proposed to present the same concentration on two consecutive trials, we increased the concentration in the second trial by one step if the participant had responded "no" to the previous trial, and we decreased the concentration by one step if the response in the previous trial was "yes", thus introducing a little more variability to avoid repetitive presentation of the same concentration on multiple consecutive trials, which we felt could have been more tiring for participants. In qYN, we did not add such a rule, as the algorithm itself introduced somewhat abrupt concentration changes once in a while in order to determine FAR and slope.

2.2.6. Taste recognition procedure

Participants were seated comfortably and blindfolded. At the beginning of each measurement, they were told which taste would be tested next. At the beginning of each trial they were asked to stick out the tongue and received the stimulus. Participants were required to indicate whether they recognized the taste by nodding ("yes") or shaking their head ("no") while they kept their tongue extended. Immediately after the response, the experimenter logged it into the computer, and participants rinsed their mouth with DI water. Participants received no feedback as to their performance during the experiment. The interval between consecutive stimuli was approx. 30 s.

2.2.7. Taste recognition termination

qYN experimental runs always ended after 20 trials. For QUEST, we employed the same termination criterion as in a previous study [1]: after more than 10 trials had been performed, we checked after each trial whether the 90% confidence interval of the threshold estimate was smaller than half a concentration step; if that was the case, the experimental run was finished. Otherwise, a maximum of 20 trials were performed.

2.3. Analysis

The significance level α was set to 0.05 *a priori* for all statistical tests. Greenhouse-Geisser correction was applied for violation of sphericity in repeated measures analysis of variance (rmANOVA); uncorrected degrees of freedom and corrected *p*-values are reported in this case.

2.3.1. Ratings

Ratings for taste liking and frequency of consumption were submitted to separate one-factorial rmANOVA with four levels (sour, salty, sweet, bitter). To quantify the potential link between taste preferences and sensitivity, we computed Spearman's correlation coefficients. For this, the average of all threshold estimates from QUEST and qYN for a given participant was used as robust measure for taste sensitivity. Data from only 37 participants are reported for liking and frequency of consumption because 4 participants did not complete the ratings.

2.3.2. Taste recognition parameter estimates

To retrieve the final threshold estimate, we calculated the mean of the concentrations weighted by the posterior distribution (QUEST) or marginal posterior distribution (qYN), respectively. This measure had been shown to produce an unbiased estimate of threshold in QUEST, as opposed to other metrics [12]. We limited the values of the threshold estimates to the range of stimulus concentrations used in the present study, as QUEST could – in rare cases – produce thresholds outside of this range for extremely sensitive or insensitive participants. This was the case for a single participant where QUEST produced one quinine hydrochloride threshold above the highest available concentration.

For qYN, FAR and slope were calculated as the mean of the respective parameter space weighted by the corresponding marginal posterior.

Out of the 656 obtained thresholds (41 participants \times 4 tastants \times 2 methods \times 2 sessions), 41 (6.3%; 17 QUEST and 24 qYN) were lost due to data corruption. The remaining 615 thresholds entered analysis. The threshold estimates were submitted to separate rmANOVAs for each tastant with the factors *method* (QUEST, qYN) and *session* (Test, Retest).

2.3.3. Test-retest reliability

First, we calculated Spearman's rank correlation, ρ , for each tastant to quantify the monotonic relationship between the measurement results in both sessions.

Correlation analysis does not necessarily provide a good indication of absolute *repeatability* of an experiment, as "[the] correlation coefficient is a reflection of how closely a set of paired observations

(test–retest data in this case) follow a straight line, regardless of the slope of the line" [13], and it also disregards systematic changes between measurements [13,14], such as a constant offset. We, therefore, conducted an additional analysis that focuses on the *differences* between measurements. For each method separately, we first calculated the difference between Test and Retest estimates for all participants; then, we calculated the standard deviation of these differences, s_d , and derived a coefficient of repeatability, $CR = 1.96 \times s_d$ [15].¹ The number 1.96 is a z-score and corresponds to the 97.5 % quantile of the normal distribution. If a participant were measured repeatedly using the same procedure, we would then expect 95 % of the absolute measurement differences not to exceed CR. This provides a straightforward, single-number representation of the magnitude of measurement variation to expect. Lastly, we calculated the mean of the differences between sessions, \bar{d} , and estimated the 95 % *limits of agreement* (LoA) as $LoA = \bar{d} \pm CR$ [15]. These limits correspond to the 95 % confidence interval of the differences, and, consequently, narrower LoAs suggest better measurement repeatability.

Because the calculations of the mean difference and LoAs are based on an experimental sample, they are estimates that naturally have a certain amount of uncertainty associated with them. We therefore also derived 95 % confidence intervals (CIs) of these estimates. The mean difference was assumed to be normally distributed with mean \bar{d} and SD s_d / \sqrt{n} ([2]; with the number of paired samples, n), and hence the CI corresponded to the 2.5 % and 97.5 % quantiles of this distribution. CIs of the LoAs were calculated via the "exact paired" method [16].

For visual comparison, we plotted the differences between Test and Retest over the mean of both measurements (which serves as our best estimate of the "true" value) and added the mean difference \bar{d} and LoAs as horizontal lines, producing so-called *Bland-Altman* or *Tukey mean difference plots*. These plots allow for a quick and straightforward inspection of measurement differences, exposing systematic biases ($\bar{d} \neq 0$) the degree of measurement differences and their variability.

The calculation of CRs and LoAs requires the differences between sessions to be approximately normally distributed to work correctly. If that is not the case – e.g., if the difference changes with the magnitude of the session mean – a logarithmic transformation can be carried out prior to analysis [14]. Here, we elected to use the decadic logarithm, \log_{10} . After the data had been transformed, \bar{d} , CR, LoAs, and CIs were calculated following the procedure described above. To plot the results that were calculated in log space, x_{\log} , in their original coordinate system for intuitive visual assessment, a back-transformation must be applied [17]: $y_{\text{back}} = 2M(10^{y_{\log}} - 1) / (10^{y_{\log}} + 1)$, where M is a given value on the abscissa (i.e., any given session mean). The result of the back-transformation, therefore, describes a *line*, not a single value, in the original coordinate system. In other words: the value of the back-transformed parameter, x_{back} , is conditional on M . For the data presented here, applying the \log_{10} transform was only necessary for FAR measures.

As the FAR determines the lower asymptote, changes in FAR necessarily lead to changes in steepness of the psychometric function if the threshold is assumed to remain unchanged: as FAR increases, the slope must decrease, and vice versa. We thus calculated Spearman's rank correlation between FARs and slopes, pooled across all tastants, in order to explain potential fluctuations of FAR and slope across sessions, which could be explained through the covariance of both variables.

2.3.4. Software

Stimulus presentation and data collection were guided by a Python computer program based on PsychoPy 1.85.4 [18] on Windows 7 (Microsoft Corp., Redmond, WA/USA). Statistical analyses were carried out with JASP 0.11.1 (<https://jasp-stats.org/>) and pingouin 0.2.9 [19]. CIs for the Bland-Altman

¹ Note that it has been suggested [2,13,14] to calculate CR as $\sqrt{2} \times 1.96 s_w$, where s_w is the *within-participant standard deviation*, i.e., the square-root of the averaged within-participant variances of measurement repetitions. We found that with our data, this approach produced very similar results (deviating only in the second decimal place) to the simpler formula $1.96 \times s_d$, which directly and intuitively corresponds to the 95 % limits of agreement in the Bland-Altman plots. Therefore, we elected to follow this simpler approach, and report CRs based on the SD of measurement differences between sessions here.

plots were calculated via pyCompare (<https://github.com/jaketmp/pyCompare>). Plots were created using matplotlib 3.1.1 (<https://matplotlib.org>) and seaborn 0.9.0 (<https://seaborn.pydata.org>).

3. Results

3.1. Ratings

The preference for different tastes varied significantly ($F_{3,36} = 33.68$, $p < 0.001$, $\eta_p^2 = 0.483$), as expected, with bitter taste being significantly less liked than salty ($t_{36} = -6.95$, $p_{\text{holm}} < 0.001$), sour ($t_{36} = -7.17$, $p_{\text{holm}} < 0.001$), and sweet ($t_{36} = -9.65$, $p_{\text{holm}} < 0.001$); and with sweet being significantly more liked than salty ($t_{36} = 2.77$, $p_{\text{holm}} = 0.018$) and sour ($t_{36} = 4.2$, $p_{\text{holm}} < 0.001$). Accordingly, scores (min=1, max=5) were highest for sweet ($M = 4.11$, $SD=0.83$), followed by salty ($M = 3.41$, $SD=1.05$), sour ($M = 3.05$, $SD=1.09$), and bitter ($M = 1.89$, $SD = 0.98$).

The preference for sour taste correlated significantly with the sucrose ($\rho = 0.403$, $p = 0.013$) and marginally with the citric acid ($\rho = 0.305$, $p = 0.066$) taste threshold. No further correlations between taste preference and thresholds were found (all $p > 0.19$).

The frequency by which participants reported to consume foods with a certain taste was higher for sweet compared to sour ($t_{36} = 2.84$, $p_{\text{holm}} = 0.037$) and salty ($t_{36} = 4.5$, $p_{\text{holm}} < 0.001$), resulting in a main effect ($F_{3,36} = 6.15$, $p < 0.003$, $\eta_p^2 = 0.146$). While the pattern of results is plausible, the reported mean frequencies raise doubts as to their validity, because participants reported to consume food and beverages only approx. once per week (score=4) for all taste qualities: sweet, $M = 4.7$ ($SD=1.03$); salty, $M = 3.64$ ($SD=1.33$); sour, $M = 4.19$ ($SD=1.1$); and bitter, $M = 4.0$ ($SD=1.24$).

3.2. Taste recognition thresholds

The distributions of threshold estimates averaged across both sessions and split by methods are shown in Fig. 1; to facilitate visual comparison between methods, only participants for whom datasets of both sessions and methods were available are included for each individual tastant. The rmANOVAs revealed a main effect of *method* – indicating that threshold estimates were systematically lower for qYN compared to QUEST – for citric acid ($F_{1,32} = 17.475$, $p < 0.001$, $\eta_p^2 = 0.353$), sodium chloride ($F_{1,32} = 44.728$, $p < 0.001$, $\eta_p^2 = 0.583$), and sucrose ($F_{1,31} = 11.198$, $p < 0.01$, $\eta_p^2 = 0.265$), but not for quinine hydrochloride ($F_{1,33} = 3.241$, $p = 0.08$, $\eta_p^2 = 0.089$). Thresholds did not differ significantly between sessions for sodium chloride ($F_{1,32} = 0.342$, $p = 0.56$), quinine hydrochloride ($F_{1,33} = 1.195$, $p = 0.28$), and sucrose ($F_{1,31} = 0.219$, $p = 0.64$); however, we found a main effect of *session* for citric acid ($F_{1,32} = 4.492$, $p = 0.042$, $\eta_p^2 = 0.123$). No interactions between *method* and *session* were found (all $p > 0.13$). The threshold estimates, their respective minimum and maximum values, and their standard deviations are listed in Tab. 1.

On average, QUEST needed 14.6 trials – corresponding to 6:50 min – to converge to a threshold. For qYN, the duration of an experimental run was always approx. 9:30 min, as the number of trials was fixed to 20.

3.3. Test-retest repeatability

Test-retest comparisons naturally included only participants for whom data from both sessions in a given method and tastant were available. Test and Retest threshold estimates correlated significantly for all tastants in both QUEST (citric acid: $\rho_{35} = 0.62$, sodium chloride: $\rho_{36} = 0.63$, quinine hydrochloride: $\rho_{36} = 0.80$, sucrose: $\rho_{36} = 0.67$; all $p < 0.01$) and qYN (citric acid: $\rho_{35} = 0.71$, sodium chloride: $\rho_{34} = 0.60$, quinine hydrochloride: $\rho_{34} = 0.76$, sucrose: $\rho_{33} = 0.76$; all $p < 0.01$); see Fig. 2. To gain a better understanding of the nature of the individual differences between Test and Retest thresholds, we constructed Bland-Altman plots for each tastant in both methods (Fig. 3). The 95% confidence intervals of the mean differences always included 0, providing no evidence of systematic differences between sessions. We then derived the coefficients of repeatability (CR); the results of two measurements would be expected to differ no more than CR in 95 % of the cases. The respective CRs for QUEST and qYN

299 were, in \log_{10} mM: 0.97 and 0.84 for citric acid (corresponding to 3.6 and 3.1 concentration steps), 0.98
 300 and 0.95 for sodium chloride (3.6 and 3.5 steps), 1.07 and 1.29 for quinine hydrochloride (4.7 and 5.6
 301 steps), and 1.10 and 1.03 for sucrose (3.7 and 3.4 steps). Accordingly, the mean CR across all tastants
 302 was 3.90 for both procedures.

Table 1. Results of the threshold measurements during Test and Retest for QUEST and qYN.

Method	Substance	N	Session	Threshold in \log_{10} mM			
				mean	min	max	SD
QUEST	Citric Acid	37	Test	-0.141	-1.564	1.350	0.621
			Retest	-0.256	-1.540	1.385	0.666
	Sodium Chloride	38	Test	1.140	-0.417	2.495	0.631
			Retest	1.069	-0.432	2.100	0.544
	Quinine-HCl	38	Test	-1.737	-3.514	0.496	1.101
			Retest	-1.889	-3.514	0.339	0.953
	Sucrose	38	Test	1.054	-0.592	2.414	0.705
			Retest	1.089	-0.563	2.194	0.660
qYN	Citric Acid	37	Test	-0.446	-1.807	0.812	0.563
			Retest	-0.558	-1.508	0.574	0.598
	Sodium Chloride	36	Test	0.785	-0.457	2.052	0.607
			Retest	0.831	-0.369	2.222	0.571
	Quinine-HCl	35	Test	-1.974	-3.409	0.239	0.952
			Retest	-1.980	-3.369	0.091	0.962
	Sucrose	36	Test	0.871	-0.613	2.243	0.677
			Retest	0.765	-0.742	1.881	0.669

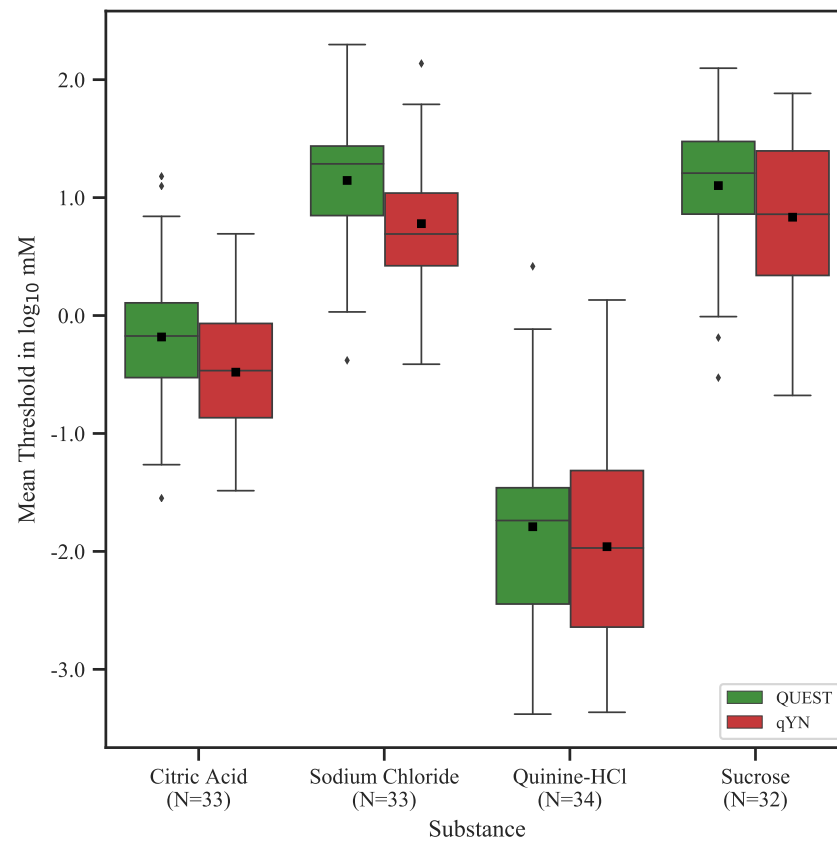


Figure 1. Distributions of the means of Test and Retest threshold estimates, split by tastant and method procedure. Squares indicate the mean, and whiskers correspond to $1.5 \times$ inter-quartile range. Only participants for whom threshold data for both sessions and methods was available are shown for each tastant; the number of included participants is given below the respective abscissa labels.

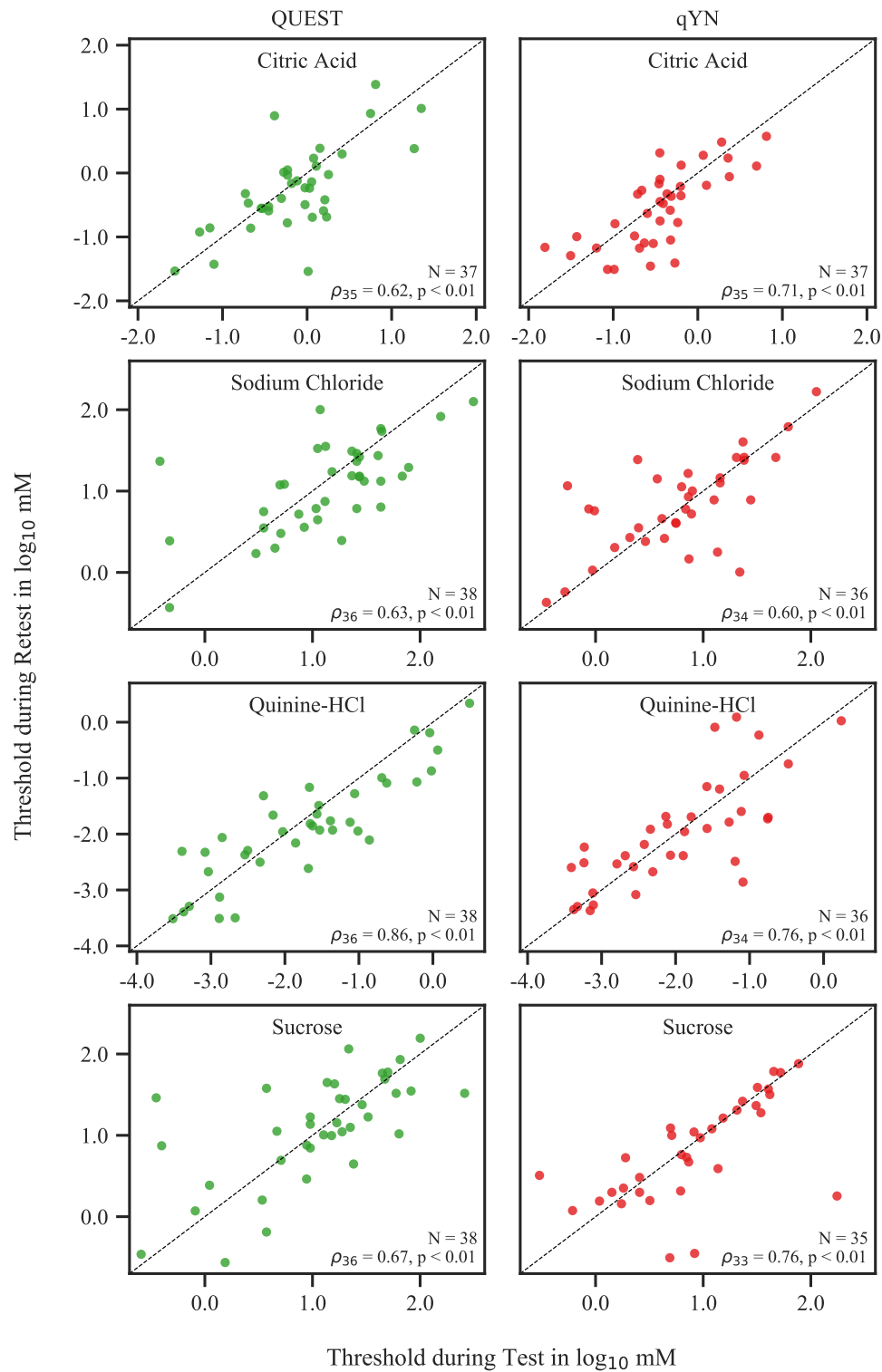


Figure 2. Correlation between Test and Retest threshold estimates for QUEST and qYN. Each point represents one participant; the dashed line is the identity line.

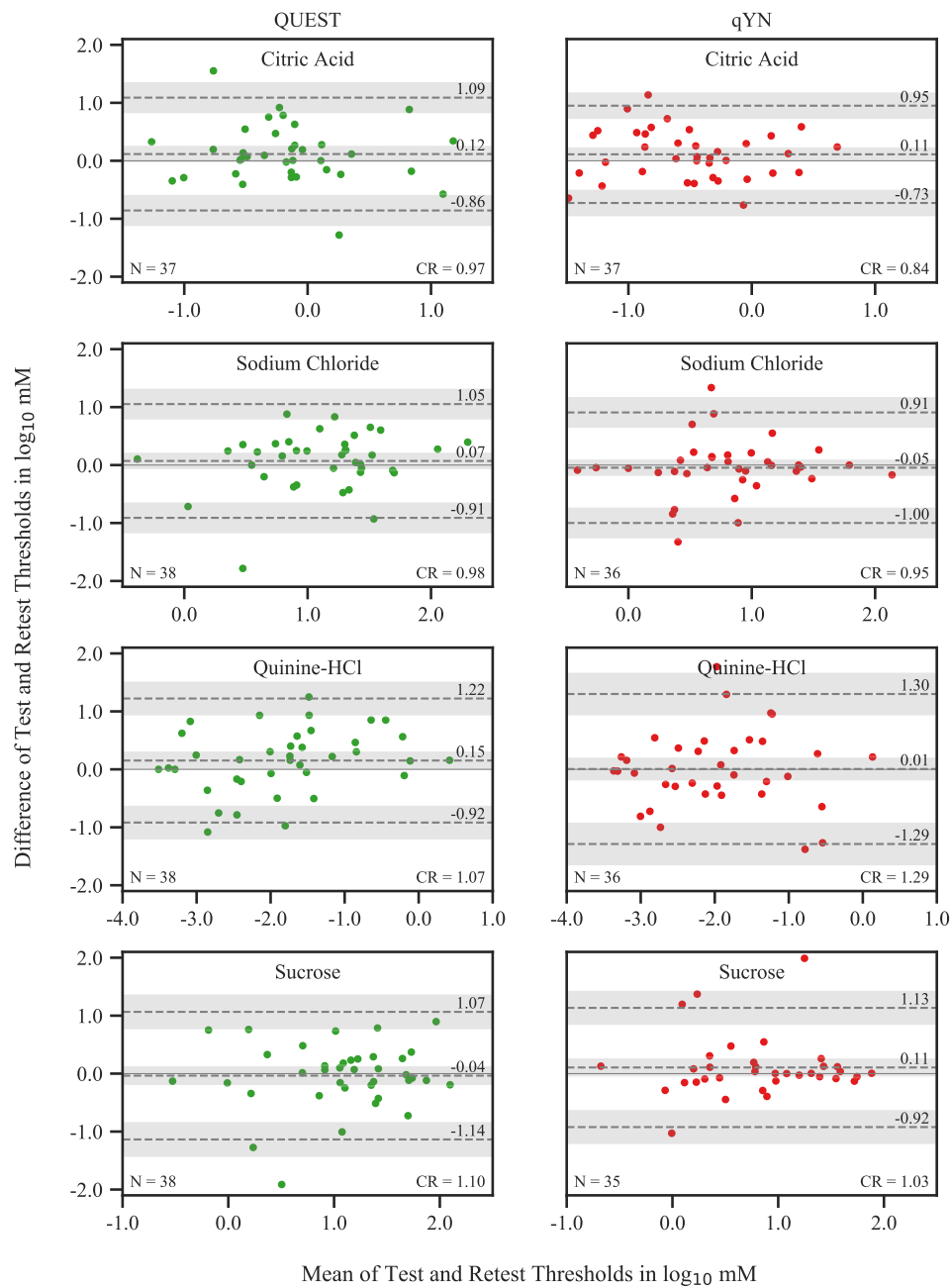


Figure 3. Bland-Altman plots showing differences between Test and Retest thresholds for QUEST and qYN, mean difference \bar{d} , and limits of agreement (LoA) corresponding to 95% confidence intervals (CIs) as $\bar{d} \pm 1.96 \times \text{SD}$. The shaded areas represent 95% CIs of these estimates.

3.3.1. False-alarm rates and psychometric function slopes

To investigate potential differences between tastants and across sessions in estimated FARs and psychometric function slopes, we conducted two rmANOVAs with the factors *tastant* and *session*, with FAR and slope as the respective dependent variables.

For FARs, we found no main effects and no interaction of the factors (all $p > 0.25$), indicating there was no evidence that response criteria would systematically vary with tastants or across sessions. We therefore decided to pool all data points, yielding a mean FAR of 0.060 (SD=0.034) spanning across a large range (1.7%–21.6%). The Bland-Altman plot revealed an increasing variability of the differences between Test and Retest as the magnitude of FARs increased (Fig. 4). This finding can be interpreted such that some participants would expose a relatively high FAR in one session, but a small FAR in the other. For the majority of participants, however, FARs varied within a relatively narrow range. Because the FAR differences between sessions were obviously not normally distributed, we \log_{10} -transformed the data for the calculations of mean difference, CR, LoAs, and the corresponding CIs. The results were then back-transformed to the original scale of the data [17]. As can be seen in Fig. 4, the back-transformation does not yield a single value, but a line spanning across the session means. The resulting CR was $0.969 \times \text{session mean}$.

Mean d' slopes pooled across sessions were 1.51 (SD=0.54) for citric acid, 1.66 (SD=0.60) for sodium chloride, 1.47 (SD=0.57) for quinine hydrochloride, and 1.44 (SD=0.46) for sucrose. We found a significant main effect of *tastant*, albeit with a small effect size ($F_{3,96} = 3.05$, $p = 0.04$, $\eta_p^2 = 0.09$). This finding suggests that the ability to discriminate between stimuli of adjacent concentration steps systematically shifted with the presented tastants. A post-hoc t -test revealed that the *tastant* effect was driven by a significant difference between sodium chloride and sucrose slopes ($t_{33} = 2.857$, $p_{\text{holm}} = 0.045$, $d = 0.497$). There was no effect of *session* and no interaction between the factors (both $p > 0.19$). Bland-Altman plots for all tastants are shown in Fig. 5. In agreement with the rmANOVA results, differences between sessions were not significantly different from zero, as indicated by the confidence intervals spanning across 0. CRs ranged from 0.99 (sucrose) to 1.32 (sodium chloride), which is large, considering the mean slopes.

There was a significant correlation between FARs and slopes ($\rho_{302} = -0.741$, $p < 0.001$), indicating that higher FARs were associated with reduced steepness of the d' sensitivity function.

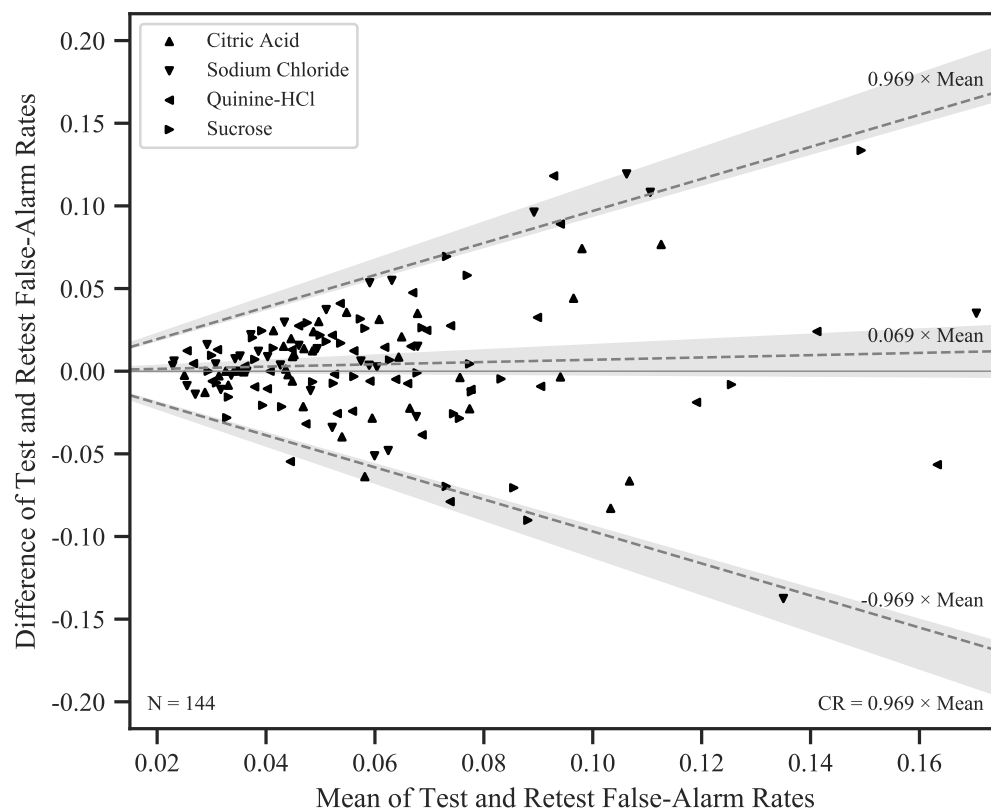


Figure 4. Bland-Altman plot for qYN false-alarm rates (FAR). Markers show differences between Test and Retest FARs. Since the variability of session differences increases with session means, the data was \log_{10} -transformed before calculating mean difference \bar{d} , limits of agreement (LoA) corresponding to 95 % confidence intervals (CIs) as $\bar{d} \pm 1.96 \times \text{SD}$, and 95 % CIs of these estimates (shaded areas). The figure shows the *back-transformed* parameter values, plotted on the original scale of the data. Due to the back-transformation, the lines representing \bar{d} and LoAs have a slope $\neq 0$, i.e., they are not parallel to the abscissa, and we provide their respective formulas. Note that the intercepts of all lines were 0, and are therefore omitted.

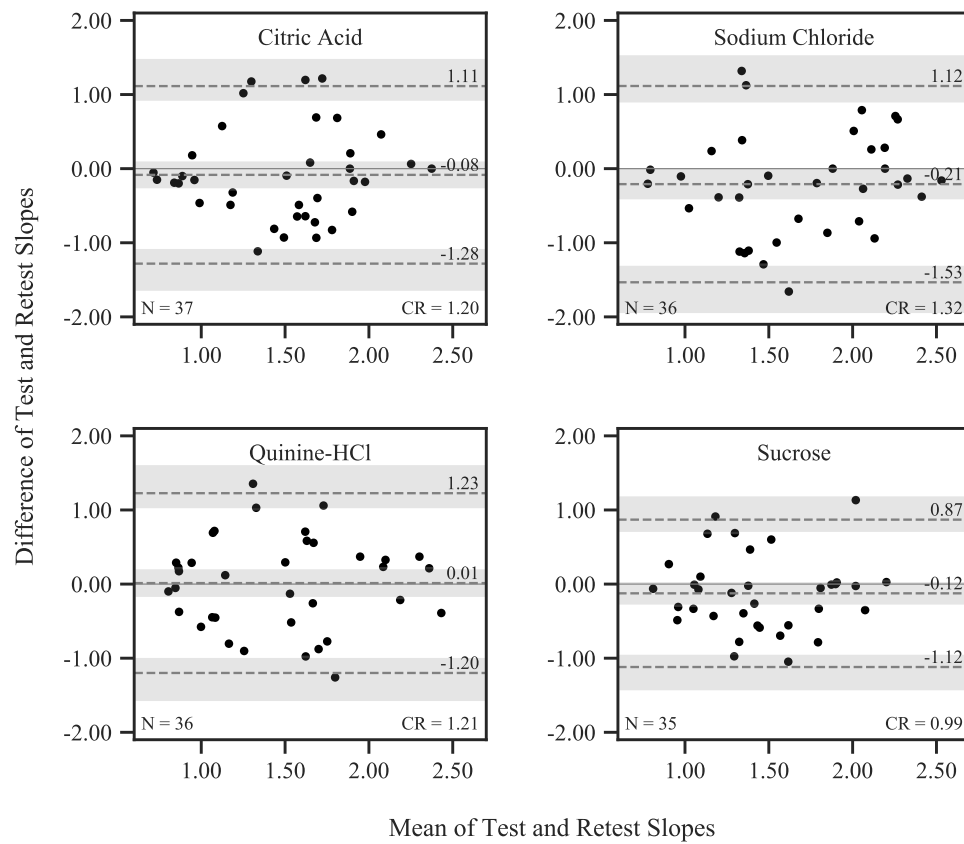


Figure 5. Bland-Altman plots showing differences between the slopes of the estimated sensitivity functions (d') in qYN Test and Retest, mean difference \bar{d} , and limits of agreement (LoA) corresponding to 95% confidence intervals (CIs) as $\bar{d} \pm 1.96 \times \text{SD}$. The shaded areas represent 95% CIs of these estimates.

4. Discussion

Using two Bayesian methods, based on *QUEST* [3] and *quick Yes-No* (qYN, [9]), we explored the impact of false-alarm rate and psychometric function slope estimation on the precision and accuracy of taste sensitivity measurements.

4.1. Taste recognition thresholds

The comparison of the session means for each procedures showed slightly, but systematically higher thresholds for *QUEST* compared to qYN for citric acid, sodium chloride, and sucrose; the difference was not significant for quinine hydrochloride, which we believe can be attributed to the larger variability of bitter threshold measurements compared to the other taste qualities. Since the exact shape and parametrization of the psychometric function and the definition of "threshold" differed between both methods, we expected the estimated threshold values to differ.

4.2. Threshold repeatability

QUEST and qYN, both, thresholds showed good monotonic relationships across sessions, as indicated by test-retest correlations ranging from $\rho = 0.62$ to $\rho = 0.86$ for *QUEST* and from $\rho = 0.60$ to $\rho = 0.76$ for qYN. No method produced consistently higher correlations than the other: qYN correlations were stronger for citric acid and sucrose, while *QUEST* showed higher correlations for sodium chloride and quinine hydrochloride. These values compare very well with a past applications of the *QUEST* method [1] and also with a modified Harris-Kalmus [20] and a forced-choice staircase [21] procedure. The latter two approaches required notably longer testing times than *QUEST*. In contrast, much smaller correlations have been observed, for example with the relatively quick three-drop method and with taste strips [22].

The observed correlation coefficients do not, however, account for systematic changes occurring between sessions, and do not necessarily honor the spread of the data and the slope associated with their relationship [13]. Therefore, we a) created Bland-Altman plots [2,14,23] to visualize the distributions of differences; and b) calculated coefficients of repeatability (CR; [15]) as an estimate of the expected measurement differences between sessions in an individual participant. We found that thresholds did not vary systematically across sessions. qYN produced smaller CRs than *QUEST* – indicating better agreement between measurement repetitions – for all tastants except quinine hydrochloride. The CR averaged across tastants was identical for both procedures at an equivalent of 3.90 concentration steps. Using a *QUEST* procedure in a 3-AFC task to estimate smell thresholds [8], we previously observed a CR corresponding to approx. 5.3 concentration steps on a \log_{10} grid with a step width of 0.300, which is similar to the step width used here for sucrose, and larger than the step width for all other tastants. Neglecting the task differences (yes-no in the present study versus 3-AFC in [8]), the *QUEST* procedure seems to perform better for taste than for smell measurements.

We would like to emphasize a discrepancy between the estimated correlation coefficients and CRs. In *QUEST*, the highest correlation was found for quinine hydrochloride; yet, repeatability according to CR was better for sodium chloride and citric acid. In fact, repeatability was best for citric acid, yet the corresponding correlation was the *worst* of all of the four tastants. Similarly, in qYN the highest correlations were found for quinine hydrochloride and sucrose, but the respective CRs were highest, i.e. repeatability was lowest, for these tastants. The highest repeatability in the entire study was found for citric acid in qYN, yet the associated correlation was only found to be in a medium range.

Correlation coefficients are commonly adopted to quantify repeatability in the chemical senses literature, and should therefore be calculated to enable comparisons with previously published studies. Yet, thorough examination of the correlated data is required to ensure that the conclusions drawn from these analyses are not inadvertently erroneous. We, therefore, suggest to always visualize the data in a scatter plot and the identity line to uncover systematic changes between measurements, which can occur even if the data points are highly correlated. In order to better understand the spread

and pattern of measurement differences, Bland-Altman plots and coefficients of repeatability (CR) should be derived [2,13–15,23]. CRs indicate the magnitude of differences to expect when applying an experimental procedure repeatedly, and help guide the decision whether that procedure is suitable for a particular investigation, e.g., a clinical assessment.

4.3. False-alarm rates and d' slopes

False-alarm rates (FARs) and d' slopes were assessed in qYN runs. FARs did not differ between tastants, indicating that participants' response criterion was not taste-specific. Furthermore, FARs were generally low ($\leq 6\%$), suggesting that most participants complied with the instruction to be conservative in their response behavior. This finding was further substantiated through inspection of FAR differences between Test and Retest, which revealed little variation in most participants. Yet, for a few participants, the variance of session differences grew as the magnitude of FARs increased, leading to a relatively high FAR in one session, but a much smaller FAR in the other.

Slopes of the sensitivity function (d') only differed between sodium chloride and sucrose. CRs for slopes were relatively large. While this parameter of the psychometric functions is known to be notoriously difficult to estimate, especially if the number of trials is small [9], slope and FAR are also directly linked: for a given sensitivity threshold, a lower FAR will lead to a steeper psychometric function while a higher FAR demands a reduction in steepness. We found evidence for this dependence, as FAR and slope were strongly negatively correlated.

Differences in FAR between sessions can be interpreted such that participants followed different cognitive strategies in the two measurements, i.e., they changed their response criterion. These changes in FAR, then, would inevitably affect the slope as well and *vice versa*. Inspection of the trial sequences of experimental with the highest FARs (in the 90th percentile and above) revealed that, here, participants had indeed responded "yes" to stimuli of very low concentrations that were clearly below threshold. This shows that it is more likely that FAR changes lead to slope adjustments, than *vice versa*. Overall, the results support our premise in the QUEST procedure that FARs are low and stable in the majority of participants.

4.4. Measurement duration

On average, QUEST finished 2:40 min quicker than qYN, thanks to its dynamic termination criterion that ends the experimental run when the confidence interval around the threshold estimate reaches a predefined low limit. qYN, on the other hand, always completes 20 trials because no termination criterion was set in order to ensure sufficient data for the simultaneous estimation of the three parameters, threshold, FAR, and slope. Whether the amount of testing time required for qYN could be reduced by employing a similar dynamic stopping criterion needs to be tested in future studies.

4.5. Sensitivity and food preference

Additionally to taste psychometric functions, we also assessed food and taste quality preferences. The data revealed no clear link between taste preference and taste sensitivity with the exception of a positive association between sour liking and sucrose as well as sour threshold, though the latter association did not reach significance. This means that sour liking was higher in participants with lower sucrose and citric acid sensitivity (higher threshold). Whether this association has the potential to shape food preferences and intake remains unanswered in the present study, as the reported food frequencies appeared to be unrealistically low and could, therefore, not be used for further analysis. The observation that obese children with a low sour taste preference along with poor sour taste abilities improved through the course of a weight loss intervention, indicates, however, that dietary changes may influence preference as well as taste function, at least to some extent [24].

5. Conclusions

We compared the repeatability of taste recognition threshold estimates produced by two adaptive procedures, QUEST and quick Yes-No (qYN), for citric acid, sodium chloride, quinine hydrochloride, and sucrose. Both procedures select stimulus concentrations such that – based on a participant's entire response history – the expected information gain about the true parameter(s) of a psychometric function is maximized. While QUEST only assesses the threshold, qYN also adjusts false-alarm rate (FAR) and slope. Our analysis consisted of the widely adopted calculation of correlation coefficients between repeated measurements, and the estimation of coefficients of repeatability (CR) to assess the expected difference between two measurements of the same participant. The magnitudes of test-retest correlations were generally good and not clearly in favor of either threshold method. The CRs, however, revealed slightly better repeatability of qYN thresholds for citric acid, sodium chloride, and sucrose, compared to QUEST. The good agreement between both methods together with the low FARs observed in qYN suggest that, overall, participants applied a conservative response criterion as instructed.

6. Data and software availability

The data reported in this paper along with graphical representations of individual thresholds are available for download from <https://doi.org/10.5281/zenodo.3540535>.

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References

1. Höchenberger, R.; Ohla, K. Rapid Estimation of Gustatory Sensitivity Thresholds with SIAM and QUEST. *Frontiers in Psychology* **2017**, *8*. doi:10.3389/fpsyg.2017.00981.
2. Bland, J.M.; Altman, D. Statistical methods for assessing agreement between two methods of clinical measurement. *The Lancet* **1986**, *327*, 307–310. doi:10.1016/s0140-6736(86)90837-8.
3. Watson, A.B.; Pelli, D.G. Quest: A Bayesian adaptive psychometric method. *Perception & Psychophysics* **1983**, *33*, 113–120. doi:10.3758/bf03202828.
4. Hardikar, S.; Höchenberger, R.; Villringer, A.; Ohla, K. Higher sensitivity to sweet and salty taste in obese compared to lean individuals. *Appetite* **2017**, *111*, 158–165. doi:10.1016/j.appet.2016.12.017.
5. Green, D.M.; Swets, J.A. *Signal detection theory and psychophysics*; Wiley: New York, NY, 1966.
6. Macmillan, N.A.; Creelman, C.D. *Detection Theory: A User's Guide*, 2nd ed.; Lawrence Erlbaum Associates, Inc.: Mahwah, NJ, 2010.
7. Klein, S.A. Measuring, estimating, and understanding the psychometric function: A commentary. *Perception & Psychophysics* **2001**, *63*, 1421–1455. doi:10.3758/bf03194552.
8. Höchenberger, R.; Ohla, K. Estimation of Olfactory Sensitivity Using a Bayesian Adaptive Method. *Nutrients* **2019**, *11*, 1278. doi:10.3390/nu11061278.
9. Lesmes, L.A.; Lu, Z.L.; Baek, J.; Tran, N.; Doshier, B.A.; Albright, T.D. Developing Bayesian adaptive methods for estimating sensitivity thresholds (d') in Yes-No and forced-choice tasks. *Frontiers in Psychology* **2015**, *6*. doi:10.3389/fpsyg.2015.01070.
10. van Strien, T.; Frijters, J.E.R.; Bergers, G.P.A.; Defares, P.B. The Dutch Eating Behavior Questionnaire (DEBQ) for assessment of restrained, emotional, and external eating behavior. *International Journal of Eating Disorders* **1986**, *5*, 295–315. doi:10.1002/1098-108x(198602)5:2<295::aid-eat2260050209>3.0.co;2-t.
11. Nagl, M.; Hilbert, A.; de Zwaan, M.; Braehler, E.; Kersting, A. The German Version of the Dutch Eating Behavior Questionnaire: Psychometric Properties, Measurement Invariance, and Population-Based Norms. *PLOS ONE* **2016**, *11*, 1–15. doi:10.1371/journal.pone.0162510.
12. King-Smith, P.E.; Grigsby, S.S.; Vingrys, A.J.; Benes, S.C.; Supowit, A. Efficient and unbiased modifications of the QUEST threshold method: Theory, simulations, experimental evaluation and practical implementation. *Vision Research* **1994**, *34*, 885–912. doi:10.1016/0042-6989(94)90039-6.
13. Vaz, S.; Falkmer, T.; Passmore, A.E.; Parsons, R.; Andreou, P. The Case for Using the Repeatability Coefficient When Calculating Test–Retest Reliability. *PLoS ONE* **2013**, *8*, e73990. doi:10.1371/journal.pone.0073990.
14. Bland, J.M.; Altman, D.G. Measuring agreement in method comparison studies. *Statistical Methods in Medical Research* **1999**, *8*, 135–160. doi:10.1191/096228099673819272.
15. Bland, J.M.; Altman, D.G. Applying the right statistics: analyses of measurement studies. *Ultrasound in Obstetrics and Gynecology* **2003**, *22*, 85–93. doi:10.1002/uog.122.
16. Carkeet, A. Exact Parametric Confidence Intervals for Bland-Altman Limits of Agreement. *Optometry and Vision Science* **2015**, *92*, e71–e80. doi:10.1097/oxp.0000000000000513.
17. Euser, A.M.; Dekker, F.W.; le Cessie, S. A practical approach to Bland-Altman plots and variation coefficients for log transformed variables. *Journal of Clinical Epidemiology* **2008**, *61*, 978–982. doi:10.1016/j.jclinepi.2007.11.003.
18. Peirce, J.; Gray, J.R.; Simpson, S.; MacAskill, M.; Höchenberger, R.; Sogo, H.; Kastman, E.; Lindeløv, J.K. PsychoPy2: Experiments in behavior made easy. *Behavior Research Methods* **2019**, *51*, 195–203. doi:10.3758/s13428-018-01193-y.
19. Vallat, R. Pingouin: statistics in Python. *Journal of Open Source Software* **2018**, *3*, 1026. doi:10.21105/joss.01026.
20. Wise, P.M.; Breslin, P.A.S. Individual Differences in Sour and Salt Sensitivity: Detection and Quality Recognition Thresholds for Citric Acid and Sodium Chloride. *Chemical Senses* **2013**, *38*, 333–342. doi:10.1093/chemse/bjt003.
21. Mattes, R.D. Reliability of psychophysical measures of gustatory function. *Perception & Psychophysics* **1988**, *43*, 107–114. doi:10.3758/bf03214187.

- 498 22. Mueller, C.; Kallert, S.; Renner, B.; Stiassny, K.; Temmel, A.F.; Hummel, T.; Kobal, G. Quantitative
499 assessment of gustatory function in a clinical context using impregnated "taste strips". *Rhinology* **2003**,
500 *41*, 2–6.
- 501 23. Altman, D.G.; Bland, J.M. Measurement in Medicine: The Analysis of Method Comparison Studies. *The*
502 *Statistician* **1983**, *32*, 307. doi:10.2307/2987937.
- 503 24. Sauer, H.; Ohla, K.; Dammann, D.; Teufel, M.; Zipfel, S.; Enck, P.; Mack, I. Changes in Gustatory
504 Function and Taste Preference Following Weight Loss. *The Journal of Pediatrics* **2017**, *182*, 120–126.
505 doi:10.1016/j.jpeds.2016.11.055.

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